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Self-Selection and the Efficiency of Tournaments

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Self-Selection and the Efficiency of Tournaments

ABSTRACT

The literature has shown that the overall efficiency of exogenously imposed tournaments is reduced by a high variance in performance. This paper reports results from an experiment analyzing whether allowing subjects to self-select into different payment schemes is reducing the variability of performance in tournaments. We show that when the subjects choose to enter a tournament instead of a piece-rate payment scheme, the average effort is higher and the between-subject variance is substantially lower than when the same payment scheme is imposed. Mainly based on risk aversion, sorting is efficiency-enhancing since it increases the homogeneity of the contestants.

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I. INTRODUCTION

The use of promotion tournaments is fairly widespread especially in the higher ranks of firms and organizations. The incentive property of tournaments has been studied early and extensively in the theoretical literature by Lazear and Rosen (1981), Green and Stokey (1983), Nalebuff and Stiglitz (1983), O'Keeffe, Viscusi, and Zeckhauser (1984); for a survey see McLaughlin (1988). The empirical studies, based on survey or experimental data, are fewer and many survey analyses use sports rather than business data (Prendergast 1999). These studies have confirmed that the efficiency of tournaments depends on the spread between the winner's and the loser's prizes, the number of prizes at stake, the size of the tournament, and the degree of uncertainty faced by the employees.ⁱ

However, both theoretical models and empirical studies also point to some factors that limit the incentive effect of tournaments, such as collusion among employees or employees sabotaging each other, as studied by Lazear (1989) in a theoretical analysis and experimentally by Harbring and Irlenbusch (2005). More generally, most laboratory experiments have provided evidence of tournaments being associated with a high variance in effort [see in particular Bull, Schotter, and Weigelt (1987), Harbring and Irlenbusch (2003), van Dijk, Sonnemans, and van Winden (2001)]. This variance of effort, which is found to be larger in tournaments than in an equivalent piece-rate scheme, reduces the overall efficiency of tournaments.

The principal aim of this paper is to show that previous experimental evidence regarding the variability of effort in tournaments is misleading because the experiments have not accounted for sorting, that is, that agents typically choose to participate in a tournament. The large

variability observed in earlier studies is explained by Bull, Schotter and Weigelt (1987) by the game nature of the tournament, which requires the agents to elaborate a strategy that is more cognitively demanding than the maximizing behavior required by a piece-rate system. Indeed, in addition to the stochastic technology of production, the agents have to cope with strategic uncertainty. Bull, Schotter and Weigelt (1987) show that the variance of effort diminishes when the strategic uncertainty is reduced, for example when the subjects know they are faced with automatons that always select the same level of effort that is also common knowledge. The variance remains high, however, indicating that the discontinuities in the payoff functions themselves contribute to the difficulty of the maximization program and to the high variance of effort. More recent papers, such as Vandegrift and Brown (2003), have shown that the use of high-variance strategies may be related to both the difficulty of the task and the ability of the individuals. The hypothesis that we test in this paper is that the variability of effort may be reduced – and thus the efficiency of tournament increased - by allowing people to choose their payment scheme, i.e., providing them with a choice to enter the competition or not. More precisely, we suggest that the observed high variance of effort may be due to the fact that in previous experiments a competitive payment scheme is imposed on very risk averse or under-confident subjects. For example, facing uncertainty, some of the subjects drop out, i.e., they choose the minimum effort, securing the loser's prize without bearing any cost of effort, whereas others choose the maximum effort, securing the winner's prize but at an inefficiently high cost of effort. Had the subjects been given the choice, like in flexible labor markets where people can choose to enter or shy away from competitive occupations, very risk averse subjects would probably not have entered the competition and the overall variance of effort would be lower.

By testing whether the performance variability is reduced by the *ex ante* sorting effect of tournaments, our paper contributes to a very recent literature about the importance of both incentive and sorting effects in the determination of payment schemes' efficiency, initiated by Lazear (2000).ⁱⁱ This literature shows that sorting influences economic behavior. Earlier, the sorting function of tournaments has mainly been documented with respect to their ability to select *ex post* the best performers. However, their *ex ante* sorting effect is considerably less studied and none of the previous empirical studies have been concerned with the impact of *ex ante* sorting on the variability of performance.ⁱⁱⁱ

To study the *ex ante* sorting effect of tournaments and its impact on the variability of effort, we have designed a laboratory experiment based on the comparison between a Benchmark treatment and a Choice treatment, and involving 120 student-subjects. In the Benchmark treatment, half of the subjects are paid according to a piece-rate payment scheme and the other half enters pair-wise tournaments. This treatment consists of a one-stage game in which the subjects choose their level of effort knowing their payment scheme and the uncertainty of the environment. We find, in line with earlier experiments, that in this treatment, the variance of effort is substantially higher in the tournament than in the piece-rate payment scheme. In the Choice treatment, we add a preliminary stage in which the subjects choose between a piece-rate scheme and a tournament. Those who choose the tournament are paired together. In the second stage, each subject decides on his level of effort. In both treatments, the individual outcome depends on both the effort level and an *i.i.d.* random shock. The difference between the two payment schemes emanates from the strategic uncertainty associated with the tournament setting.

By comparing the subjects' behavior in the two treatments, we can identify the impact of sorting on the average level and the variance of effort. We also seek to identify determinants

of self-selection. The equilibrium effort level is higher in the tournament than under the piece-rate scheme but the expected utility of both compensation schemes is the same. Hence, risk-neutral subjects should be indifferent between the two schemes. For their part, risk averse subjects can adopt a less risky scheme by choosing the piece-rate scheme. We measure the subjects' risk aversion by using the lottery procedure proposed by Holt and Laury (2002).

Our experiment delivers three main findings. First, the key novel finding is that the employees' choice of pay schemes contributes to a considerable reduction in the variance of effort among contestants in the tournament. This result is confirmed by a robustness test in which the subjects are only allowed to choose their payment scheme in the first period of the game and for its whole duration. Second, the average effort is higher when the subjects can select their payment scheme in each period, which suggests that the sorting effect reinforces the incentive effect of both tournaments and variable pay schemes. Third, the subjects self select according to their degree of risk aversion. A cluster analysis identifies a category of under-confident subjects and a category of hesitant ones who both tend to shy away from competition when they can choose their payment scheme. The resulting greater homogeneity of contestants improves the overall efficiency of tournaments. We conclude that in order to understand the origin of the high variance of effort in tournaments, and more generally, the efficiency of a payment scheme, recognition of heterogeneity of preferences is key.

The remainder of the paper is organized as follows. Section 2 presents the theoretical framework and the experimental design. Section 3 gives the experimental procedures. Section 4 describes and analyzes the experimental evidence. Section 5 discusses the results and concludes.

II. THEORY AND EXPERIMENTAL DESIGN

The model

Consider an economy with identical, risk-neutral agents. Agent i has the following utility function, separable in payment and in effort:

$$U_i(e_i) = u(p_i) - c(e_i) \quad (1)$$

with $u(p_i)$ concave and $c(e_i)$ convex.

The production technology is stochastic and output is increasing in the agent's

$$\text{effort: } y_i = f(e_i) + \varepsilon_i \quad (2)$$

with $f(e_i) = e_i$ for the sake of simplicity and ε_i is an i.i.d. random shock distributed over the interval $[-z, +z]$. Only individual outcomes are observable; individual effort is not observable, neither by the principal nor the other agents. The cost function is increasing and

$$\text{convex } c(e_i) = e_i^2 / s \quad (3)$$

with $s > 0$, $c(0) = 0$, $c'(e_i) > 0$ and $c''(e_i) > 0$.^{iv}

In the labor market, some firms pay the agents a piece-rate compensation scheme and other firms use tournaments. If there is a perfect mobility in the labor market at no cost, in the first stage the agents choose their firm (i.e. their payment scheme) and, in the second stage, they decide on their level of effort. Let us first solve the equilibrium effort levels under each mode.

