# Supplement to "When Kahneman meets Manski: Using dual systems of reasoning to interpret subjective expectations of equity returns" 

by Fabian Gouret and Guillaume Hollard

October 11, 2010

This appendix presents tables of results for the waves 13 and 14 and supplementary materials. The appendix is organized as follows. Section A describes some results for the waves 13 and 14 , similar to those presented for the wave 12 in Section 4.1. Section B describes alternative measures of coherence and the corresponding results. Section C proposes Spearman's rank correlation coefficients between the different measures of coherence to see if the rankings of respondents are similar. Section D proposes Spearman's rank correlation coefficients between $\mu$ and $\sigma$ for the respondents categorized as Coherent 1 and those categorized as Coherent 3. Sections E and F check if there is a relationship between the measures of coherence and individual attributes. The analysis in Section E is realized considering the final sample of 979 respondents (those for whom we can find a unique fitted value for $(\mu, \sigma)$ ). Section $F$ checks if there is a relationship between the measures of coherence and individual attributes considering all the respondents for whom we can compute a measure of coherence.

## A Results for the waves 13 and 14 (respondents ranked according to $d^{98}$ )

Table A1: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $d^{98}$ ) -waves 13 and 14 -

|  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $[1]$ Coherent 1 $\mu$ | $[2]$ Coherent 2 $\mu$ | $[3]$ <br> Coherent 3 <br> $\mu$ | $[4]$ <br> All types <br> $\mu$ | $[5]$ <br> Coherent 1 <br> $\mu$ | $[6]$ Coherent 2 $\mu$ | $[7]$ Coherent 3 $\mu$ | $[8]$ All types $\mu$ |
| Intercept $\sigma$ | $\begin{gathered} 0.07^{* *} \\ (0.03) \\ 1.18^{* * *} \\ (0.07) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.03) \\ 0.76^{* * *} \\ (0.06) \\ \hline \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.03) \\ 0.05 \\ (0.06) \end{gathered}$ | $\begin{gathered} 0.16^{* * *} \\ (0.04) \\ 0.35^{* * *} \\ (0.09) \\ \hline \end{gathered}$ | $\begin{gathered} 0.09^{* *} \\ (0.04) \\ 1.02^{* * *} \\ (0.11) \\ \hline \end{gathered}$ | $\begin{gathered} 0.02 \\ (0.02) \\ 0.70^{* * *} \\ (0.05) \\ \hline \end{gathered}$ | $\begin{gathered} -0.08^{* *} \\ (0.03) \\ 0.15^{* * *} \\ (0.04) \\ \hline \end{gathered}$ | $\begin{gathered} 0.14^{* * *} \\ (0.02) \\ 0.26^{* * *} \\ (0.06) \\ \hline \end{gathered}$ |
| $R^{2}$ <br> Observations (Observations excluded) | $\begin{gathered} 0.72 \\ 80 \\ 8 \\ \hline \end{gathered}$ | $\begin{gathered} 0.71 \\ 81 \\ 7 \\ \hline \end{gathered}$ | $\begin{gathered} 0.01 \\ 80 \\ 8 \\ \hline \end{gathered}$ | $\begin{gathered} 0.15 \\ 241 \\ 23 \\ \hline \end{gathered}$ | $\begin{gathered} 0.51 \\ 117 \\ 8 \\ \hline \end{gathered}$ | $\begin{gathered} 0.64 \\ 118 \\ 7 \\ \hline \end{gathered}$ | $\begin{gathered} 0.13 \\ 114 \\ 11 \\ \hline \end{gathered}$ | $\begin{gathered} 0.10 \\ 349 \\ 26 \\ \hline \end{gathered}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
$N$ iii. To test whether the intercept of the group "Coherent 1 " is equal to the risk-free rate in wave $13\left(R F_{13}\right)$, i.e. $H_{0}: \gamma_{1,13}=R F_{13}$, we have computed a $95 \%$ confidence interval for $\gamma_{1,13}$, and have obtained the following interval: [0.016; 0.137 ]. Our proxy for the $R F_{13}$ is the one-year U.S. LIBOR rate. During the wave 13 (February-May 2000), this interest rate reached a minimum of 0.067 (February 28) and a maximum of 0.075 (May 31); so $R F_{13} \in[0.067 ; 0.075]$. Given the confidence interval obtained for $\gamma_{1,13}$, we cannot reject the null hypothesis $\gamma_{1,13}=R F_{13}$. We have proceeded similarly for the wave 14 . The $95 \%$ confidence interval for $\gamma_{1,14}$ is [0.012; 0.174$]$. During the wave 14 (September 2000-March 2001), $R F_{14}$ reached a minimum of 0.045 (March 22, 2001) and a maximum of 0.069 (September 1 , 2000), so $R F_{14} \in[0.045 ; 0.069]$; hence we cannot reject the null hypothesis.
iv. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.15^{*}+1.15^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.60 \text { and } N=88\right) \\
\mu=0.02+0.71^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.61 \text { and } N=88\right) \\
\mu=0.02+0.06 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.03 \text { and } N=88\right) \\
\mu=0.22^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=0.17^{* * *}+0.78^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.66 \text { and } N=125\right) \\
\mu=-0.09+0.92^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.76 \text { and } N=125\right) \\
\mu=-0.13^{* * *}+0.20^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.21 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$

## B Alternative measures of coherence

This Section presents more details on the alternative measures of coherence discussed in Section 2 of the main text. Part of this Section repeats in the main text to keep this online appendix self-contained.

## B. 1 Alternative measures of coherence: $d_{i}^{90}$ and $d_{i}^{80}$

All the results presented in Section 3 and 4 of the main text are based on the measure of coherence $d^{98}$. This criterion of coherence was based on Assumptions 1-2, and to perform the empirical analysis, we considered $\alpha=0.01, \forall i$. It implied that the probability that $\widetilde{R}_{i}$ lied within the interval $\left[R_{i, \min }, R_{i, \max }\right.$ ] was $98 \%$, whatever the respondent. This subsection presents results based on two measures of coherence similar to $d^{98}$, except that $\alpha=0.05$ for the first measure $\left(d^{90}\right)$ and $\alpha=0.10$ for the second one $\left(d^{80}\right)$. These two measures of coherence are computed at the first threshold, like $d^{98}$.

For each wave, we have ranked the respondents according to these measures, and carried out the same exercise as we did in Section 3-4 of the main text. Tables B1-B2 present the results.

Table B1: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $d^{90}$ ) -waves 12,13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | $\begin{aligned} & 0.06^{*} \\ & (0.03) \end{aligned}$ | $\begin{gathered} 0.01 \\ (0.03) \end{gathered}$ | $\begin{aligned} & -0.04 \\ & (0.03) \end{aligned}$ | $\begin{gathered} 0.20^{* * *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.07 * * \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.07 \\ (0.05) \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.17^{* * *} \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.14^{* * *} \\ (0.04) \end{gathered}$ | $\begin{aligned} & -0.00 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & -0.05 \\ & (0.03) \end{aligned}$ | $\begin{gathered} 0.13^{* * *} \\ (0.03) \end{gathered}$ |
| $\sigma$ | $1.29^{* * *}$ | $0.75^{* * *}$ | $-0.00$ | $0.08$ | $1.12^{* * *}$ | $0.72^{* * *}$ | $0.03$ | $0.29^{* * *}$ | $0.80^{* * *}$ | $0.73^{* * *}$ | $0.14^{* * *}$ | $0.28^{* * *}$ |
|  | $(0.11)$ | $(0.06)$ | (0.02) | (0.06) | $(0.07)$ | $(0.08)$ | $(0.06)$ | $(0.09)$ | $(0.11)$ | $(0.04)$ | $(0.05)$ | (0.07) |
| $R^{2}$ | 0.72 | 0.67 | 0.00 | 0.01 | 0.65 | 0.52 | 0.00 | 0.10 | 0.39 | 0.66 | 0.11 | 0.13 |
| Observations | 101 | 106 | 110 | 317 | 80 | 82 | 81 | 243 | 117 | 118 | 114 | 349 |
| (Observations excluded) | 12 | 8 | 3 | 23 | 8 | 6 | 7 | 21 | 8 | 7 | 11 | 26 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
ii. Heteroskedasticity-robust standard errors are in parentheses.
$\star$ iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.07+1.40^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.75 \text { and } N=113\right) \\
\mu=0.02+0.68^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.59 \text { and } N=114\right) \\
\mu=-0.05+0.03 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.01 \text { and } N=113\right) \\
\mu=0.18^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=340\right) \\
\mu=0.08+1.26^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.65 \text { and } N=88\right) \\
\mu=0.06+0.70^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.56 \text { and } N=88\right) \\
\mu=0.03+0.06 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.04 \text { and } N=88\right) \\
\mu=0.21^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=0.16^{* * *}+0.79^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.60 \text { and } N=125\right) \\
\mu=-0.08^{*}+0.79^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.78 \text { and } N=125\right) \\
\mu=-0.11^{* *}+0.19^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.19 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$

Table B2: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $d^{80}$ ) -waves 12,13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | $\begin{gathered} 0.06^{* *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.01 \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.16^{* * *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.08^{* * *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.10^{* *} \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.16^{* * *} \\ (0.05) \end{gathered}$ | $\begin{gathered} 0.15^{* * *} \\ (0.04) \end{gathered}$ | $\begin{aligned} & -0.02 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & -0.02 \\ & (0.03) \end{aligned}$ | $\begin{aligned} & 0.07^{*} \\ & (0.04) \end{aligned}$ |
| $\sigma$ | $\begin{gathered} 1.18^{* * *} \\ (0.08) \\ \hline \end{gathered}$ | $\begin{gathered} 0.70^{* * *} \\ (0.06) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.07 \\ (0.04) \\ \hline \end{array}$ | $\begin{aligned} & 0.18^{* *} \\ & (0.09) \\ & \hline \end{aligned}$ | $\begin{gathered} 1.01^{* * *} \\ (0.06) \\ \hline \end{gathered}$ | $\begin{gathered} 0.71^{* * *} \\ (0.09) \\ \hline \end{gathered}$ | $\begin{array}{r} 0.043 \\ (0.06) \\ \hline \end{array}$ | $\begin{gathered} 0.32^{* * *} \\ (0.12) \\ \hline \end{gathered}$ | $\begin{gathered} 0.75^{* * *} \\ (0.11) \\ \hline \end{gathered}$ | $\begin{gathered} 0.75^{* * *} \\ (0.03) \\ \hline \end{gathered}$ | $\begin{aligned} & 0.13^{* *} \\ & (0.05) \\ & \hline \end{aligned}$ | $\begin{gathered} 0.40^{* * *} \\ (0.10) \\ \hline \end{gathered}$ |
| $R^{2}$ | 0.73 | 0.59 | 0.02 | 0.04 | 0.76 | 0.48 | 0.00 | 0.11 | 0.44 | 0.76 | 0.08 | 0.25 |
| Observations | 103 | 104 | 109 | 316 | 79 | 82 | 81 | 242 | 116 | 121 | 115 | 352 |
| (Observations excluded) | 10 | 10 | 4 | 24 | 9 | 6 | 7 | 22 | 9 | 4 | 10 | 23 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
cr ii. Heteroskedasticity-robust standard errors are in parentheses.
Cr iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.06+1.32^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.70 \text { and } N=113\right) \\
\mu=0.01+0.65^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.55 \text { and } N=114\right) \\
\mu=-0.00-0.02 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.00 \text { and } N=113\right) \\
\mu=0.18^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=340\right) \\
\mu=0.13^{* *}+1.03^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.68 \text { and } N=88\right) \\
\mu=0.09+0.63^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.50 \text { and } N=88\right) \\
\mu=0.05+0.07 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.04 \text { and } N=88\right) \\
\mu=0.22^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=0.13^{* *}+0.79^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.65 \text { and } N=125\right) \\
\mu=-0.03+0.79^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.61 \text { and } N=125\right) \\
\mu=-0.08^{*}+0.19^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.17 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$

## B. 2 Alternative measure of coherence: $\alpha_{i}$

Previous criteria of coherence are based on Assumptions 1-2. Assumption 1 says that the probability that $\widetilde{R}_{i}$ lies within the interval $\left[R_{i, \min }, R_{i, \max }\right]$ is $(1-2 \alpha)$, and the remaining part of the distribution, $2 \alpha$, is equally affected on the left and on the right of the distribution. Under Assumption 1, $\alpha$ is not individual-specific. An alternative measure of coherence can be obtained by assuming that respondents are in fact coherent, but the value of $\alpha$ varies across individuals. $\alpha_{i}$ is defined as the value that minimizes the distance between the calibrated probability $\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, 1}\right)$ and the reported probability $Q_{i, 1}$.

We thus replace Assumption 1 by Assumption 1]:

## Assumption 1'

$$
\begin{equation*}
\operatorname{Prob}\left(\widetilde{R}_{i}<R_{i, \min }\right)=\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, \max }\right)=\alpha_{i} \tag{B1}
\end{equation*}
$$

where

$$
\begin{equation*}
\alpha_{i}=\arg \inf _{\alpha_{i} \in[0,1]}\left|Q_{i, 1}-\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, 1}\right)\right| \tag{B2}
\end{equation*}
$$

We have to find $\alpha_{i}$ that satisfies Equations B1, B2.
i. Under Assumptions 1 2, and when $Q_{i, 1} \in(0,1), \alpha_{i}$ is given by:

$$
\begin{equation*}
\alpha_{i}=\Phi\left(\Phi^{-1}\left(1-Q_{i, 1}\right) \times \frac{R_{i, \min }-\widetilde{\mu}_{i}}{R_{i, 1}-\widetilde{\mu}_{i}}\right) \tag{B3}
\end{equation*}
$$

Remark that $\alpha_{i}$ is not uniquely defined when $R_{i, 1}=\widetilde{\mu}_{i}$. Four respondents are concerned. These respondents have been dropped.
ii. When $Q_{i, 1}=1$ (176 respondents), three cases arise. $\alpha_{i}=0$ if $Q_{i, 1}=1, R_{i, 1}<\widetilde{\mu}_{i}$ and $R_{i, \min }<\widetilde{\mu}_{i}$ (164 respondents). $\alpha_{i}=1$ if $Q_{i, 1}=1, R_{i, 1}>\widetilde{\mu}_{i}$ and $R_{i, \min }<\widetilde{\mu}_{i}(1$ respondent). The last case is when $Q_{i, 1}=1, R_{i, 1} \neq \widetilde{\mu}_{i}$ and $R_{i, \text { min }}=\widetilde{\mu}_{i}$ (11 respondents). The calibrated probability distribution is degenerated, so Equation B2 does not depend on $\alpha_{i}$. Hence, we have dropped these 11 respondents.

Consequently, 15 observations are lost; so, our final sample is composed of 964 respondents (333 in wave 12, 262 in wave 13 and 369 in wave 14). For each wave, we have ranked the respondents according to $\alpha$, and carried out the same exercise as we did in Section $3-4$ of the main text. Table B 3 presents the results.

