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▶ To cite this version:

Christophe Amat, Tomasz Michalski, Gilles Stoltz. Fundamentals and exchange rate forecastability with machine learning methods. 2016. halshs-01003914v5

HAL Id: halshs-01003914 https://shs.hal.science/halshs-01003914v5

Preprint submitted on 21 Dec 2016 (v5), last revised 28 May 2018 (v6)

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FUNDAMENTALS AND EXCHANGE RATE FORECASTABILITY WITH MACHINE LEARNING METHODS*

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December 20, 2016

Abstract

Using methods from machine learning we show that fundamentals from simple exchange rate models (PPP or UIRP) consistently allow to improve exchange rate forecasts for major currencies over the floating period era 1973–2014 at a 1 month forecast and allow to beat the no-change forecast. "Classic" fundamentals hence contain useful information and exchange rates are forecastable even for short forecasting horizons. Such conclusions cannot be obtained when using rolling or recursive OLS regressions as in the literature. The methods we use – sequential ridge regression and the exponentially weighted average strategy both with discount factors – do not estimate an underlying model but combine the fundamentals to directly output forecasts.

JEL codes: C53, F31, F37

Keywords: exchange rates, forecasting, machine learning, purchasing power parity, uncovered interest rate parity, monetary exchange rate models

^{*}We would like to thank Philippe Bacchetta, Charles Engel, Refet Gürkaynak, Robert Kollmann, Jose Lopez, Evren Örs, Cavit Pakel, Martin Puhl, Romain Rancière, Helene Rey, Michal Rubaszek and the seminar participants at Bilkent University, HEC Paris, National Bank of Poland, OECD, UN, EEA 2015 conference, Royal Economic Society 2015, 8th Financial Risks International Forum 2015 (Institut Louis Bachelier) for helpful comments and discussions. All remaining errors are ours. We would like to thank Investissements d'Avenir (ANR-11-IDEX-0003/Labex Ecodec/ANR-11-LABX-0047) for financial support. Corresponding author: Gilles Stoltz, HEC Paris, 1, rue de la Libération, F-78351 Jouy-en-Josas Cedex. Phone: +33(0)139677260. E-mail: stoltz@hec.fr.

1. Introduction

We show that fundamentals from "classic" exchange rate models are useful in forecasting short-term changes in exchange rates for major currencies. Fundamentals from the purchasing power parity models (PPP), uncovered interest rate parity (UIRP) and a standard monetary model that the traditional exchange rate forecasting literature (based on OLS methods) failed to find effective (see Rossi [2013] for a recent comprehensive review) in fact carry valuable information for out-of-sample forecasting even at short-term horizons.

This has important implications both for economic theory and policy. For example Alvarez et al. [2007] argue that if exchange rates are random walks, then monetary policy concentrated on short-term interest rates does not have the typically assumed nominal or real effects, but rather affects only risk premia. More generally, it has been frustrating for researchers in international economics that the forces stipulated by the simplest and intuitive models about the most important prices – exchange rates – did not seem to be empirically detectable.

We show that armed with prediction techniques borrowed from machine learning that come with guarantees on their performance – sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors – we are able broadly to improve forecasting forward the 1–month exchange rate for the floating rate period 1973–2014 using the standardly employed root mean square error (RMSE) criterion. We obtain similar success with predicting the direction of change of the exchange rates, several economic criteria and conduct different robustness checks. Our conclusion is therefore that "classic" fundamentals contain useful information about exchange rates even at short forecasting horizons¹ when proper machine-learning tools directly geared towards maximizing the quality of predictions are used.

Our findings are contrary to the consensus in the literature started by the famous Meese and Rogoff [1983] result of the inability of short-term forecasts based on fundamentals to outperform the most demanding no-change / random walk prediction². The improvements in the forecasts using fundamentals we document (up to 7% in terms of RMSE) are not large – the fundamentals we consider do not help to add a lot of predictive content relative to the one of "no change" from the previous observed exchange rate – but are sizeable (and statistically significant at conventionally

¹Forecastability of the exchange rates over longer time periods, such as 2-year forward forecasts and beyond was established in the literature, see Section 6.

²The rare exceptions are that of Clark and West [2006] that demonstrate the predictability of a UIRP-based model or Wright [2008] that uses Bayesian Model Averaging methods. See a longer discussion below.

accepted levels) in comparison to the existing literature. The observations of Engel and West [2005] may well be valid: current and past fundamentals may have low correlations with future exchange rate realizations. Fundamentals may have little predictive power against an exchange rate that can be approximated by a random walk although they may add *some* useful information.

Machine learning methods at hand. The common feature for the methods we use is the following. They do not aim to estimate the coefficients of some underlying model but rather treat the coefficients as numbers to be picked (not estimated) over time in order to form good predictions. Indeed, these methods do not rely on some stochastic modeling of the exchange rates (e.g., in terms of a linear combination of the fundamentals plus some stochastic errors). They have appealing features that we describe in Section 3.3 as theoretical guarantees (bounds) on their performance can be proven. In particular, their RMSEs converge, as the number of instances to be predicted increases, to the RMSE of the best fixed linear forecasting equation. The latter is weakly smaller than the RMSE of the no-change / random walk prediction, so asymptotically our methods should beat the random walk if fundamentals indeed drive the exchange rate. One can also show that no such guarantee can be achieved for conventional (rolling) OLS. This is a crucial difference between our work and that of Li et al. [2015], another study forecasting exchange rates with different machine learning techniques: our methods have these important guarantees and hence are less arbitrary. Moreover, we apply them to fundamentals precisely identified by simple economic models and do not run "kitchen-sink" regressions. It is to be stressed that both the discount factors and the parameters of the methods described above (a learning rate for the exponentially weighted average strategy with discount factors, a regularization factor for sequential ridge regression with discount factors) are computed from the data up to the period for which the forecast is given rather than exogenously imposed³. Our methodology is hence entirely data-driven and does not rely on some educated guesses about important parameters as for example in Wright [2008] for Bayesian model averaging methods. Note finally that the methods we implement are general forecasting methods that were not specifically designed for exchange rate forecasting. They proved to work well in other problems (forecasting of air quality, electricity consumption or the production of oil reservoirs; see Cesa-Bianchi and Lugosi [2006] as well as Mauricette et al. [2009] and Stoltz [2010]).