In the piece-rate system, the agent's payment depends only on his own outcome. The payment consists of a fixed wage, denoted by a , corresponding to an input-based payment and a linear

piece-rate, denoted by b , corresponding to an output-based payment. Under this compensation scheme, the agent's utility function becomes:

$$U_i^{PR}(e_i) = a + b.y_i - e_i^2 / s \quad (4)$$

The first order condition is:

$$\delta U_i^{PR} / \delta e_i = b - c'(e_i) = 0$$

Thus, the equilibrium effort of each agent under the piece-rate payment scheme depends positively on the incentive, b , as well as the cost scaling factor, s :

$$e^{PR*} = b.s / 2 \quad (5)$$

In the firms practicing tournaments, the agents play a non-cooperative game with incomplete information like in Lazear and Rosen (1981). In pair-wise tournaments, two prizes are distributed: W is the winner's prize allocated to the agent whose outcome is the highest and L is the loser's prize, allocated to the other agent, with $W > L$. The magnitude of the difference between the two outcomes does not affect the determination of the winner of the tournament.

The agent's utility is:

$$U_i^T(e_i, e_j) = \begin{cases} W - c(e_i) & \text{if } y_i > y_j \\ L - c(e_i) & \text{if } y_i < y_j \end{cases} \quad (6)$$

The agents being symmetric, the probability to win the tournament, $pr(e_i, e_j)$, reduces to the probability that the difference in individual random terms exceeds the difference between individual effort levels: $pr(e_i, e_j) = pr(\varepsilon_i - \varepsilon_j > e_j - e_i)$. Agent i 's expected utility of the tournament is:

$$EU_i^T(e_i, e_j) = L + [pr(e_i, e_j)(W - L)] - e_i^2 / s \quad (7)$$

The maximization program yields the following first order condition:

$$\delta EU_i^T(e_i, e_j) \delta e_i = \delta pr(e_i, e_j) \delta e_i (W - L) - 2e_i / s = 0 \quad (8)$$

We obtain a pure symmetric Nash equilibrium, where effort increases with the prize spread and decreases with both the cost of effort and the size of the shock distribution:

$$e_i^{T*} = e_j^{T*} = (W - L)s / 4z \quad (9)$$

Having determined the equilibrium effort level under each payment scheme, we now turn to the first stage problem. The agent chooses his firm by comparing his expected utility under each payment scheme. He is thus indifferent between the two schemes when:

$$0.5(W + L) - \left[\left[(W - L)s / 4z \right]^2 / s + a + (b^2 \cdot s) / 2 - \left[(b \cdot s) / 2 \right]^2 / s \right] \quad (10)$$

For equation (10) to hold, the expected utility of the tournament must increase as the fixed payment, a , and the variable payment, b , in the piece-rate scheme increase, other things equal. It must decrease when s decreases, i.e., when the marginal cost of effort increases. Indeed, if the marginal cost of effort decreases, with more equilibrium effort, utility from the tournament decreases and compensation with a piece-rate rises faster than the cost of effort, and consequently, the utility from the piece-rate increases. Moreover, a simple comparative-static exercise shows that the tournament should be preferred to the piece-rate scheme if the loser's prize, L , increases, ceteris paribus, or as the variance of the random term becomes large.

The experimental design

The instructions have been kept as close as possible to Bull, Schotter and Weigelt (1987) (see Appendix).

Two treatments. In the Benchmark treatment, after being informed of their compensation schemes, and knowing the cost of each effort level and the distribution of the random term, the subjects have to choose their level of effort. An important difference from the set-up in Bull et

al. (1987) is that in a session, half of the subjects are exogenously and randomly attributed a piece-rate payment scheme and the other half a tournament scheme. The proportion is unknown to the subjects but they are aware of the coexistence of two modes of payment. In contrast, Bull et al. (1987) organized separate sessions in which players were paid either a piece-rate or according to a tournament. Our motivation was to keep the social environment comparable with that of the Choice Treatment in which both schemes coexist in the same session in unknown proportions. The game is repeated 20 times.

The Choice treatment is similar to the Benchmark except that in the first stage of each period, the subjects choose to be paid according to either a piece-rate scheme or a tournament scheme. Those who have opted for the tournament are pooled together and paired. In case of an uneven number of contestants, one subject is randomly chosen and paid according to a piece-rate scheme; he is informed of this before deciding on his level of effort. There is no mobility cost, i.e., the subjects are free to move to the other payment scheme in each new period at no cost. In the second stage of the game, the subjects choose their level of effort.

The design of the game enables between-subject but not within-subject comparisons since each treatment is played by different subjects. The latter would have required submitting all the subjects to the exogenous piece-rate scheme, next to an exogenous tournament, and then to the choice treatment. It would then have become necessary to alternate between the various treatments to control for potential order effects within the Benchmark and between the Benchmark and the Choice treatments. Our design is simpler and allows the subjects to play more repetitions of a same treatment.^v

Matching protocol. Unlike in most experiments on tournaments, we adopt a stranger matching protocol. This is motivated by the constraint of the Choice treatment: if we had used a partner

matching protocol, a subject who is willing to choose the tournament but is paired with a person who always chooses the piece-rate, would be prevented from competing throughout the game. A consequence of our matching protocol is, however, that we reinforce the complexity of the tournament game due to conjectural variations, making it harder to make inferences about the opponent's behavior. Could the use of a random matching process have an impact on the variance of effort? On the one hand, if errors in inferences were the source of the greater variability of effort in tournaments, this should entail a greater variability of effort in our design than in games with fixed pairs. On the other hand, if using a random matching hinders within-pair coordination on either a minimum effort or a maximum effort, this may result in a lower between-group variance of effort than when pairs are fixed. We can, however, disregard both effects; in addition, this should not affect the comparison between treatments since both have been conducted with the same matching protocol.^{vi}

Parameters. Effort can take any integer value in the set: $e_i \in \{0, 1, \dots, 100\}$. In the cost function, $s = 150$, so that $c(e_i) = e_i^2 / 150$. The random shocks vary in the interval $[-40, +40]$. In the tournament, the winner's prize has been set at $W = 96$ and the loser's prize at $L = 45$. In the piece-rate scheme, the fixed wage, a , amounts to 45 and the piece-rate, b , is equal to 0.52, meaning that each unit of outcome gives 0.52 to the agent. These values ensure that the certain payment is the same under both schemes. Without such a fixed wage equal to the loser's prize, it could be rational for a risk-averse agent to choose the tournament and a minimum effort in order not to bear the consequences of a negative random shock on wages under a purely linear piece-rate scheme. Therefore, with our design, only the strategic uncertainty makes a difference between the two schemes.

Given these values, and assuming the agents to be risk neutral and rational, those who are paid according to a piece-rate scheme should provide the effort $e_i^{PR*} = 39$, according to equation (5); those who enter the tournament should provide the effort $e_i^{T*} = 48$, according to the pure strategy Nash equilibrium in equation (9). The players should be indifferent between the two payment schemes since the expected utility of both is the same ($EU_i^{PR} = EU_i^T = 55$), but they have to work harder if they choose the tournament.

Elicitation of risk aversion. The above predictions hold for risk-neutral subjects. One would expect that risk averse subjects i) reduce their effort level under each mode of payment when this is exogenous, and ii) are more likely to stay out of the tournament to avoid the strategic uncertainty due to competition. To elicit the risk aversion of the subjects, we used the lottery procedure proposed by Holt and Laury (2002).

At the end of the sessions (in order not to influence the game), the subjects filled out a questionnaire with 10 decisions (see the instructions in Appendix). Each decision consists of a choice between two paired lotteries, “option A” and “option B”. The payoffs for options A are either €2 or €1.60, whereas the riskier options B pay either €3.85 or €0.10. In the first decision, the probability of the high payoff for both options is 1/10. In the second decision the probability increases to 2/10. Similarly, the chance of receiving the high payoff for each decision increases as the number of the decision increases. When the probability of the higher payoff is large enough, subjects should cross over from option A to option B. Risk neutrality corresponds to a switch at the fifth decision, while risk loving subjects are expected to move earlier and risk averse subjects as from the sixth decision. The subjects made 10 decisions but only one was selected at random for payment.

