



**HAL**  
open science

# The Social Effects of Ethnic Diversity at the Local Level: A Natural Experiment with Exogenous Residential Allocation

Yann Algan, Camille Hémet, David Laitin

► **To cite this version:**

Yann Algan, Camille Hémet, David Laitin. The Social Effects of Ethnic Diversity at the Local Level: A Natural Experiment with Exogenous Residential Allocation. 2013. halshs-00843173

**HAL Id: halshs-00843173**

**<https://shs.hal.science/halshs-00843173v1>**

Preprint submitted on 10 Jul 2013

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

# The Social Effects of Ethnic Diversity at the Local Level: A Natural Experiment with Exogenous Residential Allocation

Yann Algan  
Camille Hémet  
David Laitin

WP 2013 - Nr 38

# The Social Effects of Ethnic Diversity at the Local Level: A Natural Experiment with Exogenous Residential Allocation

YANN ALGAN\*, CAMILLE HÉMET†, DAVID LAITIN‡

July 10, 2013

## Abstract

This paper demonstrates the effects of ethnic diversity on social relationships and the quality of public spaces at a very fine neighborhood level. We use detailed block level data on diversity and housing quality from a representative survey on housing in France. We show how and to what extent diversity within a neighborhood can directly affect household well-being and the quality of the common spaces, whereas the previous literature looks at more aggregate outcomes through voting channels. Our identification strategy relies on the exogeneity of public housing allocations with respect to ethnic characteristics in France, to address the bias due to endogenous residential sorting. Diversity is shown to have a negative effect on the quality of local public goods, either due to vandalism, not deterred by other-regarding preferences and social policing, or due to collective action failure to ensure effective property management. However, we find that diversity has no robust effect on public safety at a local level and, if anything, is more related to social anomie.

JEL Classification: H10, H41

Keywords: diversity, neighborhood effects, living conditions, public housing

---

\*Department of Economics, Sciences Po

†Aix-Marseille University (Aix-Marseille School of Economics), CNRS & EHESS and Sciences Po

‡Department of Political Science, Stanford University

The research leading to these results has received funding from the European Research Council for Yann Algan under the European Community's Seventh Framework Program (FP7/2007-2013) / ERC grant agreement n°240923. The authors acknowledge financial support from the France-Stanford Center for enabling this cross Atlantic collaboration. We are grateful to Alberto Alesina, Ernesto Dal Bo, Larry Katz, Asim Ijaz Khwaja, Edward Miguel, Leah Platt Boustan, Thierry Mayer, Daniel Posner, Andrei Shleifer and Mathias Thoenig for valuable comments and discussions. We also thank participants at the NBER Political economy seminar, at Stanford University, at Sciences Po, at the French Research Center in Economics and Statistics, at the French Association for Developing Research in Economics and Statistics, at the University of California Los Angeles, at Institute of Economics of Barcelona, at the European Economic Association, at Aix-Marseille School of Economics and at the National Bureau of Economic Research for useful comments. We are grateful to Corinne Prost, head of the Employment Division at INSEE (French National Statistical Institute) for allowing access to key data through a convention between INSEE and Sciences Po.

# 1 Introduction

Recent research has drawn an ominous picture of the implications of cultural heterogeneity on social peace and economic growth. A large literature shows a negative relationship, though not always robust, between ethnic diversity and the quality of public goods (Alesina et al., 1999; Alesina and La Ferrara, 2000; Miguel, 2004; Miguel and Gugerty, 2005), welfare spending (Luttmer, 2001), civil conflict and trust (Fearon and Laitin, 2000; Putnam, 2007; Alesina and La Ferrara, 2002) and economic growth (Alesina, Baqir and Easterly, 1999). The leading explanations of why ethnic fragmentation affects those outcomes are the heterogeneity of preferences and the free-rider problem which undermines collective action. The literature thus views the problem of fractionalization in terms of voting behavior on aggregate outcomes such as public good provision at the country or county level. Yet, little is known on how diversity could affect directly social relationships and well-being at the neighborhood level. Our paper fills this gap by looking at the effect of ethnic diversity on social relations and the quality of common spaces within local communities at the housing block level. Besides we provide a new identification strategy to overcome the endogeneity problem raised by residential self-selection. We rely on a natural experiment of exogenous spatial allocation in the French public housing sector to identify the causal effect of diversity on those outcomes.

The main contribution of our paper is to identify the effect of ethnic diversity on social relationships and the quality of public goods at a very local block level. We use micro data on housing conditions where the units of observation are public housing blocks made up of twenty adjacent households on average. This is a key improvement for the analysis of how diversity shapes social relationships compared to the previous literature which is based on aggregated data at the county, regional or country levels. Diversity might matter for various reasons at different levels and the channels through which diversity operates are likely to depend on the size of the unit of observation. By focusing on the provision of public goods at an aggregate level, the previous literature is mainly interested in the effect of diversity on collective action through lobbying or patronage (see Alesina and La Ferrara (2005) for a survey). Instead, we analyze in this paper how diversity within a small community affects individual well-being and satisfaction with housing conditions through relationships among neighbors. Indeed, the survey we use (French Housing Survey) reports specific information about the neglect and voluntary degradations of the public areas, the quality of the housing, and direct interpersonal conflicts. These data enable us to identify various effects of diversity on local social relationships and public good outcomes, and to explore the possible channels explaining this link.

When residents of more diverse blocks report that neglect and voluntary degradation is rife in their housing unit, we interpret this as a result of the residents' failure to develop social norms and other regarding preferences. When they report the breakdown and the poor quality of basic facilities (such as heating and soundproofing), we interpret this as a result of a diminished capacity for collective action for social improvement. Those goods are of course not directly degraded by diversity. But diversity might be associated with lower ability for collective action, explaining the irregularity of maintenance and the absence of repairs in more diverse blocks. In this case, the result

could be supported in equilibrium if the housing directorate reckons that it can neglect facilities in ethnically heterogeneous housing projects, knowing that it will not face collective action from its residents demanding better services. Finally, when residents report incidents of direct interpersonal conflicts, we can interpret this as an effect of diversity on cultural enmity. We test these channels by using indirect objective measures of the quality of social relationships and common spaces, such as the number of repairs and the upkeep of the security equipment. We find that diversity decreases the quality of local common spaces, but has no effect on public safety. Instead, individuals are more likely to report the absence of any social relationship with their neighbors rather than interpersonal conflicts. Thus diversity leads to social anomie, preventing the emergence of social norms and collective action, rather than antipathy at the local neighborhood level.

In order to make unbiased causal inferences, we provide a new strategy for identifying the causal effect of diversity on economic and social outcomes. The general concern in this literature is that the endogenous residential sorting of individuals on ethnic grounds biases the estimate of the impact of diversity. We address this issue by using a natural experiment in which households in France are allocated to public housing blocks without taking their ethnic origin or their preference for diversity into account. Due to a strongly republican ideology, the French public housing system allocates state planned moderate cost rental apartments (HLMs - *Habitations à Loyer Modéré*) to natives and immigrants without concern for their cultural and ethnic background, mixing people indiscriminately. Some HLM neighborhoods are consequently quite diverse, and others quite homogeneous. Furthermore, HLM inhabitants rarely move, as the rents are much lower than market rates, and moving between HLM blocks is quite difficult. Consequently, residents cannot choose whether to live near people like themselves. Rather, they accept their placement, whether next to co-ethnics or strangers. Methodologically, this means that we can take the degree of diversity in any one HLM block as exogenous, connect the level of diversity with the housing situation, and examine whether greater heterogeneity leads to poorer provision of public goods or more troubled social relationships in French communities. We extensively document the actual process of allocation of households within the public housing sector. We show that legal rules prohibit housing allocation based on ethnic backgrounds and that in practice, the characteristics of the public housing sector make it very complicated to bypass the law. Then, we conduct a variety of formal statistical tests to verify the absence of self-sorting on ethnic characteristics. In particular, we show that the observed spatial distribution of residents across public housing blocks is not statistically different from a random distribution generated by Monte Carlo simulations. We perform a variety of alternative tests. We show that while households moving into a new neighborhood tend to self-segregate in the unconstrained private housing market, there is no such evidence in the public housing market. We also examine potential self-selection prior to the move and show that households that have refused an offer end up living in public housing blocks that display the same ethnic diversity as those who accepted their first offer. Thus even if households try to be choosy with respect to the ethnic composition of their neighborhoods, they cannot self-segregate in the public housing sector due to the allocation process and the tight supply constraints of dwellings.

Naturally, this paper is not the first one to try to overcome this identification issue. But pre-

vious attempts to establish causality rely mainly on instrumental variables.<sup>1</sup> However convincing the instruments might be, this strategy cannot overcome the concern as to whether the instruments fulfill the exclusion restriction and do not have a direct effect on public goods. For instance, Miguel (2004) and Miguel and Gugerty (2005) use the pre-colonial patterns of settlement as instruments, assuming that these variables have no direct impact on present-day ethnic relations. More recently Glennerster et al. (2010) have also relied on historical data of fractionalization as an instrument. But since past settlement patterns are likely to have at least some direct impact on present-day ethnic relations, the exclusion restriction might still be technically violated. Using a natural experiment with exogenous allocation of ethnic groups is thus an alternative strategy to deal with these traditional caveats. The paper which is the closest to ours is Dahlberg et al. (2011), which uses a nation-wide policy intervention program that exogenously placed refugees coming to Sweden across the Swedish municipalities. However, their paper examines in-group bias in preferences for redistribution rather than the effect of diversity on local public goods and social relationships.

Our paper is related to the large literature on the effects of ethnic diversity on economic and social outcomes. In US cities, higher ethnic diversity has been found to be associated with lower social capital (Putnam, 2007; Alesina and La Ferrara, 2000 and 2002), lower welfare spending (Luttmer, 2001), and poorer quality of public goods (Alesina et al., 1999). In Western Kenya, the greater the mixing of tribes, the less people have public spiritedness, and the lower the contributions to public goods (Miguel, 2004; Miguel and Gugerty, 2005). In cross-national surveys, diversity correlates with low growth in GDP and low quality of institutions (Easterly and Levine, 1997; Alesina et al., 2003). Alesina and Zhuravskaya (2011) show that islands of homogeneity amid a broadly diverse country do not decrease the negative effects of diversity on the quality of government. Theoretical contribution, in particular on ethnic conflicts, can be found in Caselli and Coleman (2012) and Esteban and Ray (2011).<sup>2</sup> These findings are depressing, in a normative sense, for those who herald gains from diversity (Page, 2007); and depressing, in an empirical sense, as in our globalized world, local cultural diversity is increasingly common (Dancygier, 2010). However, the robustness of the relationship and the channels at work remain to be determined. Putnam (2007) is careful to underline that his data allow him only to claim short run correlation between diversity and trust. Miguel (2004) finds no diversity impacts on local outcomes in Tanzania, a country in which the ruling authorities have sought to ameliorate ethnic cleavages by promoting a common language. Posner (2004) shows that changed electoral rules can create broader ethnic identities thereby reducing frag-

---

<sup>1</sup>In their seminal contribution to the literature, Alesina et al. (1999) provide a first attempt to deal with this endogeneity issue by collecting data at different levels of aggregation (cities, metropolitan areas and counties). Their assumption is that different levels of aggregation allow for the correction of the potential biases introduced by Tiebout sorting.

<sup>2</sup>The magnitude of the relationship between those outcomes and ethnic diversity is substantial. Putnam (2007) finds that the difference between living in a highly homogeneous city (Bismarck, North Dakota) and the heterogeneous Los Angeles is as great as the difference between an area with a poverty rate of 7 percent and one with a poverty rate of 23 percent. Alesina et al. (1999) show that moving from complete homogeneity to complete heterogeneity is associated with a reduction in spending on roads by nine percentage points. Luttmer (2001) finds that interpersonal preferences based on negative exposure and racial group loyalty of recipients are associated with 33 percent of the cross-state variation in the support for welfare spending. Alesina et al. (2003) show that moving from perfect homogeneity to maximum heterogeneity would be associated with a reduction in a country's growth rate by two percentage points per year.

mentation. Dunning and Harrison (2010) show that inter-tribal polarization in Mali is reduced with cross-cutting cleavages. Glennerster, Miguel and Rothenberg (2010) also argue that the presence of strong chiefs at the local level, although reinforcing the salience of ethnicity, translates into effective inter-ethnic cooperation. Finally, Alesina and La Ferrara (2005) document the potential positive effect of diversity on productivity through complementarity in skills.

Our paper is also incidentally related to empirical works examining neighborhood effects on social and economic outcomes. So far, the literature has mainly focused on the neighborhood effects on physical and mental health, economic self-sufficiency, risky and criminal behavior, or educational outcomes (see among many others Katz et al., 2001; Oreopoulos, 2003; Goux and Maurin, 2007 and Kling et al., 2007). In particular, Katz et al. (2001) and subsequent contributions use the Moving to Opportunity social experiment to estimate the externalities from neighbors. To avoid the problem of endogenous neighborhood selection, those authors use data from a randomized experiment in which some families living in high-poverty U.S. housing projects were offered housing vouchers to enable them to move to higher income areas. While our paper is not based on a randomized experiment, we also avoid the inferential issues of residential endogenous selection by using the exogenous spatial allocation of households with respect to ethnic characteristics. We enlarge the dimensions analyzed in this literature by looking at how immediate neighborhood diversity affects well-being and the quality of the local environment.

The remainder of this paper is organized as follows. Section 2 presents the data. Section 3 documents our identifying assumption on the absence of residential self-sorting in public housing. We conduct various tests to show that the spatial allocation of households across public housing blocks within localities is exogenous with respect to ethnic characteristics. Section 4 shows our main results. We document the effects of ethnic diversity on satisfaction with housing condition, local public goods quality and social relationships. We discuss the various dimensions and channels through which diversity might matter for households' well-being at the finite local level. Section 5 provides tests of the validity of self-reported outcomes, and we perform a series of robustness checks on our results in section 6. Section 7 documents additional tests on the exogeneity of the residential allocation in the public housing sector. Section 8 concludes.

## **2 Presentation of the data**

### **2.1 Data sets**

The analysis is based on two representative French national surveys. We use the French Labor Force Survey (Enquête Emploi, INSEE, hereafter the LFS) to test our identification assumption that spatial allocation in the public housing market can be considered as quasi-random and exogenous relative to ethnic characteristics. The survey covers the period 2003-2007 and provides all the relevant information about ethnic background, economic characteristics and geographic location of individuals to test our identification strategy. The sampling unit of this database is at the

housing block level, which consists in twenty adjacent households on average.<sup>3</sup> Over the 2003 to 2007 period, more than 10,000 different housing blocks were sampled. All the households within a randomly selected housing block were surveyed and, within each household, all persons aged fifteen or over were interviewed. Using these data, we can work on real neighborhoods at a very small geographic level. Moreover, we have information on whether the respondent was living in a public housing unit, whether he or she has been living in his or her current public housing for at least one year or whether he or she has just moved into the neighborhood. These particular features enable us to compute the level of ethnic diversity prevailing within each public housing block and to test for the absence of self-sorting on ethnic background across public housing blocks.

We use the French Housing Survey 2002 (*Enquête Logement*, INSEE, hereafter the HS), to estimate the relationship between ethnic diversity and the quality of public space within the housing block. We identify the causal effect of diversity and control for self-sorting by focusing on the public housing sector. The HS provides detailed information on the intensity and quality of social interactions within housing block, including social relationships and the quality of local public spaces, ranging from vandalism in the common areas, to housing quality and conflicts in the neighborhood. The HS also reports detailed information about the ethnic, economic and social backgrounds of individuals within the neighborhood.<sup>4</sup> In the HS, all the individuals living in a given housing block are not systematically surveyed and are randomly drawn instead. We thus cannot compute ethnic fractionalization on a representative sample at the housing block from the HS. We overcome this concern by using the 1999 French Population Census. Each HS sample is drawn from the most recent Census and the geographical units of the HS are a subsample of those of the Census. As the Census provides variables such as birth country or nationality at birth, it allows us to compute representative fractionalization indices at the housing block level and then to match them with the corresponding housing block in the HS. Table 18 in Appendix A presents descriptive statistics for the housing blocks in the Census and the Housing Survey.

## 2.2 Fractionalization indexes

We use the standard ethno-linguistic fractionalization (hereafter ELF) index used in the literature (e.g. Alesina et La Ferrara for a survey, 2005) to construct our measure of ethnic diversity.<sup>5</sup> This index reflects the probability that two randomly drawn individuals from a given population belong to different groups (previous studies looked at ethno-linguistic or religious groups). More formally, the basic fractionalization index is computed as one minus the Herfindahl index of ethno-linguistic group shares:

$$ELF_j = \sum_{i=1}^{i=N} s_{ij} (1 - s_{ij}) = 1 - \sum_{i=1}^{i=N} s_{ij}^2 \quad (1)$$

---

<sup>3</sup>INSEE has chosen this sampling strategy so as to reduce the traveling expenses of those who administer the survey.

<sup>4</sup>Some of the key variables for our study are not public. The French Statistical Institute (INSEE) made their access possible as part of a convention between the INSEE and Sciences Po. We were required to make use of the "sensitive" data within the confines of the INSEE.

<sup>5</sup>We have also tried alternative indexes such as polarization indicators, yielding similar results.



where  $s_{ij}$  is the share of group  $i$  ( $i=1, \dots, N$ ) in area  $j$ . If the population living in area  $j$  is fully homogeneous,  $ELF_j$  equals 0 and it converges to 1 as the population heterogeneity increases. Note that  $ELF_j$  can increase for two reasons: it will increase with the number of ethno-linguistic groups, and it will increase the more equal the size of the groups. As mentioned above, the Census data and the LFS provide information about the country of birth and the nationality at birth of individuals, allowing us to construct two different measures of diversity. In the remainder of the paper, we focus on diversity as measured by nationality at birth, computed at the block level. The distribution of diversity in housing blocks is presented in Appendix A (Figure 1 and Table 19).<sup>6</sup> Unsurprisingly, given that immigrants and second generation French are more likely to be eligible for public housing dwellings on income criteria than native French, the public housing neighborhoods are characterized by higher levels of diversity than other neighborhoods.