³Unsurprisingly, variations of these methods that we call "oracles" – acting as if these optimal parameters were known beforehand – perform better.

More precisely, our methods aim to avoid a problem known as overfitting past data, which consists of being able to reconstruct well with the forecasting equations past data (have a good "in-sample" prediction error) but provide poor future predictions (have a poor "out-of-sample" prediction error). Such problems permeate the exchange-rate prediction literature for the last 35 years for example from Meese and Rogoff [1983] to Fratzscher et al. [2015].

Our first method – the exponentially weighted average strategy with discount factors – allows to weigh fundamentals by their past performance: the coefficients picked over time are convex and come in proportion to the (exponential of the) average past performance of each fundamental. They vary over time as the performance of each fundamental evolves. If some fundamentals always perform poorly they are quickly almost discarded, but, by construction all coefficients are non-zero (though some may be close to zero). This algorithm has a natural interpretation as identifying the "most fashionable" fundamentals considered by traders in some periods as in the scapegoat model of exchange rates of Bacchetta and Wincoop [2004]. When we scrutinize the weights (see Appendix C, Figures 1–6) we see that quite often there is one prevailing fundamental that carries most information for forecasting exchange rates.

Our second method – sequential ridge regression with discount factors⁴ – resembles OLS regressions, but it prevents the in-sample-overfitting issue they encounter by including a regularization term in an otherwise standard ordinary least squares (OLS). Indeed, this regularization term corresponds to an additional term proportional to the Euclidean norm of the linear coefficients in the defining minimization formulation of the chosen coefficients. This prevents them to get the smallest possible in-sample error (in order not to suffer from a large out-of-sample error) as the former error is usually achieved with coefficients with large values, while such large values are penalized by the added regularization term. The elastic net method considered by Li et al. [2015] proceeds in a similar fashion, with a combination of a Euclidian and absolute-values regularization; it however comes with no clear theoretical performance guarantees, unlike what can be proved for the ridge regression.

For both methods, we consider variants of the basic machine-learning formulations that include discount factors so that most recent forecasting errors matter more than older ones. This allows for accommodating structural breaks that might be present in the data (such as changes in

⁴The ridge regression was introduced by Hoerl and Kennard [1970] in a stochastic and non-sequential setting. But it turns out that the exact same method can be analyzed in a sequential non-stochastic setting, leading to guarantees of another nature, namely, in the form (12). We will discuss these new machine-learning analyses in Section 3.3.

the conduct of monetary policy, crisis times etc.) better than the original formulations; this is also true for any changes in the behavior of market participants⁵. It should be noted that our discounts are of a polynomial and not of a geometrical form (used in most of the economic literature). This is because it can be proved that theoretical guarantees can be associated with the former (see section 3.3) but not for the latter in our context (see Theorems 2.7 and 2.8 of Cesa-Bianchi and Lugosi [2006] for a discussion of performance for different discounting methods).

The regularization term and the discounts play a complementary role and are both useful to the good performance of our methods. The discounts give more importance to the most recent past observations and thus allow our methods to adapt quickly (but not too fast in view of their polynomial nature) to structural breaks which are frequent for exchange rate series; while the regularization term is useful to avoid overfitting and prevents the methods from sticking too much to the in-sample data.

The data used and the scope of our study. In our study, apart from the particular machine-learning methods, we used data and forecast evaluation methods that are standard in the literature as discussed in Rossi [2013]. We use lagged fundamentals and do not detrend, filter or seasonally adjust the data in any way. We use single-equation methods and our benchmark for comparison is the no change in the exchange rate prediction. We study the most important floating currency pairs for the entire period 1973–2014 since the breakup of the Bretton Woods fixed exchange rate system. We use two samples as encountered in the literature: the end-of-month exchange rate sample (with one reading of the exchange rate at the end of the month) and a sample with average monthly exchange rates as used by for example Molodtsova and Papell [2009] to allow for comparability with the Taylor-rule models they suggest.

We conduct several robustness checks – such as trying out different (sub)samples, variations on the exact form of the fundamentals included, etc., and our results are upheld. Furthermore, investigation of statistical properties of the errors shows that our improvement is uniform rather than driven by a few well predicted instances. There are several themes discussed by Rossi [2013] that also hold in our study. Traditional linear estimation methods (rolling or recursive OLS) with our fundamentals fail to forecast exchange rates better than the no change in exchange rates prediction. With the addition of several fundamentals (aggregating information) we do not get gains in

⁵Consider the comparison of forecasts in the time period around the Plaza agreement shown in Table 10 and discussed in Appendix C.

the precision of predictions – that is, parsimonious forecasting equations work best. Fundamentals contemporaneous with the forecast period (actually realized values, as used in Meese and Rogoff [1983]) perform poorly.

Our study is not comprehensive nor exhaustive in nature. Our goal is not to provide the best forecasts of exchange rates possible by finding fundamentals from a wide menu that was considered in the vast literature that have the strongest predictive power (or aggregating the information contained therein for example in hundreds of different series). We are interested both in showing that we can detect fundamentals driving the exchange rates even at short horizons and argue for the usefulness of machine learning methods in applications to well-known and hard-to-crack economic problems such as predicting inflation or output.

1.1. Organization of the paper

First, in Section 2, we lay out the fundamentals that we shall consider. Then, in Section 3, we discuss the methods of the analysis and put them in perspective with the standard linear regressions. Next, in Section 4 we present the data while Section 5 discusses the results. Section 6 contains a literature review and 7 concludes. An Appendix contains further details regarding the methods and more results tables.

2. Considered fundamentals

We consider the "classic" fundamentals stemming from the simple exchange rate models of the 1970s and the Taylor-rule based models. The reasons are fourfold, as surveyed in Rossi [2013]: (i) these fundamentals come from exchange rate models that involve basic relationships in standard international economics, (ii) they have been extensively used in the literature, (iii) studies have shown that at short-time horizons forecasting models based on these were not successful against the benchmark of no-change forecast⁶, (iv) the data is widely available for long periods for a wide set of countries and does not require any transformation (like, for example, data on productivity).

⁶To remind our reader, no other set of fundamentals was shown as of yet to beat such forecasts in a consistent manner either. See Footnote 33 on the predictability of exchange rates with Taylor-rule model-based fundamentals.