III. EXPERIMENTAL PROCEDURES

The experiments have been conducted at the GATE laboratory, Lyon, France. The experiment was computerized, using the REGATE software designed by Zeiliger (2000). We recruited 120 under-graduate students from three business or engineering schools, trying to guarantee a fair gender distribution in each session (45,83 per cent of male participants in total). Six sessions with 20 subjects in each were organized; 3 for the Benchmark treatment and 3 for the Choice treatment. Thanks to the 20 repetitions of the game, we collected a total of 2400 observations.

Upon arrival, each subject was randomly assigned a computer. Instructions were distributed and read aloud. Attached to the instructions was a sheet displaying the decision costs associated with each possible effort from 0 to 100. Questions were answered in private. The participants had had to answer a series of questions about the computation of payoffs under each payment scheme. The experiment started once all the participants answered correctly. No communication was allowed.

In the Benchmark treatment, at the beginning of the session and for its whole duration 10 subjects were attributed the piece-rate scheme and 10 the tournament scheme. In the Choice treatment, in each period they had to tick either the “mode X” (piece-rate) box or the “mode Y” (tournament) box to choose their payment scheme for the current period. In both treatments, they selected their effort (“decision number”) by means of a scrollbar. This being done, they had to click a button to generate their “personal random number” that was added to their effort choice to constitute their individual outcome (“result”). Under the tournament

scheme, the computer program compared the outcomes of the two contestants in each pair and determined who was to receive the winner's prize ("the fixed payment M") and who to get the loser's prize ("the fixed payment L"). In case of a tie, a fair random draw determined the allocation of prizes among the pair members. At the end of the period, each subject received a feedback on his payoff and in case of a tournament, on the difference between his outcome and his competitor's outcome. In each new period, the pairs involved in a tournament were randomly reconstituted.

After the completion of the 20 periods, the risk aversion post-experimental questionnaire was distributed and read aloud. Subjects noted on a sheet of paper the option they chose for each of the 10 lottery decisions. After all participants had made their decisions, each subject had to throw a ten-sided die twice: once to select the decision to be considered and a second time to determine her payoff for the option chosen, A or B.

All the transactions, except the lottery, were conducted in points, with conversion into Euros at a rate of 80 points = €1. Payment consisted of the sum of payoffs during each period plus the lottery payment and a €3 show-up fee. On average, the subjects earned €17.40. The sessions lasted approximately one hour, excluding the lottery draw and payment that was made in private in a separate room for confidentiality.

IV. EXPERIMENTAL RESULTS

We first analyze the mean and the variance of effort, before studying the determinants of the payment scheme choice. Last, we examine the heterogeneity in individuals' behavior regarding both the choice of the tournament and the decision of effort.

Mean and Variance of Effort

Table 1 displays summary statistics about the mean and the distribution of effort by payment scheme and by treatment.

First, we check whether we observe both a higher mean and a greater variance of effort under the tournament than under the piece-rate pay scheme, like in the previous experiments. In our Benchmark treatment, we find that the average effort is 46.48 under the piece-rate scheme and it is 53.28 in the tournaments. Both numbers are significantly above the equilibrium effort levels (39 and 48, respectively; t-test, $p=0.000$). As predicted by the model, the agents exert more effort in a competitive setting (Mann-Whitney U test, $p=0.000$). As regards the variance of effort, our results corroborate those of previous experiments. Averaging over all the periods, the total variance is 368.88 under the piece-rate scheme and 652.26 in tournaments. Thus, the variability of effort is clearly higher under the competitive scheme.

Next, we turn to consider the influence on the mean and the variance of effort of the possibility given to the subjects to choose their payment scheme. Table 1 and Figure 1 reveal a substantial increase of the average effort under the tournaments from 53.28 in the Benchmark treatment to 61.57 in the Choice treatment. Interestingly, average effort also increases from 46.48 in the Benchmark to 50.45 for the agents who choose to be paid a piece-rate. As a consequence, the differences relative to the equilibrium effort values are even larger when agents self-select. As for the tournaments, we note that while the subjects on average play the equilibrium effort in the last four periods in the Benchmark treatment, this behavior cannot be observed in the Choice treatment although there is a slight decline in effort over time. The choice of the payment scheme is associated with a slower convergence to equilibrium.

Table 1 and Figure 2 also show a dramatic change in the variability of effort when agents self-select. Comparing the Benchmark with the Choice treatment, we find that the variance under the piece-rate diminishes from 368.88 to 227.87 (-38.23 per cent) and the variance in the tournament decreases from 652.26 to 258.19 (-60.42 per cent). Not only is the variability of effort lower when agents self-select, but now the tournament cannot be considered as being more unstable than the piece-rate. Levene's robust test statistic rejects the hypothesis of equality of variance between the tournament and the piece-rate in the Benchmark treatment ($z=48.929$, $p<0.000$) but accepts it in the Choice treatment ($z=0.135$).

Figure 2 displays the dispersion of effort in tournaments in each treatment. The distribution of effort in tournaments in the Choice treatment is characterized by the following. The median (indicated by the horizontal line) is slightly higher than in the benchmark. The distribution of effort (given by the quartiles, the grey bars) is more concentrated around the median when agents can self-select, whereas effort is more dispersed below the median when they cannot. The adjacent values (the vertical lines) are closer to the median, meaning in particular that contestants chose a zero effort less often (7 observations out of 564) than in the Benchmark treatment (45 observations out of 600).

The variability of effort across the game may be explained by both a time-varying behavior (learning dimension) and time-invariant inter-individual characteristics (heterogeneity dimension). To gauge the relative importance of these two dimensions, in Table 2 we decompose the variance into its within and between components.

The between-subject variance of effort in tournaments accounts for two thirds of the total variance in the Benchmark treatment. It is four times lower and it accounts for less than 40% of the total variance in the Choice treatment.^{vii} Consequently, when people self-select, the

population of voluntary contestants is more homogeneous in terms of exerted effort and the variability of effort is mainly due to an intra-individual component. The within-subject variance of effort in tournaments shows that in the Choice treatment, the variability of effort is lower than in the Benchmark: the subjects learn less or they are less hesitant. Similar differences in the between-subject and within-subject variances are observed for the piece-rate scheme: both are lower when subjects can self-select.

Would we also observe a decrease in the variance of effort if the subjects were allowed to choose which incentive scheme they prefer only once at the beginning of the game and for its whole duration instead of choosing each period? In such an environment, the selection decision is more constraining since the subject cannot switch during the session. To answer this question and to carry out a robustness test, we have designed an additional treatment (the Single-Choice treatment) that replicates the Benchmark treatment, except that at the beginning of the first period, each subject must choose between the piece-rate payment scheme and the tournament for the whole duration of the session. To be consistent with the previous treatments, we have kept a stranger matching protocol. The theoretical predictions remain the same. We have implemented this treatment in one session involving 20 participants.^{viii}

The results of this additional treatment confirm that introducing self-selection reduces the variance of effort in both schemes. Indeed, in the Single-Choice treatment, the variance under the piece-rate is 270.79, which is 26.59 per cent below its level in the Benchmark treatment (368.88), and 18.84 per cent higher than in the Choice treatment (227.87). Similarly, in this new treatment the variance in the tournament is 327.11, which is 49.85 per cent lower than in the Benchmark (652.26) and 26.69 per cent higher than in the Choice treatment (258.19). Levene's robust test statistic does not reject the hypothesis of equality of variance between the tournament and the piece-rate in the Single-Choice treatment ($z=2.570$). In addition, whereas

the between-subject variance represents 66.63 per cent of the total variance of effort in tournaments when no self-selection is allowed and 39.43 per cent in the Choice treatment, the corresponding percentage in the Single-Choice treatment is only 24.67 per cent. We conclude from this additional treatment that allowing people who dislike competition to opt out contributes significantly to the reduction of the variance in tournaments and to a greater homogeneity of contestants, also when the choice is made once and for all future periods.

The descriptive statistics shown above refer to averages. Next, we account for individual characteristics and for the longitudinal character of the data. Table 3 gives the results of random-effects Tobit regressions of the effort decisions, accounting for the censoring of the observations. The results for the Benchmark treatment and the Choice treatment are displayed in the first and the second column, respectively. The independent variables include a time trend to capture learning, a payment scheme dummy to capture the impact of competition, the random shock in the previous period and individual characteristics (the degree of risk aversion and gender). The risk aversion variable (coded from 1 to 10) corresponds to the number of the decision where the subject crosses over from the safer to the riskier option in the lottery test: the higher this number, the more risk averse the subject.