Table B3: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $\alpha_{i}$ )-waves 12,13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | $\begin{aligned} & 0.07^{*} \\ & (0.03) \end{aligned}$ | $\begin{gathered} 0.05^{* *} \\ (0.03) \end{gathered}$ | $\begin{gathered} -0.05 \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.15 * * * \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.07^{* * *} \\ (0.02) \end{gathered}$ | $\begin{gathered} 0.09 * * * \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.00 \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.24^{* * *} \\ (0.05) \end{gathered}$ | $\begin{gathered} 0.11^{* * *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.07 * * \\ (0.03) \end{gathered}$ | $\begin{aligned} & -0.02 \\ & (0.03) \end{aligned}$ | $\begin{gathered} 0.08 \\ (0.05) \end{gathered}$ |
| $\sigma$ | $1.04{ }^{* * *}$ | $0.75 * * *$ | -0.05 | 0.16 * | $1.11{ }^{* * *}$ | 0.69 *** | 0.02 | 0.14 | 0.76 *** | $0.64 * * *$ | 0.02 | 0.36 *** |
|  | (0.11) | (0.06) | (0.04) | (0.09) | (0.06) | (0.06) | (0.01) | (0.09) | (0.03) | (0.06) | (0.04) | (0.12) |
| $R^{2}$ | 0.61 | 0.67 | 0.01 | 0.03 | 0.83 | 0.70 | 0.01 | 0.05 | 0.70 | 0.57 | 0.00 | 0.20 |
| Observations | 103 | 103 | 106 | 312 | 78 | 82 | 80 | 240 | 118 | 115 | 111 | 344 |
| (Observations excluded) | 8 | 8 | 5 | 21 | 9 | 6 | 7 | 22 | 5 | 8 | 12 | 25 |

Notes: i. $*, * *$ and $* * *$ represent 10, 5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
$\checkmark$ iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.04+1.23^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.69 \text { and } N=111\right) \\
\mu=0.07+0.78^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.60 \text { and } N=111\right) \\
\mu=-0.05-0.00 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.00 \text { and } N=111\right) \\
\mu=0.17^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=333\right) \\
\mu=0.16^{* * *}+1.04^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.65 \text { and } N=87\right) \\
\mu=0.06+0.72^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.71 \text { and } N=88\right) \\
\mu=0.01+0.03 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.00 \text { and } N=87\right) \\
\mu=0.21^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.14 \text { and } N=262\right) \\
\mu=0.14^{* * *}+0.74^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.70 \text { and } N=123\right) \\
\mu=0.00+0.80^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.62 \text { and } N=123\right) \\
\mu=-0.12^{* *}+0.16^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.13 \text { and } N=123\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=369\right)
\end{gathered}
$$

## B. 3 Alternative measure of coherence: $d_{i, 1000}$

The preceding measures of coherence $\left(d_{i}^{98}, d_{i}^{90}, d_{i}^{80}\right.$ and $\left.\alpha_{i}\right)$ were all computed at the first investment value threshold $R_{i, 1}$. This threshold value can be different from one respondent to another. If the midpoint of the lowest and highest values elicited at the preliminary questions is between 0 and 899, then $R_{i, 1}=500$, but if the midpoint is between 900 and $999, R_{i, 1}=800$ (see Table B4). Whereas the threshold value $R_{i, 1}$ for the first reported probability $Q_{i, 1}$ can vary across respondents, remark that each respondent's sequence of probabilistic questions included a $\$ 1000$ threshold. As shown by Table B4, it is $R_{i, 3}$ if the midpoint is lower than $999, R_{i, 2}$ if it is between 1000 and 1099, and it is $R_{i, 1}$ if it is higher than 1100. This Subsection presents a measure of coherence, $d_{i, 1000}$, such that the distance is computed at the $\$ 1000$ threshold for all the respondents.

Table B4: Algorithm for selection of investment thresholds

| Midpoint | $R_{i, 1}$ | $R_{i, 2}$ | $R_{i, 3}$ | $R_{i, 4}$ |
| :---: | :---: | :---: | :---: | :---: |
| 0 to 899 | 500 | 900 | 1000 | 1100 |
| 900 to 999 | 800 | 900 | 1000 | 1100 |
| 1000 to 1099 | 900 | 1000 | 1100 | 1200 |
| 1100 to 1299 | 1000 | 1100 | 1200 | 1500 |
| 1300 or more | 1000 | 1200 | 1500 | 2000 |

Note: The four thresholds are determined by the respondent's answer to the two preliminary questions asking for the lowest and highest one-yearahead investment value. The midpoint of the lowest and highest values permits to determine the four investment value thresholds according to the algorithm presented in this Table.
$d_{i, 1000}$ is based on Assumptions 1-2 and we consider $\alpha=0.01$ to compute the parameter of the calibrated probability distribution, like for $d_{i}^{98}$. Hence, there is only one difference with $d_{i}^{98}$ : the threshold considered.

Let $Q_{i, 1000}=\operatorname{Prob}\left(R_{i}>1000\right)$ be the reported probability at the $\$ 1000$ threshold. Then $d_{i, 1000}$ is defined as follow:

$$
d_{i, 1000}=\left|Q_{i, 1000}-\operatorname{Prob}\left(\widetilde{R}_{i}>1000\right)\right|
$$

For each wave, we have ranked the respondents according to $d_{1000}$, and carried out the same exercise as we did in Subsection 3-4 of the main text. Table B5 presents the results.

Table B5: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $d_{1000}$ ) -waves 12,13 and 14 -

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | 0.02 | 0.01 | 0.03 |  | $0.24 * * *$ |  | $0.07^{* * *}$ | $0.22^{* * *}$ | 0.09*** | 0.00 | $-0.11^{* * *}$ | 0.09*** |
|  | (0.04) | (0.02) | (0.03) | $(0.03)$ | (0.07) | (0.03) | $(0.02)$ | (0.04) | (0.03) | (0.03) | (0.04) | (0.03) |
| $\sigma$ | $1.34^{* * *}$ | $0.73^{* * *}$ | $-0.03$ | $0.22^{* *}$ | $0.79^{* * *}$ | $0.72^{* * *}$ | $-0.00$ | $0.24^{* * *}$ | $1.07^{* * *}$ | $0.61^{* * *}$ | $0.24^{* * *}$ | $0.34^{* * *}$ |
|  | $(0.08)$ | $(0.06)$ | (0.04) | $(0.09)$ | $(0.16)$ | $(0.06)$ | (0.05) | $(0.09)$ | $(0.05)$ | $(0.07)$ | $(0.05)$ | $(0.08)$ |
| $R^{2}$ | 0.72 | 0.57 | 0.00 | 0.05 | 0.28 | 0.66 | 0.00 | 0.07 | 0.77 | 0.42 | 0.28 | 0.20 |
| Observations | 102 | 106 | 106 | 314 | 82 | 83 | 80 | 245 | 118 | 116 | 119 | 353 |
| (Observations excluded) | 11 | 8 | 7 | 26 | 6 | 5 | 8 | 19 | 7 | 9 | 6 | 26 |

[^0]© iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:
\[

$$
\begin{gathered}
\mu=0.04+1.44^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.76 \text { and } N=113\right) \\
\mu=-0.01+0.67^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.47 \text { and } N=114\right) \\
\mu=0.00+0.04 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.01 \text { and } N=113\right) \\
\mu=0.18^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=340\right) \\
\mu=0.10+1.07^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.40 \text { and } N=88\right) \\
\mu=0.06+0.72^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.67 \text { and } N=88\right) \\
\mu=0.03+0.06^{*} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.04 \text { and } N=88\right) \\
\mu=0.26^{* *}+0.22^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=0.10^{*}+1.02^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.68 \text { and } N=125\right) \\
\mu=-0.01+0.62^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.53 \text { and } N=125\right) \\
\mu=-0.19^{* * *}+0.37^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.41 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$
\]

## B. 4 Alternative measure of coherence: $\alpha_{i, 1000}$

This Subsection discusses a measure of coherence, called $\alpha_{i, 1000}$, which allows for individualspecific $\alpha$. It is close to $\alpha_{i}$, except that the distance considered is between the calibrated probability $\operatorname{Prob}\left(\widetilde{R}_{i}>1000\right)$ and the reported probability $Q_{i, 1000}$.

Let's consider Assumption 1"?

## Assumption 1"

$$
\begin{equation*}
\operatorname{Prob}\left(\widetilde{R}_{i}<R_{i, \min }\right)=\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, \max }\right)=\alpha_{i, 1000} \tag{B4}
\end{equation*}
$$

where

$$
\begin{equation*}
\alpha_{i, 1000}=\arg \inf _{\alpha_{i, 1000} \in[0,1]}\left|Q_{i, 1000}-\operatorname{Prob}\left(\widetilde{R}_{i}>1000\right)\right| \tag{B5}
\end{equation*}
$$

We have to find $\alpha_{i, 1000}$ that satisfies Equations B4 B5.
i. Under Assumptions 1 1 2 , and when $Q_{i, 1000} \in(0,1), \alpha_{i, 1000}$ is given by:

$$
\begin{equation*}
\alpha_{i, 1000}=\Phi\left(\Phi^{-1}\left(1-Q_{i, 1000}\right) \times \frac{R_{i, \min }-\widetilde{\mu}_{i}}{1000-\widetilde{\mu}_{i}}\right) \tag{B6}
\end{equation*}
$$

$\alpha_{i, 1000}$ is not uniquely defined when $\widetilde{\mu}_{i}=1000.54$ respondents are concerned. These respondents have been dropped.
ii. When $Q_{i, 1000}=1$ (160 respondents), three cases arise. $\alpha_{i, 1000}=0$ if $Q_{i, 1000}=1,1000<\widetilde{\mu}_{i}$ and $R_{i, \text { min }}<\widetilde{\mu}_{i}$ (148 respondents). $\alpha_{i, 1000}=1$ if $Q_{i, 1000}=1,1000>\widetilde{\mu}_{i}$ and $R_{i, \text { min }}<\widetilde{\mu}_{i}$ (two respondents). The last case is when $Q_{i, 1000}=1, \widetilde{\mu}_{i} \neq 1000$ and $R_{i, \min }=\widetilde{\mu}_{i}$ (10 respondents). The calibrated probability distribution is degenerated, so Equation B5 does not depend on $\alpha_{i, 1000}$. Hence, we have dropped these 10 respondents.

Hence, 64 observations are lost; so our final sample is composed of 915 respondents ( 326 in wave 12, 249 in wave 13 and 340 in wave 14). For each wave, we have ranked the respondents according to $\alpha_{i, 1000}$, and carried out the same exercise as we did in Section 3-4 of the main text. Table B6 presents the results.

Table B6: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $\alpha_{i, 1000}$ ) -waves 12,13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | 0.04 |  |  | $0.15{ }^{* * *}$ | 0.09*** | $0.12{ }^{* * *}$ | 0.03 | $0.27^{* * *}$ | 0.09*** | 0.10** | -0.04 | $0.12^{* * *}$ |
|  | (0.04) | (0.03) | $(0.03)$ | $(0.04)$ | (0.03) | (0.03) | (0.02) | (0.04) | (0.03) | (0.04) | (0.04) | (0.03) |
| $\sigma$ | $1.06^{* * *}$ | $0.73^{* * *}$ | $-0.05$ | $0.19^{* *}$ | $1.04^{* * *}$ | $0.67^{* * *}$ | $0.01$ | $0.12$ | $0.85^{* * *}$ | $0.47^{* * *}$ | $0.09$ | $0.26^{* * *}$ |
|  | $(0.11)$ | $(0.06)$ | (0.04) | (0.09) | $(0.06)$ | $(0.06)$ | (0.01) | $(0.08)$ | $(0.10)$ | $(0.10)$ | (0.08) | $(0.08)$ |
| $R^{2}$ | 0.56 | 0.60 | 0.02 | 0.05 | 0.74 | 0.69 | 0.00 | 0.04 | 0.46 | 0.38 | 0.03 | 0.09 |
| Observations | 101 | 98 | 104 | 303 | 73 | 77 | 77 | 227 | 106 | 106 | 104 | 316 |
| (Observations excluded) | 8 | 10 | 5 | 23 | 10 | 6 | 6 | 22 | 7 | 8 | 9 | 24 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
$\rightleftharpoons$ iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.01+1.26^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.65 \text { and } N=109\right) \\
\mu=0.05+0.75^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.51 \text { and } N=108\right) \\
\mu=-0.00-0.03 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.00 \text { and } N=109\right) \\
\mu=0.17^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=326\right) \\
\mu=0.12^{*}+1.06^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.55 \text { and } N=83\right) \\
\mu=0.12^{* *}+0.65^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.66 \text { and } N=83\right) \\
\mu=0.05+0.00 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.00 \text { and } N=83\right) \\
\mu=0.23^{* * *}+0.25^{*} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.12 \text { and } N=249\right) \\
\mu=0.10^{* *}+0.94^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.59 \text { and } N=113\right) \\
\mu=0.07+0.58^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.49 \text { and } N=114\right) \\
\mu=-0.24^{* *}+0.41^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.36 \text { and } N=113\right) \\
\mu=0.04+0.50^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.35 \text { and } N=340\right)
\end{gathered}
$$

## B. 5 Alternative measure of coherence: $\bar{d}_{i}$

$\bar{d}_{i}$ is a measure of coherence based on the four thresholds. It is based on Assumptions 1-2, and we consider $\alpha=0.01 \forall i$ to compute the parameter of the calibrated probability distribution. This criterion of coherence averages the absolute deviations over the four thresholds:

$$
\bar{d}_{i}=\frac{1}{4} \sum_{k=1}^{4}\left|Q_{i, k}-\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, k}\right)\right|
$$

We have ranked the respondents according to $\bar{d}$ for each wave, and carried out the same exercise as we did in Subsection 3-4 of the main text. Table B7 presents the results.

## B. 6 Alternative measure of coherence: $d_{i}^{\text {Kol }}$

$d_{i}^{K o l}$ is a measure of coherence based on the four thresholds. It is based on Assumptions 1-2, and we consider $\alpha=0.01 \forall i$ to compute the parameter of the calibrated probability distribution. This criterion of coherence is a Kolmogorov distance, i.e. the highest distance among the four distances:

$$
d_{i}^{K o l}=\max _{k=1, \ldots, 4}\left|Q_{i, k}-\operatorname{Prob}\left(\widetilde{R}_{i}>R_{i, k}\right)\right|
$$

We have ranked the respondents according to $d^{K o l}$ for each wave, and carried out the same exercise as we did in Subsection 3-4 of the main text. Table B8 presents the results.