### 2.3 Sample characteristics

We now document the characteristics of the 2002 Housing Survey. Most of the variables are given at the household level. The dataset contains 32,156 households, corresponding to 78,791 individuals. 39.6 percent of the French households are renters while 56 percent are owners. Overall, 15.77 percent of the households live in public housing units, representing 39.8 percent of the tenants.

Table 1 shows the main socio-demographic characteristics of the sample. We compare households living in the public and in the private housing sectors. Foreigners (or immigrants) are over-represented in the public housing population compared to the private housing population. Public housing neighborhoods are also characterized by a poorer socio-economic environment: the unemployment rate is around twice as high as in private housing blocks. Individuals living in public housing dwellings are less educated and earn lower incomes. Around one third of the adults have no diploma at all, and the share of individuals having achieved graduate studies is less than half the corresponding share in the private housing sector. Column 3 shows that the two populations are statistically significantly different with respect to most of their characteristics. Column 4 shows the characteristics in the private housing sector when we restrict to tenants. Still, the two populations are statistically significantly different (Column 5).

Table 2 documents the perception of housing conditions by native French and immigrants in the public housing sector. On average, natives have a much better opinion about the quality of their housing than Maghrebians or other Africans.<sup>7</sup> Table 2 shows that 13.8 percent of the native French are very satisfied with their housing conditions while this is the case for only 8.42 percent of the Maghrebians. Conversely, only 9.9 percent of the native French complain about insufficient housing conditions versus 18.21 percent of the Maghrebians. The last line of Table 2 reveals that the poorer housing conditions are associated with lower levels of income, a situation more salient

---

<sup>6</sup>On the public housing graph, we see that 6 percent of public housing blocks are perfectly homogeneous. This high frequency is to a large extent explained by the fact that in many blocks we observe only very few inhabitants, thereby increasing the chance of getting a null ELF. We keep those blocks in our main analysis, but we checked that deleting them does not affect the results.

<sup>7</sup>We observe the same pattern when we look at the various subjective and objective measures of the quality of public housing

for Maghrebian and African families. In particular, the households that are very satisfied with their housing conditions earn on average 13,300 Euros per year, while very unsatisfied households earn 10,127 Euros a year on average. From the last column, we observe that the average Maghrebian family earns even less than that (8,603 Euros).

### 3 The exogeneity of diversity in the public housing sector

This section addresses the main identification issue raised by the estimation of the effect of ethnic diversity on social interactions and the quality of public goods. The issue, common to all the literature on ethnic diversity, is that fractionalization presents a high risk of endogeneity. Individuals generally tend to self segregate: they prefer forming links with others like themselves, with whom they share common interests, and in particular people of the same ethnicity or the same social background.<sup>8</sup> If people can choose the area where they live, they would rather move into neighborhoods where people are similar to themselves. If individuals who are not constrained with respect to the location of their home choose to gather along ethnic lines, then the richest individuals will be able to move into the most homogeneous neighborhoods. Therefore, the level of diversity of the neighborhoods is probably endogenous and any estimates on the implications of diversity will be biased. In particular, if the wealthy families that live in diverse settings are those that have a taste for diversity, the true effect of diversity on social outcomes should be smaller in absolute terms.<sup>9</sup>

To identify the effect of ethnic diversity, one must therefore study individuals who are assigned to their place of residence without consideration of ethnic characteristics. The purpose of this section is to bring forth evidence that spatial allocation of households across public housing blocks in France can be considered as exogenous with respect to ethnic characteristics due to French regulation. Naturally, the sample of households that apply to public housing dwellings is endogenous with respect to economic, social or cultural characteristics. But among the pool of selected households, we show that their spatial allocation across the public housing blocks of a given department is exogenous with respect to their ethnic characteristics, conditional on their other characteristics.

We also want to stress that the mere fact of working at the block level already decreases the extent of endogenous sorting. First, although households can generally choose the neighborhood in which they move, they may not be able to select a particular block in a given neighborhood. Second, while it is possible to have an idea of the socio-economic and demographic characteristics of a given neighborhood, it is much more difficult to observe such characteristics in a specific block prior to moving. Bayer, Ross and Topa (2008) rely on this key assumption and provide empirical evidence that individuals characteristics are not correlated within blocks.

#### 3.1 An ethnically-blind allocation process built into law

We first document the actual process of allocation of households across public housing dwellings. This gives a legal basis to our identifying assumption of the absence of self-sorting on ethnic charac-

---

<sup>8</sup>Race, or ethnicity, is the most salient characteristic along which homophilious relationship form.

<sup>9</sup>Combes, Decreuse, Schmutz and Trannoy (2010) use customer discrimination theory to show that owners will tend to discriminate against ethnic minorities when renting their apartment, bringing new evidence of why any causal claim of ethnic diversity on public goods in the private housing market would be biased.

teristics in the public housing sector.<sup>10</sup> In France, the only eligibility requirements for admittance into the public housing sector are to be legally living in France (as a French citizen or migrant with a valid residence permit) and to be living under a certain threshold of income per unit of consumption. This income ceiling is usually rather high: in 2009, this threshold was between 36,748 and 50,999 Euros per year for a four-person family, depending on the region of residence (the upper figure is nearly 3,000 Euros higher than the average disposable income of four-person households in 2007). Using the 2002 Housing Survey data, Jacquot (2007) estimates that given their income, between two thirds and four fifths of households living in Metropolitan France could apply for a public housing unit. As a consequence, the population eligible for public housing is about three times as large as the available space in vacant dwellings. This implies that other criteria must be taken into account in the distribution process. Hence, in addition to the income of the household, family situation and household size are taken into account to ensure a suitable match with the characteristics of vacant dwellings, as well as the emergency of the application. These are actually the main criteria used by the commission due to the boom in housing prices in the private sector during the mid-90s and the 2000s. In particular, five priority criteria are defined by law (Article L441-1 of law relative to construction and housing - *Code pour la Construction et l'Habitat*) at the national level to ensure that vacant housing will first be distributed to households with obvious social difficulties. Households satisfying these priority criteria are those in which there is a (mentally or physically) disabled person, those living in precarious or hazardous shelter due to financial constraints, those living in a temporary accommodation, individuals living in a precarious shelter who recently found a job after a long unemployment spell, and spouse-abused individuals.

To get on the queue for a housing unit, households submit a form revealing their identity and family situation, their employment status and the resources of the household, the reasons for applying to the public housing sector (currently or soon to be homeless, or reasons related to health situation, family situation, job situation, inappropriate current housing, unpleasant environment), the type of housing looked for, whether the applicant is disabled and whether this is the first application. It is important to stress the fact that the application form contains very limited information about the ethnicity of the applicant: he or she only needs to inform about his or her nationality, which is limited to three possible categories (French, European Union, or non European Union).

We now document the selection process of the applicants. The commissions of selection in charge of allocating households to vacant public housing dwellings are held at the department level (or at the city level in the case of Paris which is both a city and a department due to its size).<sup>11</sup> The composition of the commissions is regulated by law: it includes six members of the public housing

---

<sup>10</sup>The process of allocation across public housing blocks in France is mainly inspired by theories from Le Corbusier (1887-1965). Le Corbusier insisted that France must avoid the homogeneous ghettos of the urban landscapes elsewhere, and should therefore allocate housing blind to ethnicity, not permitting family networks to grow within housing establishments. These ideas were translated into state regulation (Bernardot, 2008).

<sup>11</sup>Metropolitan France is divided into 22 large administrative areas, called *régions* (regions henceforth), and into 96 smaller administrative areas, called *départements* (departments henceforth). Each department is hence a subdivision of a region, and several departments can belong to the same region. Each department is administered by an elected General Council (*Conseil Général*) and its President, whose main areas of responsibility include the management of a number of social and welfare programs, junior high schools (*collèges*), buildings and technical staff, local roads, schools, rural buses, and municipal infrastructure.

offices board, a representative of associations for social and economic insertion (appointed by the head of the department -*préfet*), mayors of the cities (or districts) in which vacant housings are to be attributed, as well as a representative of any association defending tenant rights. In addition, another department representative may attend the commission. For each vacant housing unit, at least three households must be considered by the commissioners, who finally decide which household will be allocated to which housing unit, according to the eligibility and priority criteria detailed above. Other criteria such as the number of children in the household are also taken into account in order to allocate suitable dwellings.<sup>12</sup>

With the allocation process regulated by legal rules at the national level, it seems unlikely that households can be allocated according to their origin. The main concern of the commissions is to favor socially endangered households, as shown by the priority criteria. Finally and most importantly perhaps, any decision based on the origin of an applicant, i.e. discriminating on this basis, is prohibited in France. Public housing offices are also regularly audited. If evidence of discrimination is detected, they are judged and punished accordingly. This is why the lawyers Rouquette and Lipietz (1991) stress that the rules of allocation of public housing units that prohibit "localism", and the high administrative barriers that effectively prohibit exchanges of lodgings except for changing spatial needs of families, make the allocation of public housing units largely exogenous with respect to the ethnic origins of the applicants.

Despite this legal process of allocation, one might still be worried about the possibility of self-sorting of households that refuse the residential allocation proposed by the commission. In theory, households can refuse up to three offers. However, self-sorting, especially on ethnic characteristics, seems unlikely to be a common practice. Residential mobility within the public housing sector is very low, due to the current strong shortage of supply of public housing dwellings. This makes it unlikely that the selected households could be really picky about the diversity of their neighborhood (see the study by Simon, 2003). In addition, rents are considerably lower in public housing than in private housing, increasing the opportunity cost of moving, so that the turnover is very low. More specifically, the mobility rate in the public housing sector is even lower than for recent

---

<sup>12</sup>Public housing allocation in Paris serves as a useful concrete example. We draw on the official audit of *Observatoire du Logement et de l'Habitat de Paris* (2011). Paris is a special case as it is, due to its size, a department as well as a city. The application form, the commission, and the allocation process thus take place in Paris, at the city level. As of January 2010, there were 186,017 public housing dwellings in Paris. Public housing buildings are scattered across all Parisian areas, with a high concentration (69 percent) in six districts (the 13th, 14th, 15th, 18th, 19th and 20th *arrondissements*). Within Paris, 48.7 percent of households are under the income ceiling and could be theoretically eligible. In practice, only households with very modest incomes apply (71 percent have an income lower than the minimum ceiling for all France, equivalent to 2,345 euros per month for a household with two children). On the 31st of December 2010, there were 121,937 ongoing applications, to be compared to 12,500 public housing units allocated over the year 2010. The breakdown of the population that were granted a public housing unit in 2010 is the following. 67.7 percent came from precarious housing, 28.8 percent came from the private rental sector, and 2.3 percent came from the public housing sector. In the latter case, those are people who moved for larger space following an increase in their household size (only 12 percent of the public housing dwellings have more than three rooms). The mobility rate (defined as the ratio of new entrants over the total number of public housing dwellings) is particularly low: it reaches 5.5 percent in 2010. It is formally possible to indicate a precise neighborhood in the application form, but in practice, very few applicants (6.6 percent) do provide this information. More than half of the 121,937 applicants (52.9 percent) did not mention any particular area at all, probably due to the fear of being rejected on this ground. Among those who indicated an area of preference, 91.2 percent mentioned the area where they were already living.

owners. Using data from the 2002 Housing Survey, Debrand and Taffin (2005) give precise measures of the mobility rate: it amounts to 10.3 percent for new owners, to 15.9 percent for tenants in the private housing sector, but only to 9.9 percent for tenants in the public housing sector. While even 9.9 percent may seem high, we show in section 7.2 that when households move, they almost never achieve a placement in a less diverse setting in the public sector. Besides, the mobility rates seem to have become even lower in recent years due to the boom of prices in the private sector, as shown in the Parisian case in footnote 11. The authors also document an increase in the gap in the mobility rates between the private and the public rental markets: there is was 6 point difference in 2002, to be compared to a 0.8 point difference in 1984. As a consequence of the size of the eligible population and of the low turnover, the waiting periods are rather long: the 2002 Housing Survey documents that over one third of the population applying for a public housing unit had been waiting for more than one year. A closer look at the distribution of waiting periods reveals a difference between natives and immigrants, but this difference is washed out once we control for household characteristics: the main determinant of a longer waiting period is household size. This is not surprising, as the public housing market in France is characterized by a shortage of large apartments. This is part of the explanation of the difference in waiting period between immigrants and native French, as the former tend to have more children than the latter, on average.

In a word, the public housing market is very tight, and highly regulated. This implies that households have very limited control over the time when they will be assigned a public housing dwelling and the precise place where it will be located. This is especially true at the block level, which is our level of analysis. This gives some initial support to our assumption that the distribution of households across public housing blocks is blind to ethnic characteristics and preferences of households.

### **3.2 Tests on quasi-random distribution of ethnic groups shares across public housing blocks**

In the remainder of this section, we provide a more formal statistical test to show that the spatial allocation of households across public housing blocks within a given locality is exogenous with respect to ethnic characteristics. We focus here on what we see as the most stringent test on the spatial distribution of residents, which consists in comparing the observed distribution with a random distribution. This allows us to test directly our assumption of quasi-randomness of the allocation of households across public housing blocks within a given area.

In Section 7, we perform a variety of alternative tests. First, we show that while households moving into a new neighborhood tend to self-segregate in the unconstrained private housing market, there is no such evidence in the public housing market. This test points out the absence of self-selection along ethnic lines among the movers. But self-selection could occur prior to the move. Thus we also look at households that have refused a public housing dwelling offer. We show that households that have refused an offer end up living in public housing blocks that display the same ELF as those who accepted their first offer. Thus even if households try to be choosy with respect to the ethnic composition of their neighborhoods, they cannot self-segregate in the public housing sector due to

the allocation process and the tight supply constraints of dwellings.

We now document our test on the quasi-random residential allocation in the public housing sector. As mentioned in section 3.1, the allocation of households across public housing blocks takes place at the department level. If the members of the commission follow the legal criteria and do not take into account the ethnic characteristics in the allocation process, we should find an equal distribution of households of a given nationality across the various public housing blocks within each department. For the sake of illustration, let us assume that 10 percent of Maghrebians live in the public housing sector in Paris. We should find the same share of 10 percent of Maghrebians within each Parisian housing block if the allocation was truly exogenous with respect to ethnic characteristics. Naturally, this equality of distribution of ethnic groups shares across housing blocks can hold only if we have a sufficiently large number of individuals within each housing block. Instead, in the Labor Force Survey, on which we perform the test, we only observe an average of 40.55 different individuals corresponding to 15.53 households in each neighborhood. This is due to the sampling strategy of the French National Institute of Statistics and Economics (INSEE) that interviews all the individuals from a given neighborhood, but consequently limits the size of the neighborhood. With such a small sample size of observations at the neighborhood level, any analytical test of equality of distribution of ethnic groups shares across blocks would fail. We thus use Monte Carlo simulation to reproduce an artificially random distribution of the population. We randomly reallocate the public housing population across the different blocks within each department, and then compare this random distribution to the actual distribution.

Let us now describe more precisely this test. We perform Monte Carlo simulations generating artificial random allocations that we later compare to the observed allocation. As mentioned above, we use the Labor Force Survey to perform this test, because its unique design allow us to work on the entire population of the surveyed neighborhoods. For each *département*, we pool the public housing population and reallocate it randomly, without replacement, across the different residential blocks of the corresponding *département*, maintaining unchanged the actual size of each block. We get a simulated random allocation of individuals with a given characteristic across blocks. We then run a Kolmogorov-Smirnov (KS) test of equality of distribution of this given characteristic across neighborhoods with its actual distribution. More precisely, we are interested here in the distribution of the share of Native French and of the share of Maghrebians across neighborhoods. We finally determine the percentage of *départements* for which the actual and simulated distributions across housing blocks are similar, i.e. those for which we cannot reject the null hypothesis of equality of the distribution at the 10% level. The Monte-Carlo simulation results presented here are based on one hundred replications of the process described above. For each draw we run the tests for the equality of distributions, and then average the results.

Table 3 shows the values of those tests averaging over 100 Monte Carlo draws. The labels in the first column indicate the ethnic characteristic under consideration by distinguishing Maghrebian origins and French nationality at birth. The second column reports the results of the Kolmogorov-Smirnov test within the public housing sector. For the sake of comparison, we run in Column (3) the same KS-test on the full sample, including both those who live in the public and private housing

sectors.

The equality of spatial distribution between the random simulated distribution and the real observed one is accepted in most *départements* in the public housing sector. In particular, the equality of distribution with respect to Arabic origin (respectively French origin) is not rejected in 80 percent (respectively 70 percent) of the departments in the public sector. In contrast, Column (3) shows that in the whole sample, the equality of distribution is not rejected in 54.3 percent of the departments for the share of Maghrebians and falls dramatically to 24.8 percent of the department for the share of Native French. This test shows that while French natives (and to a lesser extent Maghrebians) do self-segregate a lot when we consider the whole sample, this is no longer the case in the public housing sector. All in all, those tests are supportive of our identifying assumption that the allocation of households across the public housing blocks can be considered as exogenous with respect to their ethnic backgrounds.