The two treatments have several common determinants of effort. Other things equal, effort declines over time. Competition stimulates performance; the coefficient is larger in the Choice treatment than in the Benchmark treatment. Although it is common knowledge that the periods are independent, the subjects adjust their effort downwards (upwards) when they have received a positive (negative) random shock in the previous period. Alternative regressions (not reported here) in which the time trend has been omitted lead to the same conclusions (i.e., there is no multicollinearity between the time trend and the lagged random number). The main difference between the two treatments is related to the influence of risk

aversion. Risk aversion has a significant negative impact on effort when the payment schemes are imposed on the subjects: considering the uncertainty of the environment, risk-averse subjects reduce their cost of effort. This variable is not significant in the Choice treatment, suggesting that risk aversion plays a role in the sorting process, but not once the choice has been made. It is therefore important to understand what determines sorting.

Sorting

In the Choice treatment, the competitive scheme is chosen in 50 per cent of the cases. Its relative frequency declines slightly from 52.67 per cent in the first ten periods to 47.33 per cent in the subsequent ten periods.^{ix} This corresponds to the theoretical prediction since the expected utility in the tournament and the piece-rate scheme is the same.

Figure 3 displays the evolution of the frequency of the tournament choice over time. Contestants have been grouped into three categories: subjects who choose the tournament in at least 14 out of 20 periods (“tournament +”), subjects who choose the tournament in 6 periods or less (“tournament -“), and an intermediate category (“tournament =”). We find that the frequent competitors are relatively stable in their choices around a slightly increasing time trend. The least frequent users have chosen the tournament less often than other subjects since the very beginning of the game; moreover, after period 4, their frequency of tournament choice drops dramatically and remains at a very low level until the end of the game. Last, the intermediate category is the most unstable one, with a large oscillation of the frequency of the tournament choice from one set of periods to the next. We do not find any evidence of a selection strategy consisting of playing safer at the beginning of the game by choosing the piece-rate to secure a certain level of payoff, before switching to the riskier tournament scheme in the second part of the game.

If the tournament is selected in 50 per cent of the cases, does it mean that subjects choose at random their payment scheme or can we identify characteristics of subjects that predict their behavior?

A natural candidate for a determinant of sorting is risk aversion. Table 4 compares the distribution of our subjects in terms of risk attitude to the results in Holt and Laury (2002). We observe higher proportions of risk lovers and more than slightly risk averse subjects than in Holt and Laury's pool of subjects, but the differences are small. A Kolmogorov-Smirnov exact test does not reject the hypothesis of equality of distribution functions between our Benchmark and Choice treatments.

Figure 4 relates the frequency of our subjects' tournament choices to their proportion of safe choices in the ten decisions of the lottery task. Again, we consider the three categories of contestants as defined above (Tournament +, Tournament =, and Tournament -). The dashed line corresponds to the behavior of a risk neutral agent switching from option A to option B at decision 5.

Clearly, the subjects who choose the tournament less frequently are more risk averse than the other categories.^x All risk averse subjects considered together (who made at least 5 safe lottery choices) choose the tournament in 45.50 per cent of the periods, whereas the corresponding proportions are 60.38 per cent for the risk neutral subjects and 56.4 per cent for the risk lovers. A Poisson count model of the total number of tournaments chosen by the subject throughout the session has been estimated, including individual characteristics. It shows that only risk aversion exerts a significant influence and its marginal effect is important: crossing over from the safer to the riskier option one decision later in the lottery choices reduces by 77.80 per cent the number of tournament choices.

We have also conducted an econometric analysis of the choice of the tournament scheme, the results of which are reported in Table 5. Regression (1) estimates a random-effects Probit model, and regression (2) a fixed-effects Logit model .^{xi}

This analysis confirms that the degree of risk aversion is an important determinant of the choice of the competitive scheme. The tournament choice is also affected by previous outcomes. The regression shows that it declines over time and that bad luck in the previous period increases the probability to compete. This may reflect the subjects' attempts to get the winner's prize to compensate for small earnings in the previous period. Lastly, descriptive statistics indicate that if 72.86 per cent of those who won a tournament in the previous period choose to remain in the competitive scheme, this percentage decreases to 58.36 per cent among those who lost the previous competition.

Overall, these results suggest that risk aversion may be an important determinant of occupational choices. They are consistent with the survey analysis by Bonin et al. (2006) carried out on German data, which shows that risk averse employees tend to concentrate in jobs with low earnings risks. In contrast to Niederle and Vesterlund (2007) and Datta Gupta et al. (2005), we do not find evidence of a gender difference in competitiveness.

Heterogeneity of Behavior in Tournaments

We investigate the behavioral origins of the reduction of effort variability when individuals self-select into tournaments by adopting a cluster analysis that helps in identifying different types of behavior. In order to partition the sample, we retain three variables that summarize each individual's decisions: her frequency of tournament choices, her mean effort in the tournament and its standard deviation. In the Benchmark treatment, we only consider the last two variables. We apply the hierarchical Wald method based on the minimization of the intra-

group variance to identify the clusters that sum up the strategies. Clusters have been grouped so that the smallest one includes at least 10% of the subjects.

In both treatments, four main categories of tournament players are identified displaying similar characteristics; therefore we use the same denomination of clusters. The so-called “underconfident competitors” are subjects who exert an excessively high level of effort (more than 50% above the equilibrium), with relatively low standard deviation. The “motivated competitors” are subjects who exert a level of effort still higher than the equilibrium but closer to it. The “hesitant competitors” group consists of subjects who alternate levels of effort below and above the equilibrium and are characterized by the highest standard deviation of effort. Lastly, “economizing competitors” are subjects who follow a stable strategy based on the choice of a level of effort below the equilibrium.

Table 6 summarizes the statistics that characterize these behaviors in each treatment. The first column indicates the proportion of each cluster in the population. The second column represents the relative frequency with which the tournament has been played during the session. The following columns give the mean individual effort and within-individual standard deviation of effort in the tournament. The last column gives the between-subject standard deviation within each cluster.

In the Choice treatment, the analysis identifies two main categories of subjects according to their frequency of tournament choice. Frequent competitors, who compete in at least half of the periods, are characterized by a lower within-subject variance of effort than the occasional competitors, who choose the tournament in about one third of the periods.

When they can select their payment scheme, the individuals who enter more frequently into the tournament are both the motivated and the economizing competitors. The group of

motivated competitors is very homogenous as indicated by the low between-subject deviation. The relative importance of this group (40% of all the subjects involved in this treatment) contributes to explain the lower variance of effort in tournaments when individuals can self-select. In contrast, the group of economizing competitors shows the lowest within-subject and the highest between-subject variance of effort. It includes subjects who choose a minimum cost but can expect to win the tournament by chance. It also includes subjects who exert a level of effort slightly below equilibrium, possibly due to overconfidence or perception biases with respect to uncertainty, such as misconceptions of chance [Kahneman, Slovic, and Tversky (1982)] or illusion of control over external events originated in being given a choice as studied by Langer (1975).^{xiii} Whatever the explanation, their low-cost choices enable them to earn more on average than the motivated competitors (45.85 and 42.73, respectively).

The hesitant and the under-confident subjects are occasional competitors. The high within-subject variance of effort of the first group suggests that, facing the strategic uncertainty attached to the tournament, these subjects make errors both above and below the equilibrium. Entering the tournament less often reinforces the difficulties of learning the equilibrium. Consequently, they earn less on average in the tournament than the frequent competitors (40.80 points). The group of under-confident competitors is also not able to compute the equilibrium but always exerts a very high level of average effort in tournaments (73.20, i.e. 52.50 per cent above the equilibrium). As a consequence, even if they win relatively often the competition, the cost of effort is too high and thus they earn considerably less on average than under the piece-rate scheme (36.06 and 50.60 points, respectively).

