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► **To cite this version:**

Roger Collet, Elise Boucq, Jean-Loup Madre, Laurent Hivert. Long term dynamics of inequalities between French households concerning automobile. 12th World Conference on Transport Research, Jul 2010, Lisbonne, Portugal. 17p. hal-00615156

**HAL Id: hal-00615156**

**<https://hal.science/hal-00615156>**

Submitted on 18 Aug 2011

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# LONG TERM DYNAMICS OF INEQUALITIES BETWEEN FRENCH HOUSEHOLDS CONCERNING AUTOMOBILE

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## ABSTRACT

For each quartile of income per household or per consumption unit, annual time-series have been estimated from panel surveys, with annual waves of observations from 1974 to 2007:

- INSEE<sup>2</sup> Households' "Conjoncture" survey from 1974 to 1994,
- panel "Parc-Auto"<sup>3</sup> Sofres<sup>4</sup> since mid-80's.

In these data sources, household behaviour is described through:

- car ownership (percentage of households with at least one car, of which percentage of multi-car households, average number of cars per adult over 18, which is the minimum age for driving license in France),
- car use (annual mileage per household or per car).

The repeated sample structure of data has been used for improving the accuracy of time-series of variables highly correlated for subsequent years [Cochran, 1977].

In mid-70's, car ownership and use were quite low for the poorest income quartile, but the difference has much decreased with all the three higher income groups, which are more homogeneous. Thus, multi-car ownership, which is mainly structured by geographic and demographic determinants, has slowed down -but not reversed- the social diffusion of automobile.

As the curves representing car ownership (number of cars per adult) and car use (annual mileage per household) seem to become quite horizontal during the most recent period, logistic curves have been estimated according to time, then to real income. For each quartile

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<sup>3</sup> "Car fleet" panel.

<sup>4</sup> TNS-Sofres : Taylor-Nelson, *Société Française d'Etudes par Sondages*, a French private polling institute.

of the distribution of households by income per consumption unit, saturation thresholds are estimated, as well as the date or the point of inflection. Follows a discussion on the legitimacy of pooling the data. The relationship between temporal elasticities (for each quartile) and cross-sectional income elasticities, which can be considered as a measurement of inequality at each point in time, will be discussed [Gardes and Madre, 2005].

*Keywords: panel, automobile, car ownership, car use, income inequalities, saturation, France.*

## **1. INTRODUCTION**

In many countries like in France, most taxes on automobile were progressive, when they had been established. They have become neutral or even regressive because of the social diffusion of car ownership and use ([Madre, 1985], [Purwanto et al., 2002]). Has multi-car ownership made this evolution slower? The answer to this question is probably different in low density areas, where there is no alternative to automobile, and in large conurbations, where several destinations can be reached by foot or by bicycle and where public transport are available. Expenditure for car purchase are still concentrated in high income groups, as well as for toll motorways or parking, because cheaper (e.g. second hand cars) or free alternatives exist ([Madre, 1991], [Berri, 2005], [Berri et al., 2009]).

First, the data from two panel surveys, on which rely the calculation of annual time-series since the waves of year 1974 for each quartile of income distribution, will be presented in section 2. Then, the methodologies used for the calculation of these time-series (interpolated quartiles, optimised estimate of time-series from panel surveys) and the inequality indicators will be described in section 3. A descriptive overview will give the first results (section 4). Finally, modelling in terms of logistic curves ([Røed-Larsen, 2006]) will be implemented in section 5: do poor/medium/rich households follow the same trajectories for car ownership and use according to the evolution of their income? What results can be derived in terms of relationship between cross-sectional and longitudinal income-elasticities [Madre and Gardes, 2005]?

## **2. MORE THAN 30 YEARS OF ROTATING PANEL SURVEY DATA**

This research is based on two annual nationwide household surveys describing both car ownership and use:

- the "*Enquête de Conjoncture Auprès des Ménages*" (ECAM) conducted by the National Institutes of Statistics (INSEE) among a sample of dwellings drawn from the French census: 10,000 to 13,000 households responding by interview each year, of which about one third had been also interviewed one year before ; the period from 1974 to 1994 (year of end of this survey) is covered by the data files available at INRETS;

- the “*Parc-Auto*” (Car Fleet) survey is a postal survey conducted by the private marketing research company TNS-Sofres; each annual wave sample includes 6,000 to 7,000 volunteer respondent households, of which about 3/4 have already responded the year before (even if having moved, contrary to ECAM survey); data files are available at INRETS for all waves since 1984 and this survey is still on-going ([Hivert, Pean de Ponfilly, 2000], [Hivert et al.,2006]). For a short description of data, see also [Kalinowska, 2005] and [Papon, Hivert, 2007, 2008].

Despite these differences in survey methodology, we have checked that these data source show consistent results for the period 1984-94, when both datasets are available at INRETS.

Both questionnaires contain:

- the annual income of the household in about 10 brackets,
- a description of the household (socio-economics, demographic structure, place of residence, etc.),
- a description of cars (age, type of fuel, main driver, etc.) at permanent disposal of the household (up to 2 cars in ECAM, to 3 cars in Parc-Auto),
- an estimate of the annual mileage for each car described (rounded and heaped ([Hivert, 2001], [Yamamoto, 2009], but unbiased according to odometer reading), as well as an information on main purposes of vehicle use (used/non used).

Thus, this information has to be homogenised mainly for the income, which has been coded in brackets using different grids of nominal income over time (see below).

### **3. CALCULATION OF TIME-SERIES FOR EACH QUANTILE OF INCOME, METHODOLOGICAL ASPECTS**

For the ECAM surveys conducted by the National Institute of Statistics, about 70% of households give for each person detailed information on their resources, including different sources of income (wages, retirement pensions, social benefits, etc.). About 25% give a global information on a pre-coded grid, and 5 to 10% refuse. Thus, we chose to use it, particularly because the rate of non-response is much lower when using this pre-coded information.

Most of models rely on quantitative variables, but responses given by households are often rounded for income as well as for annual mileage. Moreover inflation, which was important before mid-80's, make comparisons difficult over time. That is why income grids have been revised in 1977, 1983, 1987 and 1997. The grids have been also typically revised in 2002 in order to be converted from Franc to Euro.