Table B7: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $\bar{d}$ ) -waves 12, 13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | $\begin{aligned} & 0.091^{*} \\ & (0.05) \end{aligned}$ | $\begin{gathered} 0.02 \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.12^{* * *} \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.22^{* * *} \\ (0.03) \end{gathered}$ | $\begin{gathered} 0.06 \\ (0.05) \end{gathered}$ | $\begin{gathered} 0.13^{* *} \\ (0.06) \end{gathered}$ | $\begin{gathered} 0.14^{* * *} \\ (0.05) \end{gathered}$ | $\underset{(0.05)}{0.19^{* * *}}$ | $\begin{aligned} & -0.01 \\ & (0.04) \end{aligned}$ | $\underset{(0.04)}{0.11 * * *}$ | $\begin{gathered} 0.04 \\ (0.04) \end{gathered}$ | $\begin{gathered} 0.09^{* * *} \\ (0.03) \end{gathered}$ |
| $\sigma$ | $\begin{gathered} 1.04^{* * *} \\ (0.15) \end{gathered}$ | $\begin{gathered} 0.66^{* * *} \\ (0.06) \end{gathered}$ | $\begin{aligned} & -0.01 \\ & (0.02) \\ & \hline \end{aligned}$ | $\begin{gathered} 0.08 \\ (0.06) \\ \hline \end{gathered}$ | $\begin{gathered} 0.93^{* * *} \\ (0.13) \end{gathered}$ | $\begin{gathered} 0.77^{* * *} \\ (0.11) \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.07) \\ \hline \end{gathered}$ | $\begin{gathered} 0.30^{* *} \\ (0.12) \\ \hline \end{gathered}$ | $\begin{gathered} 0.99^{* * *} \\ (0.08) \end{gathered}$ | $\begin{gathered} 0.39^{* * *} \\ (0.08) \end{gathered}$ | $\begin{gathered} 0.23^{* * *} \\ (0.05) \end{gathered}$ | $\begin{gathered} 0.34^{* * *} \\ (0.07) \end{gathered}$ |
| $R^{2}$ | 0.44 | 0.43 | 0.00 | 0.01 | 0.48 | 0.44 | 0.00 | 0.09 | 0.63 | 0.17 | 0.21 | 0.22 |
| Observations | 106 | 105 | 109 | 320 | 80 | 84 | 81 | 245 | 116 | 118 | 119 | 353 |
| (Observations excluded) | 7 | 9 | 4 | 20 | 8 | 4 | 7 | 19 | 9 | 7 | 6 | 22 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
$\longmapsto$ ii. Heteroskedasticity-robust standard errors are in parentheses.
$\mathfrak{c o}$ iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=0.00+1.28^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.60 \text { and } N=113\right) \\
\mu=0.03+0.59^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.40 \text { and } N=114\right) \\
\mu=0.13^{* * *}+0.01 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.00 \text { and } N=113\right) \\
\mu=0.18^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=340\right) \\
\mu=0.00+1.01^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.54 \text { and } N=88\right) \\
\mu=0.21^{* *}+0.53^{* *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.47 \text { and } N=88\right) \\
\mu=0.15+0.04 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.01 \text { and } N=88\right) \\
\mu=0.22^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=-0.03+1.06^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.71 \text { and } N=125\right) \\
\mu=0.13^{* *}+0.32^{* *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.15 \text { and } N=125\right) \\
\mu=-0.05+0.34^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.34 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$

Table B8: Least squares regressions of $\mu$ on $\sigma$ by type of coherence (respondents ranked according to $d^{\text {Kol }}$ ) -waves 12,13 and 14-

|  | Wave 12 |  |  |  | Wave 13 |  |  |  | Wave 14 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|  | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types | Coherent 1 | Coherent 2 | Coherent 3 | All types |
|  | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ | $\mu$ |
| Intercept | 0.03 | $0.14 * * *$ | 0.23 *** | 0.10 ** |  | 0.26 *** |  | $0.24^{* * *}$ | $-0.11^{* *}$ | $0.16^{* * *}$ | 0.12*** | $0.09^{* * *}$ |
|  | ${ }_{(0.04)}$ | (0.04) | (0.04) | (0.04) | ${ }_{(0.04)}$ | (0.04) | (0.06) | (0.03) | ${ }_{\text {(0.05 }}$ | (0.03) | (0.05) | (0.03) |
| $\sigma$ |  | $0.31^{* * *}$ | $-0.14^{*}$ | $0.39^{* * *}$ | $0.83^{* * *}$ | $0.23^{* * *}$ | $-0.14$ | $0.19^{* * *}$ | $1.01^{* * *}$ | $0.23^{* * *}$ | $0.23^{* *}$ | $0.35^{* * *}$ |
|  | $(0.08)$ | $(0.09)$ | (0.07) | $(0.10)$ | $(0.11)$ | $(0.04)$ | (0.09) | $(0.06)$ | $(0.14)$ | $(0.03)$ | (0.09) | $(0.06)$ |
| $R^{2}$ | 0.58 | 0.11 | 0.03 | 0.15 | 0.58 | 0.20 | 0.02 | 0.07 | 0.57 | 0.18 | 0.08 | 0.19 |
| Observations | 103 | 108 | 106 | 317 | 78 | 85 | 85 | 248 | 116 | 116 | 118 | 350 |
| (Observations excluded) | 10 | 6 | 7 | 23 | 10 | 3 | 3 | 16 | 9 | 9 | 7 | 25 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
$\curvearrowleft$ ii. Heteroskedasticity-robust standard errors are in parentheses.

- iii. If we do not exclude the observations that are flagged by the DFITS cutoff criterion, we obtain:

$$
\begin{gathered}
\mu=-0.02+0.93^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 12\left(R^{2}=0.50 \text { and } N=113\right) \\
\mu=0.19^{* * *}+0.15 \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 12\left(R^{2}=0.05 \text { and } N=114\right) \\
\mu=0.22^{* * *}-0.03 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 12\left(R^{2}=0.00 \text { and } N=113\right) \\
\mu=0.18^{* * *}+0.24^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 12\left(R^{2}=0.08 \text { and } N=340\right) \\
\mu=-0.03+0.94^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 13\left(R^{2}=0.72 \text { and } N=88\right) \\
\mu=0.26^{* * *}+0.19^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 13\left(R^{2}=0.15 \text { and } N=88\right) \\
\mu=0.32^{* * *}-0.06 \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 13\left(R^{2}=0.02 \text { and } N=88\right) \\
\mu=0.22^{* * *}+0.26^{* *} \sigma+\hat{\epsilon} \text { all types considered in wave } 13\left(R^{2}=0.13 \text { and } N=264\right) \\
\mu=-0.16^{* * *}+1.17^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 1 \text { in wave } 14\left(R^{2}=0.75 \text { and } N=125\right) \\
\mu=0.15^{* * *}+0.24^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 2 \text { in wave } 14\left(R^{2}=0.24 \text { and } N=125\right) \\
\mu=0.00+0.38^{* * *} \sigma+\hat{\epsilon} \text { for Coherent } 3 \text { in wave } 14\left(R^{2}=0.28 \text { and } N=125\right) \\
\mu=0.03+0.49^{* * *} \sigma+\hat{\epsilon} \text { all types considered in wave } 14\left(R^{2}=0.34 \text { and } N=375\right)
\end{gathered}
$$

## C Spearman's rank correlation coefficients between the different measures of coherence

Table C1: Spearman's rank correlation coefficients between the different measures of coherence -wave 12-

|  | $d^{98}$ | $\alpha_{i}$ | $d^{90}$ | $d^{80}$ | $\alpha_{i, 1000}$ | $d_{1000}$ | $d^{\text {Kol }}$ | $\bar{d}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $d^{98}$ | 1 | $0.84^{* * *}$ | 0.97*** | 0.91 *** | $0.72^{* * *}$ | $0.82 * * *$ | 0.30*** | 0.58*** |
| $\alpha_{i}$ |  | 1 | 0.81*** | $0.75 * * *$ | 0.80*** | 0.68*** | 0.33 *** | 0.50*** |
| $d^{90}$ |  |  | 1 | 0.96 *** | 0.71 *** | 0.83*** | $0.34 * * *$ | 0.60*** |
| $d^{80}$ |  |  |  | 1 | 0.68*** | 0.80*** | 0.35*** | 0.59*** |
| $\alpha_{i, 1000}$ |  |  |  |  | 1 | $0.84^{* * *}$ | $0.39^{* * *}$ | $0.56{ }^{* * *}$ |
| $d_{1000}$ |  |  |  |  |  | 1 | $0.41^{* * *}$ | 0.66*** |
| $d^{\text {Kol }}$ |  |  |  |  |  |  | 1 | $0.82{ }^{* * *}$ |
| $\bar{d}$ |  |  |  |  |  |  |  | 1 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.

Table C2: Spearman's rank correlation coefficients between the different measures of coherence -wave 13-

|  | $d^{98}$ | $\alpha_{i}$ | $d^{90}$ | $d^{80}$ | $\alpha_{i, 1000}$ | $d_{1000}$ | $d^{\text {Kol }}$ | $\bar{d}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $d^{98}$ | 1 | 0.89*** | $0.97^{* * *}$ | 0.90*** | 0.84*** | 0.89*** | 0.28*** | 0.55*** |
| $\alpha_{i}$ |  | 1 | 0.86*** | 0.79*** | 0.87*** | 0.74*** | 0.27*** | 0.48*** |
| $d^{90}$ |  |  | 1 | $0.94 * * *$ | $0.81 * * *$ | $0.86^{* * *}$ | 0.32*** | 0.60 *** |
| $d^{80}$ |  |  |  | 1 | $0.74 * * *$ | 0.81 *** | $0.36{ }^{* * *}$ | 0.58*** |
| $\alpha_{i, 1000}$ |  |  |  |  | 1 | 0.90*** | 0.32*** | $0.57 * * *$ |
| $d_{1000}$ |  |  |  |  |  | 1 | 0.33*** | 0.62*** |
| $d^{\text {Kol }}$ |  |  |  |  |  |  | 1 | 0.82*** |
| $\bar{d}$ |  |  |  |  |  |  |  | 1 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.

Table C3: Spearman's rank correlation coefficients between the different measures of coherence -wave 14-

|  | $d^{98}$ | $\alpha_{i}$ | $d^{90}$ | $d^{80}$ | $\alpha_{i, 1000}$ | $d_{1000}$ | $d^{\text {Kol }}$ | $\bar{d}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $d^{98}$ | 1 | $0.89 * * *$ | 0.97*** | 0.90*** | 0.78*** | 0.86*** | 0.34*** | $0.56^{* * *}$ |
| $\alpha_{i}$ |  | 1 | 0.86*** | 0.78*** | 0.78*** | 0.72 *** | 0.29 *** | 0.45 *** |
| $d^{90}$ |  |  | 1 | 0.96*** | 0.76*** | 0.81*** | 0.32 *** | 0.55*** |
| $d^{80}$ |  |  |  | 1 | 0.71 *** | 0.78*** | $0.33 * * *$ | 0.55*** |
| $\alpha_{i, 1000}$ |  |  |  |  | 1 | 0.89*** | 0.36 *** | $0.53^{* * *}$ |
| $d_{1000}$ |  |  |  |  |  | 1 | $0.41^{* * *}$ | $0.64 * * *$ |
| $d^{\text {Kol }}$ |  |  |  |  |  |  | 1 | $0.87^{* * *}$ |
| $\bar{d}$ |  |  |  |  |  |  |  | 1 |

[^1]
## D Spearman's rank correlation coefficients between $\mu$ and $\sigma$ for the Coherent 1 and the Coherent 3 by wave

Table D1: Spearman's rank correlation coefficients between $\mu$ and $\sigma$ for the Coherent 1 and the Coherent 3 by wave (all measures of coherence considered)

| Wave 12 | Wave 13 |  |  | Wave 14 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $[1]$ $[2]$ ${ }^{[3]}$ <br> $\rho_{1,12}$ $\rho_{3,12}$ $H_{0}: \rho_{1,12}=\rho_{3,12}$ | $\begin{gathered} {[4]} \\ \rho_{1,13} \\ \hline \end{gathered}$ | $\begin{gathered} {[5]} \\ \rho_{3,13} \\ \hline \end{gathered}$ | $\stackrel{[6]}{H_{0}: \rho_{1,13}=\rho_{3,13}}$ | $\begin{gathered} {[7]} \\ \rho_{1,14} \\ \hline \end{gathered}$ | $\begin{gathered} {[8]} \\ \rho_{3,14} \\ \hline \end{gathered}$ | $\begin{gathered} {[9]} \\ H_{0}: \rho_{1,14}=\rho_{3,14} \\ \hline \end{gathered}$ |
| Respondents preliminary ranked according to $d^{98}$   <br> $0.60^{* * *}$ 0.09  <br> $(7.06)$ $(0.92)$ $4.34^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.60^{* * *} \\ (6.21) \\ {[88]} \end{gathered}$ | $\begin{gathered} 0.12 \\ (1.08) \\ {[88]} \end{gathered}$ | 3.63 *** | $\begin{gathered} 0.55^{* * *} \\ (6.63) \\ {[125]} \end{gathered}$ | $\begin{gathered} 0.17^{*} \\ (1.84) \\ {[125]} \end{gathered}$ | $3.39 * * *$ |
| Respondents preliminary ranked according to $d^{90}$   <br> $0.56^{* * *}$ -0.03  <br> $(6.45)$ $(-0.31)$ $4.77^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.60^{* * *} \\ (6.21) \\ {[88]} \end{gathered}$ | $\begin{gathered} 0.11 \\ (0.99) \\ {[88]} \end{gathered}$ | 3.69*** | $\begin{gathered} 0.55^{* * *} \\ (6.63) \\ {[125]} \end{gathered}$ | $\begin{gathered} 0.15 \\ (1.62) \\ {[125]} \end{gathered}$ | $3.54 * * *$ |
| Respondents preliminary ranked according to $d^{80}$   <br> $0.58^{* * *}$ -0.1  <br> $(6.75)$ $(-1.02)$ $5.49^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.61^{* * *} \\ (6.35) \\ {[88]} \end{gathered}$ | $\begin{gathered} 0.14 \\ (1.26) \\ {[88]} \end{gathered}$ | 3.59*** | $\begin{gathered} 0.58^{* * *} \\ (7.11) \\ {[125]} \end{gathered}$ | $\begin{gathered} 0.12 \\ (1.29) \\ {[125]} \end{gathered}$ | 4.11*** |
| Respondents preliminary ranked according to $\alpha$   <br> $0.57^{* * *}$ -0.07  <br> $(6.54)$ $(-0.71)$ $5.12^{* * *}$ <br> $[111]$ $[111]$  | $\begin{gathered} 0.60^{* * *} \\ (6.17) \\ {[87]} \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.44) \\ {[87]} \end{gathered}$ | 4.05*** | $\begin{gathered} 0.53^{* * *} \\ (6.28) \\ {[123]} \end{gathered}$ | $\begin{gathered} 0.03 \\ (0.32) \\ {[123]} \end{gathered}$ | 4.21*** |
| Respondents preliminary ranked according to $d_{1000}$   <br> $0.56^{* * *}$ 0.06  <br> $(6.45)$ $(0.61)$ $4.13^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.44^{* * *} \\ (4.23) \\ {[88]} \end{gathered}$ | $\begin{gathered} 0.01 \\ (0.06) \\ {[88]} \end{gathered}$ | $2.94 * * *$ | $\begin{gathered} 0.54^{* * *} \\ (6.48) \\ {[125]} \end{gathered}$ | $\begin{aligned} & 0.18^{*} \\ & (1.95) \\ & {[125]} \end{aligned}$ | 3.20 *** |
| Respondents preliminary ranked according to $\alpha_{1000}$   <br> $0.53^{* * *}$ -0.08  <br> $(5.90)$ $(-0.80)$ $4.74^{* * *}$ <br> $[109]$ $[109]$  | $\begin{gathered} 0.51^{* * *} \\ (4.88) \\ {[83]} \end{gathered}$ | $\begin{gathered} -0.03 \\ (-0.26) \\ {[83]} \end{gathered}$ | $3.64 * * *$ | $\begin{gathered} 0.57^{* * *} \\ (6.59) \\ {[113]} \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.92) \\ {[113]} \end{gathered}$ | 4.01*** |
| Respondents preliminary ranked according to $\bar{d}$   <br> $0.35^{* * *}$ 0.01  <br> $(3.72)$ $(0.06)$ $2.59^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.39^{* * *} \\ (3.68) \\ {[88]} \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.81) \\ {[88]} \end{gathered}$ | 2.03** | $\begin{gathered} 0.38^{* * *} \\ (4.29) \\ {[125]} \end{gathered}$ | $\begin{gathered} 0.16^{*} \\ (1.73) \\ {[125]} \end{gathered}$ | 1.81* |
| Respondents preliminary ranked according to $d^{\text {Kol }}$   <br> $0.40^{* * *}$ -0.05  <br> $(4.32)$ $(-0.51)$ $3.41^{* * *}$ <br> $[113]$ $[113]$  | $\begin{gathered} 0.58^{* * *} \\ (5.93) \\ {[88]} \\ \hline \end{gathered}$ | $\begin{gathered} -0.18 \\ (-1.63) \\ {[88]} \end{gathered}$ | 5.35*** | $\begin{gathered} 0.41^{* * *} \\ (4.67) \\ {[125]} \\ \hline \end{gathered}$ | $\begin{gathered} 0.13 \\ (1.40) \\ {[125]} \end{gathered}$ | 2.31** |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. $\rho_{j, t}$ is the Spearman's rank correlation coefficient between $\mu$ and $\sigma$ for the Coherent $j$ in wave $t$. For instance, when respondents have been preliminary ranked according to $d^{98}, \rho_{1,12}=0.60$ is the Spearman's rho between $\mu$ and $\sigma$ for the Coherent 1 in Wave 12 .
iii. $z$-scores for testing the null hypothesis $H_{0}: \rho_{j, t}=0$ are in parentheses. For instance, when respondents have been preliminary ranked according to $d^{98}, \rho_{1,12}=0.60$, and 7.06 is the $z$-score of the test of the null hypothesis that $\rho_{1,12}=0.7 .06$ is higher than 2.58 , the 1 percent critical value from the standard normal distribution, so the null hypothesis is rejected (the $z$-score for $\rho_{j, t}$ approximately follows a standard normal distribution under the null hypothesis). These $z$-scores are obtained via a Fisher transformation. Let $F\left(\rho_{j, t}\right)=\operatorname{arctanh}\left(\rho_{j, t}\right)$ be the Fisher transformation of $\rho_{j, t}$, then $z=\sqrt{\frac{N_{j, t}-3}{1.06}} F\left(\rho_{j, t}\right)$ is the $z$-score for $\rho_{j, t}$.
iv. To test whether the Spearman's rhos of the Coherent 1 and the Coherent 3 in wave $t$ have different strengths, we test the null hypothesis $H_{0}: \rho_{1, t}=\rho_{3, t}$, we have Computed $z$-scores which are reported in Columns [3], [6] and [9]. For instance, when respondents have been preliminary ranked according to $d^{98}$, we have obtained a $z$-score of 4.34 for the wave 12 . This is higher than 2.58 , the 1 percent critical value from the standard normal distribution, so the null hypothesis is rejected. To obtain these $z$-scores, we have transformed $\rho_{1, t}$ and $\rho_{3, t}$ with the Fisher transformation (e.g., $\left.F\left(\rho_{1, t}\right)=\operatorname{arctanh}\left(\rho_{1, t}\right)\right)$. Then the $z$-score is the difference $z=\frac{F\left(\rho_{1, t}\right)-F\left(\rho_{3, t}\right)}{\sqrt{\frac{1.06}{N_{1, t}-3}+\frac{1.06}{N_{3, t}-3}}}$