2.1. General framework

The general forecasting equation⁷ of the exchange rate change for a currency pair is given by

$$\hat{s}_{t+1} - s_t = \alpha_t + \sum_{j=1}^{N} \beta_{j,t} f_{j,t},$$
(1)

where s_t is the logarithm of the exchange rate (domestic currency units per unit of the foreign currency) at time t, the intercept α_t and the slope coefficients $\beta_{j,t}$ are to be picked based on any and all information available up to time t, while the $f_{j,t}$ are the N fundamentals considered at time t. Throughout the paper we will force⁸ $\alpha_t = 0$. The time unit will be months.

By the exchange rate, we mean here either the end-of-the month value or the monthly average of the rate, depending on the data set.

The coefficients $\beta_{j,t}$ (and α_t when intercepts are considered, which is not the case here) are usually picked through rolling or recursive OLS regressions. Rossi [2013] states that "the literature has been focusing mainly on rolling or recursive window forecasting schemes (see West [2006]), where parameters are reestimated over time using a window of recent data", and Cheung et al. [2005] mention "the convention in the empirical exchange rate modeling literature of implementing rolling regressions established by Meese and Rogoff [1983]." This was because they proved best and no other method could consistently beat them. In this paper, we revisit this convention and consider other methods for picking these coefficients. We do not write "estimate these coefficients" as such a wording refers to some underlying model with true parameters α and β . We will be interested solely in the forecasting performance and not at all in the existence or use of a model.

Before recalling how rolling and recursive OLS regressions proceed to that end and presenting alternative methods stemming from the field of machine learning, we discuss the fundamentals of

$$s_{t+1} - s_t = \alpha_t + \sum_{j=1}^{N} \beta_j f_{j,t} + \varepsilon_{t+1} ,$$

from which, by estimation of the unknown coefficients β_j some prediction \widehat{s}_{t+1} could be obtained. Likewise, we avoid the terminology of "estimation".

⁷We avoid using the word "model" here, as we will not have to assume that some true underlying modeling of the differences exists, e.g.,

⁸Affine forecasting equations can be handled via one only of our two new forecating methods, namely, sequential ridge regression with discount factors, and did not offer an improvement upon the results shown in the paper. Details are available upon request. The other method, the exponentially weighted average strategy with discount factors, is not applicable to affine forecasting equations.

interest.

2.2. Fundamentals from the PPP, UIRP, monetary and Taylor-rule models

We are only interested in the predictive performance of our forecasting methods and we allow for different coefficients for home and foreign fundamentals relating to the same measured quantity, even if the models from which they come could call for equal weighting theoretically. There is a good reason to do this as in our empirical work two series, even if pertaining to a similar economic concept, may be measured completely differently in any two countries (given that they are provided by independent institutions using diverse methodologies). Hence, the best weights picked may differ as the elasticities of response of investors to purportedly similar fundamentals in both countries may and should differ⁹.

The first series of fundamentals is formed by inflation differentials, which are used also, e.g., in the relative purchasing power parity (PPP) model. The associated forecasting equation is given by

$$\widehat{s}_{t+1} - s_t = \beta_{1,t} \pi_t - \beta_{2,t} \pi_t^* \tag{2}$$

where π_t and π_t^* are respectively present home and foreign measures of 12-month inflation rates available (known) at time t. In what follows, a " \star " will denote the variables for the foreign country¹⁰.

The forecasting equations using the fundamentals from the uncovered interest rate parity model (UIRP) are of the form

$$\widehat{s}_{t+1} - s_t = \beta_{1,t} i_{t \to t+1} - \beta_{2,t} i_{t \to t+1}^{\star}, \tag{3}$$

where $i_{t\to t+1}$ is the short run (money-market) interest rate at home and in the foreign country respectively. These would be the interest rates at which investors could place money at time t for

⁹The most flagrant example are perhaps money stocks: for the United Kingdom and Sweden only M0 aggregates, and for Italy and the Netherlands only M2 are available for the entire period considered and not M1 as for other countries. M0, M1 and M2 money aggregates are typically correlated, but they obviously measure different things. Even for countries for which M1 measures are available, these contain different components depending on the country. Such issues unfortunately exist to some extent for every fundamental considered, and is a general problem in the literature. Due to the asymptotic properties of the methods we wanted to obtain the longest series possible for the largest number of currencies and had to – given the data available – sacrifice to some extent standardization.

The relative PPP forecasting equation would actually stipulate $\hat{s}_{t+1} - s_t = \beta_{1,t} \Delta p_{t+1} - \beta_{2,t} \Delta p_{t+1}^*$ where Δp_{t+1} is the change in the price level between periods t and t+1. However, as the change in the future price level is unknown at time t, we use past price changes (inflation) in our forecasting exercise.

the period t to t+1 and could be observed by them in real time.

The third series of fundamentals consists of variations of money stocks and outputs. The simplest monetary model (flexible price, Frenkel-Bilson model) states that exchanges rates can be modeled as linear combinations of the form

$$s_t = \alpha + \phi(m_t - m_t^{\star}) - \omega(y_t - y_t^{\star}),$$

where m_t and m_t^{\star} are the logarithms of the money stocks at time t while y_t and y_t^{\star} are the logarithms of outputs. Here, α , ϕ and ω denote some true underlying parameters for the model. As indicated above several times, we will not rely on the existence of such a monetary model and only consider it to extract the fundamentals it proposes and substitute them in our forecasting equations. Namely, after lagging the said fundamentals by a month (to account for known and present, not forward values), differencing the above equation¹¹, and allowing for decoupled fundamentals we consider¹² the forecasting equations

$$\widehat{s}_{t+1} - s_t = \beta_{1,t} \Delta m_t - \beta_{2,t} \Delta m_t^* - \beta_{3,t} \Delta y_t + \beta_{4,t} \Delta y_t^* \tag{4}$$

where $\Delta x_t = x_t - x_{t-1}$ denotes the change between periods t-1 and t for a variable x. In (4), the parameters $\beta_{1,t}, \ldots, \beta_{4,t}$ are to be picked over time, as in (2) and (3).

We also investigate whether all the fundamentals mentioned above taken together are able to predict changes in the exchange rate¹³.

Finally we also consider whether the fundamentals extracted from a Taylor-rule based exchangerate model (that has been found successful in establishing the predictability of the exchange rate changes, as discussed in Section 6) are useful in forecasting as well. We use the fundamentals

¹¹We proceeded this way as some of the methods we use, in particular, the exponentially weighted average strategy with discount factors, require direct predictions of the exchange rate change. An alternative is to find some"equilibrium" exchange rate driven by the aforementioned fundamentals and use the deviation of the current exchange rate from that "theoretical" exchange rate as the fundamental. Doing this does not change qualitatively our results.