## 4 Results

### 4.1 Specification

This section estimates the impact of diversity on social relationships and public goods at the local housing block level. We identify the effect of diversity by using data from the public housing sector where households are exogenously allocated with respect to ethnic characteristics. Let  $j$ ,  $k$  and  $l$  indicate respectively households, buildings and blocks. For each outcome, we estimate the following equation:

$$Y_{jkl} = \alpha + \beta ELF_l + \gamma X_j + \delta Z_k + \mu W_l + \varepsilon_{jkl} \quad (2)$$

where  $Y_{jkl}$  denotes the housing outcome we are interested in, as stated by household  $j$  in building  $k$  and block  $l$ ,  $ELF_l$  is the level of ethnic diversity in the block,  $X_j$  is a vector of household characteristics,  $Z_k$  a vector of building characteristics and  $W_l$  a vector of socio-economic characteristics of the block. We also control for department fixed effects since the spatial allocation of households across public housings is decided at the department level.<sup>13</sup> All results derive from OLS estimates, with robust standard errors clustered at the block level.<sup>14</sup>

We control for a large set of household characteristics: age, gender, level of education, labor market status and nationality<sup>15</sup> of the household head, as well as household size, and total household income per member. We also control for building characteristics, with the (log)-number of apartments in the housing project, and its date of construction. Indeed, the size and the number of occupants might affect the ability of the households to coordinate for improving the commons or to enforce norms, while the age of the building might explain part of the degradations observed and tenant satisfaction.

---

<sup>13</sup>Results were unaltered by the inclusion of city fixed effect to account for the fact that the mayors are members of the attribution committee.

<sup>14</sup>We have also run logistic regressions on dummy outcomes, with similar results. To ease the interpretation of the coefficients, we will report the OLS estimates henceforth.

<sup>15</sup>We distinguish between the following categories for nationalities: French at birth, naturalized French, from other European countries, Maghrebian, Sub-Saharan African, Asian and all others.

An important issue is whether the degree of fractionalization is picking up various dimensions of the environment where people are living, including the extent of inequality and the unemployment rate or the socio-economic background of the neighborhood (Alesina and La Ferrara, 2002). We therefore include a very detailed classification in 27 categories of the socio-economic environment of each neighborhood, constructed by Nicole Tabard (2002) from the INSEE. This classification characterizes each area according to the socio-professional category and the occupation of all men in the area. We use the classification that was built using the 1999 census data. This variable is the most detailed one available in French national surveys to capture the socio-economic background of an area. We also include the unemployment rate computed at the block level using the 1999 Population Census data. Finally, we include department fixed effects.

## 4.2 The effect of fractionalization on the opinion about housing conditions

We start by documenting the impact of diversity on the overall satisfaction about housing conditions. From the HS, we use the question: "In general how do you judge the quality of your housing conditions?". The variable takes on values from 1, for very good, to 5 for very bad. Over the public housing population, the average of this variable is of 2.5, with a 0.98 standard deviation. This question on well-being related to housing conditions is rather general. We will detail the different dimensions that could affect this well-being in the following subsection.

Table 4 shows that ethnic diversity has a negative effect on satisfaction with housing condition, statistically significant at the 5 percent level. A one standard deviation increase in ethnic diversity generates an increase in the dissatisfaction with housing conditions that amounts to 6.7 percent of its standard deviation. To get a better sense of the magnitude of this effect, we can say that the increase in the dissatisfaction with housing conditions generated by a one standard deviation increase in block unemployment rate corresponds to 13.1 percent of its standard deviation. Thus the effect of diversity on satisfaction is as sizeable as half the effect of the local unemployment rate. Two other variables seem to be related to household satisfaction with housing conditions: members of larger households tend to be less satisfied with their housing condition, while those living in newer buildings (constructed after 1990) have a significantly better opinion on the subject than others. Finally, older and more educated individuals also complain less than others, but to a lesser extent.<sup>16</sup>

## 4.3 The various effects of diversity

### 4.3.1 Vandalism, Housing quality and Public Safety

This section looks further at the various dimensions of the dissatisfaction with housing conditions that could be affected by ethnic fractionalization. The HS covers a large variety of questions on social relationships and the quality of the housing environment. Table 5 reports descriptive statistics of the outcomes we look at. To organize the discussion about those questions, we distinguish three main dimensions: (a) Behaviors and Public goods that are directly under the control of the tenants.

---

<sup>16</sup>The results reported in 4 also show that people of Asian nationalities tend to be more satisfied with their housing condition than native French. However, we do not give much credit to this figure given that we observe only 9 Asian individuals in the public housing sector in our dataset.



This category includes in particular the neglect or voluntary degradations that may be imputable to the tenants, (b) Behaviors and Public goods that are under the control of the landlords. This category include in particular the poor quality of housing due to a lack of maintenance and repairs by the landlords and (c) Interpersonal conflict and hatred social relationship. This category includes personal aggression and robberies to which we refer as civil conflicts or public safety. We have also run an exploratory factor analysis that yields similar, if not identical, categories. Appendix C reports the results obtained with the three indices resulting from the factor analysis.

The first dimension of housing quality refers to actions or goods that are largely under the control of the tenants. In this category, we include all the variables reporting neglect or voluntary deterioration in the common areas of the building. First, households are asked a general question on degradations: *"Were the common areas of your building (lobby, staircase, floors) vandalized or neglected (destruction, deterioration) over the last twelve months?"*. The answers are 1 for "Never", 2 for "Minor degradations" and 3 for "Major or very frequent degradations". Households are then asked to mention which kind of degradations they observed over the previous year. They can choose several possible answers from the following list: *graffiti or degradations of the walls (or on the floor), trash and litter on the floor, broken windows, broken doors, broken light bulbs, degradation of mail boxes, degradation of the entry phone or entry code, deterioration of the elevator*. For each outcome, the variable is coded as 1 in case of a degradation, and 0 otherwise. All those items refer more or less directly to a willful degradation. We will thus refer to this set of questions as the category *Vandalism*. We also include in this category a question about noise pollution:<sup>17</sup> *"How frequently are you disturbed by the noise in your housing during the day?", "During the night?"*. The answers are 1 for "Infrequently or never", 2 for "Rather frequently", and 3 for "Very frequently".

The second category we consider refers to goods that are not directly produced or altered by residents. But they might be related to diversity by the lack of maintenance and repairs by the HLM office to improve the housing quality. We will henceforth label this category *Poor Quality of Housing*. We include in this category variables corresponding to housing problems that can neither be caused nor solved by the tenants, but for which HLM offices are responsible. The households are first asked: *"How would you qualify the way the common areas of your building are maintained and taken care of (cleaning, maintenance of collective facilities: lighting, trash cans,...)?"*. The answer ranges from 1 for good, to 2 for average, and 3 for bad. More specific questions are also asked: *"How does the façade of your building look?"*,<sup>18</sup> *"What is the quality of the soundproofing of your housing?"*,<sup>19</sup> *"Was the elevator out of order during more than 24 hours over the past three months?"*,<sup>20</sup> *"Did you experience toilet issues (leaks, flush breakdown, drainage problem) over the last three months?"*, or *"Did you experience coldness in your apartment during more than 24 hours over the past twelve months?"*. We also include more detailed questions concerning the origin of

<sup>17</sup>The underlying assumption is that the source of the noise in the hallways and apartments of the building is not due to poor soundproofing. As a matter of fact, we see in the following sections that more diversity increases the disturbance related to noise, but fails to explain the quality of soundproofing.

<sup>18</sup>There are five possible answers: 1=As new, 2=Good, 3=Average, 4=Dirty, 5= Bad, with cracks, 6=Very bad, the building threatens to collapse.

<sup>19</sup>The possible answers are: 1=Good, 2=Average, 3=Bad.

<sup>20</sup>In contrast, the question mentioned in the *Vandalism* section refers to the interior status of the elevator rather than its mechanical breakdown.

coldness: *"Did you experience coldness because of a bad insulation?"*, *"Did you experience coldness because the heating equipment broke down ?"* and *"Did you experience coldness because of a poor heating equipment?"*. For all the previous questions, the variable is 1 when the answer is "Yes" and 0 otherwise.

The last category of questions refers to personal aggressions and criminality. We will label this category *Public Safety*. Three questions correspond to this category: *"Have you, or a member of your household, been a victim of or a witness to physical aggression in your neighborhood during the last twelve months?"*, *"Have you, or a member of your household, been a victim of or a witness to a robbery in your neighborhood during the last twelve months?"*, and *"Have you been victim of a burglary (or any attempt) over the past twelve months?"*. For these three questions, the variable equals 1 in case of the event, and zero otherwise.

### 4.3.2 Results

Tables 6 to 8 show the effect of ethnic fractionalization on the various outcomes corresponding to the three different dimensions: "Vandalism", "Quality of housing" and "Public Safety". For each outcome, we run a separate regression according to equation 2, using the largest set of control variables we could think of. More precisely, we report the results when we control for households characteristics, building characteristics (which may explain a large part of the degradations observed in the housing projects), neighborhood characteristics and department fixed effects.

Table 6 reports the effect of ethnic diversity on outcomes related to voluntary degradations and vandalism. We report the results in the public housing environment, with the full set of controls as detailed above. For almost all the outcomes considered, the estimated effect of ethnic diversity is statistically significant at the 1 percent level,<sup>21</sup> and is sizeable. Let us for instance look at the results for graffiti in the full-specification: a one standard deviation increase in ethnic diversity is associated with a rise by 5.6 percentage points in the probability of observing graffiti, which represents 12.8 percent of the total standard deviation of this outcome. The effect of ethnic diversity is comparable to the effect of local unemployment: a one standard deviation increase in block unemployment rate is associated with a rise by 4.56 percentage in the probability of observing graffiti. Regarding the deterioration of elevators, a one standard deviation increase in ethnic diversity induces a 4.9 percentage points increase in the probability of observing degradation of the elevator, which represents 16.7 percent of the total standard deviation of this outcome. This effect is once again as sizeable as that of the local unemployment rate. Note also that the size of the building (i.e. the number of housings) has a strong positive impact on all the outcomes related to vandalism.

The second set of regressions, presented in Table 7 shows the effect of diversity on outcomes signaling poor quality of housing. The coefficient associated with ethnic diversity is generally significant, except for the outcome related to the quality of soundproofing.<sup>22</sup> More diverse neighborhood

---

<sup>21</sup>The exception is for the indicator for broken doors and noise during the day, for which the effect of diversity is only significant at the 5% level.

<sup>22</sup>Two other variables are not explained by diversity, but these are not outcomes per se. These variables are reasons why the household experience coldness in the apartment.

are characterized by a lower care of the commons by the persons in charge, a poorer condition of the façade, more frequent concerns with heating, more frequent elevator breakdowns and toilet issues. It is worth noting that the estimated effects of diversity are much lower than those found for outcomes associated with vandalism. Consider the outcome associated with the probability that the elevator is out of order. We find that when the ELF increases by one standard deviation, the probability that the elevator was out of order during at least 24 hours over the last three months rises by 1.9 percentage points. This corresponds to only 5.24% of the standard deviation of this outcome. If we now turn to heating issues, our results indicate that a one standard deviation increase in ethnic diversity is associated with a rise by 2.43 percentage points in the probability to have experienced insufficient heat in the apartment during more than 24 hours over the past year, which represents 6.41 percent of the total standard deviation of this outcome. The date of construction of the building is also an important explanatory variable for most of the outcomes related to general housing quality, as it accounts for the general state of capital equipment under the responsibility of the public housing office (heating, façade, soundproofing,...). Living in a more recent building decreases particularly the probability to report poor condition of the outside walls, low quality of insulation or poor quality of soundproofing.

Finally, Table 8 reports the results for outcomes related to public safety, capturing direct aggression, robberies and burglaries. Remarkably, it shows that ethnic diversity does not have a significant impact on any of these outcome variables in our preferred specification. This finding is consistent with Fearon and Laitin (1996), who argue that despite inter-ethnic relations being generally more tense, in-group policing mechanisms typically keep violence off of the equilibrium path.

In sum, and taking advantage of data at a more micro level than has heretofore been available, we see that fractionalization operates with different degrees of impact for different sorts of public goods and social relationships.<sup>23</sup> To be sure, results are not significant for all of the outcomes that we examine. But overall, the results are clear that fractionalization at the local level increases vandalism by a great deal, decreases building maintenance by a moderate (but overall significant) degree, and has no effect on security. These findings allow us (as we do in the next section) to propose the various channels through which fractionalization works in the provision of public goods.

### 4.3.3 Rationalization of the channels

To rationalize our findings, we propose different interpretations of the channels through which fractionalization affects social relationships and local public goods. The category "Vandalism" refers to the neglect or voluntary degradations of the common areas of the building, such as damaging common property, graffiti, or depositing trash on the floor. These are outcomes over which public housing residents have control and for which they can be held responsible.<sup>24</sup> The category "Quality of housing" include variables such as the condition of the outside walls, quality of soundproofing or coldness in the apartment. Those variables are more the responsibility of the public housing

---

<sup>23</sup>Our findings are unchanged with regressions on aggregated indices obtained with a principal component analysis (see Appendix C), and with a mean effect analysis (see Appendix C).

<sup>24</sup>Given that residents need to enter a code in order to gain entry into their building, it is unlikely these degradations are coming from outsiders.

managers. Finally, the "Public Safety" category represents outcomes that are less under the control of local authorities than of the state police. We find that both locally controlled outcomes are negatively affected by diversity, and we provide a different rationalization of the channels for each type.

Our interpretation for the results on "Vandalism" is that diversity prevents the creation of social norms to punish defectors, as the threat of social sanctions is lower across groups. The other-regarding preferences are less effective in more diverse areas. This has been a standard result in the literature since the seminal work of Coleman (1988), and it helps explain why we observe more voluntary degradations with diversity. Supporting our intuition, many households living in the public housing sector report having "no relationship at all" with their neighbors, rather than "bad" or "very bad relationships", which can be a barrier to the creation of other-regarding social norms. The increase in graffiti in more diverse areas might also illustrate the need to mark one's territory in a context where several groups co-exist.

We understand the result on "Quality of housing" as the inability of more heterogeneous communities to undertake collective action that would pressure the public housing office into improving housing quality. This could be sustained (though we have no direct evidence to support this) by beliefs in the housing directorate that it need not maintain public goods to high standards in heterogeneous housing projects because the likelihood of collective action against it is minimal. In this sense, the resulting poor housing quality associated with ethnic diversity can be seen as an equilibrium in which the lack of expectations of collective action would fail to incentivize the housing directorate to make costly improvements.<sup>25</sup> Coming back to the results concerning heating issues, we can find some support for this assumption: we find that households living in more diverse neighborhoods not only report more heating failures, but also report that this is due to the poor quality of the heating equipment, an appliance typically under the control of the HLM office.<sup>26</sup>

Finally, we can think of two possible interpretations of the absence of any diversity effect on aggressions and robberies. First, this could result from more physical security provided by the city and state police in more diverse neighborhoods. The second explanation would be that individuals living in the public housing sector in general experience social anomie. In fact, one third (32.7 percent) of the public housing population, irrespective of diversity, declares to have no relationship at all with individuals living in their same area. In addition, we find that individuals living in a more diverse neighborhood tend to even more social anomie and fewer relationships with their neighbors.

To summarize, our interpretation is that diversity generates social anomie, i.e. the absence of common rules and social norms. As a consequence of anomie, there is (a) less other-regarding

---

<sup>25</sup>The collective action could also influence mayor's office. But the political logic of the public housing support is beyond the scope of the paper

<sup>26</sup>Another possible reason for having experienced coldness in the apartment that the household can mention is to have restricted heating in order to save money. The results are not reported here, but we find no significant effect of diversity on this outcome, in any specification.

preferences, hence more neglect and vandalism, (b) a failure to generate collective action to pressure the public housing offices into improving housing quality, and (c) fewer opportunities for violent confrontation at all levels of diversity. We also interpret the lack of an effect of diversity on violence by security provided at a higher level of administration, not subject to the constraints of local diversity.

#### 4.3.4 Interpretation of the channels based on Repairs

We bring additional evidence on the interpretation of the channels by looking at maintenance and repairs performed in the building. Note first that these outcomes add an objective dimension to the previous subjective questions. The variation in the effects of diversity on the number of repairs depending on the type of public good also helps us to tease out the different channels through which diversity operates.

The Housing Survey asks whether elevators, staircase, windows, heating equipment, security equipment, and so on, have been repaired or installed during the previous year. We build three measures of repairs, corresponding to our three general outcomes. We define a first variable tracking repairs that can be fixed internally by the tenants. Those repairs concern staircase, windows, doors and lights of the commons, i.e repairs related to neglect or voluntary degradations. A second variable indicates repairs that can be fixed only by the external intervention of the public housing office. Those repairs include revamping of the façade, or interventions to improve, among other things, the heating system or insulation quality, i.e. repairs related to the general quality of housing. Finally, we build a third variable accounting for the installation of security equipment in the building, which can be related to conflicts outcomes. We then regress each of these three variables (as well as less aggregated indicators of repairs) on the level of diversity of the block, controlling for factors that could explain the number of repairs: the number of dwellings in the building, and its date of construction. Table 9 reports these OLS estimates.

Column 1 of Table 9 reveals a positive and statistically significant correlation between the probability of repairs inside the building (windows, doors, lights... in the common areas) and local diversity: the more the diversity, the more the work for repairing the effects of vandalism. In the main regressions of the paper presented in section 4.3.2 (corresponding to the Table 6), we found that voluntary degradations increase with diversity. This implies that the larger number of repairs results from greater need due to a lack of publicly spirited social norms rather than from greater responsiveness by the housing authorities to regular maintenance.

Column 2 shows a negative and statistically significant correlation between the number of substantial works in the building (façade, heating, insulation...) and local diversity: the more the diversity, the less the work asked by the landlords for improving the general quality of housing. In the main regressions presented in section 4.3.2 (corresponding to the Table 7), we found that more diversity implies a lower quality of housing. Thus, it seems that more diverse neighborhoods are deprived of such substantial work, although the inhabitants actually complain (individually to survey enumerators) about the quality of housing. This supports our intuition that tenants in more diverse neighborhoods are unable to engage in collective action to pressure the public housing offices

into undertaking important works.