Rather than conventional methods ([Bhat, 1994]) like simulated residuals ([Lollivier and Verger, 1989]), we preferred a more robust method: interpolated quantiles from the middle of each bracket ([Madre and Purwanto, 2003]). It could be any quantiles (e.g. terciles or

quintiles) but we have first to check that the number of brackets is much larger than the number of quantiles.

Let us consider the distribution of a variable of interest (e.g. the number of cars in the household) by income bracket. In order to locate the limits of each income quantile, the distribution is interpolated, and the average number of cars per household is calculated in each quantile with the rather strong hypothesis, that car ownership is constant inside each bracket containing a limit between two quantiles (e.g. first quartile Q1, median or third quartile Q3).

This method has been tested on the respondents of the 1993-94 French National Travel Survey, which gave a precise answer in terms of income. The result is obviously better, when the upper and lower limit of a quantile are nearer to the thresholds of brackets on the grid. But surprisingly enough, splitting the sample into a larger number of brackets when dividing the middle of each class by the number of consumption units does not improve the quality of the interpolation, because even inside a bracket the income level is correlated with the composition of the household. Thus the interpolation is more precise for total household income than for the income per consumption unit, which is a more adequate measurement of the standard of living.

After obtaining homogenized income groups through quantiles, how to estimate time-series accurately enough despite of the small sample size of our panel survey data ? Referring to [Cochran, 1977], we have implemented the method optimising the accuracy of the time-series for the most recent periods. This optimisation is crucial for the estimation of models using variables affected with small annual changes (often 1 or 2%). However we had to smooth several time-series by moving averages over three consecutive years at national level and over five years for Paris region, which represents about 15% of the total household sample.

For the measurement of inequalities, we have adopted three indicators:

- the Gini index [Gini, 1921],
- the ratio "Q4/Q1", i.e. between the means of the variable of interest for the extreme quartiles (Q1 representing the poorest and Q4 the richest),
- the cross-sectional income-elasticity, with usual references (1 when the variable of interest increases proportionally with income and 0 when it is not influenced by income).

More sophisticated indicators could be considered, for example those proposed by [Atkinson, 1970], [Atkinson et al., 1980, 1982] (also cited in [Madre, 1985], [Berri, 2005]). In order to avoid heterogeneity over time due to different income grids, these indicators have been calculated from only four points, which are the mean values for each quartile.

## 4. DESCRIPTIVE ANALYSIS

For each quartile of the distribution of income per consumption unit, time-series have been estimated since 1974 for :

- the average number of cars per adult (i.e. aged > 18, which is the minimum for driving license in France),
- the proportion of equipped households,
- the proportion of multi-car households among equipped households,
- the average annual mileage per household or per car,

separately nationwide, and for the Metropolitan Area of Paris (Paris region) in order to show the specificity of a high population density area, where the diffusion of automobile has started earlier than elsewhere in France. Generally, the figures for intermediate quartiles Q2 and Q3 are quite close, thus they have been represented by a single curve.

Nationwide (figure 1), the inequalities of income per household have decreased between mid-70's and mid-80's, then have remained almost constant till 2000 except a temporary increase during the 1993-94 recession. Since 2001, high incomes are growing faster than those of medium or low income groups.

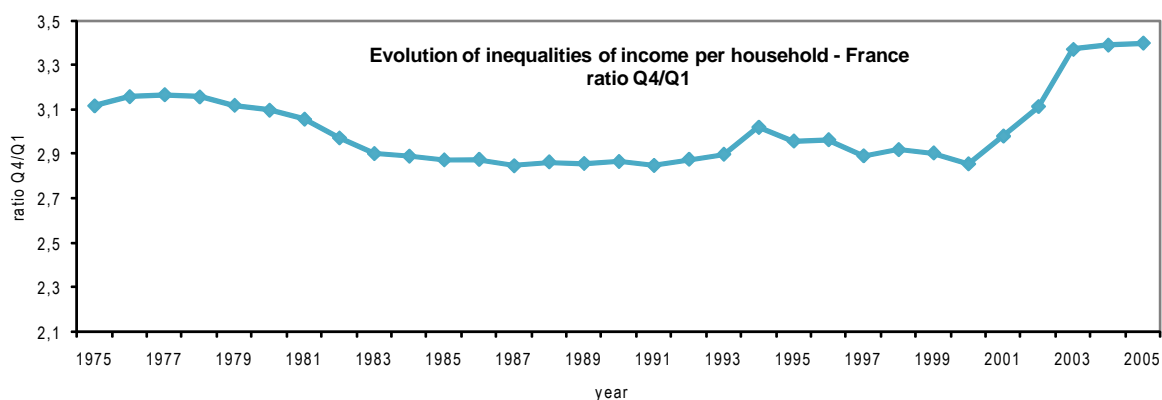


Figure 1: Evolution of inequalities for income per household (ratio of extreme quartiles Q4/Q1) - France

In Paris region (figure 2), inequalities have decreased more rapidly during the first period, then have remained constant only from mid-80's to early 90's, and increase since the 1993 recession. Because of the relative stability of households structure, the evolutions of income per consumption unit follow the same patterns, as shown since 1994 on Parc-Auto data.

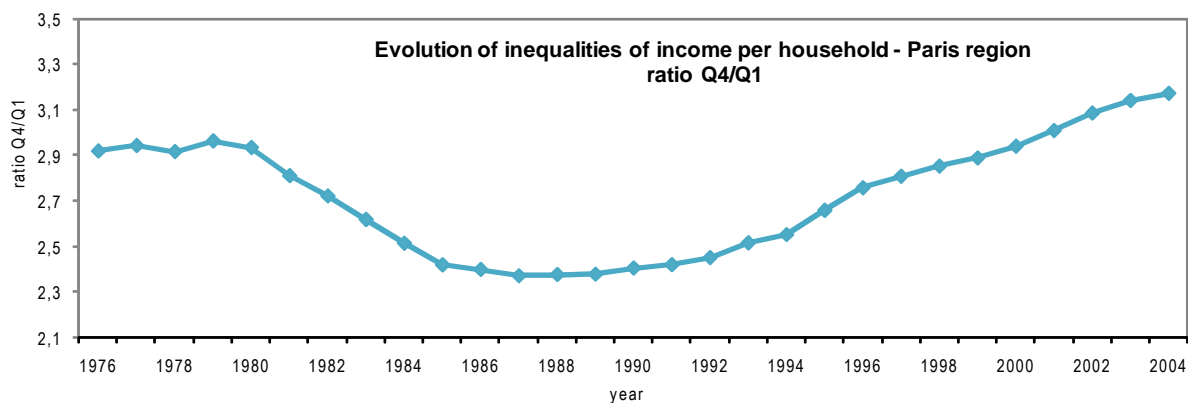


Figure 2: Evolution of inequalities for income per household (ratio of extreme quartiles Q4/Q1) - Paris region

For the whole country, the ratio Q4/Q1 was of 2.1 for the average number of cars per adult in 1974 (figure 3). It has remained almost constant around 1.7 between 1982 and 1993, then it has decreased to 1.4 in 2005. The cross-sectional income-elasticity of the number of cars per adult has followed the same pattern, from 0.6 in mid-70's to 0.2 around 2005, questioning the hypothesis that cross-sectional elasticities represent long term elasticities.