¹²We recall that the model does not need to hold: it merely helps us decide which fundamentals to consider in our forecasting equations.

¹³The form in which the fundamentals are included does not seem to matter. We produced forecasts using fundamentals from the PPP and monetary models for our main sample also using price or output indices in levels or the monetary stock: our qualitative conclusions do not change. We also investigated fundamentals extracted from the other popular monetary flexible-price model that also involves differences in interest rates, which in our forecasting equations would add terms $i_t - i_{t-1}$ and $i_t^* - i_{t-1}^*$ as predictors. We do not show the results as they were very close to the performance of the forecasting equation (4). We included these fundamentals, however, when considering forecasts with all fundamentals at hand to see whether aggregating information from many different series helps.

from the most parsimonious version of equation (7) from Molodtsova and Papell [2009] for our forecasting equation:

$$\widehat{s}_{t+1} - s_t = \beta_{1,t} \pi_t^* - \beta_{2,t} \pi_t + \beta_{3,t} \widetilde{y}_t^* - \beta_{4,t} \widetilde{y}_t + \beta_{5,t} i_{t-1 \to t}^* - \beta_{6,t} i_{t-1 \to t}$$
 (5)

where \widetilde{y}_t and \widetilde{y}_t^{\star} are present measures of the output gaps. Therefore, our Taylor-rule fundamentals are the inflation, output gaps and the lagged interest rates.

The exact data series used are further discussed in Section 4.

Coupled fundamentals. In the forecasting equations (2)–(5) fundamentals of the home and foreign countries pertaining to the same concept can also get the same weight. For example, the forecasting equation inspired by the PPP model in (2) will be then $\hat{s}_{t+1} - s_t = \beta_{1,t} (\pi_t - \pi_t^*)$. We will refer to this situation as coupled fundamentals and perform robustness checks using this forecasting equation. Note that it still falls under the general umbrella (1), only with N/2 coefficients to pick instead of N and with the $f_{j,t}$ then referring to the differences between home and foreign fundamentals.

3. Methodology

3.1. Assessing the quality of the forecasts

As indicated above, forecasts \hat{s}_{t+1} of the 1-month ahead exchange rates s_{t+1} are based on the forecasting equations (2)–(5). A forecasting method consists of a rule for picking the coefficients $\beta_{j,t}$ over time based on past and present information. Several such methods are presented in Section 3.2, namely, the conventional rolling and recursive OLS regressions as well as two other methods stemming from the field of sequential learning.

We denote by T the total number of monthly values to be forecast, from months 1 to T. As is standard in the literature (see Molodtsova and Papell [2009] or Rossi [2013]) and for essentially the same reasons, namely, the consideration of rolling OLS regressions, we allow for a training period of length $t_0 = 120$ months (10 years) and only evaluate the accuracy of the forecasts on

months $t_0 + 1 = 121$ to T. To that end, we consider the root mean square error,

RMSE =
$$\sqrt{\frac{1}{T - t_0} \sum_{t=t_0+1}^{T} (\hat{s}_t - s_t)^2}$$
 (6)
= $\sqrt{\frac{1}{T - t_0} \sum_{t=t_0+1}^{T} ((\hat{s}_t - s_{t-1}) - (s_t - s_{t-1}))^2}$,

and note that (by subtracting and adding the pivotal values s_{t-1}) this root mean square error is indifferently the one for the logarithms of exchange rates s_t or for the changes in logarithms of exchange rates $s_t - s_{t-1}$.

We will want to investigate whether the improvements in RMSE of one method over another one are statistically significant. Denote by \hat{s}_t' and \hat{s}_t the respective forecasts of two methods of interest. We aim to test the hypothesis H_0 that the difference in forecasting accuracy is not significant against the alternative hypothesis H_1 that the second method—the one outputting the forecast \hat{s}_t —is significantly better on average. To that end, the standard practice (see the two tests presented below) is to consider the instantaneous differences in accuracy

$$d_t = (\widehat{s}_t' - s_t)^2 - (\widehat{s}_t - s_t)^2.$$

We denote by

$$\overline{d}_T = \frac{1}{T - t_0} \sum_{t=t_0+1}^T d_t$$

the empirical average of the differences and by

$$\overline{\sigma}_T^2 = \frac{1}{T - t_0} \sum_{t=t_0+1}^T \left(d_t - \overline{d}_T \right)^2$$

their empirical variance.

3.1.1. Some descriptive statistics

In this section we consider the no-change prediction as the benchmark method: $\hat{s}'_t = s_{t-1}$.

To get a first feeling of whether H_0 should be rejected or not in this case, one can take a look at the quantiles of the series of the d_t . E.g., in Tables 11–12 of Appendix C, we see that in the case of

the sequential ridge regression with discount factors or the exponentially weighted average strategy with discount factors, the distribution of the differences d_t is shifted toward positive values: the 75% and 90% quantiles are larger in absolute values than the 25% and 10% quantiles, while in addition, the median is positive. This tendency is especially visible for the sample with the average exchange rates. This is not the case for rolling or recursive OLS regressions. Additional comments on these matters are provided in Section 5.1.

3.1.2. A general test for comparing the predictive accuracies

In this section we may still consider the no-change prediction as the benchmark method, $\hat{s}'_t = s_{t-1}$, but we can and actually will also consider other benchmarks, like rolling or recursive OLS regressions.

Diebold and Mariano [1995] presented a test relying on mild and direct assumptions on the behavior of the differences d_t . This test uses the differences in the forecasting errors as primitives, even though the latter can be serially correlated. We state their results with a rectangular lag, as they advocated to do so.