Finally Column 3 shows a positive and statistically significant correlation between the existence of security equipment and local diversity. In the main regressions presented in section 4.3.2 (corresponding to the Table 8), we found no impact of diversity on aggression and robberies. The presence of security equipment in more diverse neighborhood might be part of the explanation of the absence of diversity effect on burglaries. In addition, it is supportive of our idea that vandalism in the common areas of the buildings is imputable to the tenants, who have access to the building.

## 5 Robustness checks on self-reported quality of public goods

One concern in the previous analysis is related to the subjective nature of the outcome variables used in our study. Self-reported perceptions might reflect personal bias rather than be correlated with objective measures of public good provision. Perhaps people are just happier when they are surrounded by people more like themselves, and this is reflected in their answer to the quality of housing. We conduct several tests challenging this alternative explanation of personal bias.

First, as mentioned in the previous section, the HS provides information about various types of repairs and work that have been done in the building or in the housing unit over the previous year. These variables present the advantage of being objective. The lower part of Table 9 reports simple correlations between the various outcomes and the associated repairs. We find that most of our subjective outcomes are strongly and positively correlated with the existence of repairs, i.e. objective outcomes, especially for the variables related to vandalism.<sup>27</sup> This is our initial evidence of the reliability of our subjective measures of housing quality and well-being.

We then conduct more formal tests. We replicate the regression on the dissatisfaction with housing conditions (section 4.2) including interaction terms between diversity and the various ethnic groups. Those estimates reveal whether different groups react in different ways to the level of diversity of their neighborhood. Column 2 of Table 10 shows that there does not seem to be a different effect for the various groups, and the coefficient for diversity remains unchanged (see Column 1 for the baseline specification). Then we concentrate on actual differences between "pure French" households<sup>28</sup> and fully Maghrebian households' dissatisfaction with housing conditions. In particular, we interact the dummies of being in a fully native French household or being in a fully Maghrebian household with the ELF: none of the coefficients is significant (see Table 10, Column 3). Thus for any given level of diversity, there is no significant difference in the answers given by pure French and fully Maghrebian households. In other words, the idea that bad opinions of housing conditions are driven by average bad feelings due to being surrounded by foreigners can be rejected. Moreover, including these additional controls only slightly affects the magnitude of the ethnic diversity coefficient, and does not affect its direction or its significance.

---

<sup>27</sup>An exception is the condition of the outside walls, which is negatively correlated with the probability that façade work was done. This is not surprising as the assessment of the façade's condition is done at the time of the survey, while repairs concern the previous year.

<sup>28</sup>Both children and parents were born French in France.

An alternative test to show that subjective perceptions have an objective foundation is to look at how much within-housing project variation there is in perceptions. We regress self-reported perception on housing project fixed effects and individual characteristics. Once we control for housing project fixed effects, assuming there is no within-project variation in public goods, the remainder of the variation tells us if certain ethnic or socio-demographic groups are more likely to be positively or negatively biased. If perceptions have a high signal-to-noise ratio, there should be less within-project variation because perceptions would be a good signal of project level public goods. Table 11 shows the regressions of our main indicator of satisfaction about housing conditions on individual characteristics. Column 1 shows the within-housing project estimates by including housing project fixed effects, exploiting variation within public housing. Column 2 shows the between-housing project estimates. Column 1 shows that the only individual characteristics statistically significantly correlated with within-project variation in the perception of the environment are age and household size. Income, education or the country of origin of the households are uncorrelated with perceptions of the environment. We also compute the standard deviation in the perception of the quality of housing between public housing projects and within public housing projects. The standard deviation is almost twice as high across blocks (.801) than within blocks (.435), and this difference is statistically significant. In sum, low levels of within block variation on perceptions adds confidence that there is an objective foundation for tenants' subjective reports.

Finally, we also estimate the effect of diversity on the perception of the quality of public goods that are financed by the city, the department or the state rather than locally financed by the HLM offices. If there is a reporting bias in general, then, the effect on all types of public goods should be the same. If it is related to localized collective action failures, then the impact should only be on locally provided/maintained public goods. Thus this test provides both an additional robustness check on the channels through which diversity affects public goods and on the absence of a reporting bias. The local public goods we have focused on so far (except for individuals' protection) are provided or maintained by the private company that owns and manages the public housing building. We now consider public goods that are managed at the city or department level. In the HS, three public goods enter this category. The first one is the perception of the quality of roads and streets with the following question: *"What is your opinion about the maintenance of the streets, roads and public spaces in the area?"*. The second question measures the access to public transportation: *"What is your opinion about the accessibility of your area by public transportation?"*. The third item measures the accessibility of the area by private transportation: *"What is your opinion about the accessibility of your area by private vehicles (parking, congestion)?"*. The answer ranges for all three questions from 1 for good, 2 for neither good nor bad to 3 for bad. Table 12 reports the OLS estimates, controlling for all the previous household and local characteristics in addition to department fixed effects. We find that ethnic fractionalization is neither correlated with the quality of public spaces and roads in the areas (Column 1), nor with public transportation (Column 2), nor with car parking and general congestion (Column 3). Again, our confidence that the subjective reports to enumerators on housing quality have an objective foundation is increased.

## 6 Further tests: Fractionalization and ethnic shares

The basic regressions measure ethnic diversity using a standard ELF, controlling for household, building and neighborhood characteristics. Yet, as suggested by Vigdor (2002), it might be important to control for ethnic group shares to get a more comprehensive set of covariates for diversity. Column 1 of Table 13 reports the results once we control for ethnic group shares.<sup>29</sup> The estimated impact of ELF is now even stronger than in the previous specifications, confirming the robustness of our result along this dimension.

Moreover, we run regressions replacing the fractionalization index by ethnic groups shares (Column 2 of Table 13), and by ethnic group shares and their square (Column 3 of Table 13), controlling for the usual individual and local characteristics. Only one group (Maghrebians) seems to have a significantly negative effect on the dissatisfaction with housing conditions: the higher the share of Maghrebians in a block (relative to the share of French), the more likely individuals are to complain about their housing conditions. However, this negative effect decreases with the share of Maghrebians. From this result, we infer that our measure of diversity reflects not only the actual ethnic composition of the neighborhood, but also that some ethnic groups might have different effects on self-reported perceptions of the quality of public spaces as they become a majority of the neighborhood population. However, this result does not call into question the effect of diversity per se on which we have already reported (Column 1 of Table 13).

Finally, we re-run our main regressions using an alternative fractionalization index, trying to encompass another dimension of diversity. More precisely, in order to account for communication issues potentially related to the diversity of origins, we compute an alternative fractionalization index based on a proxy for French speaking. We use information provided by the International Organisation of La Francophonie to group countries according to the share of the population that is French speaking (in 2010). Their classification allows us to distinguish among six groups of countries: countries which are not members of the organisation,<sup>30</sup> countries in which French speakers represent less than 5 percent of the population, countries in which French speakers represent between 5 and 15 percent of the population, countries in which French speakers represent between 16 and 35 percent of the population, countries in which French speakers represent between 36 and 60 percent of the population, and countries in which French speakers represent more than 60 percent of the population.

We then assign one of the six French-speaking levels to each individual (from the 1999 census dataset), according to his or her nationality at birth. This sorting of individuals captures the probability that they actually speak French. Finally, we compute a standard ELF for each block (using the same methodology as for our main index of diversity) relying on the shares of the block population belonging to one of the six groups.

We replicate the regressions of the paper<sup>31</sup> using this alternative index instead of the one based

---

<sup>29</sup>In Table 13, we aggregate the various nationalities at birth into six different categories. The results are similar when we work with more detailed shares for all nationalities. The share of native French is the omitted category.

<sup>30</sup>For these countries, the organisation does not provide any data, but we can reasonably assume that the share of French speaking population in non-member countries is close to zero.

<sup>31</sup>We replicate the regression of the dissatisfaction with housing conditions (Columns 1) as well as of the three



on nationality at birth. The results are reported in Table 14. Each coefficient comes from a separate regression. The columns indicate the four dependent variables under study. Panel A and panel B respectively correspond to the measure of diversity used in each regression. Our results are unaffected when we use the new index based on French speaking origin. This is not very surprising given that the correlation between the two indices is very large (98.45 %).

## 7 Additional Tests on the Exogeneity of Residential allocation in the Public housing sector

### 7.1 Absence of self-sorting on ethnic backgrounds

Our first set of alternative tests consists in showing that while households tend to self-segregate in the unconstrained private housing market, there is no such evidence in the public housing market. We test this using the LFS and focusing on individuals who recently moved into an area (within the previous year).

We first estimate the correlation between the hourly wages of the movers and the level of diversity of the area into which they just moved. Without prior beliefs over agents' preferences, if individuals have a taste for or against homogeneity, there should be a significant relationship between the level of ethnic diversity prevailing in their neighborhood and their wages when their choice is not constrained by legal rules. Indeed, in an unconstrained market (e.g. the private housing market), the richer the individual, the easier it should be for him or her to choose his or her neighborhood. Therefore, if the level of diversity of the area enters one's preferences, there should be a correlation between individual wealth and the level of diversity in the area. In the public housing market as well, some public housing units are more expensive than others, depending on the location and the date of construction. The wealthiest inhabitants could thus have some control over the diversity of their neighborhood, in theory.

To test this, we compute the fractionalization index of the area to which a household recently moved taking into account only the neighbors who had been living there for more than one year. We thus calculate the fractionalization indices at stake prior to the move. For the private housing market, we compute the fractionalization index of the whole area, including both the population living in private and public housing dwellings within this area. We follow this strategy since there are a few areas with both public and private housing units in the LFS. It is reasonable to think that it is the level of diversity of the whole neighborhood that will matter in the mobility decision in the private market.<sup>32</sup> Regarding the public housing sector, we compute the level of diversity including residents of the public housing only, our identification assumption being that households do not have control over the level of diversity of their neighborhood *within* the public housing sector.

We run OLS estimates of the hourly wage of newly arrived individuals on the level of diver-

---

aggregate indices we obtained with the principal component analysis presented in Appendix B.

<sup>32</sup>The results are unchanged if we consider only the population living in the private housing sector: the magnitude of the correlation decreases marginally, but remains statistically significant.

sity of the area in which they just moved, controlling for the department of residence.<sup>33</sup> First, we focus on individuals having just moved into a private dwelling. We find a very strong negative relationship between income and diversity (the estimated coefficient is -0.14 and is significant at the 1 percent level).<sup>34</sup> Then we look at the sample of individuals having moved into a public housing dwelling within the past year. In this case, the simple OLS regression reveals that there is no significant correlation between the income of individuals moving into a public housing dwelling and pre-existing diversity of nationalities within the neighborhood.<sup>35</sup> These correlations show that while the wealthiest households tend to self-segregate in less diverse areas in the (unregulated) private housing market, it does not seem to be the case in the (regulated) public housing market. In other words, although diversity enters households' preferences as revealed by the private housing market result, the location in the public housing market seems to be unaffected by such preferences.

Our second test uses the same methodology and estimates the link between the origin (nationality) of individuals moving into a new area and the share of the area's "long term" population of the same origin.<sup>36</sup> We expect a significant relationship in the private housing market where location choice is relatively unconstrained but not in the public housing sector. Table 15 reports the results from an OLS regression of the share of neighbors from the same origin as new movers on new movers' characteristics: nationality group, public housing dummy, quadratic function of age, hourly wage (log) education, socio-economic category, department fixed effects, and interaction of individual characteristics with the public housing dummy. We consider seven different nationality groups: native French, naturalized French, Europeans, Maghrebians, other Africans, Asians, and other nationalities, which is taken as the reference group.

Three facts are worth noting here. First, there is indeed evidence that on average native French are significantly more likely to move in neighborhoods where the share of natives is higher, compared to households from other nationalities. This is not surprising given the fact that natives make up a large majority in the French population. The second interesting point is that the coefficient for living in the public housing sector is negative and significant at the 5% level. More precisely, it reveals that HLM households move in areas where the share of individuals from the same origin is on average 15.5% lower than for households in the private housing sector. This result strengthens the idea that the extent to which households in the public housing sector live close to their co-ethnics is lower than in the private sector. Finally, when we interact nationalities with the public housing dummy, none of the coefficients but one are significant. This comforts us with the idea that there is no particular self-segregation along ethnic lines in the public housing sector. The only coefficient that is significant is for naturalized French individuals, which is not a clear ethnic group. However, when we control for the share of one's own ethnic group in one's department in column (2)<sup>37</sup>, even

---

<sup>33</sup>If we reverse the dependent and the explanatory variables, the sign and significance level of the coefficient remains the same.

<sup>34</sup>This is powerful evidence of the bias introduced with endogenous sorting

<sup>35</sup>The results are not displayed but are available upon request.

<sup>36</sup>A similar test was proposed by Goux and Maurin (2007) to show that the educational achievement of the children of newcomers in public housing is uncorrelated with that of the current residents. Individuals do not self-select in public housing neighborhoods according to the educational achievement of the neighbors' children. By contrast, the authors find a strong self-selection on the educational characteristics in the private housing sector.

<sup>37</sup>The purpose of including this variables is to account for the fact that part of the ethnic sorting is likely to be due

this interaction term turns insignificant.<sup>38</sup>

We conduct the same kind of test on other individual characteristics, and reach similar conclusions. We find that in the private sector, highly educated (respectively low skilled) individuals are very likely to move into neighborhoods with higher levels of highly educated (respectively low skilled) people. This is not surprising and illustrates self segregation along education level in the private sector. On the contrary, such segregation does not appear in the public housing sector. The only characteristic for which we find a positive correlation between the new and the old inhabitants in public housing blocks is the fact of being a factory worker. This is perhaps not too surprising either given that factory workers represent more than 30 percent of the public housing population, and due to the history of public housing, which was initially (and over several decades) dedicated to factory workers.

We also regress the probability of having moved in a new HLM dwelling in the past year (dummy equal to one in this case and to zero if the individual was already living in the same HLM apartment one year before) on individual characteristics (nationality, age, gender, wage, education, socio professional group), and the interaction of these characteristics with the ethnic diversity among the public housing population of the block.<sup>39</sup> As would be expected in the absence of sorting, the coefficients on the interaction terms are not significantly different from zero. The only exception is for the interaction of ELF with the dummy for African nationality, for which the coefficient is negative and significant at the 10 percent level.

Finally, we build on the information about relationships provided by the Housing Survey to bring additional evidence on the absence of sorting. More specifically, respondents to the HS are asked the following questions. "*How would you qualify the relationships with your close neighbors?*" and "*How would you qualify the relationships in your neighborhood?*", where the first question refers to the direct neighborhood, while the second one refers to a larger neighborhood. The possible answers are 1 for "Good", 2 for "Average", 3 for "Bad" and 4 for "No relationship". From this, we build two dummy variables indicating whether the individual reports any relationship (good, average or bad) in the direct or in the larger neighborhood, respectively. Simple descriptive statistics and correlations reveal that living in the public housing sector is associated with a significantly higher probability to have no relationship at all with your neighbors, the more so the larger the neighborhood considered. However, when we control for neighborhood, households and building

---

to an over-representation of some groups in given departments and in particular in the public housing sector.

<sup>38</sup>In a previous version of the paper, we regressed individuals' origin on the share of the population of his or her new neighborhood from each nationality in each housing sector, rather than pooling individuals from all origins. The results were the following. In the private housing sector, a significant relationship between one's nationality and the share of same-origin neighbors showed up for most of the nationality groups. By contrast, in the public housing sector, there was no statistically significant relationship between the nationality of the individual and the share of the "long term" population in the area having the same nationality. The correlation was close to zero for households with African origins, and was around three times as low as in the private sector for households with Maghrebian origins. The only significant relationship showed up for immigrants from Europe, who represent a marginal share of the whole immigrant population compared to immigrants from the Maghreb and Africa. However, because the standard errors were quite large in the public housing sector due to a low number of observations, we decided to run this alternative test with pooled data.

<sup>39</sup>The results are not displayed in the paper but are available from authors upon request.

characteristics, there is no significant difference between the private and the public housing sectors in terms of existence of relationships. In addition, for individuals reporting having relationships, we built two variables characterizing the quality of the relationship in the direct or in the larger neighborhood, respectively. In this case, simple correlations reveal that individuals living in the public housing sector tend to have worse relationship with their neighbors, if any. This relationship remains significant when we control for neighborhood, households and building characteristics.<sup>40</sup> To summarize, public housing neighborhoods are characterized by an absence of any relationship, and by a decrease in the quality of relationship when they exist. This brings additional support to our assumption that public housing tenants do not choose their neighbors, otherwise we would expect them to report more frequent relationship, of better quality.

## 7.2 Tests on the refusal rate of public housing offers

The previous tests point out the absence of self-selection along ethnic lines among the movers. But self-selection could occur prior to the move. In this case the sample of movers that we observe in the database would be biased. We address this issue by looking at households that have refused a public housing dwelling offer. Actually, a disturbing fact for our assumption is that a non-negligible share of households waiting to be allocated into a public housing unit report to have declined at least one offer. In the Housing Survey, 24.2 percent of households currently living in a public housing dwelling report to have rejected at least one proposal before finding their current place. Besides, 16.5 percent of the households that are still waiting for an offer at the time of the survey - whether they are already living in a public housing dwelling or not- have previously turned down at least one offer. An additional concern is that 47.9 percent of the households in public housing that had declined an offer at the time they were on the waiting list declared that one of the major reasons for this decision was they found the local environment unpleasant.<sup>41</sup> The corresponding figure for the households still waiting for an answer amounts to 57 percent. The answer "unpleasant local environment" is hard to interpret at this stage, since it could refer to diversity as well as the proximity to public transport and infrastructure, lack of green spaces and so on.

Yet, we show that even if households declined at least one offer, possibly due to the ethnic diversity of the neighborhood, they were still unable to choose the level of diversity of the area in which they end up living, and would not be able to do so for any neighborhood to which they would receive an allocation in the future. To put it another way, although households may have a distaste for diversity, we find evidence that this is not taken into account in their allocation process by the attribution commissions. In principle, households can decline up to three offers. But due to the strong shortage of public housing dwellings, we find that households that have declined an offer in the past cannot self-select into less diverse neighborhoods in the future. We provide evidence of that fact in what follows.