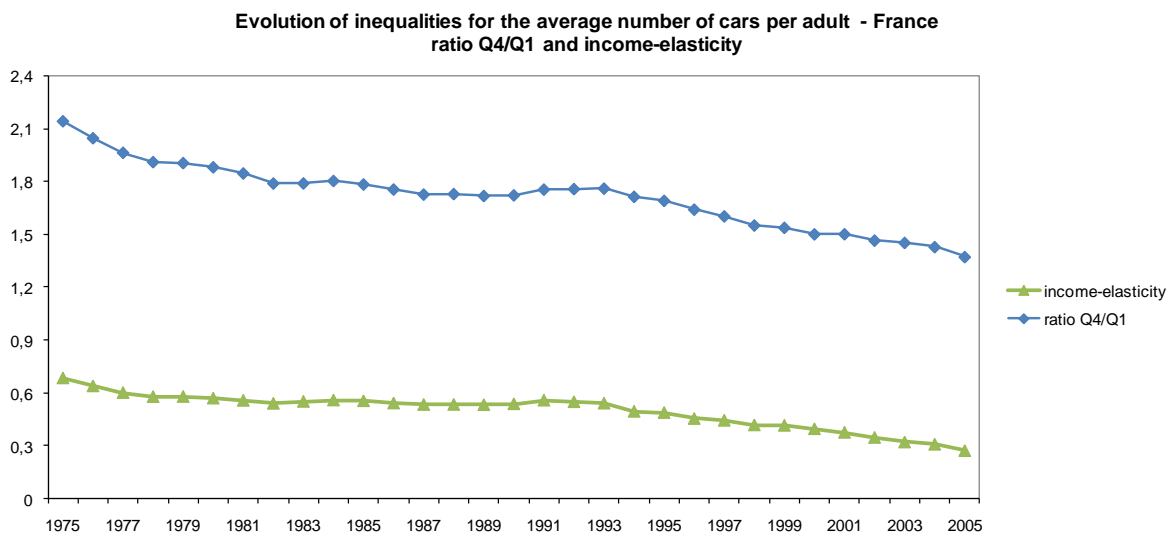


Figure 3: Evolution of inequalities for the average number of cars per adult - France

In the special case of Paris region, the ratio Q4/Q1 was only 1.8 in the mid-70's and has remained almost constant at 1.5 since 1982. Indeed, car ownership has been higher in the conurbation of Paris than in the rest of France since the mid-60's, but high population density with good public transports limits the expansion of car ownership.

Thus, multi-car ownership has not reversed the social diffusion of automobile, but it has stopped it in Paris region since the early 80's and nationwide between mid-80's and mid-90's.

As seen in figure 4, the decrease in inequality for car ownership during the first period is mainly due to the poorest quartile (Q1) catching up with the upper ones: the proportion of households without car has dropped from 55% in mid-70's to 35% in mid-90's. During the same decades, the proportion of multi-car households among equipped households has

developed more slowly in Q1 than in richer quartiles, especially Q2 and Q3 catching up with Q4.

Then, figure 4bis presents the evolution of ownership in terms of average number of cars per adult. Taking into account that average number of adults per household remains higher in lower quartiles (e.g. from 2.29 in 1994 to 1.90 in 2006 for Q1, versus from 1.77 to 1.61 in Q4), the average number of cars per households is finally almost the same in 2006 (1.05 in Q1 versus 1.06 in Q4), so the Q4/Q1 ratio decreased from 1.24 in 1994 to 1.0 in 2006.

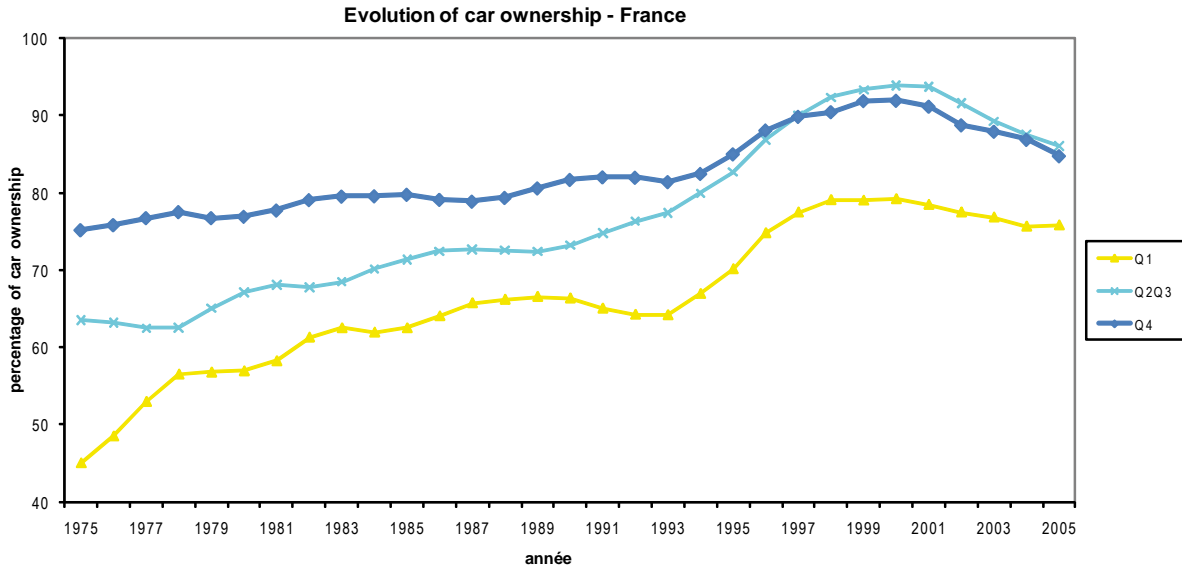


Figure 4: Evolution of car ownership with time: share of equipped households by quartile of income per consumption unit (Q2 and Q3 together, smoothed by mobile averages MA3) – France