[^2]E Who's who in term of coherence considering alternative measures of coherence

F Who's who in term of coherence considering all the respondents for whom we can compute a measure of coherence

Table E1: Who's who in term of coherence $\left(d^{90}\right)$


Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E2: Who's who in term of coherence $\left(d^{80}\right)$

|  | OLS |  |  | Tobit |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[2]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[3]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[4]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[5]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[6]} \\ d^{80} \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.016 |  | -0.020 | -0.014 |  | -0.018 |
|  | (0.052) |  | (0.051) | (0.040) |  | (0.041) |
| Attended college | -0.047 |  | -0.057 | -0.044 |  | -0.055 |
|  | (0.046) |  | (0.046) | (0.036) |  | (0.036) |
| Associated degree | -0.053 |  | -0.064 | -0.051 |  | -0.062 |
|  | (0.050) |  | (0.050) | (0.041) |  | (0.041) |
| BA/BS | -0.053 |  | -0.064 | -0.049 |  | -0.060* |
|  | (0.047) |  | (0.046) | (0.036) |  | (0.036) |
| MS or PhD | -0.080* |  | -0.091* | -0.075* |  | -0.087** |
|  | (0.048) |  | (0.047) | (0.039) |  | (0.038) |
| Professional | -0.012 |  | -0.014 | -0.006 |  | -0.009 |
|  | (0.066) |  | (0.064) | (0.050) |  | (0.050) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | 0.009 | -0.013 |  | 0.006 | -0.016 |  |
|  | (0.031) | (0.031) |  | (0.028) | (0.028) |  |
| $20000 \leq$ Income $<30000$ | -0.027 | $-0.047^{*}$ |  | $-0.030$ | $-0.050^{*}$ |  |
|  | (0.026) | (0.027) |  | (0.027) | (0.027) |  |
| $30000 \leq$ Income $<40000$ | -0.024 | $-0.049^{*}$ |  | $-0.030$ | $-0.054^{* *}$ |  |
|  | (0.028) | $(0.028)$ |  | $(0.027)$ | $(0.027)$ |  |
| $40000 \leq$ Income $<50000$ | -0.031 | -0.061** |  | -0.033 | $-0.063^{* *}$ |  |
|  | (0.028) | (0.028) |  | (0.029) | (0.028) |  |
| $50000 \leq$ Income $<60000$ | -0.059** | -0.079*** |  | -0.061* | -0.081** |  |
|  | (0.026) | (0.029) |  | (0.032) | (0.031) |  |
| Income $\geq 60000$ | 0.005 | -0.025 |  | 0.002 | -0.028 |  |
|  | (0.025) | (0.026) |  | (0.026) | (0.025) |  |
| Did not know/refused | $-0.037$ | $-0.067^{* *}$ |  | $-0.041$ | $-0.071^{* * *}$ |  |
|  | $(0.026)$ | $(0.026)$ |  | $(0.027)$ | $(0.026)$ |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.301 | -0.439* | -0.341 | -0.293 | -0.420* | -0.336 |
|  | (0.26) | (0.24) | (0.25) | (0.23) | (0.22) | (0.22) |
| $(\text { Age } / 100)^{2}$ | 0.004 | 0.005** | 0.004 | 0.004 | 0.005** | 0.004* |
|  | (0.003) | (0.003) | (0.003) | (0.002) | (0.002) | (0.002) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.010 | -0.006 | -0.014 | -0.009 | -0.004 | -0.012 |
|  | (0.012) | (0.012) | (0.012) | (0.012) | (0.012) | (0.012) |
| Race |  |  |  |  |  |  |
| White | $-0.013$ |  | $-0.010$ | $-0.016$ | $-0.021$ | $-0.013$ |
|  | $(0.023)$ | $(0.023)$ | $(0.023)$ | $(0.023)$ | $(0.023)$ | $(0.023)$ |
| Black | 0.017 | 0.036 | 0.015 | 0.019 | $0.039$ | 0.018 |
|  | (0.036) | (0.037) | (0.036) | (0.033) | $(0.033)$ | (0.033) |
| Asian | 0.012 | -0.002 | 0.019 | 0.011 | -0.002 | 0.019 |
|  | (0.043) | (0.043) | (0.042) | (0.042) | (0.043) | (0.043) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | 0.011 | 0.016 | 0.015 | 0.012 | 0.017 | 0.016 |
|  | (0.022) | (0.021) | (0.022) | (0.024) | (0.024) | (0.024) |
| Roman Catholic |  |  | $0.020$ | $0.021$ | $0.032$ | $0.021$ |
|  | $(0.020)$ | $(0.020)$ | $(0.021)$ | $(0.022)$ | $(0.023)$ | $(0.023)$ |
| Protestant | $-0.001$ | $0.006$ | $-0.001$ | $-0.004$ | $0.002$ | $-0.004$ |
|  | $(0.020)$ | $(0.019)$ | $(0.020)$ | $(0.022)$ | $(0.023)$ | $(0.022)$ |
| Christian | 0.002 | 0.009 | $0.001$ | $-0.001$ | $0.006$ | $-0.002$ |
|  | (0.019) | (0.019) | $(0.020)$ | $(0.022)$ | $(0.022)$ | $(0.022)$ |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.003 | -0.003 | 0.001 | 0.003 | -0.003 | 0.002 |
|  | (0.015) | (0.015) | (0.015) | (0.016) | (0.015) | (0.016) |
| Wave 14 | 0.007 | -0.003 | 0.008 | 0.008 | -0.003 | 0.010 |
|  | (0.014) | (0.013) | (0.014) | (0.015) | (0.014) | (0.015) |
| Intercept | $0.305^{* * *}$ | $0.312 * * *$ | $0.305^{* * *}$ | $0.303^{* * *}$ | $0.311^{* * *}$ | $0.301 * * *$ |
|  | (0.067) | (0.057) | (0.067) | (0.062) | (0.055) | (0.060) |
| Observations | 907 | 960 | 907 | 907 | 960 | 907 |
| $R^{2}$ | 0.04 | 0.04 | 0.03 |  |  |  |
| Significance of the regression: P-value | 0.02** | 0.04** | 0.16 | 0.08* | 0.02** | 0.18 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P-value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E3: Who's who in term of coherence $\left(\alpha_{i}\right)$

|  | OLS |  |  | Tobit |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] |
|  | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ |
| Education $\longrightarrow$ |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.009 |  | -0.015 | -0.001 |  | -0.009 |
|  | (0.064) |  | (0.063) | (0.059) |  | (0.059) |
| Attended college | -0.046 |  | -0.055 | -0.047 |  | -0.058 |
|  | (0.060) |  | (0.058) | (0.052) |  | (0.052) |
| Associated degree | -0.052 |  | -0.062 | -0.049 |  | -0.061 |
|  | (0.064) |  | (0.063) | (0.060) |  | (0.059) |
| BA/BS | -0.077 |  | -0.087 | -0.077 |  | -0.088* |
|  | (0.060) |  | (0.059) | (0.053) |  | (0.053) |
| MS or PhD | -0.114* |  | -0.124** | -0.102* |  | -0.114** |
|  | (0.062) |  | (0.061) | (0.056) |  | (0.056) |
| Professional | -0.079 |  | -0.078 | -0.071 |  | -0.070 |
|  | (0.076) |  | (0.074) | (0.073) |  | (0.073) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.040 | -0.066** |  | -0.056 | -0.086** |  |
|  | (0.035) | (0.033) |  | (0.040) | (0.039) |  |
| $20000 \leq$ Income $<30000$ | -0.051 | -0.068** |  | -0.067* | $-0.084^{* *}$ |  |
|  | (0.032) | (0.030) |  | (0.039) | (0.037) |  |
| $30000 \leq$ Income $<40000$ | -0.048 | $-0.077^{* *}$ |  | -0.066* | -0.096** |  |
|  | (0.033) | (0.031) |  | (0.039) | (0.037) |  |
| $40000 \leq$ Income $<50000$ | -0.052 | -0.086*** |  | -0.071* | -0.109*** |  |
|  | (0.035) | (0.033) |  | (0.041) | (0.040) |  |
| $50000 \leq$ Income $<60000$ | -0.067* | -0.094*** |  | -0.085* | $-0.112^{* * *}$ |  |
|  | (0.037) | (0.035) |  | (0.046) | (0.044) |  |
| Income $\geq 60000$ | -0.010 | -0.044 |  | -0.014 | -0.048 |  |
|  | (0.031) | (0.029) |  | (0.037) | (0.035) |  |
| Did not know/refused | -0.055* | -0.080** |  | -0.062 | $-0.088^{* *}$ |  |
|  | (0.033) | (0.031) |  | (0.038) | (0.037) |  |
| Age ${ }^{\text {a }}$ |  |  |  |  |  |  |
| Age/100 | -0.188 | -0.393 | -0.204 | -0.229 | -0.439 | -0.253 |
|  | (0.300) | (0.27) | (0.29) | (0.32) | (0.31) | (0.32) |
| $\left(\right.$ Age /100) ${ }^{2}$ | 0.003 | 0.005* | 0.003 | 0.003 | 0.005 | 0.003 |
|  | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.021 | -0.012 | -0.023 | -0.020 | -0.012 | -0.023 |
|  | (0.015) | (0.015) | (0.015) | (0.018) | (0.017) | (0.017) |
| Race |  |  |  |  |  |  |
| White | -0.032 | -0.051* | -0.029 | -0.025 | -0.047 | -0.026 |
|  | (0.030) | (0.029) | (0.029) | (0.033) | (0.032) | (0.033) |
| Black | 0.011 | 0.011 | 0.007 | 0.016 | 0.016 | 0.012 |
|  | (0.039) | (0.039) | (0.039) | (0.047) | (0.046) | (0.047) |
| Asian | 0.067 | 0.031 | 0.076 | 0.085 | 0.050 | 0.097 |
|  | (0.057) | (0.056) | (0.056) | (0.061) | (0.061) | (0.061) |
| Other |  | Base group |  |  | Base group |  |
| Religion $0.010{ }^{\text {c }}$ |  |  |  |  |  |  |
| No religion |  |  |  |  | $0.027$ | 0.011 |
|  | $(0.028)$ | $(0.028)$ | $(0.028)$ | $(0.034)$ | $(0.034)$ | (0.034) |
| Roman Catholic | 0.033 | $0.045^{*}$ | $0.032$ | $0.041$ | $0.055^{*}$ | 0.039 |
|  | (0.027) | $(0.026)$ | $(0.027)$ | $(0.032)$ | $(0.032)$ | (0.032) |
| Protestant | -0.009 | 0.009 | -0.012 | -0.008 | 0.015 | -0.012 |
|  | (0.025) | (0.025) | (0.026) | (0.031) | (0.031) | (0.032) |
| Christian | -0.008 | 0.009 | -0.009 | -0.007 | 0.012 | -0.010 |
|  | (0.026) | (0.026) | (0.026) | (0.032) | (0.032) | (0.032) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | -0.019 | -0.020 | -0.020 | -0.027 | -0.026 | -0.027 |
|  | (0.019) | (0.018) | (0.019) | (0.023) | (0.022) | (0.023) |
| Wave 14 | -0.009 | -0.016 | -0.007 | -0.010 | -0.015 | -0.006 |
|  | (0.018) | (0.017) | (0.018) | (0.021) | (0.020) | (0.021) |
| Intercept | 0.422*** | $0.432^{* * *}$ | 0.395*** | $0.418^{* * *}$ | 0.429*** | $0.382^{* * *}$ |
|  | (0.083) | (0.066) | (0.081) | (0.088) | (0.078) | (0.086) |
| Observations | 895 | 946 | 895 | 895 | 946 | 895 |
| $R^{2}$ | 0.05 | 0.04 | 0.04 |  |  |  |
| Significance of the regression: P-value | 0.00*** | 0.00*** | 0.01** | 0.07* | 0.02** | 0.09* |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P-value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E4: Who's who in term of coherence $\left(d_{1000}\right)$