First, if there were self-selection upon diversity, we should expect households that turned down

---

<sup>40</sup>The precise figures corresponding to the previous results are not reported in the paper but are available from the authors upon request.

<sup>41</sup>The other possible answers were: inconvenient place, rent too expensive, low quality building, and apartment not corresponding to household needs.

proposals before being allocated to their current public housing dwelling to end up living in less diverse neighborhoods. To test this conjecture, we run OLS regressions of a variable indicating whether the household declined at least one offer (during the latest application process) on the level of diversity of the neighborhood in which it now lives.<sup>42</sup> Panel A-I of Table 16 shows various estimates of the effect of ethnic diversity on the probability of having turned down offers. Column 1 shows the correlation without any additional control variables. In Column 2, we control for household characteristics. We add up the characteristics of the housing project in Column 3. Column 4 finally includes neighborhood characteristics and department fixed effects since the allocation of a public housing dwelling takes place at the department level. In each specification, the coefficient on ELF is not significantly different from zero, showing that households having declined offers during their past allocation process do not end up living in neighborhoods with significantly different levels of diversity.

We explore further the validity of this conjecture by focusing on the reasons adumbrated by households for refusing an offer. If public housing residents were to sort themselves on the basis of their (dis)taste for diversity, those who declined "because of the local environment" should now live in significantly less diverse neighborhoods. We thus regress a dummy variable indicating whether an "unpleasant environment" was the reason why the household declined at least one offer (during the past application process) on the level of diversity of its current neighborhood. Panel A-II of Table 16 reports the estimates on the level of diversity, using the same specifications as above. Here again, none of the coefficients is significant. Instead, household characteristics such as the labor market status of the head of household and the size of the household are the only ones that matter in these regressions.

Alternatively, we perform these tests on the subsample of individuals currently waiting for an HLM offer. Panel B of Table 16 shows the regressions of the refusal dummy (B-I) and the "refusal due to unpleasant environment" dummy (B-II), for individuals who are currently applying for public housing on the diversity in their neighborhood. We still control for household, building and neighborhood characteristics. Once again, we find that the ethnic diversity of a block is uncorrelated with households wait-listed for an HLM assignment having turned down offers since the beginning of their request (B-I). This suggests that the current level of diversity in the block does not rush households out of the area, as their propensity to decline an offer is independent of the ELF in the current neighborhood. The high refusal rates of HLM offers do not therefore seem driven by a hope to reduce diversity by waiting.

Let us now focus on individuals who left their previous housing unit because they did not like the environment. In the Housing Survey, 5 percent of households that moved over the past four years mention an unpleasant environment as one of the main reason they moved. In this question, the phrase "unpleasant environment" explicitly refers to troubles such as "noise, lifestyle or insecurity". Again, this could be related to high levels of diversity. If this is true, and if households can actually select the block to which they move, then we expect that those households having moved because

---

<sup>42</sup>In this paper, we always rely on OLS estimations, even when the dependent variables are dummies. Using probit estimates does significantly affect our results.

they disliked their environment ended up living in less diverse neighborhoods than the households that moved for a different reason.

We perform OLS regressions of a variable indicating whether the household left its previous housing due to an unpleasant environment, on the level of diversity of its current neighborhood. Table 17 shows the coefficients on diversity in the specification controlling for household, building and neighborhood characteristics, and including department fixed effects. Column 1 shows the results for households that moved within the private housing market. As expected, households that left their previous housing to escape from an unpleasant environment now live in blocks where the diversity is significantly lower. Column 2 shows that this result does not hold for households that moved within the public housing market. This result suggests once again that in the public housing sector, households do not have control over the diversity of the block to which they are allocated.

A potential concern with the previous result is due to the small sample of observations (only 627 in the public housing case), generating large standard errors. Therefore, we replicate this test on a larger subsample. Instead of focusing on households that have moved within a housing sector, we now concentrate on households having moved into each sector, no matter the sector in which they were living prior to their move.<sup>43</sup> As previously, we see that for households living in the private housing sector, the probability that they left their previous housing due to an unpleasant environment is negatively correlated with the diversity in the current neighborhood (Column 3). However, no such significant relationship shows up for households living in a public housing dwelling (Column 4), and the estimates are now more precise than in Column 2. We can infer from those tests that households tend to self-select in low-diversity neighborhoods in the private housing sector, but are unable to do so in the public housing sector.

## 8 Conclusion

This paper exploits French public housing policy as a natural experiment to identify the causal effect of diversity on well-being, social relationships and the quality of local public goods. The French Housing Survey provides a unique micro level of analysis of social interaction between adjacent neighbors within housing blocks. We provide a detailed analysis of the channels through which diversity operates at the local level while the previous literature focused so far on aggregate outcomes and channels. We use the exogenous allocation of households within public housing with respect to ethnic characteristics in France to address the bias from endogenous residential sorting that reduces the confidence in previous empirical findings on fractionalization. We find that fractionalization has a negative impact on other-regarding preferences, leading to higher neglect and vandalism in the housing commons. Fractionalization also undermines collective action for the improvement of the quality of housing. But in our context, fractionalization has no effect on public safety, diversity being associated with social anomie within the housing blocks rather than violent confrontations among neighbors.

---

<sup>43</sup>To summarize, Columns 1 and 2 report the results for households moving from a housing dwelling in the private (1) and public (2) sectors into a housing dwelling in the same sectors. Columns 3 and 4 report the results for households moving from any housing sector into the private (3) and public (4) sectors.

This natural experiment calls for future research on the specific role of national, local and informal institutions in mitigating or magnifying the effect of ethnic diversity on the provision of public goods. France is a country with a republican tradition that resolutely refuses to reify ethnic identification as a strategy to prevent the ethnification of everyday life. Yet we find a significant negative effect of diversity on local public goods in its public housing sector, comparable to the association found in the US localities where multiculturalist institutions regulate ethnic relations (Putnam, 2007) and in cases where public institutions are weak (Alesina and La Ferrara, 2005). However, on issues of physical security in French public housing, the costs to ethnic diversity disappear. This may be due to the emergence of informal institutions (such as in-group policing as in Fearon and Laitin, 1996) or the supremacy of state-level institutions in which local diversity plays no role in the supply of order. In any event, the results raise a puzzle, to be addressed in future research, on the general power of institutional arguments in overcoming the negative implications of ethnic heterogeneity on the provision of public goods.

Table 1: Public Housing and Private Housing population characteristics (households heads, Housing Survey 2002)

	Public Housing (HLM) (1)	Private Housing (2)	p-val (1)/(2) (3)	Private Rental Housing market (4)	p-val (1)/(4) (5)
<b>Birth Country</b>					
France	78.63	88.34	0.000	86.16	0.000
Portugal	1.72	1.23	0.028	1.05	0.004
Spain	1.12	0.78	0.089	0.68	0.033
Italy	0.72	1.10	0.007	0.51	0.210
Other E.U. country	0.71	1.05	0.091	1.11	0.086
Turkey	1.24	0.30	0.000	0.51	0.000
Other European country	0.74	0.75	0.767	0.89	0.457
Maghreb	11.06	4.14	0.000	5.00	0.000
Other African country	2.66	1.06	0.000	2.17	0.327
Cambodia, Vietnam, Laos	0.69	0.41	0.016	0.51	0.227
Other countries	0.72	0.84	0.381	1.40	0.001
<b>Nationality</b>					
French at birth	82.07	91.53	0.000	88.98	0.000
French by acquisition	5.81	3.72	0.000	3.47	0.000
Portuguese	1.33	1.00	0.090	1.02	0.202
Spanish	0.62	0.31	0.004	0.44	0.265
Italian	0.47	0.43	0.865	0.20	0.012
Other E.U. nationality	0.21	0.63	0.002	0.68	0.00
Turkish	1.01	0.19	0.000	0.41	0.000
Other European nationality	0.33	0.32	0.959	0.53	0.157
Maghrebian	6.34	1.02	0.000	2.10	0.000
Other African nationality	1.50	0.38	0.000	1.03	0.096
Cambodian, Vietnamese, Laotian	0.17	0.07	0.004	0.14	0.354
Other nationalities	0.14	0.40	0.016	0.99	0.000
<b>Employment status</b>					
Employed	58.19	56.11	0.090	63.74	0.000
Unemployed	10.82	4.08	0.000	7.75	0.000
Inactive	30.99	39.81	0.000	28.51	0.004
<b>Level of education</b> (highest diploma obtained, individuals above 25 years old)					
No diploma	28.26	14.85	0.000	14.40	0.000
Lower education	50.62	48.33	0.009	37.38	0.000
Baccalaureate	9.37	12.44	0.000	16.26	0.000
Higher education	11.74	24.38	0.000	31.96	0.000
<b>Socio professional group</b>					
Farmer	0.18	1.96	0.000	0.67	0.000
Craftsman, Shopkeeper	1.50	5.03	0.000	3.99	0.000
Executive or other high position	3.64	13.03	0.000	14.14	0.000
Intermediate occupation	12.01	14.29	0.000	16.50	0.000
Employee	20.18	9.93	0.000	15.44	0.000
(Factory) Worker	31.10	16.02	0.000	20.35	0.000
<b>Age (mean)</b>	47.09	51.71	0.000	41.55	0.000
<b>Annual income (mean)</b>	12,226	18,041	0.000	15,902	0.000

Column 3 reports the p-value from a t-test for the null hypothesis that the mean of a given variable is the same for the public housing (Column 1) and private housing (Column 2) populations. Column 5 reports the p-value from a t-test for the null hypothesis that the mean of a given variable is the same for the public housing population (Column 1) and for the population of tenants in the private housing market (Column 4).



Table 2: Dissatisfaction with housing conditions by income level and ethnic origin in the Public Housing sector

	Dissatisfaction with Housing Conditions					Mean Annual Income (in euros)
	Very satisfying	Satisfying	Average	Insufficient	Very insufficient	
<b>Ethnic origin</b>						
French born	13.8	44.01	28.54	9.9	3.75	12,758
Naturalized French	10.68	43.51	29.01	13.8	2.99	10,459
Other European	10.71	39.58	31.21	13.14	5.35	12,292
Maghrebian	8.42	33.27	34.83	18.21	5.26	8,603
African	7.82	20.77	41.29	25.14	4.99	7,865
Asian	0.00	60.64	11.25	28.11	0.00	12,892
<b>Mean Annual Income</b>	13,300	12,856	11,842	10,288	10,127	

Table 3: Monte-Carlo Test of Random Allocation

% departments without residential sorting relative to households' characteristics

	Public Housing	Total Sample
Household's characteristics		
Nationality from Maghreb/Middle East	80.08	54.36
French Nationality at Birth	70.23	24.89

Note: Comparison between the actual and simulated distributions by ethnic groups shares across public housing blocks (Col. 1) and across the whole sample of housing blocks (Col. 2). Percentage of *départements* where equality is not rejected at the 10 percent level using a Kolmogorov-Smirnov test.

Table 4: Ethnic diversity and dissatisfaction with housing condition

Dependent Variable:	<b>Dissatisfaction with housing condition</b>	
	Coefficient	(Std error)
<b>Ethnic Diversity</b>	0.368**	(0.129)
<b>Household characteristics:</b>		
Gender	-0.018	(0.032)
Age	-0.005***	(0.001)
Level of Education	-0.013*	(0.007)
Income (log)	-0.043	(0.031)
Household size	0.105***	(0.013)
<i>Employment status (ref: Employed)</i>		
Unemployed	0.056	(0.055)
Inactive	-0.049	(0.047)
<i>Nationality (ref: French at birth)</i>		
Naturalized French	-0.047	(0.066)
European	0.040	(0.089)
Maghrebian	-0.097	(0.067)
African	0.109	(0.143)
Asian	-0.623**	(0.311)
Other nationality	0.557	(0.633)
<b>Building characteristics:</b>		
Nb of dwellings (log)	0.020	(0.013)
<i>Date of construction of the building (ref: before 1948)</i>		
1949 ≤ t < 1974	0.017	(0.077)
1975 ≤ t < 1981	-0.094	(0.081)
1982 ≤ t < 1989	-0.109	(0.082)
1990 ≤ t < 1998	-0.423***	(0.085)
1999 ≤ t	-0.751***	(0.166)
<b>Neighborhood characteristics:</b>		
Block unemployment rate	1.003***	(0.176)
Socio-economic background (Tabard)	Yes	
Department Fixed Effects	Yes	
Intercept	4.377***	(0.394)
R-squared	0.128	
Observations	4379	

Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 5: Descriptive statistics for each outcome in the public housing sector

	Mean	std dev	Values
<b>1. Degradation of the common areas due to vandalism</b>			
Damaging the premises	1.637	0.778	1 - 3
Graffiti	0.257	0.437	0 - 1
Garbage on the floor	0.188	0.391	0 - 1
Broken windows	0.136	0.343	0 - 1
Broken doors	0.127	0.333	0 - 1
Broken light bulbs	0.094	0.291	0 - 1
Broken mailboxes	0.154	0.361	0 - 1
Vandalism on the elevator	0.085	0.279	0 - 1
Noise in daytime	1.595	0.748	1 - 3
Noise in night time	1.374	0.627	1 - 3
<b>2. Poor quality of housing due to low maintenance</b>			
Care of the common areas	1.593	0.752	1 - 3
Condition of the outside walls	2.433	0.962	1 - 5
Cold in the apartment	0.175	0.380	0 - 1
Cold due to bad insulation	0.065	0.246	0 - 1
Cold due to breakdown in heating equipment	0.045	0.207	0 - 1
Cold due to poor equipment	0.059	0.236	0 - 1
Quality of soundproofing	1.981	0.823	1 - 3
Breakdown of the elevator	0.155	0.362	0 - 1
Toilet malfunction	0.153	0.360	0 - 1
<b>3. Public Safety</b>			
Robberies	0.095	0.293	0 - 1
Aggressions	0.081	0.273	0 - 1
Burglary (or attempt)	0.041	0.198	0 - 1

Depending on the questions, we have between 4,310 and 5,189 observations

Table 6: Diversity and Public goods: Degradation of the common areas due to vandalism (To be continued)

Outcomes:	Damaging the premises	Graffiti	Garbage on the floor	Broken windows	Broken doors	Broken light bulbs	Broken mailboxes	Vandalism on the elevator	Noise in daytime	Noise in Nighttime
<b>Ethnic Diversity</b>	0.630*** (0.127)	0.313*** (0.063)	0.298*** (0.060)	0.200*** (0.053)	0.151** (0.051)	0.271*** (0.048)	0.330*** (0.058)	0.168*** (0.047)	0.288** (0.106)	0.313*** (0.091)
<b>Household characteristics:</b>										
Gender	0.006 (0.028)	0.006 (0.015)	0.003 (0.013)	0.004 (0.011)	0.007 (0.012)	0.012 (0.010)	0.004 (0.012)	0.016* (0.010)	-0.057** (0.026)	-0.018 (0.022)
Age	-0.000 (0.001)	-0.000 (0.001)	-0.000 (0.001)	-0.000 (0.000)	-0.000 (0.000)	-0.000 (0.000)	-0.001** (0.000)	0.000 (0.000)	-0.005*** (0.001)	-0.004*** (0.001)
Level of Education	-0.001 (0.006)	-0.004 (0.003)	0.004 (0.003)	-0.001 (0.003)	-0.003 (0.003)	-0.000 (0.002)	-0.009*** (0.003)	-0.002 (0.002)	-0.008 (0.005)	-0.014** (0.005)
Income (log)	0.002 (0.028)	-0.014 (0.015)	0.003 (0.014)	0.000 (0.012)	-0.018 (0.012)	-0.006 (0.010)	-0.005 (0.012)	-0.008 (0.009)	-0.011 (0.026)	-0.034 (0.021)
Household size	0.044*** (0.010)	0.015** (0.005)	0.006 (0.005)	0.013** (0.004)	0.009** (0.004)	0.018*** (0.004)	0.014** (0.005)	0.005 (0.004)	0.031*** (0.009)	0.016** (0.008)
<i>Employment status (ref: Employed)</i>										
Unemployed	0.047 (0.045)	0.017 (0.023)	0.012 (0.021)	0.021 (0.020)	-0.009 (0.019)	-0.013 (0.016)	0.001 (0.020)	-0.011 (0.015)	-0.043 (0.039)	-0.023 (0.034)
Inactive	-0.044 (0.039)	-0.031 (0.020)	-0.020 (0.018)	-0.011 (0.017)	-0.028* (0.017)	0.012 (0.015)	0.013 (0.018)	-0.022* (0.013)	0.043 (0.037)	0.047 (0.032)
<i>Nationality (ref: French at birth)</i>										
Naturalized French	-0.088 (0.057)	-0.026 (0.030)	-0.037 (0.026)	0.007 (0.027)	0.001 (0.024)	-0.001 (0.022)	0.009 (0.027)	0.012 (0.024)	-0.019 (0.055)	-0.051 (0.044)
European	0.003 (0.078)	-0.049 (0.039)	-0.027 (0.035)	0.060* (0.035)	0.032 (0.037)	-0.009 (0.029)	0.041 (0.037)	-0.037 (0.024)	-0.024 (0.065)	0.078 (0.063)
Maghrebian	-0.179*** (0.054)	-0.054* (0.030)	-0.072** (0.026)	-0.048** (0.024)	-0.058** (0.025)	-0.062** (0.022)	-0.057** (0.027)	-0.063** (0.019)	-0.073 (0.049)	-0.076* (0.043)
African	0.019 (0.112)	-0.063 (0.064)	-0.041 (0.059)	-0.027 (0.054)	0.037 (0.059)	-0.027 (0.052)	-0.029 (0.057)	0.006 (0.051)	0.102 (0.100)	0.056 (0.090)
Asian	-0.111 (0.231)	-0.009 (0.156)	-0.258** (0.121)	0.228 (0.156)	0.002 (0.148)	0.025 (0.158)	0.134 (0.151)	0.305* (0.167)	0.070 (0.231)	0.107 (0.194)
Other nationality	-0.311 (0.306)	-0.338*** (0.073)	-0.235** (0.084)	0.023 (0.152)	0.027 (0.161)	0.093 (0.158)	0.039 (0.155)	-0.086** (0.039)	0.233 (0.306)	0.552 (0.391)

Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 6: Diversity and Public goods: Degradation of the common areas due to vandalism (Continued)

Outcomes:	Damaging the premises	Graffiti	Garbage on the floor	Broken windows	Broken doors	Broken light bulbs	Broken mailboxes	Vandalism on the elevator	Noise in daytime	Noise in Nightime
<b>Building characteristics:</b>										
Nb of dwellings (log)	0.105*** (0.018)	0.068*** (0.006)	0.055*** (0.005)	0.036*** (0.005)	0.037*** (0.004)	0.029*** (0.004)	0.047*** (0.004)	0.040*** (0.004)	.027** (0.010)	0.025** (0.008)
<i>Date of construction of the building (ref: before 1948)</i>										
1949 ≤ t < 1974	0.253*** (0.061)	0.115*** (0.025)	0.081*** (0.022)	0.054** (0.020)	0.058** (0.021)	0.017 (0.015)	0.047** (0.020)	0.023 (0.017)	0.041 (0.059)	-0.024 (0.051)
1975 ≤ t < 1981	0.277*** (0.070)	0.118*** (0.028)	0.088*** (0.026)	0.053** (0.023)	0.049** (0.024)	0.034* (0.018)	0.043** (0.022)	0.038** (0.019)	-0.079 (0.062)	-0.059 (0.053)
1982 ≤ t < 1989	0.357*** (0.073)	0.134*** (0.029)	0.081** (0.026)	0.060** (0.023)	0.061** (0.025)	0.051** (0.020)	0.087*** (0.024)	0.053** (0.021)	-0.106* (0.063)	-0.048 (0.055)
1990 ≤ t < 1998	0.261*** (0.071)	0.093*** (0.027)	0.066** (0.025)	0.038* (0.021)	0.047** (0.023)	0.009 (0.017)	0.049** (0.021)	0.025 (0.018)	-0.171** (0.063)	-0.105* (0.055)
1999 ≤ t	0.210 (0.154)	0.073 (0.067)	0.039 (0.075)	0.113 (0.082)	0.071 (0.067)	-0.037 (0.028)	-0.039 (0.034)	-0.034 (0.030)	-0.321** (0.140)	-0.341*** (0.066)
<b>Neighborhood characteristics:</b>										
Block unempl. rate	1.063*** (0.166)	0.351*** (0.068)	0.415*** (0.072)	0.207*** (0.058)	0.230*** (0.055)	0.263*** (0.045)	0.294*** (0.063)	0.100* (0.056)	0.478*** (0.129)	0.414*** (0.112)
Socio-economic background (Tabard)	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Department	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Fixed Effects										
Intercept	0.588** (0.298)	-0.034 (0.155)	-0.186 (0.140)	-0.132 (0.126)	0.085 (0.128)	-0.095 (0.105)	-0.096 (0.128)	-0.100 (0.093)	1.899*** (0.269)	1.852*** (0.222)
R-squared	0.136	0.157	0.134	0.091	0.072	0.086	0.129	0.093	0.091	0.081
Observations	3661	4379	4379	4379	4379	4379	4379	4379	4379	4379

Each column corresponds to a different regression, for each dependent variable, as reported in the first line. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 7: Diversity and Public goods: Poor Quality of Housing due to low Maintenance (To be continued)

Outcomes:	Care of the common areas	Condition of the outside walls	Cold in the apartment	Cold due to bad insulation	Cold due to breakdown of heating equipment	Cold due to poor heating equipment	Quality of soundproofing	Breakdown of the elevator	Toilet malfunction
<b>Ethnic Diversity</b>	0.384** (0.121)	0.260* (0.145)	0.136** (0.059)	0.021 (0.040)	0.007 (0.037)	0.084** (0.036)	0.021 (0.110)	0.106** (0.051)	0.133** (0.051)
<b>Household characteristics:</b>									
Gender	-0.027 (0.028)	-0.030 (0.032)	-0.029** (0.014)	-0.013 (0.009)	-0.019** (0.008)	-0.022** (0.009)	-0.009 (0.028)	-0.002 (0.012)	-0.017 (0.013)
Age	-0.003** (0.001)	-0.003** (0.001)	-0.000 (0.001)	-0.000 (0.000)	-0.000 (0.000)	-0.000 (0.000)	-0.006*** (0.001)	0.001* (0.000)	-0.001** (0.000)
Level of Education	-0.000 (0.006)	-0.010 (0.007)	-0.000 (0.003)	-0.000 (0.002)	0.001 (0.002)	-0.001 (0.002)	-0.007 (0.006)	-0.000 (0.003)	0.001 (0.003)
Income (log)	0.051* (0.029)	0.024 (0.032)	-0.009 (0.014)	-0.004 (0.008)	0.009 (0.008)	-0.010 (0.008)	0.004 (0.029)	0.011 (0.012)	-0.022 (0.013)
Household size	0.051*** (0.011)	0.041*** (0.011)	0.013** (0.005)	0.009** (0.003)	0.004 (0.003)	0.006* (0.003)	0.027** (0.010)	0.013** (0.004)	0.020*** (0.005)
<i>Employment status (ref: Employed)</i>									
Unemployed	0.041 (0.044)	-0.030 (0.049)	0.008 (0.022)	0.006 (0.014)	-0.003 (0.011)	0.012 (0.015)	-0.019 (0.042)	0.000 (0.019)	0.021 (0.021)
Inactive	0.026 (0.039)	0.026 (0.043)	-0.015 (0.019)	0.005 (0.013)	-0.001 (0.010)	-0.004 (0.012)	-0.011 (0.038)	-0.014 (0.016)	0.004 (0.019)
<i>Nationality (ref: French at birth)</i>									
Naturalized French	-0.060 (0.056)	-0.075 (0.059)	0.046 (0.029)	0.010 (0.018)	0.016 (0.017)	0.020 (0.020)	-0.027 (0.060)	-0.047** (0.023)	-0.033 (0.025)
European	0.155** (0.079)	-0.094 (0.087)	-0.022 (0.036)	0.007 (0.025)	-0.002 (0.020)	0.006 (0.024)	0.037 (0.072)	-0.026 (0.034)	-0.021 (0.033)
Maghrebian	-0.080 (0.054)	-0.247*** (0.061)	0.068** (0.029)	0.002 (0.020)	0.016 (0.016)	0.051** (0.021)	-0.170** (0.052)	0.002 (0.025)	-0.061** (0.025)
African	0.122 (0.101)	-0.033 (0.116)	0.171** (0.063)	-0.001 (0.040)	0.055 (0.038)	0.035 (0.043)	0.068 (0.094)	0.016 (0.056)	0.071 (0.059)
Asian	0.162 (0.343)	0.033 (0.301)	-0.241*** (0.033)	-0.104*** (0.021)	-0.072*** (0.021)	-0.094*** (0.020)	0.172 (0.223)	0.088 (0.147)	0.057 (0.143)
Other nationality	0.177 (0.329)	0.487 (0.393)	-0.164*** (0.041)	-0.082*** (0.022)	-0.030 (0.027)	-0.050** (0.018)	0.038 (0.440)	-0.143** (0.060)	0.104 (0.185)

Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 7: Diversity and Public goods: Poor Quality of Housing due to low Maintenance (Continued)

Outcomes:	Care of the common areas	Condition of the outside walls	Cold in the apartment	Cold due to bad insulation	Cold due to breakdown of heating equipment	Cold due to poor heating equipment	Quality of soundproofing	Breakdown of the elevator	Toilet malfunction
<b>Building characteristics:</b>									
Nb of dwellings (log)	-0.026 (0.016)	-0.025* (0.015)	0.003 (0.005)	0.001 (0.003)	-0.000 (0.003)	-0.001 (0.003)	0.042*** (0.011)	0.069*** (0.005)	-0.007 (0.005)
<i>Date of construction of the building (ref: before 1948)</i>									
1949 ≤ t < 1974	0.049 (0.070)	0.041 (0.077)	0.017 (0.027)	-0.003 (0.019)	0.004 (0.016)	0.029** (0.014)	-0.004 (0.063)	0.028 (0.022)	-0.004 (0.027)
1975 ≤ t < 1981	-0.005 (0.075)	0.138* (0.081)	0.046 (0.029)	0.013 (0.020)	0.006 (0.017)	0.036** (0.016)	-0.161** (0.068)	0.097*** (0.025)	0.007 (0.029)
1982 ≤ t < 1989	0.030 (0.079)	0.048 (0.084)	0.004 (0.030)	-0.011 (0.021)	0.002 (0.018)	0.022 (0.016)	-0.388*** (0.067)	0.061** (0.026)	-0.050* (0.030)
1990 ≤ t < 1998	-0.055 (0.080)	-0.439*** (0.082)	-0.050* (0.029)	-0.041** (0.020)	-0.001 (0.018)	-0.003 (0.014)	-0.574*** (0.068)	0.046* (0.026)	-0.089** (0.030)
1999 ≤ t	0.039 (0.162)	-1.055*** (0.152)	-0.078 (0.057)	-0.071** (0.024)	0.026 (0.049)	-0.009 (0.020)	-0.849*** (0.150)	-0.081** (0.039)	-0.009 (0.094)
<b>Neighborhood characteristics:</b>									
Block unempl. rate	1.044*** (0.158)	0.827*** (0.166)	0.216*** (0.065)	0.104** (0.042)	0.054 (0.046)	0.063* (0.038)	0.641*** (0.126)	0.074 (0.057)	0.084 (0.060)
Socio-economic background (Tabard)	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Department	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Fixed Effects									
Intercept	0.812** (0.296)	2.321*** (0.330)	0.226 (0.142)	0.047 (0.084)	0.033 (0.083)	0.074 (0.077)	1.939*** (0.296)	-0.312** (0.127)	0.332** (0.138)
R-squared	0.085	0.097	0.040	0.018	0.008	0.023	0.132	0.151	0.040
Observations	3661	4379	4379	4379	4379	4379	4379	4379	4379

Each column corresponds to a different regression, for each dependent variable, as reported in the first line. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 8: Diversity and Public goods: Public safety in the neighborhood (To be continued)

Outcomes:	Robberies	Aggressions	Burglary (or attempt)
<b>Ethnic Diversity</b>	0.043 (0.039)	-0.024 (0.038)	-0.001 (0.027)
<b>Household characteristics:</b>			
Gender	-0.002 (0.010)	-0.012 (0.011)	-0.010 (0.007)
Age	-0.000 (0.000)	-0.001* (0.000)	-0.000 (0.000)
Level of Education	0.005** (0.002)	0.006** (0.002)	-0.001 (0.002)
Income (log)	-0.004 (0.010)	-0.005 (0.010)	-0.004 (0.006)
Household size	0.010** (0.004)	0.012** (0.004)	-0.001 (0.003)
<i>Employment status (ref: Employed)</i>			
Unemployed	0.008 (0.016)	0.037** (0.017)	0.009 (0.011)
Inactive	-0.002 (0.014)	0.008 (0.015)	0.008 (0.010)
<i>Nationality (ref: French at birth)</i>			
Naturalized French	-0.007 (0.019)	-0.024 (0.018)	0.024 (0.017)
European	-0.025 (0.026)	-0.018 (0.024)	0.025 (0.021)
Maghrebian	-0.043** (0.019)	-0.034* (0.020)	0.008 (0.015)
African	-0.029 (0.041)	-0.056 (0.037)	0.037 (0.040)
Asian	0.005 (0.129)	0.016 (0.113)	-0.036** (0.014)
Other nationality	-0.075** (0.027)	-0.106*** (0.031)	-0.051** (0.018)

Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1



Table 8: Diversity and Public goods: Public safety in the neighborhood (Continued)

Outcomes:	Robberies	Aggressions	Burglary (or attempt)
<b>Building characteristics:</b>			
Nb of dwellings (log)	-0.000 (0.004)	0.004 (0.004)	-0.002 (0.003)
<i>Date of construction of the building (ref: before 1948)</i>			
1949 ≤ t < 1974	0.050** (0.019)	0.034* (0.020)	0.012 (0.015)
1975 ≤ t < 1981	0.071** (0.022)	0.020 (0.021)	0.019 (0.015)
1982 ≤ t < 1989	0.071*** (0.021)	0.047** (0.022)	0.011 (0.015)
1990 ≤ t < 1998	0.071** (0.022)	0.022 (0.022)	0.011 (0.015)
1999 ≤ t	0.027 (0.052)	-0.033 (0.025)	0.076 (0.064)
<b>Neighborhood characteristics:</b>			
Block unempl. rate	0.019 (0.046)	0.200*** (0.051)	0.042 (0.031)
Socio-economic background (Tabard)	Yes	Yes	Yes
Department	Yes	Yes	Yes
Fixed Effects			
Intercept	0.007 (0.097)	0.006 (0.098)	0.092 (0.069)
R-squared	0.020	0.024	0.002
Observations	4379	4379	4379

Each column corresponds to a different regression, for each dependent variable, as reported in the first line. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 9: Type of repairs done

	<i>Vandalism</i>	<i>Poor Housing Quality</i>	<i>Public Safety</i>
	Work in the commons: staircase, doors, lights, glass	Major works: façade, heating, elevator, toilets	Security Equipment: entry code, locks
	(1)	(2)	(3)
Ethnic Diversity	0.134** (0.054)	-0.213*** (0.069)	0.141*** (0.041)
R-squared	0.024	0.012	0.010
N	2220	2220	2220

## Correlation (in %) with perception of degradations

**1. Vandalism**

Damaging the premises	5.79***	
Graffiti	16.94***	
Garbage on the floor	14.35***	
Broken windows	11.86***	
Broken doors	13.74***	
Broken light bulbs	12.24***	
Broken mailboxes	13.10***	
Vandalism on the elevator	13.54***	12.72***

**2. Poor Housing Quality**

Condition of the outside walls		-3.46***
Cold in the apartment		3.89***
Cold due to bad insulation		1.29
Cold due to breakdown in heating equipment		4.70***
Cold due to poor equipment		0.012
Breakdown of the elevator		-0.001
Toilet malfunction		4.31***

**3. Public Safety**

Robberies			2.52***
Aggressions			4.15***
Burglary (or attempt)			2.35***

Robust standard errors adjusted for block clustering are in parentheses

\*\*\* p&lt;0.01, \*\* p&lt;0.05, \* p&lt;0.1

Table 10: Are results driven by some major ethnic groups disliking being around foreigners ? (to be continued)

	Dissatisfaction with Housing Conditions		
	(1)	(2)	(3)
<b>Ethnic Diversity</b>	0.368** (0.129)	0.359** (0.141)	0.313** (0.154)
<b>Household characteristics:</b>			
Gender	-0.018 (0.032)	-0.018 (0.033)	-0.018 (0.032)
Age	-0.005*** (0.001)	-0.005*** (0.001)	-0.005*** (0.001)
Education	-0.013* (0.007)	-0.013* (0.007)	-0.012* (0.007)
Income (log)	-0.043 (0.031)	-0.045 (0.031)	-0.038 (0.031)
Household size	0.105*** (0.013)	0.105*** (0.013)	0.094*** (0.015)
<i>Employment status(ref: Employed)</i>			
Unemployed	0.056 (0.055)	0.057 (0.055)	0.055 (0.055)
Inactive	-0.049 (0.047)	-0.049 (0.047)	-0.051 (0.047)
<i>Nationality (ref: French at birth)</i>			
Naturalized French	-0.047 (0.066)	-0.030 (0.144)	
European	0.040 (0.089)	0.043 (0.197)	
Maghrebian	-0.097 (0.067)	-0.095 (0.186)	
African	0.109 (0.143)	-0.239 (0.324)	
Asian	-0.623** (0.311)	-0.150 (0.814)	
Other nationality	0.557 (0.633)	0.788 (1.203)	
<i>Interaction terms: ELF * origin</i>			
ELF * naturalized French		-0.052 (0.373)	
ELF * European		-0.007 (0.542)	
ELF * Maghrebian		-0.001 (0.401)	
ELF * African		0.894 (0.751)	
ELF * Asian		-1.123 (1.340)	
ELF * Other nationality		-1.136 (4.552)	

Each coefficient is estimated from a separate regression, according to equation 2. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 10: Continued

Dependent Variable:	<b>Dissatisfaction with housing condition</b>		
	(1)	(2)	(3)
<b><i>Major groups in HLM: Native French and Maghrebians</i></b>			
Native French household			-0.043 (0.056)
ELF * Native French household			0.027 (0.169)
Maghrebian household			0.138 (0.259)
ELF * Maghrebian household			-0.300 (0.559)
<b>Building characteristics:</b>			
Nb of dwellings (log)	0.020 (0.013)	0.020 (0.013)	0.019 (0.013)
<i>Date of construction of the building (ref: before 1948)</i>			
1949 ≤ t < 1974	0.017 (0.077)	0.017 (0.077)	0.022 (0.077)
1975 ≤ t < 1981	-0.094 (0.081)	-0.094 (0.082)	-0.087 (0.082)
1982 ≤ t < 1989	-0.109 (0.082)	-0.109 (0.082)	-0.104 (0.082)
1990 ≤ t < 1998	-0.423*** (0.085)	-0.424*** (0.086)	-0.421*** (0.086)
1999 ≤ t	-0.751*** (0.166)	-0.753*** (0.166)	-0.748*** (0.165)
<b>Neighborhood characteristics:</b>			
Block unemployment rate	1.003*** (0.176)	0.999*** (0.176)	0.995*** (0.175)
Socio-economic background (Tabard)	Yes	Yes	Yes
Department	Yes	Yes	Yes
Fixed Effects			
Intercept	3.991*** (0.360)	4.010*** (0.363)	4.004*** (0.353)
R-squared	0.128	0.127	0.127
Observations	4379	4379	4379