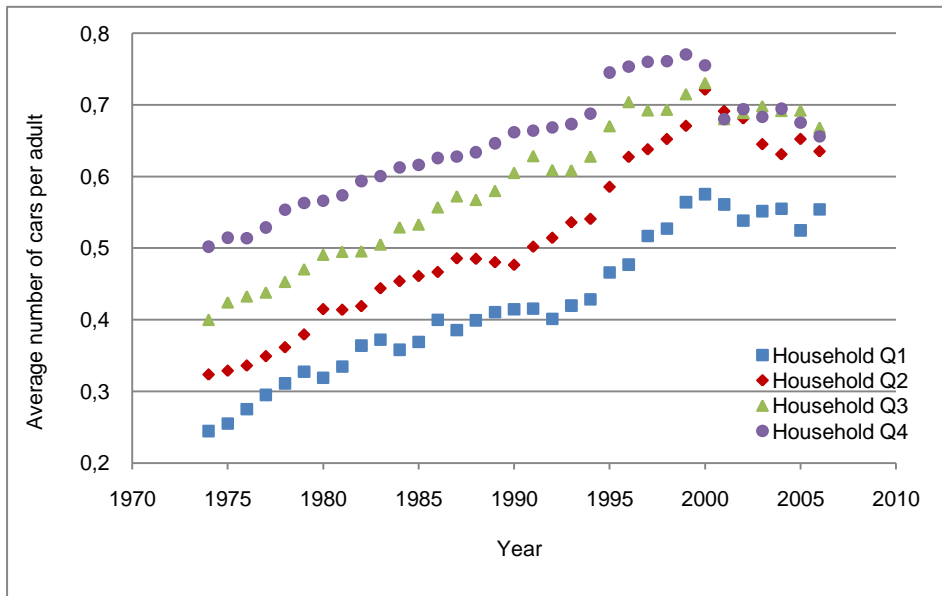


Figure 4bis: Evolution of car ownership with time: average number of cars per adult by quartile of income per consumption unit - France

From mid-90's, different zones can be identified. In the most densely populated areas, the income-elasticity of multi-car ownership dropped from 0.7 during the late 90's to 0.1 in 2005:



Q1 curve is increasing, while those for Q2 to Q4 are decreasing. In Paris region at least since 1995, there is almost no differentiation of multi-car ownership according to income. In low density areas the elasticity decreases more continuously, as well as the ratio Q4/Q1.

Mid-90's, the increase of ownership in terms of cars per adult (especially visible for Q4) can be explained by incentives introduced by two successive Governments in France (see below).

A new grant for scrappage ("*prime à la casse*") has been recently introduced in France (early 2009). It has succeeded to previous incentives for scrapping old cars implemented in the middle of the nineties :

- decided by the Balladur Government, a scrap bonus of 5,000 francs (762 euros) was implemented from February 1994 to June 1995 to purchase a new car, when scrapping a vehicle more than ten years old. This "balladurette" measure aimed to help a depressed market and boost consumption, but also cleaned up somehow the French car fleet.
- under the next Juppé Government, from October 1995 to October 1996, another measure allowed any owner to benefit a bonus when scrapping a vehicle more than 8 years old and under 3.5 tonnes. This premium amounted to 5,000 francs to buy a new car of economic or low range, and to 7,000 francs (1,067 euros) for higher-end vehicles. This premium, known as "Juppette" was also intended to support the Government decision of toughening automobile technical inspections (originally introduced in 1992).

The annual mileage per household (figure 5) has increased till 2000, then decreased, due to a lower mileage per vehicle induced by a more expensive fuel (peak of fuel price in 2000, then a more continuous increase from 2004 to mid-2008). This inflection is described in detail for example in [Hivert et al., 2008] and in [Kemel et al., 2009]. This decline is more important in urban areas than in low density areas, especially for medium and high income groups (Q2 to Q4).

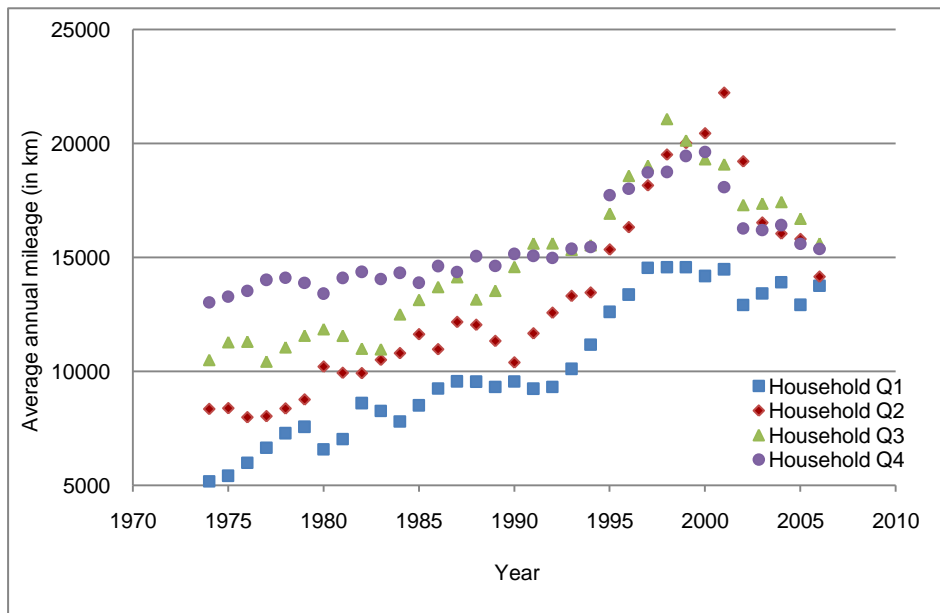


Figure 5: Evolution of average annual mileage per household by quartile of income per consumption unit – France

In terms of annual mileage per vehicle, the ratio Q4/Q1 has declined from 1.8 in 1984 to 1.1 in 1995, because of a slower increase of car use in high than in low income groups. This period corresponds to decreasing fuel price: sharp decline of oil price between mid-1985 and mid-1986, followed by an increasing proportion of diesel vehicles using a 30% cheaper fuel. Between mid-90's and 2005, the ratio Q4/Q1 has decreased in low density areas, while it has increased in urban areas, because car use can be reduced by low income households only where an alternative exists (public transport or slow modes).