|  | $\begin{gathered} {[1]} \\ d_{1000} \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { OLS } \\ {[2]} \\ d_{1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[3]} \\ d_{1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[4]} \\ d_{1000} \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Tobit } \\ {[5]} \\ d_{1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[6]} \\ d_{1000} \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Education $\leq$ Grade 11 | Base group |  |  | Base group |  |  |
| High school graduate | 0.008 |  | 0.003 | 0.015 |  | 0.008 |
|  | (0.054) |  | (0.054) | (0.049) |  | (0.049) |
| Attended college | -0.022 |  | -0.032 | -0.026 |  | -0.036 |
|  | (0.048) |  | (0.048) | (0.043) |  | (0.043) |
| Associated degree | -0.046 |  | -0.057 | -0.051 |  | -0.062 |
|  | (0.052) |  | (0.052) | (0.050) |  | (0.050) |
| BA/BS | -0.036 |  | -0.045 | -0.035 |  | -0.044 |
|  | (0.048) |  | (0.048) | (0.044) |  | (0.044) |
| MS or PhD | -0.049 |  | -0.059 | -0.047 |  | -0.056 |
|  | (0.050) |  | (0.049) | (0.047) |  | (0.047) |
| Professional | -0.003 |  | -0.005 | -0.003 |  | -0.003 |
|  | (0.065) |  | (0.064) | (0.061) |  | (0.061) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.006 | -0.026 |  | -0.014 | -0.034 |  |
|  | (0.036) | (0.035) |  | (0.034) | (0.033) |  |
| $20000 \leq$ Income $<30000$ | -0.043 | -0.062** |  | -0.046 | -0.066** |  |
|  | (0.030) | (0.030) |  | (0.033) | (0.032) |  |
| $30000 \leq$ Income $<40000$ | -0.030 | -0.051 |  | -0.039 | $-0.062^{*}$ |  |
|  | (0.032) | (0.032) |  | (0.034) | (0.032) |  |
| $40000 \leq$ Income $<50000$ | -0.051 | -0.077** |  | -0.057 | -0.080** |  |
|  | (0.032) | (0.032) |  | (0.035) | (0.034) |  |
| $50000 \leq$ Income $<60000$ | -0.059* | -0.074** |  | -0.060 | -0.074** |  |
|  | (0.033) | (0.034) |  | (0.040) | (0.038) |  |
| Income $\geq 60000$ | 0.005 | -0.020 |  | 0.011 | -0.016 |  |
|  | (0.030) | (0.029) |  | (0.032) | (0.030) |  |
| Did not know/refused | -0.039 | -0.062** |  | -0.032 | -0.062* |  |
|  | (0.031) | (0.030) |  | (0.033) | (0.032) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.273 | $-0.435^{*}$ | -0.301 | -0.353 | -0.518** | -0.366 |
|  | (0.28) | (0.26) | (0.27) | (0.28) | (0.26) | (0.27) |
| $\left(\right.$ Age/100) ${ }^{2}$ | $0.003$ | $0.005^{*}$ | $0.004$ | $0.004$ | $0.006^{* *}$ | $0.004$ |
| Gender 1 if male | -0.017 | -0.015 | -0.020 | -0.017 | -0.014 | -0.019 |
|  | (0.014) | (0.014) | (0.014) | (0.015) | (0.015) | (0.015) |
| Race ${ }^{\text {a }}$ |  |  |  |  |  |  |
| White | -0.014 | -0.019 | -0.013 | -0.014 | -0.018 | -0.011 |
|  | (0.027) | (0.026) | (0.027) | (0.028) | (0.028) | (0.028) |
| Black | 0.007 | 0.015 | 0.002 | 0.009 | 0.017 | 0.004 |
|  | (0.039) | (0.038) | (0.039) | (0.040) | (0.039) | (0.040) |
| Asian | -0.008 | -0.018 | 0.002 | -0.011 | -0.021 | 0.002 |
|  | (0.045) | (0.044) | (0.044) | (0.052) | (0.052) | (0.052) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | 0.007 | 0.014 | 0.011 | 0.008 | 0.015 | 0.011 |
|  | (0.025) | (0.025) | (0.025) | (0.029) | (0.029) | (0.029) |
| Roman Catholic | 0.023 | 0.033 | 0.023 | 0.024 | 0.033 | 0.023 |
|  | (0.024) | (0.024) | (0.024) | (0.027) | (0.027) | (0.028) |
| Protestant | 0.005 | 0.016 | 0.003 | 0.003 | 0.014 | 0.000 |
|  | (0.023) | (0.022) | (0.023) | (0.027) | (0.027) | (0.027) |
| Christian | -0.000 | 0.006 | -0.001 | 0.000 | 0.005 | -0.001 |
|  | (0.023) | (0.023) | (0.024) | (0.027) | (0.027) | (0.027) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | -0.014 | -0.013 | -0.015 | -0.018 | -0.015 | -0.018 |
|  | (0.017) | (0.017) | (0.017) | (0.019) | (0.018) | (0.019) |
| Wave 14 | 0.007 | 0.004 | 0.008 | 0.004 | 0.002 | 0.006 |
|  | (0.016) | (0.016) | (0.016) | (0.018) | (0.017) | (0.018) |
| Intercept | 0.332*** | $0.356^{* * *}$ | 0.322*** | 0.347*** | $0.372^{* * *}$ | 0.330*** |
|  | (0.074) | (0.064) | (0.075) | (0.075) | (0.066) | (0.073) |
| Observations | 907 | 960 | 907 | 907 | 960 | 907 |
| $R^{2}$ | 0.03 | 0.03 | 0.02 |  |  |  |
| Significance of the regression: P-value | 0.21 | 0.07* | 0.53 | 0.32 | 0.10 | 0.63 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E5: Who's who in term of coherence $\left(\alpha_{i, 1000}\right)$

| Education | $\begin{gathered} {[1]} \\ \alpha_{i}^{1000} \\ \hline \end{gathered}$ | OLS $[2]$ $\alpha_{i, 1000}$ | $\begin{gathered} {[3]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[4]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} \text { Tobit } \\ {[5]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[6]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |
| $\leq$ Grade 11 <br> High school graduate | 0.027 | Base group | 0.018 | 0.039 | Base group | 0.028 |
|  | (0.055) |  | (0.055) | (0.066) |  | (0.067) |
| Attended college | -0.018 |  | -0.027 | -0.021 |  | -0.031 |
|  | (0.047) |  | (0.047) | (0.059) |  | (0.059) |
| Associated degree | -0.037 |  | -0.046 | -0.037 |  | -0.047 |
|  | (0.054) |  | (0.053) | (0.067) |  | (0.067) |
| BA/BS | -0.030 |  | -0.038 | -0.027 |  | -0.035 |
|  | (0.048) |  | (0.048) | (0.059) |  | (0.059) |
| MS or PhD | -0.056 |  | -0.064 | -0.045 |  | -0.054 |
|  | (0.053) |  | (0.052) | (0.064) |  | (0.063) |
| Professional | -0.036 |  | -0.026 | -0.030 |  | -0.018 |
|  | (0.070) |  | (0.068) | (0.084) |  | (0.084) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.055 | -0.072** |  | -0.076* | -0.096** |  |
|  | (0.038) | (0.036) |  | (0.046) | (0.044) |  |
| $20000 \leq$ Income $<30000$ | -0.067** | -0.075** |  | -0.089** | -0.096** |  |
|  | (0.033) | (0.031) |  | (0.044) | (0.042) |  |
| $30000 \leq$ Income $<40000$ | -0.052 | -0.069** |  | -0.071* | -0.089** |  |
|  | (0.036) | (0.034) |  | (0.044) | (0.042) |  |
| $40000 \leq$ Income $<50000$ | -0.057 | $-0.078^{* *}$ |  | -0.078* | $-0.103 * *$ |  |
|  | (0.038) | (0.036) |  | (0.046) | (0.044) |  |
| $50000 \leq$ Income $<60000$ | -0.072* | -0.081** |  | -0.089* | $-0.097^{* *}$ |  |
|  | (0.040) | (0.038) |  | (0.052) | (0.049) |  |
| Income $\geq 60000$ | -0.006 | -0.013 |  | 0.005 | -0.014 |  |
|  | (0.034) | (0.032) |  | (0.041) | (0.039) |  |
| Did not know/refused | -0.057 | $-0.075^{* *}$ |  | -0.065 | $-0.084^{* *}$ |  |
|  | (0.037) | $(0.034)$ |  | (0.043) | $(0.041)$ |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.166 | -0.351 | -0.137 | -0.239 | -0.426 | -0.202 |
|  | (0.31) | (0.28) | $(0.31)$ | (0.37) | (0.34) | (0.36) |
| $(\text { Age } / 100)^{2}$ | $0.003$ | $0.005$ | $0.003$ | $0.003$ | $0.005$ | $0.003$ |
|  | $(0.003)$ | $(0.003)$ | (0.003) | $(0.003)$ | $(0.003)$ | $(0.003)$ |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.011 | -0.008 | -0.010 | -0.012 | -0.009 | -0.011 |
|  | (0.017) | (0.017) | (0.017) | (0.020) | (0.020) | (0.020) |
| Race |  |  |  |  |  |  |
| White | -0.021 | -0.037 | -0.018 | -0.016 | -0.036 | -0.013 |
|  | (0.035) | (0.034) | (0.035) | (0.037) | (0.036) | (0.037) |
| Black | 0.002 | -0.003 | -0.004 | 0.001 | -0.004 | -0.006 |
|  | (0.043) | (0.041) | (0.043) | (0.052) | (0.051) | (0.053) |
| Asian | 0.031 | 0.004 | 0.046 | 0.049 | 0.023 | 0.069 |
|  | (0.060) | (0.057) | (0.058) | (0.069) | $(0.068)$ | (0.069) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | 0.006 | 0.017 | 0.007 | 0.015 | 0.031 | 0.015 |
|  | (0.032) | (0.031) | (0.032) | (0.038) | (0.038) | (0.038) |
| Roman Catholic | 0.046 | 0.055* | 0.046 | 0.061* | 0.072** | 0.059 |
|  | (0.031) | (0.030) | (0.032) | (0.036) | (0.036) | (0.036) |
| Protestant | 0.021 | 0.039 | 0.017 | 0.029 | 0.053 | 0.025 |
|  | (0.030) | (0.029) | (0.031) | (0.036) | (0.035) | (0.036) |
| Christian | 0.010 | 0.020 | 0.007 | 0.017 | 0.030 | 0.013 |
|  | (0.030) | (0.029) | (0.030) | (0.036) | (0.035) | (0.036) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | -0.045** | -0.039** | -0.045** | -0.055** | -0.047** | -0.056** |
|  | (0.021) | $(0.019)$ | $(0.021)$ | $(0.023)$ | (0.022) | (0.024) |
| Wave 14 | 0.002 | 0.004 | 0.005 | 0.000 | 0.005 | 0.005 |
|  | (0.021) | (0.019) | (0.021) | (0.023) | (0.022) | (0.024) |
| Intercept | 0.378*** | 0.409*** | 0.336*** | 0.378*** | $0.410^{* * *}$ | $0.321^{* * *}$ |
|  | (0.083) | (0.070) | (0.082) | (0.100) | (0.086) | (0.098) |
| Observations | 848 | 900 | 848 | 848 | 900 | 848 |
| $R^{2}$ | 0.04 | 0.04 | 0.03 |  |  |  |
| Significance of the regression: P -value | 0.03** | 0.00*** | 0.13 | 0.09* | 0.01** | 0.36 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P-value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E6: Who's who in term of coherence $(\bar{d})$

|  | OLS |  |  | Tobit |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\frac{[1]}{d}$ | $\stackrel{[2]}{\bar{d}}$ | $\frac{[3]}{\bar{d}}$ | $\stackrel{[4]}{\bar{d}}$ | $\stackrel{[5]}{\bar{d}}$ | $\stackrel{[6]}{\bar{d}}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.046 |  | -0.050 | -0.046 |  | -0.050 |
|  | (0.047) |  | (0.047) | (0.034) |  | (0.034) |
| Attended college | -0.061 |  | -0.069 | -0.061** |  | -0.069** |
|  | (0.043) |  | (0.043) | (0.030) |  | (0.030) |
| Associated degree | -0.076 |  | -0.085* | -0.075** |  | -0.085** |
|  | (0.048) |  | (0.048) | (0.034) |  | (0.034) |
| BA/BS | -0.088** |  | $-0.097 * *$ | $-0.088^{* * *}$ |  | $-0.097^{* * *}$ |
|  | (0.043) |  | (0.043) | (0.030) |  | (0.030) |
| MS or PhD | -0.091** |  | -0.101** | -0.091*** |  | -0.101*** |
|  | (0.044) |  | (0.044) | (0.033) |  | (0.032) |
| Professional | -0.039 |  | -0.043 | -0.040 |  | -0.043 |
|  | (0.056) |  | (0.056) | (0.043) |  | (0.043) |
| Income |  |  |  |  |  |  |
| Income $<10000$ |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.012 | -0.033 |  | -0.012 | -0.033 |  |
|  | (0.029) | (0.028) |  | (0.024) | (0.023) |  |
| $20000 \leq$ Income $<30000$ | -0.040 | $-0.054^{* *}$ |  | -0.040* | -0.054** |  |
|  | (0.028) | (0.027) |  | (0.023) | (0.022) |  |
| $30000 \leq$ Income $<40000$ | -0.025 | -0.047* |  | -0.025 | -0.047** |  |
|  | (0.027) | (0.027) |  | (0.023) | (0.022) |  |
| $40000 \leq$ Income $<50000$ | -0.047* | -0.069*** |  | -0.047* | -0.069*** |  |
|  | (0.027) | (0.026) |  | (0.024) | (0.023) |  |
| $50000 \leq$ Income $<60000$ | -0.058** | -0.079*** |  | -0.058** | -0.079*** |  |
|  | (0.028) | (0.027) |  | (0.027) | (0.026) |  |
| Income $\geq 60000$ | -0.017 | -0.040 |  | -0.016 | -0.040* |  |
|  | (0.026) | (0.025) |  | (0.022) | (0.021) |  |
| Did not know/refused | -0.040 | -0.066** |  | -0.041* | $-0.067 * * *$ |  |
|  | (0.027) | (0.026) |  | (0.023) | (0.022) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.252 | -0.462** |  | -0.249 | -0.460** |  |
|  | (0.19) | (0.19) | (0.19) | (0.19) | (0.18) | (0.19) |
| $(\text { Age } / 100)^{2}$ | 0.002 | $0.004^{* *}$ | 0.002 | 0.002 | $0.004^{* *}$ | $0.003$ |
|  | (0.002) | $(0.002)$ | (0.002) | (0.002) | $(0.002)$ | $(0.002)$ |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.024** | -0.018* | $-0.028^{* * *}$ | -0.024** | -0.018* | -0.028*** |
|  | (0.011) | (0.010) | (0.010) | (0.011) | (0.010) | (0.010) |
| Race ${ }^{\text {c }}$ |  |  |  |  |  |  |
| White | -0.001 | -0.007 | -0.001 | -0.001 | -0.007 | -0.001 |
|  | (0.021) | (0.021) | (0.021) | (0.019) | (0.019) | (0.019) |
| Black | 0.048 | 0.047 | 0.045 | 0.048* | 0.047* | 0.045 |
|  | (0.032) | (0.030) | (0.032) | (0.028) | (0.027) | (0.028) |
| Asian | -0.006 | -0.017 | -0.002 | -0.006 | -0.017 | -0.002 |
|  | (0.032) | (0.031) | (0.031) | (0.036) | (0.036) | (0.036) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | 0.005 | 0.010 | 0.008 | 0.005 | 0.010 | 0.008 |
|  | (0.021) | (0.020) | (0.021) | (0.020) | (0.020) | (0.020) |
| Roman Catholic | 0.026 | 0.037* | 0.027 | 0.026 | 0.037* | 0.027 |
|  | (0.019) | (0.019) | (0.019) | (0.019) | (0.019) | (0.019) |
| Protestant | 0.015 | 0.024 | 0.015 | 0.015 | 0.024 | 0.015 |
|  | (0.018) | (0.017) | (0.018) | (0.019) | (0.18) | (0.018) |
| Christian | 0.005 | 0.010 | 0.005 | 0.005 | 0.010 | 0.005 |
|  | (0.018) | (0.018) | (0.018) | (0.019) | (0.019) | (0.019) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.014 | 0.010 | 0.012 | 0.014 | 0.010 | 0.012 |
|  | (0.013) | (0.018) | (0.013) | (0.013) | (0.013) | (0.013) |
| Wave 14 | 0.028** | 0.020* | 0.029** | 0.028* | 0.020* | 0.029** |
|  | (0.012) | (0.011) | (0.012) | (0.012) | (0.012) | (0.012) |
| Intercept | 0.415*** | $0.409^{* * *}$ | 0.409*** | $0.415^{* * *}$ | 0.409*** | 0.408*** |
|  | (0.062) | (0.050) | (0.068) | (0.052) | (0.046) | (0.051) |
| Observations | 907 | 960 | 907 | 907 | 960 | 907 |
| $R^{2}$ | 0.06 | 0.05 | 0.05 |  |  |  |
| Significance of the regression: P-value | 0.00*** | 0.00*** | 0.00*** | $0.00^{* * *}$ | $0.00^{* * *}$ | $0.00^{* * *}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table E7: Who's who in term of coherence ( $\left.d^{\text {Kol }}\right)$


Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P-value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihood ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F1: Who's who in term of coherence $\left(d^{98}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ d^{98} \end{gathered}$ | $[2]$ $d^{98}$ | $\begin{gathered} {[3]} \\ d^{98} \end{gathered}$ | $\begin{gathered} {[4]} \\ d^{98} \end{gathered}$ | $\begin{gathered} {[5]} \\ d^{98} \end{gathered}$ | $\begin{gathered} {[6]} \\ d^{98} \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.049 |  | -0.049 | -0.044 |  | -0.044 |
|  | (0.052) |  | (0.051) | (0.042) |  | (0.042) |
| Attended college | -0.115** |  | $-0.122^{* * *}$ | -0.118*** |  | -0.126*** |
|  | (0.046) |  | (0.045) | (0.036) |  | (0.036) |
| Associated degree | -0.112** |  | -0.122** | -0.112*** |  | -0.123*** |
|  | (0.050) |  | (0.049) | (0.042) |  | (0.042) |
| BA/BS | -0.139*** |  | $-0.148^{* * *}$ | -0.138*** |  | -0.148*** |
|  | (0.046) |  | (0.045) | (0.037) |  | (0.037) |
| MS or PhD | -0.154*** |  | -0.165*** | -0.149*** |  | -0.160*** |
|  | (0.048) |  | (0.047) | (0.041) |  | (0.040) |
| Professional | -0.094 |  | -0.095 | -0.095* |  | -0.095* |
|  | (0.068) |  | (0.067) | (0.057) |  | (0.057) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | 0.023 | 0.001 |  | 0.014 | -0.007 |  |
|  | (0.034) | (0.033) |  | (0.033) | (0.032) |  |
| $20000 \leq$ Income $<30000$ | -0.006 | -0.033 |  | -0.011 | -0.038 |  |
|  | (0.031) | (0.030) |  | (0.032) | (0.031) |  |
| $30000 \leq$ Income $<40000$ | -0.013 | -0.045 |  | -0.024 | -0.057* |  |
|  | (0.032) | (0.031) |  | (0.032) | (0.032) |  |
| $40000 \leq$ Income $<50000$ | -0.034 | -0.066** |  | -0.042 | -0.073** |  |
|  | (0.031) | (0.031) |  | (0.031) | (0.031) |  |
| $50000 \leq$ Income $<60000$ | -0.057* | -0.087*** |  | -0.065* | -0.094** |  |
|  | (0.030) | (0.031) |  | (0.038) | (0.038) |  |
| Income $\geq 60000$ | 0.004 | -0.032 |  | 0.004 | -0.033 |  |
|  | (0.028) | (0.028) |  | (0.030) | (0.030) |  |
| Did not know/refused | -0.035 | -0.068** |  | -0.038 | $-0.071^{* *}$ |  |
|  | (0.029) | (0.028) |  | (0.032) | $(0.031)$ |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.190 | -0.416 | -0.270 | -0.215 | -0.447* | -0.293 |
|  | (0.27) | (0.26) | (0.26) | (0.26) | (0.25) | (0.26) |
| $\left(\right.$ Age /100) ${ }^{2}$ | 0.003 | 0.005* | 0.003 | 0.003 | 0.005* | 0.003 |
|  | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) | (0.003) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.021 | -0.011 | -0.026* | -0.017 | -0.008 | -0.022 |
|  | (0.014) | (0.014) | (0.014) | (0.015) | (0.015) | (0.015) |
| Race ${ }^{\text {c }}$ |  |  |  |  |  |  |
| White | -0.010 | -0.022 | -0.009 | -0.006 | -0.019 | -0.004 |
|  | (0.025) | (0.026) | (0.025) | (0.027) | (0.027) | (0.027) |
| Black | 0.066 | 0.071* | 0.064 | 0.072* | 0.078** | 0.070* |
|  | (0.041) | (0.040) | (0.041) | (0.038) | (0.037) | (0.038) |
| Asian | 0.043 | 0.019 | 0.048 | 0.046 | 0.023 | 0.053 |
|  | (0.052) | (0.052) | (0.051) | (0.051) | (0.052) | (0.052) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | -0.015 | -0.003 | -0.011 | -0.014 | -0.003 | -0.010 |
|  | (0.025) | (0.025) | (0.025) | (0.028) | (0.028) | (0.028) |
| Roman Catholic | -0.000 | 0.017 | 0.000 | -0.004 | 0.012 | -0.004 |
|  | (0.024) | (0.024) | (0.024) | (0.027) | (0.027) | (0.027) |
| Protestant | -0.021 | -0.001 | -0.020 | -0.025 | -0.006 | -0.025 |
|  | (0.024) | (0.023) | (0.024) | (0.026) | (0.026) | (0.026) |
| Christian | -0.014 | -0.000 | -0.012 | -0.017 | -0.005 | -0.016 |
|  | (0.025) | (0.024) | (0.025) | (0.027) | (0.027) | (0.027) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.006 | 0.007 | 0.004 | 0.005 | 0.006 | 0.003 |
|  | (0.017) | (0.017) | (0.017) | (0.019) | (0.018) | (0.019) |
| Wave 14 | 0.006 | 0.002 | 0.006 | 0.007 | 0.003 | 0.007 |
|  | (0.016) | (0.015) | (0.016) | (0.017) | (0.017) | (0.017) |
| Intercept | 0.405*** | 0.365*** | 0.420*** | 0.408*** | 0.369*** | $0.417^{* * *}$ |
|  | (0.073) | (0.062) | (0.071) | (0.070) | (0.064) | (0.068) |
| Observations | 1044 | 1105 | 1044 | 1044 | 1105 | 1044 |
| $R^{2}$ | 0.06 | 0.04 | 0.05 |  |  |  |
| Significance of the regression: P -value | $0.00^{* * *}$ | 0.00*** | $0.00^{* * *}$ | 0.00*** | 0.01** | $0.00^{* * *}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likeliぬ $\mathrm{b}^{\circ}$ od ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F2: Who's who in term of coherence $\left(d^{90}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ d^{90} \end{gathered}$ | $\begin{gathered} {[2]} \\ d^{90} \end{gathered}$ | $\begin{gathered} {[3]} \\ d^{90} \end{gathered}$ | $\begin{gathered} {[4]} \\ d^{90} \end{gathered}$ | $\begin{gathered} {[5]} \\ d^{90} \\ \hline \end{gathered}$ | $\begin{gathered} {[6]} \\ d^{90} \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.046 |  | -0.046 | -0.038 |  | -0.038 |
|  | (0.050) |  | (0.050) | (0.039) |  | (0.039) |
| Attended college | $-0.112^{* *}$ |  | $-0.121^{* * *}$ | $-0.107 * * *$ |  | -0.115*** |
|  | (0.044) |  | (0.044) | (0.034) |  | (0.034) |
| Associated degree | -0.109** |  | -0.121** | -0.104*** |  | -0.115*** |
|  | (0.048) |  | (0.047) | (0.039) |  | (0.039) |
| BA/BS | -0.132*** |  | $-0.143^{* * *}$ | -0.128*** |  | -0.138*** |
|  | (0.044) |  | (0.044) | (0.034) |  | (0.034) |
| MS or PhD | -0.155*** |  | $-0.167^{* * *}$ | -0.153*** |  | -0.164*** |
|  | (0.046) |  | (0.045) | (0.038) |  | (0.038) |
| Professional | -0.082 |  | -0.084 | -0.077 |  | -0.079 |
|  | (0.066) |  | (0.064) | (0.052) |  | (0.052) |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | 0.028 | 0.007 |  | 0.026 | 0.0042 |  |
|  | (0.033) | (0.032) |  | (0.030) | (0.030) |  |
| $20000 \leq$ Income $<30000$ | -0.000 | -0.027 |  | -0.003 | -0.031 |  |
|  | (0.030) | (0.029) |  | (0.030) | (0.029) |  |
| $30000 \leq$ Income $<40000$ | -0.006 | -0.038 |  | -0.010 | -0.042 |  |
|  | (0.031) | (0.030) |  | (0.030) | (0.029) |  |
| $40000 \leq$ Income $<50000$ | -0.037 | -0.067** |  | -0.036 | -0.066** |  |
|  | (0.030) | (0.030) |  | (0.032) | (0.031) |  |
| $50000 \leq$ Income $<60000$ | -0.056* | -0.086*** |  | -0.053 | -0.085** |  |
|  | (0.029) | (0.030) |  | (0.036) | (0.035) |  |
| Income $\geq 60000$ | 0.003 | -0.032 |  | 0.003 | -0.034 |  |
|  | (0.027) | (0.026) |  | (0.028) | (0.027) |  |
| Did not know/refused | -0.038 | -0.072*** |  | -0.037 | $-0.071^{* *}$ |  |
|  | (0.027) | (0.027) |  | (0.030) | $(0.029)$ |  |
| Age ${ }^{\text {a }}$ |  |  |  |  |  |  |
| Age/100 | -0.143 | -0.360 | -0.233 | -0.132 | -0.349 | -0.215 |
|  | (0.26) | (0.25) | (0.25) | (0.24) | (0.23) | (0.24) |
| $(\text { Age } / 100)^{2}$ | 0.002 | 0.004* | 0.003 | 0.002 | 0.004* | 0.003 |
|  | (0.003) | (0.003) | (0.003) | (0.002) | (0.002) | (0.002) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.019 | -0.009 | -0.025* | -0.019 | -0.009 | -0.024* |
|  | (0.013) | (0.013) | (0.013) | (0.014) | (0.014) | (0.013) |
| Race |  |  |  |  |  |  |
| White | -0.011 | -0.023 | -0.010 | -0.008 | -0.019 | -0.007 |
|  | (0.024) | (0.025) | (0.024) | (0.025) | (0.025) | (0.025) |
| Black | 0.067 | 0.073* | 0.065 | 0.071** | 0.078** | 0.069** |
|  | (0.040) | (0.039) | (0.040) | (0.035) | (0.035) | (0.035) |
| Asian | 0.038 | 0.014 | 0.042 | 0.041 | 0.016 | 0.045 |
|  | (0.052) | (0.052) | (0.051) | (0.048) | (0.048) | (0.048) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | $-0.013$ | -0.004 | -0.008 | -0.011 | -0.002 | -0.007 |
|  | (0.024) | (0.024) | (0.024) | (0.026) | (0.026) | (0.026) |
| Roman Catholic | 0.003 | 0.019 | $0.004$ | -0.000 | 0.017 | 0.000 |
|  | (0.023) | (0.023) | $(0.023)$ | (0.025) | (0.025) | (0.025) |
| Christian | -0.007 | 0.005 | -0.005 | -0.009 | 0.003 | -0.007 |
|  | (0.024) | (0.024) | (0.024) | (0.025) | (0.025) | (0.025) |
| Protestant | -0.017 | 0.001 | -0.015 | -0.019 | -0.001 | -0.0177 |
|  | (0.023) | (0.022) | (0.023) | (0.024) | (0.024) | (0.024) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.007 | 0.007 | 0.005 | 0.008 | 0.008 | 0.007 |
|  | (0.017) | (0.017) | (0.016) | (0.018) | (0.017) | (0.018) |
| Wave 14 | 0.006 | 0.002 | 0.006 | 0.008 | 0.004 | 0.009 |
|  | (0.015) | (0.015) | (0.015) | (0.016) | (0.016) | (0.016) |
| Intercept | $0.366^{* * *}$ | $0.327^{* * *}$ | $0.386^{* * *}$ | $0.356^{* * *}$ | 0.322*** | $0.373^{* * *}$ |
|  | (0.071) | (0.060) | (0.069) | (0.065) | (0.059) | (0.063) |
| Observations | 1044 | 1105 | 1044 | 1044 | 1105 | 1044 |
| $R^{2}$ | 0.07 | 0.04 | 0.05 |  |  |  |
| Significance of the regression: P-value | 0.00 *** | 0.00*** | $0.00^{* * *}$ | 0.00*** | $0.00^{* * *}$ | $0.00^{* * *}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likeliłøod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F3: Who's who in term of coherence $\left(d^{80}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[2]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[3]} \\ d^{80} \end{gathered}$ | $\begin{gathered} {[4]} \\ d^{80} \end{gathered}$ | $\begin{array}{r} {[5]} \\ d^{80} \end{array}$ | $\begin{gathered} {[6]} \\ d^{80} \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | $-0.048$ |  | -0.048 | $-0.048$ |  | -0.047 |
|  | (0.049) |  | (0.048) | (0.037) |  | (0.038) |
| Attended college | $-0.112^{* * *}$ |  | $-0.120 * * *$ | -0.111*** |  | -0.119*** |
|  | (0.042) |  | (0.042) | (0.032) |  | (0.032) |
| Associated degree | -0.108** |  | $-0.119^{* * *}$ | -0.105*** |  | -0.117*** |
|  | (0.046) |  | (0.045) | (0.037) |  | (0.037) |
| BA/BS | -0.126*** |  | -0.136*** | -0.123*** |  | -0.134*** |
|  | (0.042) |  | (0.042) | (0.033) |  | (0.033) |
| MS or PhD | -0.150*** |  | $-0.162^{* * *}$ | -0.148*** |  | -0.160*** |
|  | (0.044) |  | (0.043) | (0.036) |  | (0.036) |
| Professional |  |  |  |  |  | -0.073 |
|  | $(0.063)$ |  | $(0.062)$ | $(0.050)$ |  | (0.050) |
| Income |  |  |  |  |  |  |
| Income $<10000$ |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | 0.032 | 0.012 |  | 0.026 | 0.006 |  |
|  | (0.032) | (0.031) |  | (0.029) | (0.028) |  |
| $20000 \leq$ Income $<30000$ | 0.004 | -0.021 |  | 0.000 | -0.025 |  |
|  | (0.029) | (0.028) |  | (0.028) | (0.028) |  |
| $30000 \leq$ Income $<40000$ | -0.000 | -0.030 |  | -0.007 | -0.036 |  |
|  | (0.030) | (0.029) |  | (0.029) | (0.028) |  |
| $40000 \leq$ Income $<50000$ | -0.039 | -0.066** |  | -0.042 | -0.069** |  |
|  | (0.028) | (0.028) |  | (0.030) | (0.030) |  |
| $50000 \leq$ Income $<60000$ | $-0.053^{*}$ | $-0.081^{* * *}$ |  | $-0.055$ | $-0.082^{* *}$ |  |
|  | $(0.027)$ | $(0.029)$ |  | $(0.034)$ | $(0.033)$ |  |
| Income $\geq 60000$ | 0.005 | -0.028 |  | 0.002 | -0.032 |  |
|  | (0.026) | (0.026) |  | (0.027) | (0.026) |  |
| Did not know/refused | -0.039 | -0.071*** |  | -0.044 | -0.076*** |  |
|  | (0.026) | (0.026) |  | (0.028) | (0.028) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.129 | -0.326 | -0.222 | -0.117 | -0.309 | -0.210 |
|  | $(0.25)$ | (0.24) | (0.24) | (0.23) | (0.22) | (0.23) |
| $(\text { Age } / 100)^{2}$ | $0.002$ | $0.004^{*}$ | $0.003$ | $0.002$ | $0.004^{*}$ | $0.003$ |
|  | $(0.003)$ | $(0.003)$ | $(0.003)$ | $(0.002)$ | $(0.002)$ | (0.002) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.015 | -0.006 | -0.021* | -0.014 | -0.005 | -0.019 |
|  | (0.013) | (0.013) | (0.012) | (0.013) | (0.013) | (0.013) |
| Race |  |  |  |  |  |  |
| White | -0.010 | -0.021 | -0.010 | -0.012 | -0.021 | -0.011 |
|  | (0.023) | (0.024) | (0.023) | (0.024) | (0.024) | (0.024) |
| Black | 0.069* | 0.075** | 0.067* | 0.073** | 0.079** | 0.071** |
|  | (0.038) | (0.037) | (0.038) | (0.033) | (0.033) | (0.033) |
| Asian | 0.031 | 0.008 | 0.034 | 0.028 | 0.007 | 0.032 |
|  | (0.050) | (0.051) | (0.049) | (0.045) | (0.046) | (0.046) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | -0.013 | -0.005 | -0.008 | -0.012 | -0.005 | -0.007 |
|  | (0.023) | (0.023) | (0.023) | (0.025) | (0.025) | (0.025) |
| Roman Catholic | 0.003 | 0.018 | 0.005 | 0.000 | 0.015 | 0.002 |
|  | (0.022) | (0.022) | (0.022) | (0.024) | (0.024) | (0.024) |
| Christian | -0.004 | 0.007 | -0.000 | -0.007 | 0.003 | -0.004 |
|  | (0.023) | (0.022) | (0.023) | (0.024) | (0.024) | (0.024) |
| Protestant | -0.012 | 0.001 | -0.010 | -0.016 | -0.002 | -0.014 |
|  | (0.022) | (0.021) | (0.022) | (0.023) | (0.023) | (0.023) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.008 | 0.008 | 0.006 | 0.008 | 0.007 | 0.007 |
|  | (0.016) | (0.016) | (0.016) | (0.017) | (0.016) | (0.017) |
| Wave 14 | 0.007 | 0.003 | 0.008 | 0.008 | 0.003 | 0.009 |
|  | (0.014) | (0.014) | (0.014) | (0.015) | (0.015) | (0.015) |
| Intercept | 0.338*** | 0.298*** | $0.361 * * *$ | $0.337^{* * *}$ | $0.297 * * *$ | $0.356^{* * *}$ |
|  | (0.068) | (0.058) | (0.067) | (0.062) | (0.057) | (0.060) |
| Observations | 1044 | 1105 | 1044 | 1044 | 1105 | 1044 |
| $R^{2}$ | 0.07 | 0.04 | 0.06 |  |  |  |
| Significance of the regression: P-value | 0.00 *** | $0.00^{* * *}$ | 0.00*** | 0.00*** | 0.00*** | 0.00*** |