Each coefficient is estimated from a separate regression, according to equation 2. Robust standard errors adjusted for block clustering are in parentheses

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 11: Variation in Perception of Housing quality: Within and Between Public Housing Blocks

	Dissatisfaction with Housing conditions	
	Within correlation	Between correlation
Gender	0.01 (0.04)	-0.03 (0.04)
Age	-0.00* (0.00)	-0.01*** (0.00)
Education	0.01 (0.01)	-0.03*** (0.01)
Income (log)	-0.07* (0.04)	-0.09** (0.04)
Unemployed	0.05 (0.07)	0.08 (0.06)
Inactive	-0.08 (0.06)	0.02 (0.06)
Household size	0.09*** (0.02)	0.13*** (0.01)
Naturalized French	-0.11 (0.08)	-0.00 (0.09)
European	-0.00 (0.11)	0.19* (0.11)
Maghrebien	-0.09 (0.09)	-0.01 (0.08)
African	-0.05 (0.15)	0.44** (0.20)
Asian	-0.34 (0.52)	-0.56 (0.53)
Other nationality	-0.11 (0.79)	0.14 (0.53)
Building size (log)		0.05*** (0.01)
Housing Project FE	Yes	No
R-squared	0.056	0.172
Observations	5105	5105

Robust standard errors adjusted for block clustering are in parentheses

\*\*\* p&lt;0.01, \*\* p&lt;0.05, \* p&lt;0.1

Table 12: Ethnic Diversity and Distant public goods

	Maintenance of streets	Accessibility to public transports	Accessibility to private transports	
Ethnic Diversity	0.158 (0.096)	-0.050 (0.134)	0.142 (0.102)	
Gender	0.001 (0.022)	-0.036 (0.028)	0.015 (0.024)	
Age	-0.000 (0.001)	-0.002** (0.001)	-0.003** (0.001)	
Education	-0.004 (0.005)	0.001 (0.006)	0.006 (0.005)	
Income (log)	0.013 (0.024)	0.010 (0.030)	0.038* (0.023)	
Unemployed	0.012 (0.036)	0.059 (0.041)	-0.027 (0.037)	
Inactive	0.039 (0.032)	0.104** (0.039)	0.008 (0.032)	
Household size	0.018** (0.008)	0.002 (0.010)	0.011 (0.009)	
Naturalized French	-0.096** (0.045)	-0.009 (0.055)	0.030 (0.048)	Robust
European	0.040 (0.070)	0.023 (0.067)	-0.099* (0.060)	
Maghrebian	-0.057 (0.045)	-0.005 (0.051)	-0.069 (0.046)	
African	0.037 (0.095)	0.119 (0.097)	-0.032 (0.101)	
Asian	-0.105 (0.183)	-0.128 (0.282)	0.176 (0.246)	
Other nationality	-0.330*** (0.073)	0.112 (0.343)	-0.376*** (0.083)	
Block unemployment rate	0.386*** (0.114)	-0.102 (0.144)	-0.041 (0.112)	
Intercept	0.814** (0.276)	4.365*** (0.471)	0.783*** (0.233)	
Department Fixed Effects	Yes	Yes	Yes	
Socio economic backgrounds	Yes	Yes	Yes	
R-squared	0.053	0.283	0.101	
N	4451	4451	4451	

standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 13: Ethnic Diversity and Dissatisfaction with Housing conditions – Robustness Checks (to be continued)

	Dissatisfaction with Housing Conditions		
	(1)	(2)	(3)
Ethnic Diversity	1.392** (0.530)		
<b>Household characteristics:</b>			
Gender	-0.018 (0.032)	-0.019 (0.032)	-0.021 (0.032)
Age	-0.005*** (0.001)	-0.005*** (0.001)	-0.005*** (0.001)
Education	-0.013* (0.007)	-0.014* (0.007)	-0.014* (0.007)
Income (log)	-0.041 (0.031)	-0.041 (0.031)	-0.040 (0.031)
Household size	0.105*** (0.013)	0.104*** (0.013)	0.105*** (0.013)
<i>Employment status (ref: Employed)</i>			
Unemployed	0.050 (0.055)	0.059 (0.055)	0.051 (0.055)
Inactive	-0.050 (0.047)	-0.049 (0.047)	-0.051 (0.047)
<b>Nationality (ref: French at birth)</b>			
Naturalized French	-0.044 (0.066)	-0.043 (0.066)	-0.036 (0.066)
European	0.059 (0.088)	0.058 (0.089)	0.062 (0.089)
Maghrebian	-0.107 (0.068)	-0.106 (0.068)	-0.108 (0.068)
African	0.092 (0.143)	0.098 (0.143)	0.097 (0.143)
Asian	-0.615* (0.316)	-0.607* (0.318)	-0.605* (0.315)
Other nationality	0.577 (0.630)	0.565 (0.632)	0.569 (0.634)
<b>Building characteristics:</b>			
Nb of dwellings (log)	0.017 (0.013)	0.020 (0.013)	0.019 (0.013)
<i>Date of construction of the building (ref: before 1948)</i>			
1949 ≤ t < 1974	0.007 (0.077)	0.022 (0.077)	0.012 (0.076)
1975 ≤ t < 1981	-0.115 (0.082)	-0.101 (0.081)	-0.110 (0.081)
1982 ≤ t < 1989	-0.134 (0.082)	-0.122 (0.082)	-0.129 (0.082)
1990 ≤ t < 1998	-0.444*** (0.085)	-0.438*** (0.086)	-0.439*** (0.085)
1999 ≤ t	-0.769*** (0.168)	-0.770*** (0.168)	-0.751*** (0.170)

Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 13: Continued

Dependent Variable:	<b>Dissatisfaction with housing condition</b>		
	(1)	(2)	(3)
<b>Neighborhood characteristics:</b>			
Block unemployment rate	0.909*** (0.175)	0.954*** (0.175)	0.911*** (0.174)
Socio-economic background (Tabard)	Yes	Yes	Yes
Department	Yes	Yes	Yes
Fixed Effects			
<b><i>Ethnic group shares:</i></b>			
% European	-3.249** (1.007)	-0.928* (0.493)	-2.273** (0.961)
% Maghrebian	-1.092 (0.707)	0.633** (0.277)	2.014*** (0.521)
% African	-1.203 (1.101)	0.941 (0.823)	1.218 (1.428)
% Asian	-2.186* (1.323)	0.167 (0.974)	0.315 (1.838)
% Other nationality	-1.488* (0.855)	0.260 (0.501)	-0.263 (0.864)
<b><i>Squared ethnic group shares:</i></b>			
(% European) <sup>2</sup>			7.587* (4.332)
(% Maghrebian) <sup>2</sup>			-3.767** (1.220)
(% African) <sup>2</sup>			-4.225 (7.839)
(% Asian) <sup>2</sup>			-4.792 (14.523)
(% Other nationality) <sup>2</sup>			1.351 (2.895)
Intercept	4.009*** (0.361)	4.071*** (0.362)	4.023*** (0.363)
R-squared	0.131	0.129	0.131
Observations	4379	4379	4379

Robust standard errors adjusted for block clustering are in parentheses

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1



Table 14: A proxy for language fractionalization

Dependent Variable:	Dissatisfaction with housing conditions (1)	Neglect of the commons (2)	Quality of housing (3)	Insecurity (4)
<i>Panel A:</i>				
Diversity based on nationality at birth	0.368*** (0.129)	1.532*** (0.422)	0.727*** (0.263)	0.0252 (0.183)
Observations	4,379	1,693	3,797	4,379
R-squared	0.156	0.201	0.192	0.063
<i>Panel B:</i>				
Diversity based on share of the population speaking French in country of origin	0.366*** (0.141)	1.560*** (0.472)	0.741** (0.290)	-0.0377 (0.199)
Observations	4,365	1,689	3,788	4,365
R-squared	0.157	0.202	0.193	0.064

Standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1 Controls: head of household characteristics (gender, age, education, activity status, aggregated nationality), household characteristics (log of income, household size), building characteristics (date of construction, log number of housing units), socio-economic background of the neighborhood (unemployment rate, Nicole Tabbarð classification), department fixed-effects. Each coefficient comes from a separate regression. The columns indicate the four dependent variables under study. Panel A and panel B respectively correspond to the measure of diversity used in each regression.

Table 15: Correlation between new inhabitants' nationality and share of the area population of the same nationality

Dep Var: Share of population of the same ethnic group as new movers in a given area		
	(1)	(2)
Share of the department population of the the same ethnic group		0.890*** (0.034)
<b>Nationality</b> (reference group: Other nationalities)		
Native	0.816*** (0.010)	0.067** (0.030)
Naturalized French	-0.044*** (0.012)	-0.037** (0.012)
European	-0.022* (0.012)	-0.007 (0.011)
Maghrebian	0.001 (0.014)	0.007 (0.013)
African	-0.025 (0.015)	-0.016 (0.015)
Asian	-0.085* (0.049)	-0.019 (0.047)
<b>Public Housing (HLM)</b>	-0.158** (0.068)	-0.184** (0.066)
<b>Nationality * HLM</b>		
HLM * Native	-0.038 (0.034)	0.040 (0.033)
HLM * Naturalized	0.095** (0.037)	0.051 (0.036)
HLM * European	-0.018 (0.038)	0.010 (0.037)
HLM * Maghrebian	0.055 (0.037)	0.024 (0.036)
HLM * African	0.046 (0.040)	0.017 (0.039)
HLM * Asian	0.083 (0.080)	0.027 (0.078)
<b>Household characteristics</b>		
Age	0.002** (0.001)	0.002** (0.001)
Age squared	-0.000** (0.000)	-0.000** (0.000)
Hourly wage (log)	-0.001 (0.002)	-0.001 (0.002)
Education	Yes	Yes
HLM * Education	Yes	Yes
Socio-Econ Category	Yes	Yes
HLM * SEC	Yes	Yes
Department Fixed Effect	Yes	Yes
Intercept	0.085** (0.036)	0.070** (0.035)
R-squared	0.856	0.864
N	11519	11519

\* p<0.10, \*\* p<0.05, \*\*\* p<0.001

Table 16: Rejection of HLM offers and Ethnic diversity

Rows: Dependent Variables	Coefficient associated with Ethnic Diversity			
	(1)	(2)	(3)	(4)
<b>Panel A:</b> Sample of households who currently live in public housing:				
I. Probability of having declined at least one HLM offer during the previous application process	0.058 (0.058)	0.069 (0.063)	0.017 (0.067)	0.123 (0.0886)
N	1,779	1,779	1,748	1,744
R <sup>2</sup>	0.001	0.021	0.023	0.089
II. Probability that the reason for having declined an HLM offer during the previous application was "unpleasant environment"	0.162 (0.144)	0.061 (0.158)	0.017 (0.171)	-0.0310 (0.258)
N	417	417	415	414
R <sup>2</sup>	0.003	0.035	0.050	0.308
<b>Panel B:</b> Sample of households who are currently applying to public housing:				
I. Probability of having declined at least one HLM offer during the current application process	-0.063 (0.057)	-0.043 (0.064)	-0.088 (0.071)	-0.116 (0.103)
N	1,192	1,192	1,173	1,171
R <sup>2</sup>	0.001	0.014	0.024	0.121
II. Probability that the reason for having declined an HLM offer during the current application was "unpleasant environment"	0.004 (0.194)	-0.007 (0.237)	-0.104 (0.250)	-0.122 (0.506)
N	198	198	195	194
R <sup>2</sup>	0.000	0.083	0.115	0.590

Each of the coefficients is estimated from a separate regression of each of the four dependent variables described in the first column on ethnic diversity. Column 1 does not include any control. Column 2 includes households characteristics (gender, age, education, employment status and nationality of the head of household, total income (in log) of the household per unit of consumption, and household size). Column 3 adds up the characteristics of the building (number of apartments (in log) and construction date). On top of that, column 4 includes neighborhood characteristics (socio-economic background (Tabard index), and local unemployment rate), as well as department fixed effects. In addition, a dummy variable indicating whether the household already lives in the public housing sector is included in specifications 2 to 4 of Panel B. The coefficients for all the controls are available from authors upon request. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 17: Do households having left their previous housing due to an “unpleasant environment” now live in less diverse neighborhoods?

Dependent Variable:	Main reason for leaving previous housing: unpleasant environment (noise, lifestyle or insecurity)			
	Households who moved <b>within</b> the		Households who moved <b>toward</b> the	
	Private	Public	Private	Public
	Housing sector	Housing sector	Housing sector	Housing sector
	(1)	(2)	(3)	(4)
Ethnic Diversity	-0.073** (0.030)	0.083 (0.140)	-0.061* (0.032)	0.016 (0.052)
Observations	5,955	627	6,560	1,793
R-squared	0.030	0.207	0.031	0.079

In each regression, we control for household characteristics (gender, age, education, income (in log), employment status, nationality, household size), building characteristics (number of apartments and construction date), neighborhood characteristics (socio-economic background (Tabard index), and local unemployment rate) and department fixed effects. The coefficients for all the controls are available from authors upon request.

Robust standard errors adjusted for block clustering are in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

## Appendix

### A. Fractionalization index and French blocks

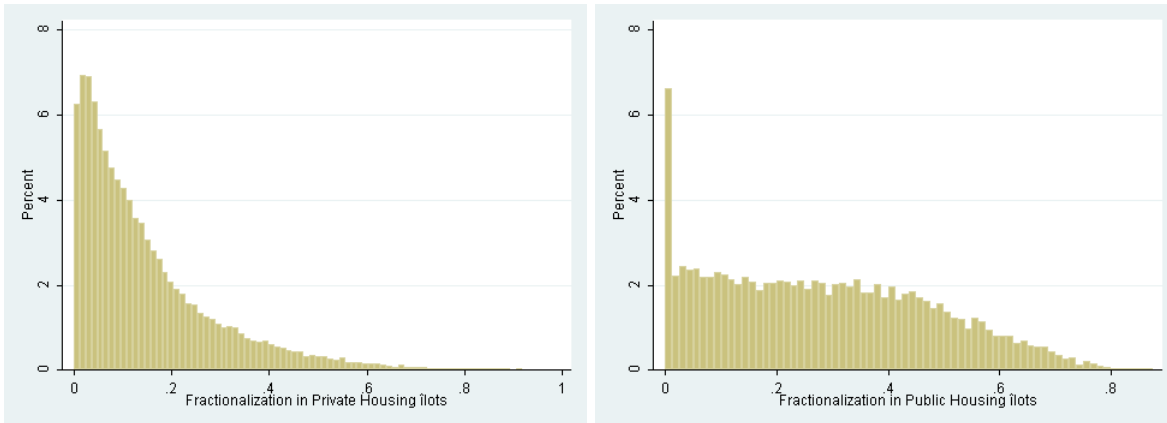
Table 18: Housing blocks in the Census 1999

	Number of blocks sampled per department		Number of individuals sampled per block	
	All France	HLM Population	All France	HLM Population
Mean	2,894.5	932.9	24.6	18.4
Median	2,236	740.5	15	8

Table 19: Fractionalization by nationality at birth in housing blocks

	1999 Census			2002 Housing Survey	
	Whole France	Private Housing	HLM Population	Whole France	HLM Population
Mean	16.65	14.29	27.68	16.23	25.33
Median	11.82	10.29	25.18	11.98	23.37
Std Dev	15.33	13.36	18.75	14.2	17.94
Minimum	0	0	0	0	0
Maximum	91.83	91.83	91.3	84.94	80.26
N	6,643,287	5,027,235	1,616,052	28,744	4,465

Figure 1: Fractionalization by nationality at birth within private and public housing blocks, Census 1999



## B. Principal component analysis

To decompose the various effects of fractionalization, we alternatively run an exploratory analysis to extract the main dimensions with which the various questions reported in the HS correlate the most. We then interpret those factors as different dimensions of the quality of public spaces that could be affected by ethnic diversity. The principal component analysis lets the correlation patterns among the various questions emerge endogenously from the data, rather than grouping them in an arbitrary way. We select (following the Kaiser criterion) three main factors with eigenvalues higher than one that emerged from the principal component analysis of the relevant survey questions. Table 20 reports those three factors and the rotated matrix of correlations between those factors and each question. Three main patterns of correlation emerge that refer broadly speaking to three dimensions of the quality of the public space. Table 5 reports descriptive statistics of the various questions. For each variable, a lower value represents a better outcome (e.g. greater care of the commons, less graffiti, better soundproofing...).

We then create summary indices from the three groups of questions identified in the previous section. We run a principal component analysis on each group of questions, and take the first principal component of each. We refer to those indices as "Neglect of the public areas", "Quality of housing" and "Public Safety". The higher the indices, the more unfavorable are the outcomes. We also check the robustness of the results by looking at alternative summary indices, taking the sum of the questions belonging to each group, or performing a mean effect analysis for each group. The estimates for these alternative indices are reported in Appendix C, yielding similar results.

Table 20: Principal component analysis

	Factor 1	Factor 2	Factor 3
Care of the commons	0.142	0.225	0.053
Voluntary degradations of the commons	0.675	0.172	0.091
Graffiti on the walls	0.209	-0.126	0.023
Trash in the commons	0.247	0.047	0.085
Broken doors in the commons	0.591	0.166	0.123
Broken lights in the commons	0.564	0.072	0.031
Degradation of mail boxes	0.528	0.130	0.034
Broken elevators	0.528	-0.031	-0.041
Quality of the building's façade	-0.038	0.239	-0.030
Problem with heating in the building	-0.007	0.336	0.029
Quality of soundproofing	0.042	0.703	0.004
Noise disturbance during the day in the housing	0.060	0.831	0.052
Noise disturbance at night in the housing	0.113	0.807	0.105
Victim or witness of aggression in the neighborhood	0.098	0.136	0.746
Victim or witness of robbery in the neighborhood	-0.006	0.028	0.810

Table 21 shows the effect of ethnic fractionalization on those three different dimensions: "Neglect of the public areas", "Quality of housing" and "Public Safety".<sup>44</sup> For each index, we run separate regressions on ethnic diversity controlling for the usual household, building and neighborhood characteristics as specified at the bottom of each column in Table 21.