The comparison of treatments indicates that the reduction of the effort variability in tournaments when agents can self-select is due to the fact that the most extreme categories in

terms of average effort and the most unstable agents tend to stay out of the competition. The experiment points to a potential limitation of sorting. The motivated competitors provide an over-supply of effort and their net earnings are not very high. These subjects do not enter into a rat race (that is, effort does not increase over time), but nevertheless, sorting reinforces a tendency to exert excess effort from some employees. A potential explanation of this observation is that in the Choice treatment subjects who have chosen to compete know that they are likely to face other subjects who are also eager to win.

V. DISCUSSION AND CONCLUSION

In a one-shot game environment, our results confirm that both the average level and the variance of effort are higher under a tournament than under a piece-rate payment scheme. This higher variability of effort has long been considered an important disadvantage since the employers have to bear uncertainty as to how the agents behave in relative performance compensation schemes. However, by analyzing an experimental setting that accounts for a key feature of markets, that the agents can choose their payment scheme repeatedly, our results paint a fundamentally different picture. A major finding of this paper is that when the subjects enter the tournament voluntarily, the average effort is higher and the variance of effort is substantially lower compared to situations in which the same payment scheme is imposed.

In our experiment, average effort in the freely chosen tournament is 32.47 per cent higher than in the exogenously imposed piece-rate scheme. This differential can be decomposed into an incentive and a sorting effect. The difference between effort levels in the imposed piece-rate and in the imposed tournament is an estimate of the incentive effect of tournaments: here, this is of the magnitude of a 14.63 per cent increase in effort. The difference of 17.84 per cent between the total increase in effort and the estimated incentive effect can be attributed to the

sorting effect of tournaments. The sorting effect makes up a little more than half of the total increase in effort; this is comparable to Lazear's (2000) corresponding estimates in connection with the switch from a fixed pay to a variable pay scheme. This confirms the importance of taking sorting into account when evaluating the efficiency of compensation schemes.

Another important and new result is that sorting significantly decreases the variance of effort in tournaments. When agents freely enter the tournament, the between-subject variance is four times (and even five times in the Single-Choice treatment) smaller than when this scheme is imposed and it is even lower than the variance of effort under an imposed piece-rate scheme. It is worth noting that we obtain this result in spite of the increased complexity of the task to be performed as compared to previous studies. Consequently, our experiment does not lead to the same recommendations as Bull et al. (1987), who suggest that to attract contestants, an employer should offer them a higher expected utility than under a piece-rate scheme. Our conclusion is rather that labor market flexibility, in particular the absence of restrictions on mobility between firms, is a key condition for a higher efficiency of relative performance pay.

Our results indicate that the efficiency-enhancing effect of sorting derives from the resulting greater homogeneity of contestants. In the Choice treatment, we have seen that since tournaments involve higher uncertainty than the piece-rate scheme, risk averse subjects choose them to a lesser extent. Under-confident subjects also prefer the piece-rate scheme since they exert too much effort in the tournament, entailing an excessive cost of effort. Hesitant subjects, alternating between above and below equilibrium effort levels, are not attracted by the tournament either. On the other hand, individuals who are motivated to work hard do not hesitate to choose the tournament in which equilibrium effort is higher. Among frequent contestants, the motivated competitors represent the biggest and the most stable category. Thus, the homogeneity of the contestants is higher when the tournament is chosen and this

contributes to the lower variance of effort. More homogeneity does not, however, give rise to collusion. Our interpretation is that these motivated subjects anticipate that they will face other subjects who like themselves are eager to win, too. Beyond this, our results suggest that introducing more competitive payment schemes in some occupations would sort employees and that the attitude toward risk may be an important driver of mobility between firms or sectors.

Having demonstrated that sorting has profound implications for the level and variance of effort in tournaments, we think further work should focus on how sorting is affected both by the prize structure and by differences in individuals' skills and social preferences: if people care about the negative externalities imposed on others by their individual effort, they may try not only to reduce their level of effort in relative pay schemes [Bandiera, Barankay, and Rasul (2005)], but also to stay out of a competition. Our work suggests more generally the importance of reconsidering the influence of sorting in many economic decisions.

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FOOTNOTES

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ⁱ Survey data analyses include Bognanno (2001), Ehrenberg and Bognanno (1990b), Eriksson (1999), Knoeber and Thurman (1994), Main, O'Reilly III, and Wade (1993), and experiments

include Bull, Schotter, and Weigelt (1987), Harbring and Irlenbusch (2003), Nalbantian and Schotter (1997), Orrison, Schotter, and Weigelt (2004), Schotter and Weigelt (1992).

ⁱⁱ In the Safelite study by Lazear (2000), half of the productivity gain associated with the introduction of variable pay is attributed to its ability to sort the most skilled employees. Experimental tests include Cadsby, Song, and Tapon (2007), and Eriksson and Villeval (2004), on sorting and incentives; Lazear, Malmendier, and Weber (2005) on sorting and social preferences; Bohnet and Kübler (2005) on sorting and cooperation.

ⁱⁱⁱ In the theoretical literature, Fullerton and McAfee (1999) propose an auction design in order to limit the entry into tournaments to selected highly qualified contestants. Hvide and Kristiansen (2003) show, however, that improving the quality of the contestants pool does not necessarily increase the selection efficiency of tournaments. In the empirical literature, Ehrenberg and Bognanno (1990a) show that higher winners' prizes attract better players and Knoeber and Thurman (1994) propose setting minimum standards to get rid of the poor performing competitors. Niederle and Vesterlund (2007) and Datta Gupta, Poulsen, and Villeval (2005) identify the importance of gender in the ex ante sorting effect of tournaments. Bonin, Dohmen, Falk, Huffman, and Sunde (2006) use survey data and show that risk averse individuals are more likely to be sorted into occupations with low earnings risk. In a field experiment, Bellemare and Shearer (2006) show that a firm that uses high-intensity incentive contracts and operate in a risky environment attracts more risk-tolerant individuals than other firms. Dohmen and Falk (2006) observe that risk averse subjects prefer fixed payments over piece-rate or tournament schemes and that tournaments attract subjects with different personalities, abilities, self-assessment and preferences than the other payment schemes.

^{iv} In this design, s is assumed to be similar for all agents, for the sake of simplicity.

^v We took care of using similar pools of subjects in both treatments. The average age of the subjects is 20.95 in the Benchmark treatment and 20.92 years in the Choice treatment; the average safety indexes are 5.30 and 5.35 respectively; Kolmogorov-Smirnov tests indicate that the distribution of age and safety index are not different between treatments. The proportion of men is 45 per cent in the Benchmark and 46.67 per cent in the Choice treatment. A Wilcoxon test accepts the equality between treatments.

^{vi} Regarding the first possible effect, Bull, Schotter and Weigelt (1987) reject the errors in inference explanation of the variance: in fixed pairs, giving the subjects a feedback on the effort chosen by their opponent does increase the variance of effort in comparison with situations in which the subjects are only informed about their rank or on the total outcome. Regarding the second and opposite effect, Harbring and Irlenbusch (2003) observe opposite strategies between pairs in a game where both the minimum and the maximum effort levels are two equilibria in pure strategies. There is, however, no reason to observe this multiple convention outcome when the game has a unique interior equilibrium as in our game. In addition, Bull et al. (1987) mention that the high variance they observe with an interior equilibrium is not due to outlier pairs and that there is no tendency for variance to decline as the experiment progresses, which should be the case if the formation of conventions was at stake.

^{vii} If we remove the observations with a level of effort of 0 or 100 (73 and 41, respectively), the between-subject variance in the tournament still represents 64.43 per cent (270.60/419.96) of the total variance in the Benchmark and 42.04 per cent (76.12/181.05) in the Choice treatment. Thus the structure of the variance remains the same as when all the contestants are considered.

^{viii} Since an uneven number of participants chose the tournament (11), one of them has been randomly drawn and informed that he would be paid according to the piece-rate scheme.

^{ix} The proportion of subjects choosing the tournament was 55 per cent in the Single Choice treatment.

^x We checked that although made at the end of the sessions, the lottery decisions do not result from the behavior in the main game instead of explaining it. In fact, there is no correlation between the number of safe choices in the lotteries and the individual's net payoff.