Thus, after a period of status quo between mid-80's and mid-90's, the social diffusion of automobile continues like during the 70's, especially in low density areas, whose inhabitants are more car dependant even if their income is low.

The two next figures present, for each quartile by consumption unit, the evolution of car ownership (per adult) and car use by household income, in constant 2006 Euros (considering all the annual waves together).

Globally, the scatter of points in figure 6 shows that, as expected, household car ownership has increased with their income. Indeed, the number of cars per adult has been about 0.45 for a real income of €10,000 while it is steady about 0.7 for incomes over €25,000 (constant 2006 Euros). The shape of the scatter of points suggests a concave augmentation of car equipment with income. However, the slope seems different when comparing each of the four quartile. Indeed, it is decreasing when the household has a higher social position. Particularly, car ownership is increasing faster with real income for the household Q1 than for the household Q4. Over the years, this has induced a reduction of social inequalities concerning car ownership, as it is shown in figure 3.

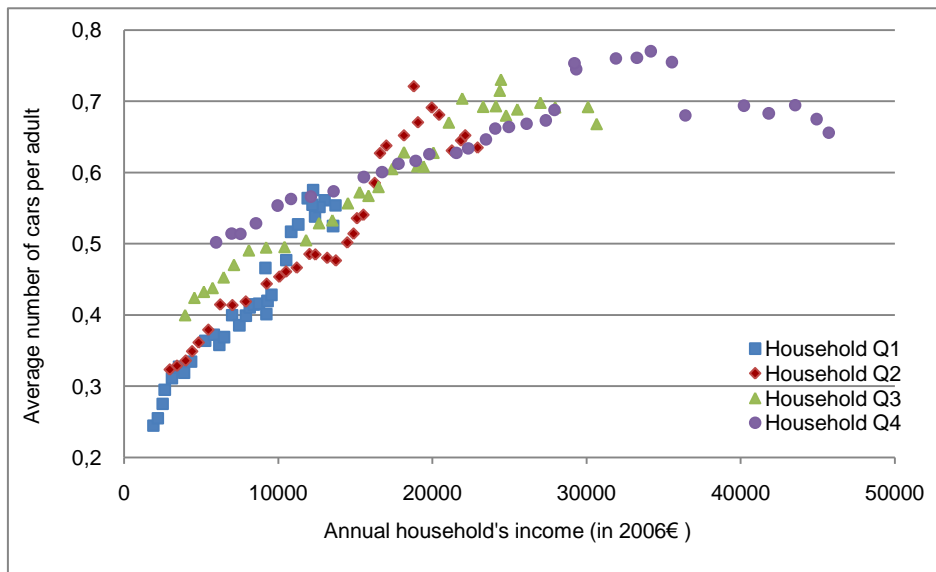


Figure 6: Evolution of car ownership (average number of cars per adult) with household income - France

Considering car use in figure 7, analog conclusions roughly emerge. The average mileage of households has been about 12,500 km/year for an annual income of €10,000. Over €25,000, it has globally ranged between 15,000 and 20,000 km/year. Even if it seems less obvious than in figure 6, the scatter of points in figure 7 also suggests a concave growth of household car use with their real income. Regarding the quartile specific households, the slope of the annual mileage over real income also seems to decrease with their position in the income scale.

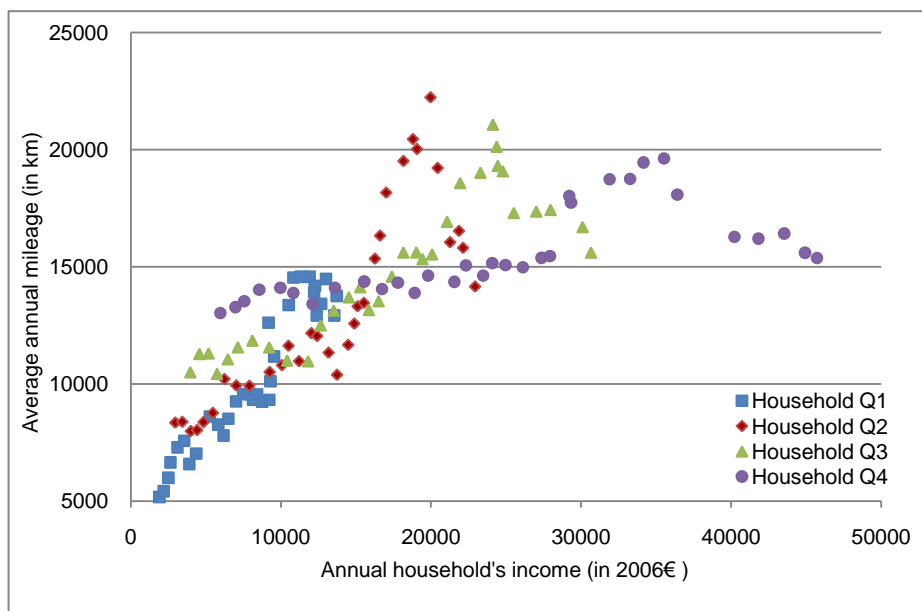


Figure 7: Evolution of car use (annual mileage per household) with household income - France

The concave shapes in figures 6 and 7 suggest that the diffusion of automobile can reach saturation thresholds for both car ownership and use when the households are getting wealthy. Theoretically, the social diffusion of a good (either in time or in the income scale) can be represented using a sigmoid curve, ended by a saturation level. Referring to figures 6 and 7, our data seem to cover only the upper part of an assumed sigmoid. It is modelled in

the following section. In comparison, time (figures 4bis and 5) is clearly less relevant than income to fit a sigmoid curve. However, it is also tested as explanatory variable.

## 5. ESTIMATION OF LOGISTIC CURVES TO MODEL CAR OWNERSHIP AND USE

The average households for each of the four quartiles of annual income per consumption unit are observed annually, during 33 years from 1974 to 2006. Let Q1, Q2, Q3 and Q4 refer to these households by increasing order of resource. For both car ownership and use, the figures in section 4 suggest the upper part of a sigmoid curve. In this section, we assume that they are generated by logistic functions. The logistic model, which is exposed below, is applied on the data of each quartile specific household and on the pooled dataset.