More precisely, they showed that under an assumption of covariance stationarity and of short memory of the differences d_t , for a properly chosen truncation lag denoted by $H \geqslant 0$, the test statistics

$$S_{\mathrm{DM},H} = \sqrt{T - t_0} \, \frac{\overline{d}_T}{\sqrt{\overline{\sigma}_T^{H,2}}} \,,$$

where

$$\overline{\sigma}_{T}^{H,2} = \frac{1}{T - t_0} \sum_{t=t_0+1}^{T} \left(d_t - \overline{d}_T \right)^2 + \frac{2}{T - t_0} \sum_{\tau=1}^{H} \sum_{t=\tau+t_0+1}^{T} \left(d_t - \overline{d}_T \right) \left(d_{t-\tau} - \overline{d}_T \right),$$

converge to a $\mathcal{N}(0,1)$ distribution under H_0 while converging to $+\infty$ under H_1 . Diebold [2012] insists on how general the method is to compare the predictive accuracy of forecasts between any two methods, as long as the mild assumption stated above, namely, covariance stationarity of the differences in accuracy d_t , holds. He explains in Section 2.2 of the mentioned reference why this assumption is natural and often met in practice.

The choice of the truncation lag H was partially left open; Diebold and Mariano suggested to pick it as a function of the length of the short memory (of the autocovariance degree). Estimating

by \widehat{H} a proper H based on our data then substituting its value would lead to considering the test statistic $S_{\mathrm{DM},\widehat{H}}$, which would not be guaranteed anymore to converge to a $\mathcal{N}(0,1)$ distribution under H_0 . We instead take a more robust approach to reject H_0 , that is, we build a more conservative test than the original test based on some good, a priori value of H. Namely, we consider

$$S_{\text{DM}} = \sqrt{T - t_0} \frac{\overline{d}_T}{\sqrt{\max_{H \in \{0, 1, \dots, 20\}} \overline{\sigma}_T^{H, 2}}},$$
 (7)

which is smaller than any of the corresponding original statistics $S_{\mathrm{DM},H}$. The limiting distribution under H_0 is smaller than a $\mathcal{N}(0,1)$ distribution (while a convergence to $+\infty$ is still achieved under H_1 , at a slower rate). Yet, we compute the p-values using quantiles of the normal distribution, which is very conservative. (I.e., the size and the power of the test are smaller than usual.) Despite all, we will be able to reject the hypothesis H_0 of equal accuracy abilities in many circumstances.

Note that the maximal value 20 for H was set on our data set because it corresponds roughly to the maximal value of $\sqrt{T-t_0}$ on our data. The p-values associated with the test thus constructed will be reported in the tables in columns labeled "DM p-value"; also, in the sequel, this test will be referred to as the DM test.

3.1.3. A test for the case of nested forecasting equations?

The tests by West [1996], Clark and West [2006, 2007] are designed for comparing models, while we are interested here only in comparing forecasts and are not ready to assume the existence of models. See Diebold [2012], Rossi [2013] for discussions. Further, the tests by Clark and West have been designed and studied only for the case of rolling and recursive OLS regressions, while in our case we obviously are interested in alternative methods. This is underlined in the original papers and pointed out again by Rogoff and Stavrakeva [2008].

Nonetheless, for the sake of completeness and honesty¹⁴, we compute p-values associated with the test developed by Clark and West for nested models (even though we consider "forecasting equations" only and not models). These p-values should merely be considered as some statistical indicators of predictability. In our eyes, the significant tests being performed in our study are the DM tests mentioned above. The readers will note that in our tables the test by Clark and West is

¹⁴We would prefer to drop the results related to the tests by West [1996], Clark and West [2006, 2007] in the final version of this paper; we provide them in this version only for the reviewers to see what their results are.

systematically less conservative than the DM test (as in all other empirical and theoretical studies).

Mathematical description of the test. For this test we restrict our attention to the no-change prediction as the benchmark method: $\hat{s}'_{t+1} = s_t$.

This no-change prediction is nested into the larger forecasting equation (1). In this case, assuming the existence of underlying models, Clark and West [2006] argued that the differences in forecasting abilities d_t were an unfair measure and advocated to use instead some adjusted differences, which in our context would equal

$$a_t = d_t + (\hat{s}_t - s_{t-1})^2$$

where \hat{s}_t is the prediction associated with the larger forecasting equation (1). We denote by

$$\overline{a}_T = \frac{1}{T - t_0 + 1} \sum_{t = t_0 + 1}^{T} a_t \tag{8}$$

the empirical average of the adjusted differences a_t and by

$$\overline{S}_T^2 = \frac{1}{T - t_0 + 1} \sum_{t=t_0+1}^{T} \left(a_t - \overline{a}_T \right)^2 \tag{9}$$

their empirical variance. The test statistic equals

$$S_{\text{CW}} = \sqrt{T - t_0 + 1} \frac{\overline{a}_T}{\sqrt{\overline{S}_T^2}}.$$
 (10)

The hypotheses tested are in terms of an underlying model (if one is ready to assume that some exists). Namely, H_0' : the true underlying model is the no-change model versus H_1' : the true underlying linear model uses at least one fundamental of (1). Here again, we see that the test is about predictability, not forecastability. Under an assumption of residuals of the models forming a martingale difference sequence, and provided that rolling and recursive OLS regressions are considered to pick the coefficients in (1), Clark and West [2006] showed that S_{CW} converges to a $\mathcal{N}(0,1)$ distribution under H_0' while converging to $+\infty$ under H_1' . The p-values associated with the test thus constructed will be reported in the tables in columns labeled "CW p-value"; also, in the sequel, this test will be referred to as the CW test.

We recall that as indicated several times above, these p-values are merely indicators, and indicators of predictability rather than of forecastability. We include them for the sake of completeness.

3.2. Forecasting methods (1/2): classical ones

Getting back to the forecasting equations (2)–(5), we now present the different forecasting methods considered in this paper. We will do so in some generality, encompassing all these equations under the umbrella (1):

$$\hat{s}_{t+1} - s_t = \sum_{j=1}^{N} \beta_{j,t} f_{j,t}$$
.

We recall that in our view, the coefficients $\beta_{j,t}$ are to be picked according to some rule; they do not need to be understood as estimating some unknown underlying true value.

The no-change forecasting method. The first strategy consists of choosing $\beta_{j,t} = 0$ for all j at each round t, that is, of forecasting \hat{s}_{t+1} by s_t . We call it the no-change forecasting method.