Notes: i. $*, * *$ and $* * *$ represent 10, 5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelidgod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F4: Who's who in term of coherence $\left(\alpha_{i}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | [1] | [2] | [3] | [4] | [5] | [6] |
|  | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ | $\alpha_{i}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.044 |  | -0.047 | -0.041 |  | -0.044 |
|  | (0.051) |  | (0.051) | (0.051) |  | (0.051) |
| Attended college | -0.110** |  | -0.120*** | -0.121*** |  | -0.132*** |
|  | (0.047) |  | (0.046) | (0.044) |  | (0.044) |
| Associated degree | -0.087* |  | -0.101** | -0.089* |  | -0.106** |
|  | (0.052) |  | (0.051) | (0.051) |  | (0.051) |
| BA/BS | -0.151*** |  | $-0.163^{* * *}$ | -0.159*** |  | -0.173*** |
|  | (0.047) |  | (0.046) | (0.044) |  | (0.044) |
| MS or PhD | -0.171*** |  | -0.185*** | -0.166*** |  | -0.182*** |
|  | (0.050) |  | (0.048) | (0.049) |  | (0.049) |
| Professional | $-0.124^{*}$ |  | $-0.126^{*}$ | $-0.122^{*}$ |  |  |
|  | $(0.067)$ |  | $(0.065)$ | $(0.068)$ |  | $(0.068)$ |
| Income |  |  |  |  |  |  |
| Income<10000 |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.004 | -0.024 |  | -0.015 | -0.038 |  |
|  | (0.035) | (0.033) |  | (0.039) | (0.038) |  |
| $20000 \leq$ Income $<30000$ | -0.031 | -0.049* |  | -0.045 | -0.064* |  |
|  | (0.032) | (0.030) |  | (0.038) | (0.037) |  |
| $30000 \leq$ Income $<40000$ | -0.033 | -0.062** |  | -0.051 | -0.081** |  |
|  | (0.032) | (0.031) |  | (0.038) | (0.037) |  |
| $40000 \leq$ Income $<50000$ | -0.065* | -0.091*** |  | -0.087** | $-0.117^{* * *}$ |  |
|  | (0.033) | (0.032) |  | (0.040) | (0.039) |  |
| $50000 \leq$ Income $<60000$ | -0.071** | -0.099*** |  | -0.091** | -0.119*** |  |
|  | (0.036) | (0.034) |  | (0.046) | (0.044) |  |
| Income $\geq 60000$ | -0.019 | -0.052* |  | -0.025 | -0.058* |  |
|  | (0.030) | (0.028) |  | (0.036) | (0.035) |  |
| Did not know/refused | -0.063** | -0.089*** |  | -0.072* | -0.099*** |  |
|  | (0.031) | (0.029) |  | (0.038) | (0.036) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.054 | -0.324 | -0.149 | -0.072 | -0.351 | -0.190 |
|  | (0.28) | (0.26) | (0.27) | (0.31) | (0.30) | (0.31) |
| $(\text { Age } / 100)^{2}$ |  | $0.004$ | $0.002$ | $0.001$ | $0.004$ | $0.003$ |
|  | $(0.003)$ | $(0.003)$ | (0.003) | $(0.003)$ | $(0.003)$ | (0.003) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.019 | -0.009 | -0.026* | -0.018 | -0.007 | -0.025 |
|  | (0.015) | (0.015) | (0.015) | (0.017) | (0.017) | (0.017) |
| Race ${ }^{\text {a }}$ |  |  |  |  |  |  |
| White | -0.041 | -0.060** | -0.039 | -0.036 | -0.060* | -0.034 |
|  | (0.028) | (0.028) | (0.028) | (0.032) | (0.032) | (0.032) |
| Black | 0.027 | 0.019 | 0.025 | 0.029 | 0.020 | 0.028 |
|  | (0.041) | (0.039) | (0.041) | (0.045) | (0.044) | (0.045) |
| Asian | 0.031 | -0.008 | 0.037 | 0.041 | 0.002 | 0.049 |
|  | (0.057) | (0.056) | (0.056) | (0.062) | (0.062) | (0.062) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | $-0.007$ | $0.006$ | $-0.003$ | $-0.009$ | $0.008$ | $-0.005$ |
|  | $(0.027)$ | $(0.027)$ | $(0.027)$ | $(0.033)$ | $(0.033)$ | $(0.033)$ |
| Roman Catholic | 0.009 | 0.023 | 0.010 | 0.011 | 0.028 | 0.012 |
|  | (0.026) | (0.026) | (0.026) | (0.032) | (0.032) | (0.032) |
| Christian | -0.012 | 0.007 | -0.010 | -0.012 | 0.008 | -0.012 |
|  | (0.026) | (0.026) | (0.026) | (0.032) | (0.032) | (0.032) |
| Protestant | -0.020 | 0.002 | -0.018 | -0.020 | 0.0063 | -0.019 |
|  | (0.025) | (0.025) | (0.025) | (0.031) | (0.031) | (0.031) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | -0.019 | -0.016 | -0.021 | -0.026 | -0.022 | -0.029 |
|  | (0.019) | (0.019) | (0.019) | (0.022) | (0.022) | (0.023) |
| Wave 14 | -0.014 | -0.015 | -0.013 | -0.016 | -0.015 | -0.014 |
|  | (0.018) | (0.017) | (0.018) | (0.021) | (0.020) | (0.021) |
| Intercept | $0.483^{* * *}$ | $0.448^{* * *}$ | $0.482^{* * *}$ | $0.486^{* * *}$ | $0.446^{* * *}$ | $0.479^{* * *}$ |
|  | (0.078) | (0.064) | (0.075) | (0.083) | (0.076) | (0.081) |
| Observations | 1025 | 1084 | 1025 | 1025 | 1084 | 1025 |
| $R^{2}$ | 0.06 | 0.03 | 0.05 |  |  |  |
| Significance of the regression: P -value | 0.00*** | 0.01** | 0.00*** | 0.00*** | 0.02** | 0.00*** |

Notes: i. $*, * *$ and $* * *$ represent 10, 5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelidgod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F5: Who's who in term of coherence $\left(d_{1000}\right)$-All respondents for whom we can compute a measure of coherence-