### 5.1. The model

Let  $y_{it}$  refer either to the number of cars per adult or to the annual mileage in  $10^4$  kilometres for the household  $Q_i$  at period  $t$ . Both these variables are modelled separately assuming a logistic specification. The explanatory variable is denoted  $x_{it}$  and stands successively for the time (section 5.2) and the real income (section 5.3). Thus, the model is given by:

$$y_{it} = \frac{L}{1 + \exp(-\alpha x_{it})} \quad (1)$$

where  $\varepsilon_{it}$  is assumed to be i.i.d. along a  $N(0, \sigma^2)$ , and where  $\alpha, L$  are the parameters to be estimated. For  $\alpha$  positive, the formulation (1) implies that  $y_{it}$  is increasing with  $x_{it}$  along a symmetrical sigmoid, bounded by two horizontal asymptotes: the lower plateau is fixed at  $L/2$  while the upper plateau, corresponding to a saturation level, is located at  $L$ . The inflection point, for which the second derivative of  $y_{it}$  with respect to  $x_{it}$  is zero, is located at  $(x_{it} = 0; y_{it} = L/2)$ .

### 5.2. Time as explanatory variable

In this section, the explanatory variable of model (1) is the time. Precisely,  $x_{it}$  is given by the index  $t$ , with  $t=0$  for the year 1974,  $t=1$  for 1975... and  $t=33$  for 2006. Tables 1 and 2 report the estimates for car ownership and use respectively.

Table 1: Estimates for the car ownership model as a function of time

| Parameters  | Pooling              | Hh.Q1                | Hh.Q2                | Hh.Q3                | Hh.Q4                |
|---|----------------------|----------------------|----------------------|----------------------|----------------------|
|   | 0.07                 | 0.06                 | 0.06                 | 0.08                 | 0.10                 |
|   | -0.05*               | -0.66                | -0.52                | 0.04*                | 0.66                 |
|   | -0.29                | -0.27                | -0.17                | -0.26                | -0.31                |
|   | -4.83                | -7.58                | -7.03                | -7.84                | -6.98                |
| Mean log-likelihood   | 1.00                 | 2.37                 | 2.09                 | 2.40                 | 2.07                 |
| Nb. of observations   | 132                  | 33                   | 33                   | 33                   | 33                   |
| Saturation threshold<br>(# of cars per adult)<br>[90% conf. interval] | 0.74<br>[0.63, 0.87] | 0.76<br>[0.61, 0.96] | 0.84<br>[0.72; 0.98] | 0.77<br>[0.73, 0.81] | 0.74<br>[0.71, 0.77] |
| Date of inflexion   | 1975                 | 1986                 | 1982                 | 1973                 | 1967                 |

Note: the dependant variable is the number of cars per adult in households. The explanatory variable is time. All the parameters are significant at the 90% level, except those indicated by \*.

As expected, the results show that car ownership and use have increased along time. Indeed, the parameter  $\beta$  is found to be positive and significant in every case on tables 1 and 2. For both car ownership and use, the inflection has occurred in the mid 80's for the two lower quartiles (Q1 and Q2), in the 70's for Q3 and in the 60's for the higher income quartile Q4. The highest saturation level is for Q2, then it decreases till Q4 because higher income groups live in more densely populated areas. Saturation thresholds are quite low for Q1.

Considering the car ownership model, pooling the data does not seem illegitimate since the confidence intervals for saturation level overlap around 0.74 cars per adult (table 1). The same stands for the car use model. Indeed, the confidence intervals for the saturation level in table 2 overlap around an annual mileage of 20,000 km per household.

Table 2: Estimates for the car use model as a function of time

| Parameters  | Pooling                 | Hh.Q1                   | Hh.Q2                   | Hh.Q3                   | Hh.Q4                   |
|---|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
|   | 0.06                    | 0.07                    | 0.07                    | 0.07                    | 0.05                    |
|   | -0.32*                  | -0.83                   | -0.79                   | -0.19*                  | 0.66                    |
|   | 0.71                    | 0.56                    | 0.81                    | 0.74                    | 0.65                    |
|   | -2.70                   | -4.75                   | -3.27                   | -3.94                   | -4.27                   |
| Mean log-likelihood                                       | -0.06                   | 0.95                    | 0.21                    | 0.55                    | 0.71                    |
| Nb. of observations                                       | 132                     | 33                      | 33                      | 33                      | 33                      |
| Saturation threshold<br>(km/year)<br>[90% conf. interval] | 20400<br>[16300, 25600] | 17500<br>[14600, 21100] | 22600<br>[16800, 30200] | 21100<br>[17800, 24900] | 19300<br>[15600, 24000] |
| Date of inflexion   | 1979                    | 1985                    | 1985                    | 1977                    | 1961                    |

Note: the dependant variable is the annual mileage of households in kilometres. The explanatory variable is time. All the parameters are significant at the 90% level, except those indicated by \*.

However, the resources of households are a more robust factor than time to explain the diffusion of automobile. Let's now consider income growth as determinant in model (1).

### 5.3. Real income as explanatory variable

In this section, household car ownership and use are modelled using real income as explanatory variable. Particularly,  $Y_{it}$  in model (1) refers here to the real annual income of the household  $Q_i$  at period  $t$ , expressed in  $10^4$  constant 2006 Euros.

Pooling the 132 corresponding points to fit the model (1), the saturation levels are estimated at 0.72 cars per adult and 17,800 yearly km per household (tables 3 and 4). These thresholds are not very far from the levels already reached. However, when adjusting the model separately for each quartile, these estimated thresholds are higher. Indeed, the saturation threshold for the number of cars per adult ranges from 0.73 for the wealthiest household (Q4) to 1.23 for the poorest one (Q1). This later value seems quite high, but the pertained confidence interval is very wide. Indeed, the estimation of the saturation threshold is much more accurate for higher income groups, which are currently near to their estimated asymptote.

Here, pooling the data is more questionable since the confidence intervals for the four quartile households overlap just at 0.74, which is the upper limit of the quite narrow confidence interval obtained from pooled data. Thus, while the estimates of the saturation level are less accurate using income instead of time as explanatory variable, we obtain a very accurate estimate of the saturation threshold from the income model when assuming that these data can be pooled.