Rolling OLS regression. This is the most standard technique in the literature—"the convention" as e.g., Cheung et al. [2005] state. The idea is to truncate the available information to account for the most recent relationships between variables that can change through time because of policy changes (for example, a change to Taylor-rule based monetary policy), structural changes in the economy (such as shifting relationships between the money stock and inflation), etc. Because of this forecasting method, we need a training period, which we set to $t_0 = 120$ months as is standard in the literature (see Molodtsova and Papell [2009]). The rolling OLS regression picks, for months $t \ge t_0$,

$$(\beta_{1,t}, \ldots, \beta_{N,t}) = \underset{\beta_1, \ldots, \beta_N \in \mathbb{R}}{\operatorname{arg \, min}} \sum_{\tau=t-t_0+1}^t \left(s_{\tau} - s_{\tau-1} - \sum_{j=1}^N \beta_j f_{j,\tau-1} \right)^2.$$

Recursive OLS regression. This is another standard technique in the literature. It consists of choosing, for months $t \ge t_0$,

$$(\beta_{1,t}, \ldots, \beta_{N,t}) = \underset{\beta_1, \ldots, \beta_N \in \mathbb{R}}{\operatorname{arg \, min}} \sum_{\tau=1}^t \left(s_{\tau} - s_{\tau-1} - \sum_{j=1}^N \beta_j f_{j,\tau-1} \right)^2,$$

i.e., all past time instances, not only the t_0 most recent ones, are used to form the prediction.

3.3. Forecasting methods (2/2): new ones from machine learning

The new forecasting methods considered in this study stem from the field of machine learning, where they are already standard methods for the robust online prediction of quantitative phenomena by aggregation of basic predictors (fundamentals). The monograph by Cesa-Bianchi and Lugosi [2006] summarizes the research performed around them in the period 1989–2006.

These methods have been applied exactly as they are stated below (i.e., we provide no tweak on these methods) in the following fields: forecasting of the air quality (see, e.g., Mauricette et al. [2009], Mallet [2010], Debry and Mallet [2014]); the forecasting of electricity consumption (see, e.g., Devaine et al. [2013], Gaillard and Goude [2015]); the forecasting of the production of oil reservoirs (work in progress). We underline that thus, these methods are not ad hoc methods constructed solely for the problem of predicting exchanges rates.

The theoretical out-of-sample guarantees that they do come with are of the following form: for all possible bounded sequences of exchange rates and of fundamentals, their predictions are such that

$$\sum_{t=1}^{T} \left(\widehat{s}_{t} - s_{t-1} - \sum_{j=1}^{N} \beta_{j,t-1} f_{j,t-1} \right)^{2} - \inf_{\beta^{\dagger} \in \mathcal{F}} \sum_{t=1}^{T} \left(\widehat{s}_{t} - s_{t-1} - \sum_{j=1}^{N} \beta_{j}^{\dagger} f_{j,t-1} \right)^{2} \leqslant B(\mathcal{F}, T)$$
(11)

with $B(\mathcal{F},T)\ll T$ and where $\mathcal{F}\subset\mathbb{R}^N$ is some comparison class (the forecasting methods presented below are independent of the choice of \mathcal{F} , only the bound $B(\mathcal{F},T)$ is sensitive to it). We explain below what we mean by a comparison class: a set of candidates for the underlying expression of the exchange rates in terms of linear combinations of fundamentals. With the algorithms presented below, such linear combinations can be given by \mathcal{F} equal to the set of all point mass combinations (weights that equal 1 for one fundamental and are null for all others) or to some bounded Euclidean ball of \mathbb{R}^N (e.g., all linear weights with Euclidian norm bounded by some constant U, and then, U appears in the bound).

In particular, from (11), dealing separately with the training period of size t_0 (for which an error bounded by something of the order of t_0 is made in the worst case), dividing by $T - t_0$, and

taking square roots, we get the following guarantee on out-of-sample RMSEs:

$$\lim_{T \to \infty} \left\{ \sqrt{\frac{1}{T - t_0}} \sum_{t=t_0+1}^{T} \left(\widehat{s}_t - s_{t-1} - \sum_{j=1}^{N} \beta_{j,t-1} f_{j,t-1} \right)^2 - \inf_{\beta^{\dagger} \in \mathcal{F}} \sqrt{\frac{1}{T - t_0}} \sum_{t=t_0+1}^{T} \left(\widehat{s}_t - s_{t-1} - \sum_{j=1}^{N} \beta_j^{\dagger} f_{j,t-1} \right)^2 \right\} \leqslant 0.$$
(12)

Let us comment in words about the properties of these out-of-sample guarantees. First, they rely on no stochastic modeling, they are achieved for all possible bounded sequences of exchange rates and of fundamentals: these performance bounds are deterministic. The bound $B(\mathcal{F},T)$ actually also depends on the range in which the exchange rates and the fundamentals lie; but it is a uniform bound. Second, what they truly ensure is that the forecasting method has asymptotically an average performance as good as or even better than the one of the best constant linear combination in $\mathcal{F} \subset \mathbb{R}^N$, i.e., the best fixed model with coefficients in \mathcal{F} . But of course, as mentioned above, such a model does not need to exist, since no assumption of stochastic modeling is required. It just turns out that the methods presented below mimic the performance of the best model if applicable (of the best fixed linear forecasting equation otherwise). Put differently, we can interpret the infimum in (11) (the error suffered by the best pick in the comparison class) as measuring some approximation error (how well in hindsight the fundamentals can predict the exchange rates) while the $B(\mathcal{F},T)$ term is a sequential estimation error (the price to pay for facing a sequential rather than a batch problem).

One may wonder where the catch is—the type of guarantee exhibited above seems too good to be true. Actually, getting close to or slightly better than the performance of the best fixed linear forecast is not necessarily good enough for obtaining great forecasts: the best fixed linear forecast can have a poor performance. This may be in particular true in the setting of exchange rates where fundamentals are known to have a questionable value for modeling purposes. This all is in contrast with the other fields of application mentioned above such as forecasting of electricity consumption or air quality. Finally, we note that it can be shown that out-of-sample guarantees like (12) cannot be achieved by rolling OLS regression¹⁵. Performance guarantees on the latter forecasting method (if any exists) thus would require at least a stochastic modeling of the sequence of exchange rates.

¹⁵Available upon request.

Is it better to use coupled or decoupled fundamentals? The short answer will be: there is no theoretical reason to prefer one or the other. Indeed, while the approximation errors of the decoupled version are always smaller than the ones of the coupled versions (just because they correspond to an infimum taken on a larger set), the bounds $B(\mathcal{F},T)$ on the sequential estimator errors in (11) always increase with the number N of independent fundamentals, thus are larger with decoupled than with coupled fundamentals. As the total errors suffered by our forecasting methods are the sums of these two errors, their behavior as N increases is unclear—and actually, they can either increase or decrease.