^{xi} We also tested a two-step selection model with random effects, including the treatment variable in the non-selection equation and the inverse of the Mill's ratio, fitted by simulated maximum likelihood. The selection equation included risk aversion, gender, the lagged random shock and a time trend. The analysis concludes to the absence of any selection bias, so that we can proceed with separate equations.

^{xii} In the first period of the game, after the subjects have chosen their level of effort, we asked them: « *How big do you estimate your chances are that you will draw a random number that increases your payoff?* ». 14.17 percent reported a probability lower than 0.49 and 13.33 percent a probability exceeding 0.50. 61.11 percent of the optimistic subjects opted for the tournament, whereas the corresponding percentages are 47.31 for the pessimistic and 48.29 for the well-calibrated subjects. According to a Probit regression (not shown) including only the first period data and individual observable characteristics, optimism marginally but significantly (at the 10 per cent level) increases tournament entry. If all periods are considered, miscalibration is no longer significant since subjective beliefs are revised throughout the game.

TABLE 1
Summary statistics on average level and variance of effort

Periods	Average effort				Mean variance of effort			
	All	1	1-10	11-20	All	1	1-10	11-20
<i>Piece-rate</i>								
Benchmark t.	46.48	55.73	48.92	44.04	368.88	388.06	381.29	345.73
Choice t.	50.45	47.63	51.35	49.63	227.87	192.37	228.34	226.70
<i>Tournament</i>								
Benchmark t.	53.28	60.03	55.62	50.94	652.26	663.76	672.51	623.19
Choice t.	61.57	65.75	63.37	59.55	258.19	319.38	239.59	272.30

TABLE 2
Decomposition of the variance of effort

Variance	Between-subject	Within-subject	Total
<i>Benchmark treatment</i>			
Piece-rate	193.69 (52.51)	175.19 (47.49)	368.88 (100)
Tournament	434.55 (66.62)	217.71 (33.38)	652.26 (100)
<i>Choice treatment</i>			
Piece-rate	120.01 (52.67)	107.86 (47.33)	227.87 (100)
Tournament	101.79 (39.42)	156.41 (60.58)	258.19 (100)

Note: Percentages of the total variance in parentheses.

TABLE 3

Determinants of the effort decision

Treatments	Random-effects Tobit regressions	
	Benchmark (1)	Choice (2)
Time trend	-0.495*** (0.079)	-0.305*** (0.071)
Payment scheme (Tournament=1)	3.377** (1.537)	11.554*** (0.859)
Lagged random shock	-0.048*** (0.018)	-0.035** (0.017)
Risk aversion	-2.908*** (0.452)	-0.361 (0.474)
Gender (male=1)	-7.518*** (1.411)	-4.377*** (1.346)
Constant	71.877*** (3.083)	56.718*** (3.027)
Nb obs.	1140	1140
Left censored obs.	63	8
Right censored obs.	20	15
Log likelihood	-4478.911	-4576.915
Wald χ^2	124.97	227.34
Prob> χ^2	0.000	0.000

Note: Standard errors in parentheses; *** significant at the 1% level, ** at the 5% level, * at the 10% level.

TABLE 4

Distribution of risk attitudes

Number of safe choices	Holt and Laury's classification	Holt and Laury's experiment	Our experiment	
			Benchmark T.	Choice T.
0-1	Highly Risk Lover	0.01	0.05	0.00
2	Very Risk Lover	0.01	0.00	0.02
3	Risk Lover	0.06	0.05	0.10
4	Risk Neutral	0.26	0.18	0.22
5	Slightly Risk Averse	0.26	0.18	0.15
6	Risk Averse	0.23	0.32	0.30
7	Very Risk Averse	0.13	0.17	0.17
8	Highly Risk Averse	0.03	0.03	0.03
9-10	Stay in Bed	0.01	0.02	0.02

Note: The number of safe choices corresponds to the number of the decisions with the “safe” option A, and thus corresponds to the “risk aversion” variable in our econometric analysis.

TABLE 5
Determinants of the tournament choice

	Random-effects Probit regression (1)	Fixed-effects Logit regression (2)
Time trend	-0.017** (0.007)	-0.029** (0.012)
Lagged random shock	-0.003* (0.002)	-0.005* (0.003)
Risk aversion	-0.162** (0.066)	
Gender	-0.046 (0.202)	
Constant	1.047*** (0.386)	
Nb observations	1140	1102
Wald χ^2 / LR χ^2	13.92	8.60
Prob> χ^2	0.007	0.014
Log Likelihood	-691.202	-517.065

Note: standard errors in parentheses; *** significant at 1% level, ** at 5% level, * at 10% level. In the conditional fixed-effects logistic regression, the number of observations is lower than in the random-effects model because individuals have been dropped due to all positive or all negative outcomes.

TABLE 6
Behavior in tournaments

	Share in the population	Relative frequency	Mean effort	Within-subjects SD	Between-subjects SD
<i>Benchmark treatment</i>					
Under-confident Competitors	30.00	100	74.48	6.35	4.20
Motivated Competitors	30.00	100	59.93	10.78	5.37
Hesitant Competitors	30.00	100	40.66	20.48	7.09
Economizing Competitors	10.00	100	7.60	9.96	7.14
<i>Choice treatment</i>					
<i>Frequent competitors</i>					
Motivated Competitors	40.00	57.90	61.87	9.45	3.73
Economizing Competitors	18.33	50.90	44.56	7.60	15.55
<i>Occasional competitors</i>					
Hesitant Competitors	10.00	35.80	53.06	32.61	6.48
Under-confident Competitors	31.67	34.45	73.20	10.74	6.34