In the meantime, the estimated saturation mileage ranges from 18,200 km/year for the household Q4 to 24,800 km/year for the household Q2. This is higher than in the pooling case. Therefore, the pooling estimates might be affected by an heterogeneity bias. But it is not sure, because the four confidence intervals overlap, and the estimate for the pooled data (17,800 km/year) is the lower limit of the intersection of these four intervals.

Regarding the saturation levels, the estimates show a regular decrease of the number of cars per adults with the economic position of households in the income scale. The same stands for the saturation threshold of car use, except for the household Q1 which has a greater estimated threshold (22,300 km/year) than the household Q2. These results can be explained by a location effect. Indeed, richest households can afford to live in dense urban areas where public transports are an alternative to car use. At the opposite, poorest households are more frequently living in outer suburbs or rural areas, where car is often the only possible mode.

Table 3 : Estimates for the car ownership model as a function of real income

| Parameters  | Pooling              | Hh.Q1                | Hh.Q2                | Hh.Q3                | Hh.Q4                |
|---|----------------------|----------------------|----------------------|----------------------|----------------------|
|   | 1.31                 | 0.97                 | 0.84                 | 0.85                 | 0.93                 |
|   | -0.62                | -1.46                | -0.90                | -0.29                | 0.15*                |
|   | -0.32                | 0.21*                | -0.06                | -0.23                | -0.31                |
|   | -6.52                | -7.64                | -6.80                | -7.65                | -6.87                |
| Mean log-likelihood   | 1.84                 | 2.40                 | 1.98                 | 2.40                 | 2.02                 |
| Nb. of observations   | 132                  | 33                   | 33                   | 33                   | 33                   |
| Saturation threshold<br>(# of cars per adult)<br>[90% conf. interval] | 0.72<br>[0.71, 0.74] | 1.23<br>[0.56, 2.71] | 0.94<br>[0.71, 1.24] | 0.79<br>[0.74, 0.85] | 0.73<br>[0.70, 0.75] |
| Income of inflexion<br>(in 2006 Euros)                                | 4700                 | 15000                | 10700                | 3500                 | -1700                |

Note: the dependant variable is the number of cars per adult in households. The explanatory variable is their income expressed in  $10^4$  constant 2006 Euros. All the parameters are significant at the 90% level, except those indicated by \*.

Table 4: Estimates for the car use model as a function of real income

| Parameters  | Pooling                 | Hh.Q1                   | Hh.Q2                   | Hh.Q3                   | Hh.Q4                   |
|---|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
|   | 1.52                    | 1.48                    | 0.99                    | 0.78                    | 0.56                    |
|   | -0.79                   | -1.38                   | -1.16                   | -0.51                   | 0.51                    |
|   | 0.57                    | 0.80                    | 0.90                    | 0.77                    | 0.59                    |
|   | -3.45                   | -4.69                   | -3.18                   | -3.88                   | -4.18                   |
| Mean log-likelihood                                       | 0.31                    | 0.92                    | 0.17                    | 0.52                    | 0.67                    |
| Nb. of observations                                       | 132                     | 33                      | 33                      | 33                      | 33                      |
| Saturation threshold<br>(km/year)<br>[90% conf. interval] | 17800<br>[17100, 18500] | 22300<br>[14300, 34700] | 24800<br>[15900, 38600] | 21700<br>[17800, 26400] | 18200<br>[16200, 20300] |
| Income of inflexion<br>(in 2006 Euros)                    | 5200                    | 9300                    | 11800                   | 6500                    | -9200                   |

Note: the dependant variable is the annual mileage of households in kilometres. The explanatory variable is their income expressed in  $10^4$  constant 2006 Euros. All the parameters are significant at the 90% level.

For a better representation in the left side of the logistic curves, the model (1) has also been fitted using the log of real income,  $\ln(\text{real income})$ , instead of real income in level. The logistic adjustment does not perform better using this alternative specification: pooling the data, the mean log-likelihood slightly decrease to 1.76 for the car ownership model and 0.27 for the car use model. The estimated saturation levels are higher: 22,700 km per household for annual car use, and 1.05 cars per adult. These estimates are closer to those obtained from the previous household-specific adjustments using real income in level (in tables 3 and 4). However, [Hivert, 2006] showed that the number of cars in a household does not exceed the number of driving license holders in Germany and in France. The contrary is observed in the US, where there are much more specialized vehicles (SUV, vans, urban cars, etc.).

The model (1) assumes the symmetry of the adjusted logistic curves with respect to the inflection point. This can be viewed as a strong hypothesis and asymmetric curves might be more relevant in our context (for an example of application, see [Ricketts and Head, 1999]). Particularly, the estimated income of inflection for the household Q4, when assuming symmetry, is found to be negative for both car ownership and use. This unrealistic result pleads for an asymmetric shape of the logistic curves. However, the assumption of symmetry is made necessary when considering our data. As we mentioned above, the scatter of points of the annual mileage as well as the one of the number of cars per adult against the annual income when pooling observations suggest a strictly concave shape, corresponding to the upper part of a sigmoid (figures 6 and 7). In other words, the data do not cover the lower part of the sigmoid, at the left side of the real inflection point. As there is no empirical information about this part, its representation has to rely on some assumptions. A solution would have been to extend the observation period farther in the past, back to the end of World War II. Unfortunately, detailed data about car use and equipment between 1945 and 1973 are not available.

## 6. CONCLUSIONS

The social diffusion of car is a major feature of economic growth in Occidental Europe after World War II. This paper describes this phenomenon in France for four income groups from 1974 to 2006. Multi-car ownership has interrupted but not reversed this long term trend : after

a period of status quo between mid-80's and mid-90's, the social diffusion of automobile continues like during the 70's, especially in low density areas, whose inhabitants are more car dependant even if their income is low.

However, the growth of car ownership becomes slower and slower, which shows that saturation levels are not far from already reached levels. In order to determine these thresholds, we have adjusted logistic curves separately for each income quartile, then on pooled time-series. Geographic factors (as density, distance to city-centre, size of conurbation) probably explain why it is for the second quartile that saturation levels are the highest, which could be checked by estimating the same models only for the inhabitants of the Metropolitan Area of Paris. Moreover, pooling these data does not seem illegitimate because the saturation level for pooled series just belongs to the intervals determined by the four quartiles. It lets us suspect some heterogeneity, mainly due to the geographic factors mentioned above. It is only through more narrow confidence intervals obtained from longer time-series or more adequate functional forms that we could get a real proof of this heterogeneity.