EWA: the exponentially weighted average strategy with discount factors. It was introduced by Vovk [1990], Littlestone and Warmuth [1994] as early as in the beginning of the 90s and further understood and studied by, among others, Cesa-Bianchi et al. [1997], Cesa-Bianchi [1999], Auer et al. [2002]. We present a slight generalization in which past prediction instances get slightly more weight when they are more recent. This forecasting method relies on different parameters: a sequence (η_t) of positive numbers referred to as the learning rates; a nonnegative number γ referred to as the discount factor; a positive number $\kappa > 0$ referred to as the discount power. It picks the weights according to

$$\beta_{j,t} = \frac{1}{Z_t} \exp\left(-\eta_t \sum_{\tau=1}^t \left(1 + \frac{\gamma}{(t+1-\tau)^{\kappa}}\right) \left(s_{\tau} - s_{\tau-1} - f_{j,\tau-1}\right)^2\right)$$
(13)

where Z_t is a normalization factor 16 ,

$$Z_{t} = 2 \exp\left(-\eta_{t} \sum_{\tau=1}^{t} \left(1 + \frac{\gamma}{(t+1-\tau)^{\kappa}}\right) \left(s_{\tau} - s_{\tau-1}\right)^{2}\right) + \sum_{j=1}^{N} \exp\left(-\eta_{t} \sum_{\tau=1}^{t} \left(1 + \frac{\gamma}{(t+1-\tau)^{\kappa}}\right) \left(s_{\tau} - s_{\tau-1} - f_{j,\tau-1}\right)^{2}\right).$$
(14)

Because of this factor Z_t and of the exponent function in the definition equations of the $\beta_{j,t}$, this strategy is referred to as the exponentially weighted average strategy (EWA strategy in short). Note

This is the version for decoupled experts. With coupled experts, the number of summands is reduced by a factor of 2 in the defining sum over j and concomitantly the factor 2 in the first term in the definition of Z_t is to be replaced by 1. This is because this first term accounts for the no-change prediction, which in its decoupled version as in (14) corresponds to the difference of 2 zero variations (one for the value of the currency of each country), instead of 1 zero variation (of the exchange rate) in the coupled case.

that the obtained weights form a sub-convex weight vector: the components $\beta_{j,t}$ are nonnegative and sum up to something smaller than or equal to 1. (The missing mass to 1 is to be interpreted as a measure of the confidence that no change will take place between s_t and s_{t+1} .)

A study of the theoretical out-of-sample guarantees of EWA in the presence of the γ factors was offered in Stoltz [2010, Theorem 3], see also Cesa-Bianchi and Lugosi [2006, § 2.11], and can be instantiated as follows in our context. For all choices of γ and choices of κ and of non-increasing sequences (η_t) such that, as $t \to +\infty$,

$$t \eta_t \longrightarrow +\infty$$
 and $\eta_t \sum_{\tau=1}^t \frac{1}{\tau^{\kappa}} \longrightarrow 0$, (15)

the desired guarantee (12) holds, for the set $\mathcal{F}=\mathcal{M}$ of all point-mass vectors. For instance, $\eta_t=1/\sqrt{t}$ and $\kappa=2$ would be suitable choices but many other choices associated with the desired theoretical guarantees exist.

More precisely 17 the bound $B(\mathcal{M}, T)$ of (11) equals

$$B(\mathcal{M}, T) = \frac{2 \ln N}{\eta_T} + \sum_{t=1}^{T} \frac{\eta_t}{2} L^2 + L \sum_{t=1}^{T} (\exp(2L\eta_t B_{t-1}) - 1)$$

where

$$B_{t-1} = \sum_{\tau=1}^{t-1} \frac{\gamma}{\tau^{\kappa}}$$

and L is a bound on the quadratic errors $(s_t - s_{t-1})^2$ and $(s_t - s_{t-1} - f_{j,t-1})^2$ as t and j vary.

EWA in practice: how to choose its parameters. Instead of reporting in our simulation study the performance of the EWA method for several well-chosen sets of parameters, as is usual in the machine-learning literature, we re-use a more operational approach developed in Devaine et al. [2013] and only report the results obtained for one instance of the method. The latter selects at each time instance t, in a data-driven way (detailed below), parameters $\eta_t > 0$ and $\gamma_t > 0$ to be used in (13) (for $\gamma_t > 0$: instead of the fixed parameter γ initially considered). We choose $\kappa = 2$, which is common in the literature and which allows us to have a theoretical bound given condition (15).

¹⁷See Chapter 6 of the technical report by Mallet et al. [2007], which is cited as the main ingredient in the proof of Theorem 3 of Stoltz [2010]. Also, readers aware of the bounds that can be proved for the square loss in our context may note that we do not exploit its exp-concavity through a well-chosen fixed learning rate; this is for practical performance.

The parameters η_t and γ_t for the prediction of instance t+1 are respectively selected in the grids \mathcal{E}/\sqrt{t} and \mathcal{G} , where

$$\mathcal{E} = \left\{ m \times 10^k, \ m \in \{1, 2, 5\} \text{ and } k \in \{-4, -3, -2, -1\} \right\} \cup \{1\},$$
 and
$$\mathcal{G} = \left\{ 0 \right\} \cup \left\{ m \times 10^k, \ m \in \{1, 2, 5\} \text{ and } k \in \{0, 1, 2\} \right\} \cup \left\{ 1\,000 \right\}$$
$$= \left\{ 0, 1, 2, 5, 10, 20, 50, 100, 200, 500, 1000 \right\}.$$

These grids are (somewhat) evenly spaced in a logarithmic scale¹⁸. To predict instance t+1, we resort to (13) with the pair of parameters (η, γ) in the grids \mathcal{E} and \mathcal{G} whose associated EWA strategy with sequence of learning parameters η , $\eta/\sqrt{1}$, ..., $\eta/\sqrt{t-1}$ and fixed discount factor γ performed best in total on instances 1 to t.

SRidge: sequential ridge regression with discount factors. The issue with (recursive or rolling) linear regressions is that they tend to overfit past data, i.e., they lead to good in-sample predictions but poor out-of-sample ones. To prevent this overfitting, one can add what is called a regularization term to the squared error, to help controlling (reducing) the range of the components $\beta_{j,t}$ of the linear vector picked. We implement below this second path, with a forecasting method called sequential ridge regression with discount factors.