FIGURE 1

Evolution of the effort decisions by treatment and by mode of payment over time

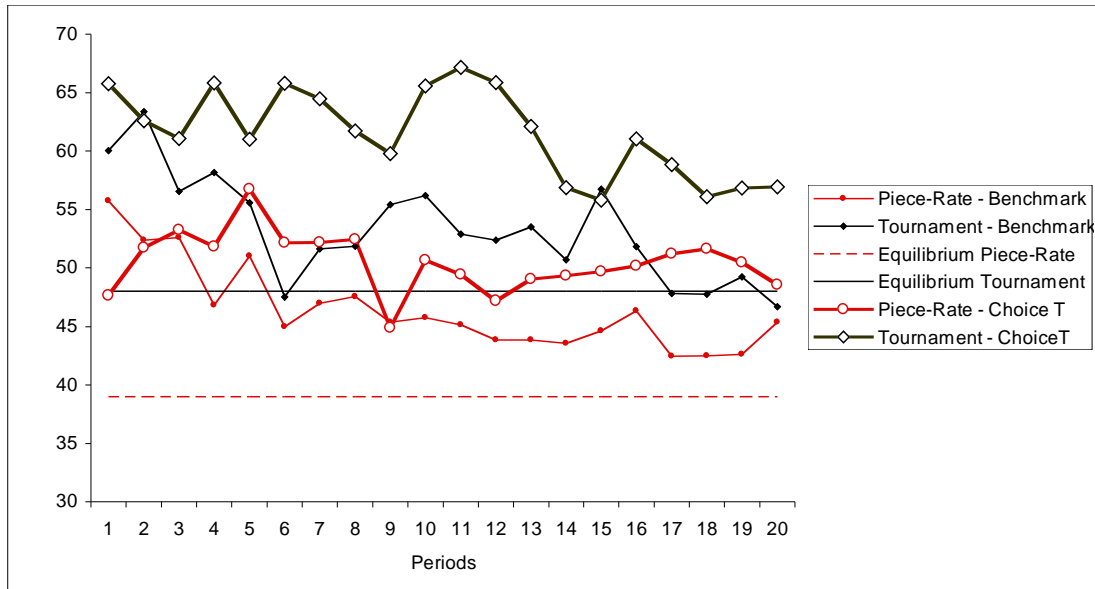


FIGURE 2

Dispersion of effort in the tournaments by treatment and category of periods

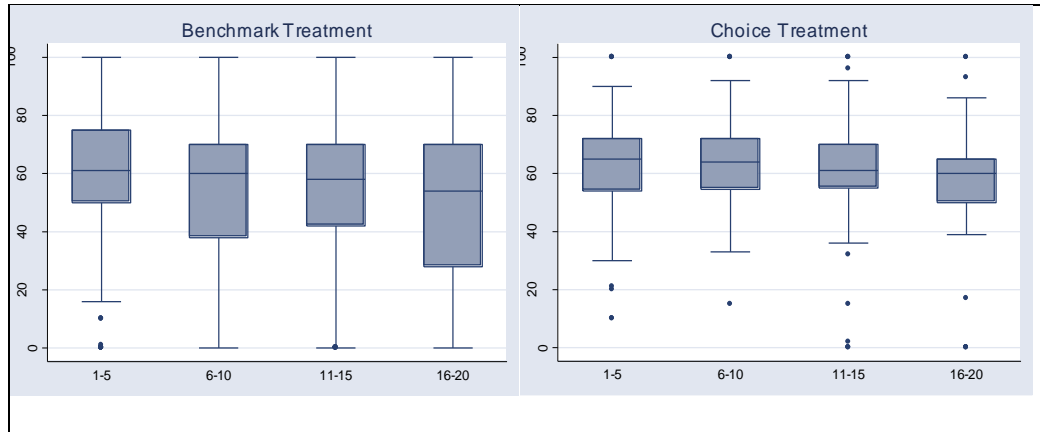


FIGURE 3

Evolution of the frequency of the tournament choice over time by category of subjects

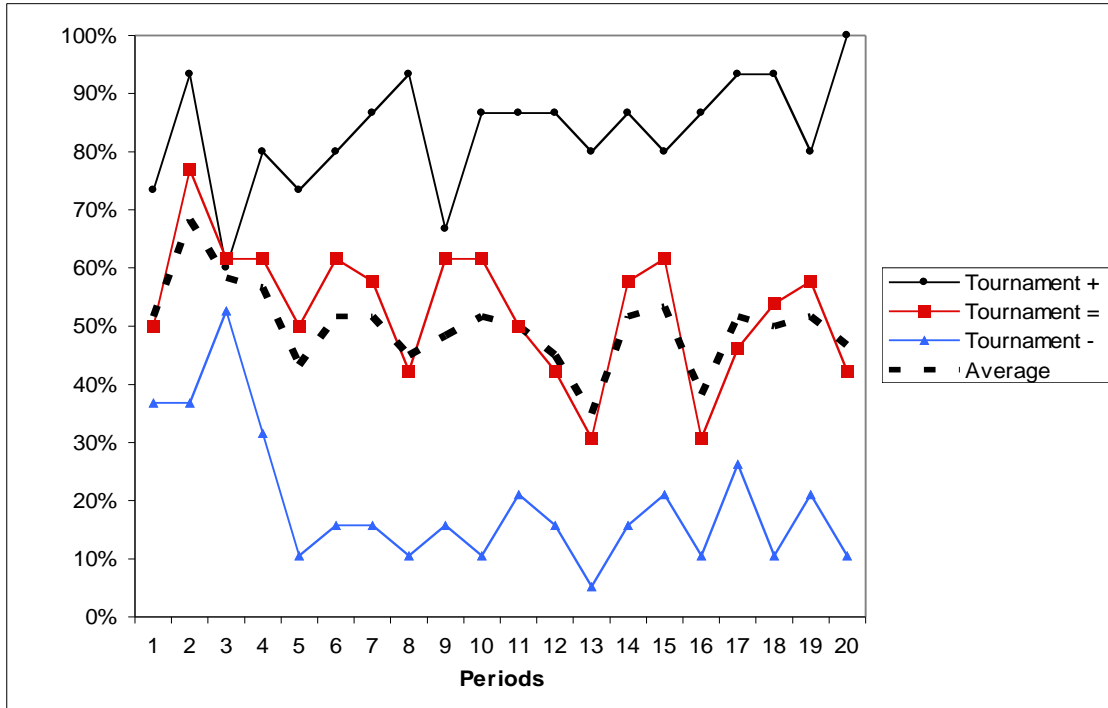
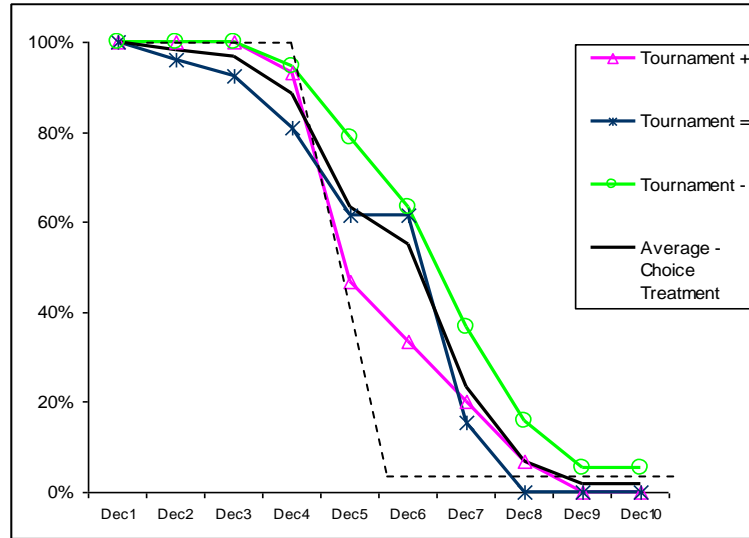


FIGURE 4

Proportion of safe lottery choices and frequency of the tournament choices



Appendix. Instructions of the Choice Treatment

You are about to participate in an experiment on decision-making organized for the GATE research institute and the Aarhus School of Business in Denmark. During this session, you can earn money. The amount of your earnings depends on your decisions and on the decisions of the participants you will have interacted with. During the session, your earnings will be calculated in points,

with 80 points = 1Euro

During the session, losses are possible. However, they can be avoided with certainty by your decisions. In addition, if a loss would occur in a period, the gains realized during the other periods should compensate this loss.

At the end of the session, all the profits you have made in each period will be added up and converted into Euros. In addition, you will receive a show-up fee of 3 Euros. You will have also an opportunity to earn additional money by participating in a decision task at the end of the session. Your earnings will be paid to you in cash in a separate room in order to preserve confidentiality.

The session consists of 20 independent periods.

Description of each period

Each period consists of two stages.

- In stage 1**, you choose between two modes of payment, mode X and mode Y.
- In stage 2**, you carry out a task.

Your profit during each period depends on the mode of payment you have chosen and on your result from the task.

Description of the task

- A table is attached to these instructions: numbers, from 0 to 100, are given in column A. In the second stage of each period, your task consists of selecting one of these numbers. This number will be called your “decision number”. Associated with each number is a cost, called “decision cost”. These decision costs are listed in column B. Note that the higher the decision number chosen, the greater is the associated cost. You make your choice by means of a scrollbar on your computer screen and you confirm this choice by clicking the “OK” button.
- Then, you have to click a button on your screen that will generate a random number. This number is called your “personal random draw number”. This number can take any value between -40 and $+40$. Each number between -40 and $+40$ is as likely to be drawn and there is one independent random draw between -40 and $+40$ for each subject in the lab.

Your “result” for the task is the sum of your decision number and your personal random draw number.

Your result = your decision number + your personal random draw number

Choice of the mode of payment and calculation of your payoff

There are two different modes of payment, mode X and mode Y. In the first stage of each period, you choose to be paid according to mode X or to mode Y. If you like, you can change the mode of payment at each new period.

- Description of mode of payment X

If you choose the mode of payment X, your result is multiplied by 0.52. You also receive a fixed amount of 45 points. Next, the decision cost associated to the choice of your decision number is subtracted. Note, the amount subtracted (your decision cost) is only a function of your decision number; that is, your personal random draw number does not affect the amount subtracted.

Your payoff thus depends on your decision number and your personal random draw number. Your net payoff under mode X is thus given by the following formula:

Your net payoff of the period under mode X = $45 + (\text{your result} * 0.52) - \text{your decision cost}$
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At the end of the period, you are informed about your result and about your net payoff for the current period.