As can be surmised from an examination of three sets of regressions on Table 21, the results relying on categories derived from the principal component analysis rather than the categories that followed from the theoretical literature on public goods, and relying on identical model specifications, are basically similar. For the effect of ethnic diversity on the synthetic index Neglect of Public Areas (see the first panel of Table 21), the effect of ethnic diversity is always statistically significant at the 1 percent level, and is substantively sizeable. For the index of Poor Housing Quality, the coefficient associated with ethnic diversity is statistically significant at the 1 percent level for the first two specifications but only at the 5 percent level when all the controls are included. However, as with the results using the theoretically inferred categorization in the main body of the paper, its effect is much lower than for the index for voluntary degradations. For the index of Public Safety, the data here show that ethnic diversity does not have a statistically significant effect on public safety in the public housing sector in our full specification, as is the case in our main results. In sum, categorization by principal components analysis yields quite similar results as to those reported with the theoretically derived categorization.

<sup>44</sup>The coefficient estimates for the control variables are not reported here but are very similar to those reported in Table 4. The full regression results are available upon request.

Table 21: Ethnic Diversity and Public Goods (public housing)

	Ethnic Diversity		
	(1)	(2)	(3)
<b>Index for Neglect of Public Areas</b>			
	1.791***	1.514***	1.532***
	(0.322)	(0.330)	(0.422)
Observations	1,700	1,693	1,693
R-squared	0.060	0.084	0.134
<b>Index for Poor Quality of Housing</b>			
	2.132***	1.382***	0.727**
	(0.187)	(0.200)	(0.263)
Observations	3,869	3,805	3,797
R-squared	0.092	0.128	0,161
<b>Index for Public Safety</b>			
	0.330**	0.273*	0.025
	(0.129)	(0.139)	(0.183)
Observations	4,464	4,388	4,379
R-squared	0.014	0.014	0.033
Household characteristics	Yes	Yes	Yes
Building characteristics	No	Yes	Yes
Neighborhood characteristics	No	No	Yes
Department fixed effects	No	No	Yes

Each coefficient is estimated from a separate regression, according to equation 2. The four dependent variables considered include the answer to the general opinion / dissatisfaction question and the three indices that were derived from principal component analysis as described in section 4.2. Each index is regressed on either ethnic diversity, controlling for the usual household and neighborhood characteristics unless otherwise indicated. Robust standard errors adjusted for block clustering are in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

### C. Alternative summary indices - Mean effect analysis

As a robustness check, we have also experimented with alternative indexes for measuring these three dimensions of housing conditions. We have first looked at basic indices defined as the sum of the outcome variables related to each dimension. For each of the three dimensions considered, we thus obtain a variable which increases with the number of adverse outcomes one faces. Table 22 reports the results of the regression of these three indices on ethnic diversity for our favorite specification. Our results are robust to these alternative indices: the effect of ethnic diversity is still strongly significant for the index of Housing Quality, and is even stronger for the index for Neglect of Public Areas. As noted previously, there is no effect of diversity on public safety.

Table 22: Diversity and Public goods: sum of the various outcomes, Public Housing

	Ethnic Diversity
	(1)
<b>1. Neglect of the Public Areas</b>	
Sum of the outcome variables	2.321*** (0.575)
<b>2. Quality of Housing</b>	
Sum of the outcome variables	1.511*** (0.360)
<b>3. Public Safety</b>	
Sum of the outcome variables	0.029 (0.059)
Socio-economic	
Background of area	Yes***
Department fixed effects	Yes***

Each coefficient is estimated from a separate regression, according to equation 2. The three dependent variables considered are the three indices reported in bold. Each index is regressed on either ethnic diversity, controlling for the usual household and neighborhood characteristics unless otherwise indicated. Robust standard errors adjusted for block clustering are in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

To be more thorough, we next perform a mean effect analysis. Following Kling, Liebman and Katz (2007), we construct summary indices aggregating information across the various related outcomes for each of the three dimensions studied above. To build the three summary indices, we first normalize each outcome using a pseudo-control group defined by individuals living in blocks characterized by a below-median fractionalization index, as in Glennerster, Miguel and Rothenberg (2010). Let  $Y_k$  be the  $k^{th}$  of  $K$  related outcomes. Each standardized outcome  $Y_k^*$  is obtained by subtracting the mean  $\mu_k$  and dividing by the standard deviation  $\sigma_k$  of the outcome variable among the low diversity pseudo-control group:  $Y_k^* = (Y_k - \mu_k)/\sigma_k$ . We then average the related standardized outcomes to form the summary index :  $Y^* = \sum_k Y_k^*/K$ . Accordingly, our summary index for neglect of the commons averages nine standardized measures including graffiti, broken mailboxes, broken elevator, low care of the commons, voluntary degradations and garbage on the floor; the index for poor housing conditions averages five standardized measures of quality of apartment's soundproofing, of efficiency of the heating system, and of the general state of the outside walls; and finally the public safety indicator averages standardized measures of robbery and personal aggression.

Table 23 presents the raw and normalized components of the three broad summary indices. The first column displays the mean of each outcome among the low-diversity group. The normalized outcomes for this pseudo-control group are displayed in column 2, with mean equal to zero by construction. Column 3 reports the difference between the mean of each outcome among the households living in high diversity neighborhoods (a treated group of sorts) and that of the low diversity population. All but one of the differences are positive indicating that the average outcome is generally worse in more heterogeneous areas. Column 4 shows the difference between the nor-



malized outcomes for treatment and control group, and allows for a more comprehensive reading. For instance, we know from column 3 that the raw difference between care of the commons in low and high diversity areas is of 0.14. Column 4 now tells us that this difference is of 0.19 standard deviations, relative to the control group standard deviation.

Table 24 reports mean effect estimates from regressing the summary indices for negligence, quality of housing and public safety on ethnic diversity and other variables, as in specification 2. The coefficient on ethnic diversity is the mean effect size. As expected, for the negligence index and the quality of housing index, mean effect estimates of ethnic diversity are strongly positive (column 1). Using summary indices also allows us to compare the mean effect of diversity on those two normalized outcomes: lines 1 and 2 of Table 24 tell us that ethnic diversity has a more adverse impact on the neglect of common areas than on the average quality of housing. This gives us an insight concerning the mechanisms at play: high levels of ethnic diversity are more likely to generate uncivic behaviors that could be avoided by higher quality social norms. By contrast, the mean effect estimate in the third line indicates that ethnic diversity plays no role on public safety.

Table 23: Components of Summary Indices,Public Housing

	Low ELF		High ELF - low ELF	
	Raw (1)	Norm (2)	Raw (3)	Norm (4)
<b>1. Neglect of the Public Areas</b>				
Care of the commons	1.49	0	0.14	0.19
Damaging in the premises	1.45	0	0.25	0.36
Graffiti	0.63	0	0.03	0.07
Garbage on the floor	0.42	0	0.08	0.16
Broken glass	0.31	0	0.06	0.13
Broken doors	0.32	0	0.01	0.01
Broken light bulbs	0.18	0	0.07	0.18
Broken mailboxes	0.33	0	0.09	0.19
Broken elevators	0.14	0	0.1	0.28
<b>2. Quality of Housing</b>				
Condition of the outside walls	2.42	0	0.01	0.01
Quality of soundproofing	1.83	0	0.23	0.28
Noisy in daytime	1.48	0	0.16	0.22
Noisy in night time	1.27	0	0.15	0.27
Cold in the apartment	0.14	0	0.08	0.23
<b>3. Public Safety</b>				
Robberies	0.08	0	0.01	0.05
Aggressions	0.06	0	0.02	0.08

Table 24: Diversity and Public goods: mean effect estimates, Public Housing

	Ethnic Diversity
	(1)
<b>1. Neglect of the Public Areas</b>	
Mean effect estimate	0.545*** (0.130)
<b>2. Quality of Housing</b>	
Mean effect estimate	0.467*** (0.107)
<b>3. Public Safety</b>	
Mean effect estimate	0.050 (0.112)
Socio-economic	
Background of area	Yes***
Department fixed effects	Yes***

Each coefficient is estimated from a separate regression, according to equation 2. The three dependent variables considered are the three summary indices reported in bold. Each index is regressed on either ethnic diversity, controlling for the usual household and neighborhood characteristics unless otherwise indicated. Robust standard errors adjusted for block clustering are in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

The broad picture drawn in the three previous sets of regressions is largely confirmed by the regressions of each separate normalized outcome. The corresponding mean effect estimates of ethnic diversity are presented in table 25. Although diversity has no significant impact on a few outcomes, such as broken glass or broken light bulbs in the commons, we still have a very strong negative effect of diversity on every other negligence or housing quality outcome in the public sector. As noted earlier, the mean effects estimates for negligence outcomes are on average larger than those for housing quality. The effect measured on broken light bulbs is the strongest, with a more than one standard deviation difference between low and high diversity neighborhoods, while the lowest is obtained for the quality of sound proofing, with a difference of about one third in terms of its standard deviation. Turning to public safety, the mean effect estimates on robberies and direct aggressions are both insignificant in the public housing sector in our favorite specification.

Table 25: Ethnic diversity and disaggregated housing outcomes : mean effects analysis, Public Housing

	Ethnic diversity	
	(1)	(2)
<b>1. Neglect of the Public Areas</b>		
Care of the commons	0.752*** (0.124)	0.549*** (0.164)
Damaging the premises	1.271*** (0.140)	1.019*** (0.183)
Graffiti	0.387** (0.151)	0.488** (0.216)
Garbage on the floor	0.668*** (0.162)	0.625*** (0.227)
Broken glass	0.475*** (0.182)	0.368 (0.238)
Broken doors	0.228 (0.176)	0.110 (0.241)
Broken light bulbs	0.878*** (0.187)	1.049*** (0.248)
Broken mailboxes	0.652*** (0.176)	0.927*** (0.239)
Broken elevators	0.694*** (0.227)	0.656** (0.288)
<b>2. Quality of Housing</b>		
Condition of the outside walls	0.629*** (0.109)	0.414*** (0.151)
Quality of soundproofing	0.963*** (0.099)	0.393*** (0.138)
Noisy in daytime	0.935*** (0.110)	0.613*** (0.148)
Noisy in night time	1.096*** (0.121)	0.676*** (0.159)
Cold in the apartment	0.634*** (0.128)	0.418** (0.184)
<b>3. Public Safety</b>		
Robberies	0.207** (0.103)	0.149 (0.140)
Aggressions	0.231** (0.104)	-0.0489 (0.149)
Socio eco. background and department. fixed effects	No	Yes

Each entry is the coefficient estimate on ethnic diversity from a separate regression.

All the regressions include controls for household characteristics.

Robust standard errors adjusted for block clustering are in parentheses. The components of the three summary indices are the variables listed below each of them. Descriptive statistics for these outcomes are in Table 5.

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

## References

- [1] Alesina A., R. Baqir and W. Easterly, 1999: "Public Goods and Ethnic Divisions", *Quarterly Journal of Economics*, Vol. 114(4), pp. 1243-1284
- [2] Alesina A., A. Devleeschauwer, W. Easterly, S. Kurlat and R. Wacziarg, 2003: "Fractionalization", *Journal of Economic Growth*, Vol. 8, pp. 155-194
- [3] Alesina, A. and E. La Ferrara, 2000: "Participation in Heterogeneous Communities", *Quarterly Journal of Economics*, Vol. 115(3), pp. 847-904
- [4] Alesina, A. and E. La Ferrara, 2002: "Who trusts others?", *Journal of Public Economics*, Vol. 85(2), pp. 207-234
- [5] Alesina, A. and E. La Ferrara, 2005: "Ethnic Diversity and Economic Performance", *Journal of Economic Literature*, Vol 43(3), pp. 762-800
- [6] Alesina, A. and E. Zhuravskaya, 2011: "Segregation and the Quality of Government in a Cross Section of Countries", *American Economic Review*, Vol. 101, pp. 1872-1911
- [7] Algan, Y., T. Mayer and M. Thoenig, 2011: "The Economic Incentives of Cultural Transmission: Spatial evidence from Naming Patterns Across France", *Sciences Po Working Paper*
- [8] Bayer, P., S. Ross and G. Topa, 2008: "Place of Work and Place of Residence: Informal Hiring Networks and Labor Market Outcomes", *Journal of Political Economy*, Vol. 116(6), pp. 1150-1196
- [9] Bernardot, M., 2008 : "Loger les immigrés", *Bellecombe-en-Bauges, Rhône-Alpes: Editions du Croquant*
- [10] Caselli, F., and Coleman, J., 2012 : "On the Theory of Ethnic conflict", *Journal of the European Economic Association*, forthcoming.
- [11] Coleman, J., 1988: "Social Capital in the Creation of Human Capital", *The American Journal of Sociology*, Vol. 94, supplement "Organizations and Institutions: Sociological and Economic Approaches to the Analysis of Social Structure", pp. S95-S120
- [12] Combes, P.P., B. Decreuse, B. Schmutz and A. Trannoy, 2010: "The Neighbor is King: Customer Discrimination in the Housing Market", *working paper*
- [13] Dahlberg, M., K. Edmark and H. Lundqvist, 2011: "Ethnic diversity and preferences for redistribution", *CESifo working paper* N° 3325
- [14] Dancygier, R., 2010 : "Immigration and Conflict in Europe", *Cambridge University Press*
- [15] Debrand, T. and C. Taffin, 2005: "Les facteurs structurels et conjoncturels de la mobilité résidentielle depuis 20 ans", *Economie et Statistique* N° 381-382.

- [16] Dunning, T. and L. Harrison, 2010: "Cross-Cutting Cleavages and Ethnic Voting: An Experimental Study of Cousinage in Mali.", *American Political Science Review*, Vol. 104(1), pp. 21-39
- [17] Easterly, W. and R. Levine, 1997: "Africa's growth tragedy: policies and ethnic divisions", *Quarterly Journal of Economics*, Vol. 112(4), pp. 1203-1250
- [18] Esteban, J. and Ray, D., 2011: "A model of ethnic conflict" , *Journal of the European Economic Association*
- [19] Fearon, J. D. and D. Laitin, 1996: "Explaining Interethnic Cooperation", *American Political Science Review*, Vol. 90(4), pp. 715-735
- [20] Fearon, J. D. and D. Laitin, 2000: "Violence and the Social Construction of Ethnic Identity", *International Organization*, Vol. 54(4), pp. 845-877
- [21] Glennerster R., E. Miguel and A. Rothenberg, 2010: "Working Together: Collective Action in Diverse Sierra Leone Communities", *Working Paper*
- [22] Goux, D. and E. Maurin, 2007: "Close neighbours matter: neighbourhood effects on early performance at school", *Economic Journal*, Vol. 117, pp. 1-24
- [23] Jacquot, A., 2007: "L'occupation du parc HLM: éclairage à partir des enquêtes logement de l'INSEE", Direction des Statistiques Démographiques et Sociales, Document de Travail N° F0708
- [24] Katz, L. F., J. R. Kling and J. B. Liebman, 2001: "Moving to Opportunity in Boston: Early Results of a Randomized Mobility Experiment," *Quarterly Journal of Economics*, Vol. 116, pp.607-654
- [25] Kling, J. R., J. B. Liebman and L. F. Katz, 2007: "Experimental Analysis of Neighborhood Effects", *Econometrica*, Vol. 75(1), pp. 83-119
- [26] Luttmer, E., 2001: "Group Loyalty and the Taste for Redistribution", *Journal of Political Economy*, Vol. 109(3), pp. 500-528
- [27] Miguel, E., 2004: "Tribe or nation?: Nation building and public goods in Kenya versus Tanzania", *World Politics*, Vol 56(3), pp. 327-362
- [28] Miguel, E. and M. K. Gugerty, 2005: "Ethnic Diversity, Social Sanctions, and Public Goods in Kenya", *Journal of Public Economics*, Vol. 89(11), pp. 2325-2368
- [29] Observatoire du Logement et de l'Habitat de Paris, 2011: "L'accès au logement social à Paris: Analyse de la demande de logement social et bilan des propositions et des attributions aux logements sociaux à Paris en 2010"
- [30] Oreopoulos, P., 2003: "The long run consequences in working in a poor neighborhood", *Quarterly Journal of Economics*, Vol. 118(4), pp. 1533-1575

- [31] Page, S.E., 2007: "The Difference: How the Power of Diversity Creates Better Groups, firms, Schools, and Societies", *Princeton University Press*
- [32] Posner, D., 2004: "The Political Salience of Cultural Difference: Why Chewas and Tumbukas are Allies in Zambia and Adversaries in Malawi", *American Political Science Review*, Vol. 98(4), pp. 529-545
- [33] Putnam, R., 2007: "E Pluribus Unum: Diversity and Community in the twenty-first Century", *Scandinavian Political Studies*, Vol. 30(2), pp. 137-74
- [34] Rouquette, R. and H. Lipietz, 1991: "Droit du Logement Social", *Editions du Moniteur*
- [35] Simon, P., 2003: "Le logement social en France et la gestion des populations à risques", *Hommes et Migrations*, No. 1246, pp. 76-91
- [36] Tabard, N., 2002: "Inégalités et disparités entre les quartiers en 1999", in Géraldine Martin-Houssart, *Données sociales, La société française*.
- [37] Vigdor, J., 2002: "Interpreting ethnic fragmentation effects", *Economic Letters*, Vol. 75, pp. 271-276