Notes: i. $*, * *$ and $* * *$ represent 10, 5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelibqod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F6: Who's who in term of coherence $\left(\alpha_{i, 1000}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[2]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[3]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[4]} \\ \alpha_{i, 1000} \end{gathered}$ | $\begin{gathered} {[5]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ | $\begin{gathered} {[6]} \\ \alpha_{i, 1000} \\ \hline \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | 0.003 |  | -0.003 | 0.012 |  | 0.005 |
|  | (0.049) |  | (0.049) | (0.057) |  | (0.057) |
| Attended college | -0.071* |  | -0.081* | -0.083* |  | -0.094* |
|  | (0.043) |  | (0.042) | (0.049) |  | (0.049) |
| Associated degree | -0.072 |  | -0.085* | -0.077 |  | -0.093 |
|  | (0.048) |  | (0.047) | (0.057) |  | (0.057) |
| BA/BS | -0.086** |  | -0.097** | -0.090* |  | -0.101** |
|  | (0.043) |  | (0.042) | (0.050) |  | (0.049) |
| MS or PhD | -0.094* |  | -0.107** | -0.089 |  | -0.104* |
|  | (0.048) |  | (0.047) | (0.055) |  | (0.055) |
| Professional | -0.063 |  | -0.056 | -0.062 |  | -0.054 |
|  | (0.065) |  | (0.063) | (0.077) |  | (0.077) |
| Income |  |  |  |  |  |  |
| Income $<10000$ |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.037 | -0.045 |  | -0.051 | -0.062 |  |
|  | (0.037) | (0.035) |  | (0.044) | (0.042) |  |
| $20000 \leq$ Income $<30000$ | -0.059* | -0.068** |  | -0.078* | -0.086** |  |
|  | (0.033) | (0.031) |  | (0.042) | (0.040) |  |
| $30000 \leq$ Income $<40000$ | -0.052 | $-0.067^{* *}$ |  | -0.072* | $-0.087^{* *}$ |  |
|  | (0.036) | (0.033) |  | (0.043) | (0.041) |  |
| $40000 \leq$ Income $<50000$ | -0.086** | -0.098*** |  | -0.111** | $-0.127^{* * *}$ |  |
|  | (0.037) | (0.035) |  | (0.045) | (0.044) |  |
| $50000 \leq$ Income $<60000$ | -0.078* | $-0.087 * *$ |  | -0.097* | -0.106** |  |
|  | (0.042) | (0.039) |  | (0.052) | (0.049) |  |
| Income $\geq 60000$ | -0.009 | -0.026 |  | -0.012 | -0.029 |  |
|  | (0.033) | (0.031) |  | (0.040) | (0.038) |  |
| Did not know/refused | -0.075** | -0.091*** |  | -0.085** | -0.101** |  |
|  | (0.036) | (0.033) |  | (0.043) | (0.040) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.066 | -0.272 | -0.110 | -0.097 | -0.311 | -0.150 |
|  | (0.31) | (0.28) | (0.30) | (0.36) | (0.33) | (0.35) |
| $\left(\right.$ Age /100) ${ }^{2}$ | 0.002 | 0.004 | 0.002 | 0.002 | 0.004 | 0.002 |
|  | (0.003) | (0.003) | (0.003) | (0.004) | (0.003) | (0.004) |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.000 | 0.004 | -0.005 | 0.001 | 0.006 | -0.004 |
|  | (0.017) | (0.016) | (0.017) | (0.020) | (0.019) | (0.020) |
| Race |  |  |  |  |  |  |
| White | -0.028 | -0.042 | -0.025 | -0.026 | -0.044 | -0.022 |
|  | (0.033) | (0.032) | (0.033) | (0.036) | (0.035) | (0.036) |
| Black | 0.037 | 0.029 | 0.034 | 0.037 | 0.028 | 0.033 |
|  | (0.044) | (0.042) | (0.044) | (0.050) | (0.048) | (0.050) |
| Asian | 0.001 | -0.023 | 0.015 | 0.010 | -0.012 | 0.029 |
|  | (0.059) | (0.057) | (0.057) | (0.069) | (0.068) | (0.070) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | -0.016 | -0.006 | -0.013 | -0.014 | 0.002 | -0.010 |
|  | (0.032) | (0.031) | (0.032) | (0.038) | (0.037) | (0.038) |
| Roman Catholic | 0.030 | 0.038 | 0.032 | 0.037 | 0.048 | 0.038 |
|  | (0.031) | (0.030) | (0.031) | (0.036) | (0.035) | (0.036) |
| Christian | 0.003 | 0.013 | 0.004 | 0.007 | 0.018 | 0.007 |
|  | (0.030) | (0.030) | (0.031) | (0.036) | (0.035) | (0.036) |
| Protestant | 0.009 | 0.027 | 0.009 | 0.016 | 0.038 | 0.015 |
|  | (0.030) | (0.029) | (0.031) | (0.035) | (0.035) | (0.035) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | -0.033 | -0.023 | -0.035* | -0.042* | -0.030 | -0.045* |
|  | (0.021) | (0.020) | (0.021) | (0.025) | (0.024) | (0.025) |
| Wave 14 | -0.005 | 0.001 | -0.002 | -0.007 | 0.001 | -0.004 |
|  | (0.020) | (0.019) | (0.021) | (0.023) | (0.022) | (0.023) |
| Intercept | $0.436 * * *$ | $0.421^{* * *}$ | $0.407^{* * *}$ | $0.439^{* * *}$ | 0.419*** | $0.400 * * *$ |
|  | (0.082) | (0.068) | (0.079) | (0.095) | (0.084) | (0.092) |
| Observations | 968 | 1028 | 968 | 968 | 1028 | 968 |
| $R^{2}$ | 0.04 | 0.03 | 0.03 |  |  |  |
| Significance of the regression: P-value | 0.01** | 0.01** | 0.04** | 0.04** | 0.04** | 0.16 |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelidgod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F7: Who's who in term of coherence $(\bar{d})$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\frac{[1]}{\bar{d}}$ | $\stackrel{[2]}{d}$ | $\frac{[3]}{\bar{d}}$ | $\frac{[4]}{\bar{d}}$ | $\stackrel{[5]}{\bar{d}}$ | $\frac{[6]}{\bar{d}}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate |  |  | $-0.071^{*}$ | -0.068** |  | -0.070** |
|  | $(0.042)$ |  | $(0.042)$ | (0.031) |  | (0.031) |
| Attended college | -0.106*** |  | -0.113*** | -0.105*** |  | $-0.113^{* * *}$ |
|  | (0.038) |  | (0.038) | (0.026) |  | (0.026) |
| Associated degree | -0.109** |  | -0.120*** | -0.108*** |  | -0.119*** |
|  | (0.042) |  | (0.042) | (0.031) |  | (0.031) |
| BA/BS | $-0.126^{* * *}$ |  | -0.134*** | $-0.126^{* * *}$ |  | -0.134*** |
|  | (0.038) |  | (0.038) | (0.027) |  | (0.027) |
| MS or PhD | -0.130*** |  | -0.140*** | -0.129*** |  | $-0.139^{* * *}$ |
|  | (0.039) |  | (0.039) | (0.029) |  | (0.029) |
| Professional | -0.059 |  | $-0.063$ | $-0.058$ |  |  |
|  | (0.052) |  | $(0.052)$ | $(0.041)$ |  | $(0.041)$ |
| Income |  |  |  |  |  |  |
| Income $<10000$ |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | -0.000 | -0.018 |  | -0.001 | -0.019 |  |
|  | (0.029) | (0.028) |  | (0.024) | (0.023) |  |
| $20000 \leq$ Income $<30000$ | -0.034 | -0.049* |  | -0.033 | -0.049** |  |
|  | (0.028) | (0.026) |  | (0.023) | (0.023) |  |
| $30000 \leq$ Income $<40000$ | -0.014 | -0.039 |  | -0.015 | -0.039* |  |
|  | (0.028) | (0.027) |  | (0.024) | (0.023) |  |
| $40000 \leq$ Income $<50000$ | -0.056** | -0.073*** |  | -0.056** | $-0.073^{* * *}$ |  |
|  | (0.027) | (0.026) |  | (0.025) | (0.024) |  |
| $50000 \leq$ Income $<60000$ | $-0.051^{*}$ | $-0.076^{* * *}$ |  | -0.051* | $-0.076^{* * *}$ |  |
|  | $(0.028)$ | $(0.027)$ |  | (0.028) | $(0.027)$ |  |
| Income $\geq 60000$ | -0.016 | -0.039 |  | -0.016 | -0.039* |  |
|  | (0.026) | (0.025) |  | (0.022) | (0.021) |  |
| Did not know/refused | -0.037 | -0.062** |  | -0.037 | -0.062*** |  |
|  | (0.027) | (0.026) |  | (0.023) | (0.022) |  |
| Age |  |  |  |  |  |  |
| Age/100 | -0.156 | -0.347* | -0.238 | -0.152 | -0.342* | -0.233 |
|  | (0.21) | (0.20) | (0.20) | (0.19) | $(0.18)$ | $(0.19)$ |
| $\left(\right.$ Age /100) ${ }^{2}$ | $0.002$ | $0.003^{*}$ | $0.002$ | $0.002$ | $0.003^{*}$ | $0.002$ |
|  | $(0.002)$ | $(0.002)$ | $(0.002)$ | $(0.002)$ | $(0.002)$ | $(0.002)$ |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.022* | -0.013 | -0.026** | -0.021** | -0.013 | -0.026** |
|  | (0.011) | (0.011) | (0.011) | (0.011) | (0.010) | (0.011) |
| Race ${ }^{\text {c }}$ |  |  |  |  |  |  |
| White | -0.007 | -0.015 | -0.007 | -0.006 | -0.015 | -0.007 |
|  | (0.021) | (0.022) | (0.021) | (0.020) | (0.019) | (0.020) |
| Black | $0.074 * *$ | $0.067^{* *}$ | 0.071 ** | $0.074^{* * *}$ | 0.068** | $0.072^{* * *}$ |
|  | (0.034) | (0.032) | (0.034) | (0.027) | (0.027) | (0.027) |
| Asian | -0.025 | -0.041 | -0.022 | -0.027 | -0.043 | -0.024 |
|  | (0.033) | (0.032) | (0.032) | (0.037) | (0.037) | (0.038) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion | -0.011 | -0.003 | -0.008 | -0.012 | -0.003 | -0.008 |
|  | (0.021) | (0.020) | (0.020) | (0.020) | (0.020) | (0.021) |
| Roman Catholic | 0.018 | 0.032 | 0.020 | 0.017 | 0.031 | 0.019 |
|  | (0.020) | (0.020) | (0.019) | (0.019) | (0.019) | (0.020) |
| Christian | 0.008 | 0.016 | 0.009 | 0.007 | 0.016 | 0.009 |
|  | (0.019) | (0.019) | (0.019) | (0.020) | (0.019) | (0.020) |
| Protestant | 0.004 | 0.019 | 0.006 | 0.004 | 0.019 | 0.006 |
|  | (0.019) | (0.018) | (0.019) | (0.019) | (0.019) | (0.019) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.014 | 0.013 | 0.012 | 0.014 | 0.013 | 0.012 |
|  | (0.013) | (0.013) | (0.013) | (0.014) | (0.013) | (0.014) |
| Wave 14 | 0.018 | 0.015 | 0.018 | 0.018 | 0.015 | 0.018 |
|  | (0.012) | (0.012) | (0.013) | (0.013) | (0.012) | (0.013) |
| Intercept | $0.438{ }^{* * *}$ | 0.390*** | $0.440^{* * *}$ | $0.436^{* * *}$ | 0.389*** | $0.438^{* * *}$ |
|  | (0.062) | (0.051) | (0.060) | (0.051) | (0.046) | (0.049) |
| Observations | 1044 | 1105 | 1044 | 1044 | 1105 | 1044 |
| $R^{2}$ | 0.08 | 0.05 | 0.07 |  |  |  |
| Significance of the regression: P -value | 0.00*** | $0.00^{* * *}$ | $0.00^{* * *}$ | 0.00*** | $0.00^{* * *}$ | $0.00^{* * *}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelihgod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.

Table F8: Who's who in term of coherence $\left(d^{K o l}\right)$-All respondents for whom we can compute a measure of coherence-

|  |  | OLS |  |  | Tobit |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} {[1]} \\ d^{K o l} \end{gathered}$ | $\begin{gathered} {[2]} \\ d^{K o l} \end{gathered}$ | $\begin{gathered} {[3]} \\ d^{K o l} \end{gathered}$ | $\begin{gathered} {[4]} \\ d^{K o l} \end{gathered}$ | $\begin{gathered} {[5]} \\ d^{K o l} \end{gathered}$ | $\begin{gathered} {[6]} \\ d^{K o l} \end{gathered}$ |
| Education |  |  |  |  |  |  |
| $\leq$ Grade 11 |  | Base group |  |  | Base group |  |
| High school graduate | -0.055 |  | -0.055 | -0.055 |  | -0.055 |
|  | (0.048) |  | (0.048) | (0.041) |  | (0.041) |
| Attended college | -0.088** |  | -0.094** | -0.088** |  | -0.094*** |
|  | (0.043) |  | (0.043) | (0.035) |  | (0.035) |
| Associated degree | -0.122** |  | -0.128*** | -0.121*** |  | $-0.128^{* * *}$ |
|  | (0.049) |  | (0.049) | (0.041) |  | (0.041) |
| BA/BS | -0.118*** |  | -0.122*** | -0.119*** |  | $-0.123^{* * *}$ |
|  | (0.043) |  | (0.043) | (0.036) |  | (0.036) |
| MS or PhD | -0.122*** |  | -0.128*** | -0.122*** |  | -0.127*** |
|  | (0.045) |  | (0.045) | (0.039) |  | (0.039) |
| Professional |  |  | $-0.059$ | $-0.057$ |  |  |
|  | $(0.062)$ |  | $(0.062)$ | $(0.055)$ |  | $(0.055)$ |
| Income |  |  |  |  |  |  |
| Income $<10000$ |  | Base group |  |  | Base group |  |
| $10000 \leq$ Income $<20000$ | 0.033 | 0.006 |  | 0.034 | 0.007 |  |
|  | (0.035) | (0.033) |  | (0.032) | (0.030) |  |
| $20000 \leq$ Income $<30000$ | -0.007 | -0.022 |  | -0.006 | -0.021 |  |
|  | (0.034) | (0.032) |  | (0.031) | (0.030) |  |
| $30000 \leq$ Income $<40000$ | 0.026 | -0.004 |  | 0.026 | -0.003 |  |
|  | (0.035) | (0.033) |  | (0.031) | (0.030) |  |
| $40000 \leq$ Income $<50000$ | $-0.011$ | -0.033 |  | $-0.010$ | $-0.032$ |  |
|  | $(0.036)$ | (0.034) |  | (0.033) | $(0.032)$ |  |
| $50000 \leq$ Income $<60000$ | -0.022 | -0.042 |  | -0.021 | -0.041 |  |
|  | (0.037) | (0.036) |  | (0.038) | (0.036) |  |
| Income $\geq 60000$ | 0.015 | -0.013 |  | 0.016 | -0.012 |  |
|  | (0.032) | (0.030) |  | (0.029) | (0.028) |  |
| Did not know/refused | -0.009 | -0.044 |  | -0.009 | -0.043 |  |
|  | (0.034) | (0.032) |  | (0.031) | (0.030) |  |
| Age |  |  |  |  |  |  |
| Age/100 | 0.086 | -0.088 | 0.054 | 0.085 | -0.089 | 0.054 |
|  | (0.27) | (0.25) | (0.26) | (0.25) | (0.24) | (0.25) |
| $\left(\right.$ Age /100) ${ }^{2}$ | $0.000$ | $0.002$ | $0.001$ | $0.000$ | $0.002$ | $0.001$ |
|  | $(0.003)$ | $(0.003)$ | $(0.003)$ | $(0.002)$ | $(0.002)$ | $(0.002)$ |
| Gender |  |  |  |  |  |  |
| 1 if male | -0.033** | -0.023* | -0.035** | -0.033** | -0.023* | -0.035** |
|  | (0.014) | (0.014) | (0.014) | (0.014) | (0.014) | (0.014) |
| Race ${ }^{\text {a }}$ |  |  |  |  |  |  |
| White | 0.013 | -0.002 | 0.011 | 0.012 | -0.002 | 0.010 |
|  | (0.025) | (0.026) | (0.025) | (0.026) | (0.025) | (0.026) |
| Black | 0.091** | $0.086^{* *}$ | 0.088** | 0.092** | 0.087** | 0.089** |
|  | (0.038) | (0.036) | (0.038) | (0.037) | (0.035) | (0.037) |
| Asian | 0.018 | -0.002 | 0.018 | 0.015 | -0.005 | 0.015 |
|  | (0.048) | (0.046) | (0.047) | (0.050) | (0.049) | (0.050) |
| Other |  | Base group |  |  | Base group |  |
| Religion |  |  |  |  |  |  |
| No religion |  |  |  |  |  |  |
|  | $(0.027)$ | $(0.026)$ | $(0.026)$ | $(0.027)$ | $(0.027)$ | $(0.027)$ |
| Roman Catholic | 0.015 | 0.031 | 0.018 | 0.014 | 0.030 | 0.017 |
|  | (0.026) | (0.026) | (0.026) | (0.026) | (0.026) | (0.026) |
| Christian | 0.016 | 0.025 | 0.018 | 0.015 | 0.024 | 0.018 |
|  | (0.025) | (0.025) | (0.025) | (0.026) | (0.025) | (0.026) |
| Protestant | 0.012 | 0.032 | 0.014 | 0.012 | 0.032 | 0.014 |
|  | (0.025) | (0.024) | (0.024) | (0.025) | (0.025) | (0.025) |
| Other |  | Base group |  |  | Base group |  |
| Wave dummies |  |  |  |  |  |  |
| Wave 12 |  | Base group |  |  | Base group |  |
| Wave 13 | 0.011 | 0.011 | 0.009 | 0.010 | 0.010 | 0.009 |
|  | (0.018) | (0.017) | (0.018) | (0.018) | (0.017) | (0.018) |
| Wave 14 | 0.011 | 0.007 | 0.011 | 0.010 | 0.007 | 0.010 |
|  | (0.017) | (0.016) | (0.017) | (0.017) | (0.016) | (0.017) |
| Intercept | 0.489*** | $0.453 * * *$ | $0.507^{* * *}$ | 0.490*** | $0.453 * * *$ | 0.508*** |
|  | (0.077) | (0.063) | (0.074) | (0.068) | (0.061) | (0.066) |
| Observations | 1044 | 1105 | 1044 | 1044 | 1105 | 1044 |
| $R^{2}$ | 0.06 | 0.04 | 0.05 |  |  |  |
| Significance of the regression: P-value | 0.00*** | $0.00^{* * *}$ | 0.00*** | $0.00^{* * *}$ | $0.00^{* * *}$ | $0.00^{* * *}$ |

Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.
ii. Heteroskedasticity-robust standard errors are in parentheses.
iii. "Significance of the regression" provides the P-value of the joint test of the hypotheses that all the coefficients except the constant term are zero. For least squares estimates it corresponds to the P -value associated to a F ratio. For Tobit estimates the P -value is the one associated to a likelighod ratio statistic which has a limiting chi-squared distribution under the null hypothesis.


[^0]:    Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively
    ii. Heteroskedasticity-robust standard errors are in parentheses.

[^1]:    Notes: i. $*, * *$ and $* * *$ represent 10,5 and $1 \%$ significance, respectively.

[^2]:    iv. Number of observations are in square-brackets.