Example of net payoff calculation under mode of payment X

For example, say that you choose a decision number of 55 and you draw a personal random number of 10. Your net payoff calculation will look like:

$$45 + [(55 + 10) * 0.52] - 20.17 = 58.63$$

□ Description of mode of payment Y

If you choose the mode of payment Y, another subject in the room, who has also chosen the mode of payment Y, is paired with you at random for the current period. This subject is called your “pair member”. The identity of your pair member will never be revealed to you.

Your pair member has an identical sheet as yours. Like you and simultaneously, he has to select a decision number and he will draw his personal random number. As for you, the “result” of your pair member is computed by adding his decision number and his personal random draw number.

Then, the computer program will compare your result and the result of your pair member.

- If your result is greater than your pair member’s result, you receive the fixed payment M, equal to 96 points.
- If your result is lower than your pair member’s result, you receive the fixed payment L, equal to 45 points.
- In case of equal results, a fair random move decides on which subject receives M and who receives L.

Whether you receive M or L as your fixed payment depends only on whether your result is greater or not than your pair member’s. It does not depend on how much bigger it is.

To determine your net payoff, the decision cost associated with the choice of your decision number is subtracted. Note, the amount subtracted is only a function of your decision number; that is, your personal random draw number does not affect the amount subtracted.

Therefore, your net payoff depends on your decision number, your personal random draw number, and your pair member’s decision number and his personal random draw number.

Your net payoff under mode Y is given by the following formula:

Your net payoff of the period under mode of payment Y = Fixed payment (M or L) – your decision cost
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At the end of the period, you are informed about your result; you are told by how much your total is greater or less than that of your pair member and you are informed about your net payoff for the current period.

Example of net payoff calculation under mode of payment Y

For example, say that pair member A chooses a decision number of 25 and draws a personal random number of 20, while pair member B selects a decision number of 55 and draws a personal random number of -5.

A’s result is: $25 + 20 = 45$

B’s result is: $55 - 5 = 50$

B’s result is larger than A’s result. Thus, B receives M (=96) and A receives L (=45).

A’s net payoff is: $45 - 4.17 = 40.83$

B's net payoff is: $96 - 20.17 = 75.83$

To sum up, in each period you make two decisions:

- In stage 1, you choose between mode of payment X and mode of payment Y. Note that if an uneven number of participants has chosen mode Y, one of these participants will be randomly chosen and paid according to mode X. To be paid according to mode Y, pairs must be formed. This participant will be informed of this before moving to stage 2.
- In stage 2, you select your decision number and you draw a personal random number. Your net payoffs for the current period are then computed.

At the end of a period, a new period starts automatically. Each period is independent. The random draws are independent from one period to the next. In each period, under mode of payment Y, pairs are composed at random among the participants who have chosen this mode of payment.

If you have any question regarding these instructions, please raise your hand. Your questions will be answered in private. Throughout the entire session, talking is not allowed. Any violation of this rule will result in being excluded from the session and not receiving payment.

Thank you for your participation.

Decision Costs Table

Column A Decision Nb	Column B Cost of Decision	Column A Decision Nb	Column B Cost of Decision	Column A Decision Nb	Column B Cost of Decision
0	0.00	35	8.17	70	32.67
1	0.01	36	8.64	71	33.61
2	0.03	37	9.13	72	34.56
3	0.06	38	9.63	73	35.53
4	0.11	39	10.14	74	36.51
5	0.17	40	10.67	75	37.50
6	0.24	41	11.21	76	38.51
7	0.33	42	11.76	77	39.53
8	0.43	43	12.33	78	40.56
9	0.54	44	12.91	79	41.61
10	0.67	45	13.50	80	42.67
11	0.81	46	14.11	81	43.74
12	0.96	47	14.73	82	44.83
13	1.13	48	15.36	83	45.93
14	1.31	49	16.01	84	47.04
15	1.50	50	16.67	85	48.17
16	1.71	51	17.34	86	49.31
17	1.93	52	18.03	87	50.46
18	2.16	53	18.73	88	51.63
19	2.41	54	19.44	89	52.81
20	2.67	55	20.17	90	54.00
21	2.94	56	20.91	91	55.21
22	3.23	57	21.66	92	56.43
23	3.53	58	22.43	93	57.66
24	3.84	59	23.21	94	58.91
25	4.17	60	24.00	95	60.17
26	4.51	61	24.81	96	61.44
27	4.86	62	25.63	97	62.73
28	5.23	63	26.46	98	64.03
29	5.61	64	27.31	99	65.34
30	6.00	65	28.17	100	66.67
31	6.41	66	29.04		
32	6.83	67	29.93		
33	7.26	68	30.83		
34	7.71	69	31.74		

Post experimental questionnaire

[Instructions for the test of risk aversion directly taken from Holt and Laury, 2002]

We thank you for filling out this form that enables you to earn additional money. The attached sheet of paper shows ten decisions. Each decision is a paired choice between “Option A” and “Option B”. You will make ten choices and record these in the column on the right, but only one of them will be used in the end to determine your additional earnings. Let us explain how these choices will affect your earnings.

Here is a ten-sided die that will be used to determine this payoff. The faces are numbered from 1 to 10 (the “0” face of the die will serve as 10). After you have made all of your choices, and when you come to the other office to receive your payment, you will throw this die twice:

- once to select one of the ten decisions to be used,
- and a second time to determine what your payoff is for the option you chose, A or B, for the particular decision selected.

Even though we ask you to make ten decisions, only one of these will end up affecting your earnings. However, you will not know in advance which decision will be used. Obviously, each decision has an equal chance of being used in the end.

- Look at Decision 1.

Option A pays 2 € if the throw of the dice is 1, and it pays 1.6 € if the throw is 2-10.

Option B yields 3.85 € if the throw of the dice is 1 and it pays 0.1 € if the throw is 2-10.

- Look at Decision 2.

Option A pays 2 € if the throw of the dice is 1 or 2, and it pays 1.6 € if the throw is 3-10.

Option B yields 3.85 € if the throw of the dice is 1 or 2 and it pays 0.1 € if the throw is 3-10.

The other decisions are similar, except that as you move down the table, the chances of a higher payoff for each option increase. In fact, for Decision 10 in the bottom row, the dice will not be needed since each option pays the highest payoff for sure, so your choice here is between 2 € and 3.85 €.

To summarize,

- you will make ten choices. For each decision row, you will have to choose between Option A and Option B. You may choose A for some decision rows and B for other rows. You may change your decisions and make them in any order.
- When you come to the other room to receive your earnings from the experiment, you will throw the ten-sided die to select which of the ten decisions will be used.
- Then, you will throw the die again to determine your money earnings for the Option you chose for that Decision.

Earnings (in Euros) for this choice will be added to your previous earnings, and you will be paid all earnings in cash.

If you have any question, please raise your hand. Your questions will be answered in private. Please do not talk with anyone.

Please indicate for each of the following 10 decisions if you choose Option A or Option B.

	Your decision
Decision 1 Option A: 1/10 of 2 € and 9/10 of 1.6 € Option B: 1/10 of 3.85 € and 9/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 2 Option A: 2/10 of 2 € and 8/10 of 1.6 € Option B: 2/10 of 3.85 € and 8/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 3 Option A: 3/10 of 2 € and 7/10 of 1.6 € Option B: 3/10 of 3.85 € and 7/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 4 Option A: 4/10 of 2 € and 6/10 of 1.6 € Option B: 4/10 of 3.85 € and 6/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 5 Option A: 5/10 of 2 € and 5/10 of 1.6 € Option B: 5/10 of 3.85 € and 5/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 6 Option A: 6/10 of 2 € and 4/10 of 1.6 € Option B: 6/10 of 3.85 € and 4/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 7 Option A: 7/10 of 2 € and 3/10 of 1.6 € Option B: 7/10 of 3.85 € and 3/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 8 Option A: 8/10 of 2 € and 2/10 of 1.6 € Option B: 8/10 of 3.85 € and 2/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 9 Option A: 9/10 of 2 € and 1/10 of 1.6 € Option B: 9/10 of 3.85 € and 1/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>
Decision 10 Option A: 10/10 of 2 € and 0/10 of 1.6 € Option B: 10/10 of 3.85 € and 0/10 of 0.1 €	Option A <input type="radio"/> Option B <input type="radio"/>