This heterogeneity opens theoretical discussions about cross-sectional versus longitudinal estimates. Another example is an important change over time of cross-sectional income elasticities, which makes impossible to consider them as a proxy for long term longitudinal elasticities. There is still much to do with these data. For instance, introducing fuel price as explanatory factor for car use should allow to derive saturation thresholds for the annual mileage at different fuel price levels, which could give an important information for building scenarios for sustainable development. However we should take into account that even fuel price elasticity depends on income level [Kemel et al., 2009].

## REFERENCES

- Atkinson, A. B. (1970). On the Measurement of Inequality, Faculty of Economics and Politics, University of Cambridge, England, *Journal of Economic Theory* 2, 244-263.
- Atkinson, A.B. and J.E. Stiglitz (1980). *Lectures and Public Economics* », London, New-York, Mc Graw-Hill.
- Atkinson, A.B. and F. Bourguignon (1982). The Comparison of Multi-dimensioned Distributions of Economic Status, *Review of Economic Studies*, vol. XLIX.
- Berri, A. (2005). *Dynamiques de la motorisation et des dépenses de transport des ménages Analyses sur données individuelles et semi-agrégées*, PhD in Economics, University Paris I -Panthéon-Sorbonne.
- Berri, A., S. Vincent Lyk-Jensen, I. Mulalic and T. Zachariadis (2009). Transport consumption inequalities and redistributive effects of taxes: a comparison of France, Denmark and Cyprus, 26èmes Journées de Microéconomie Appliquée, 4-5 june, Dijon.
- Bhat, C.R. (1994). Imputing a continuous income variable from grouped and missing income observations, *Economics Letters*, Vol. 46, pp. 311-319.
- Bureau, B., M. Glachant, E. Boucq and J.-L. Madre (2009). *Équité de scénarios de tarification de l'automobile : une analyse quantitative sur les données du panel Parc*



- Auto TNS-Sofres, PREDIT Go11 report for the French Department of Transport (MEDAD-DRAST, now MEDM-DRI).
- Cochran, W.G. (1977). Sampling Techniques, John Wiley and sons.
- Gini, C. (1921). Measurement of inequality of income, in Economic Journal 31 (1921), 22-43.
- Hivert, L. and J. Pean De Ponfily (2000). Inégalités d'accès à l'automobile, disparités d'équipement et d'usage entre les ménages les plus pauvres et les plus riches, INRETS report for the French Agency for the Environment and Energy Management (ADEME).
- Hivert, L. (2001). Le parc automobile des ménages, étude en fin d'année 1999 à partir de la source "Parc Auto" SOFRES, INRETS report for the French Agency for the Environment and Energy Management (ADEME), 180 p.
- Hivert, L. (2006). L'automobile en Europe, Changements de comportements d'équipement et d'usage, PREDIT3 Go1 report for the French Agency for the Environment and Energy Management (ADEME), 159 p.
- Hivert, L., P. Ottmann, Z. Krakutovski, D. Legros, B. Chlond, J.-L. Madre and D. Zumkeller (2008). Multimodalité et tendances démographiques - *CADMIUM : Comparative Analysis of Demographic Modifications and Impact on the longitudinal Use of Modes*, Deufrako-Predit3 report, INRETS & IfV Karlsruhe Universität for the French Agency for the Environment and Energy Management (ADEME), 202 p.
- Kalinowska, D. and L. Hivert (2005). Comparability and comparison of French and German passenger car purchase, ownership and utilization : emphasizing the diverging developments between petrol and diesel driven vehicles, mission report for COST (355) Action (ref. COST-STSM-355-01420), Arcueil, may, 22p.
- Kemel, E., R. Collet and L. Hivert (2009). How do French motorists react to a multi-annual fuel price increase ? An econometric analysis based on 1999-2007 panel data, International Association for Travel Behaviour Research, 12<sup>th</sup> ICTBR, Jaipur, 13-18 December.
- Lollivier, S. and D. Verger(1989). D'une variable discrète à une variable continue : Une application de la méthode des résidus simulés dans *Mélanges Economiques, Essais en l'honneur d'Edmond Malinvaud*. Editions de l'EHESS, 1094 p.
- Madre, J.-L. (1985). Les interventions publiques dans les transports de voyageurs et leurs effets redistributifs, PhD in Economics, University Paris I - Panthéon-Sorbonne.
- Madre, J.-L. (1991). Quelques éléments sur les effets redistributifs du péage urbain, Entretiens Jacques Cartier, Lyon.
- Madre, J.-L. and A.J. Purwanto (2003). Des quantiles interpolés pour suivre l'évolution des inégalités, XXXIX<sup>ème</sup> colloque de l'ASRDLF, Lyon, 1 -3 September.
- Madre, J.-L. and F. Gardes (2005). Why Long term dynamic elasticities differ from Cross-section elasticities ?, European Transport Conference, Strasbourg.
- Papon, F. and L. Hivert (2008). Adulterous behaviour within the car-owner couple: some analyses from French panel data on car rental and car sharing within households, 11th ICTBR, Kyoto, date, published in special issue "Long-term Dynamics", of IATSS Research Vol.32 No.2, 2008, Tokyo, Japan, 6-15.
- Purwanto, A. J. and J.-L. Madre (2002). The Dynamics of Inequality in Transportation Expenditure between Private Car Users in the Paris Region, International Conference of Automotive Engineers (FISITA), Helsinki, 2-8 june.

- Ricketts, J.H. and G. Head (1999). A five-parameter logistic equation for investigating asymmetry of curvature in baroreflex studies, *American Journal of Physiology* 277 (Regulatory Integrative and Comparative Physiology 46). R441-R454.
- Røed-Larsen, E. (2006). Distributional Effects of Environmental Taxes on Transportation : Evidence from Engel Curves in the United States, *Journal of Consumer Policy*, 29 (3), 301.
- Yamamoto, T. (2009). A random heaping model of annual Vehicle Kilometres Travelled considering heterogeneous approximation in reporting, *International Association for Travel Behaviour Research*, 12<sup>th</sup> ICTBR, Jaipur, 13-18 December.