The ridge regression was introduced by Hoerl and Kennard [1970] in a stochastic (non-sequential) setting. What follows relies on recent new analyses of the ridge regression in the machine learning community; see the original papers by Vovk [2001], Azoury and Warmuth [2001] as well as the survey in the monograph by Cesa-Bianchi and Lugosi [2006]. The sequential ridge regression (without discount factors and for a constant regularization factor $\lambda \ge 0$) picks the weights

$$(\beta_{1,t}, \dots, \beta_{N,t}) = \underset{\beta_1, \dots, \beta_N \in \mathbb{R}}{\arg\min} \left\{ \lambda \sum_{j=1}^N \beta_j^2 + \sum_{\tau=1}^t \left(s_\tau - s_{\tau-1} - \sum_{j=1}^N \beta_j f_{j,\tau-1} \right)^2 \right\}.$$
 (16)

We state the associated performance bound (11). It is in terms of the classes \mathcal{F}_U of linear weights with Euclidean norm bounded by U>0 (where the bound holds for all U>0 simultaneously). We denote by L a bound on the exchange rates and on the fundamentals: a value such that $s_t \in$

¹⁸More details on the influence of a particular choice of a grid are given in Section B.

[-L,L] and $f_{j,t} \in [-L,L]$. Then, for all U>0, the bound $B(\mathcal{F}_U,T)$ of (11) equals 19

$$B(\mathcal{F}_U, T) = \lambda U^2 + 4NL^2 \left(1 + \frac{NTL^2}{\lambda}\right) \ln\left(1 + \frac{TL^2}{N\lambda}\right).$$

In particular, when λ is well chosen (e.g., of the order of \sqrt{T}), one has

$$B(\mathcal{F}_U, T) = \mathcal{O}(\sqrt{T} \ln T) \ll T$$
.

The recursive linear regression corresponds to the special case when $\lambda=0$, but no theoretical bound is offered in this case.

As announced above, we used a common variant of the classical ridge regression presented above so as to focus more on recent observations. This is obtained via discounting: the sequential ridge regression (with discount factor γ and for a constant regularization factor $\lambda \geqslant 0$) picks the weights

$$(\beta_{1,t}, \dots, \beta_{N,t}) = \underset{\beta_1, \dots, \beta_N \in \mathbb{R}}{\arg \min} \left\{ \lambda \sum_{j=1}^{N} \beta_j^2 + \sum_{\tau=1}^{t} \left(1 + \frac{\gamma}{(t+1-\tau)^2} \right) \left(s_{\tau} - s_{\tau-1} - \sum_{j=1}^{N} \beta_j f_{j,\tau-1} \right)^2 \right\}.$$

We took above the same discount power $\kappa=2$ as in the case of EWA. Again, the parameters λ and γ are to be calibrated well: instead of reporting the performance for various pairs of fixed possible values, we resort to a more operational approach and select parameters λ_t and γ_t to be used for month t+1 based on past data. We do so in the same way as described above for EWA

$$\lambda \sum_{j=1}^{N} \beta_j^2 + N \ln \left(1 + \frac{TL^2}{N\lambda} \right) \max_{t \leqslant T} \left(\hat{s}_t - s_t \right)^2.$$

Nevertheless, this does not imply a logarithmic bound as the term $\max_{t \leqslant T} (\widehat{s}_t - s_t)^2$ may be large, as pointed out by Gerchinovitz [2011, page 66]. One can show (details upon request) that this maximum is however never larger than

$$4\max\left\{L^2,\ \frac{NTL^4}{\lambda}\right\},\,$$

hence our stated bound.

The original bound of Azoury and Warmuth [2001, Theorem 4.6], as cited by Cesa-Bianchi and Lugosi [2006, Section 11.7] in the special case $\lambda = 1/2$, reads

and consider to that end the same grid \mathcal{G} for the γ parameters and the grid for λ given by

$$\begin{split} &\Lambda = \{0\} \cup \left\{m \times 10^k, \;\; m \in \{1,2,5\} \; \text{and} \; k \in \{1,2,3\} \right\} \cup \{10\,000\} \\ &= \left\{0,\,1,\,2,\,5,\,10,\,20,\,50,\,100,\,200,\,500,\,1\,000,\,2\,000,\,5\,000,\,10\,000\right\}. \end{split}$$

4. Data and fundamentals used

Our main sample includes major floating exchange rates between March 1973–December 2014 (at most we have 502 data points per currency). We use two types of exchange rates. The first sample contains average monthly exchange rates used for example by Molodtsova and Papell [2009]. The second sample has end-of-month exchange rates. The average exchange rates have less variance than end-of-month rates; most available fundamentals (such as prices, output, but also the interest rates available from the IFS data set) are also averages. It makes hence sense to scrutinize the performance of the methods for both types of data. We also use a supplementary sample for March 1973-June 2006 taken directly from Molodtsova and Papell [2009] in order to compare our methods with the fundamentals taken from a Taylor-rule based exchange rate model that are deemed in the literature as generally being the most successful in obtaining predictability²⁰ of the exchange rate at the short 1-month horizon. For the sample 1973-2014 we try to extend the same data series for the fundamentals as used in Molodtsova and Papell [2009], but some of them were discontinued. As a result, we tried to find the closest substitutes possible for the entire period 1973–2014 from similar sources (IMF, OECD) through Datastream. The exchange rates are taken from the Federal Reserve Bank of Saint Louis database (for the average rates) and the IMF IFS database (for the end-of-month series). A detailed description of the data is given in Appendix A.

We principally study the behavior of 12 major floating currencies that are active throughout the 1973–2014 period. The introduction of the Euro in 1999 constrains the sample for some continental Europe currencies (FRF/USD, DEM/USD, ITL/USD, NLG/USD, PTE/USD). Even for the active currencies (USD/GBP, JPY/USD, CHF/USD, CAD/USD, SEK/USD, USD/AUD, DNK/USD), however, it was not possible to obtain fundamentals series for the entire 1973–2014 period (see details for each series in Appendix A).

Our fundamentals were formed as follows. The inflation differentials are calculated as 12-

²⁰As explained in Footnote 33 predictability does not imply forecastability; the Taylor models studied in the existing literature were not able most of the time to deliver the latter for short-period horizons.

month changes in consumer price indexes (CPI). We use a money market rate or 3-month interest rate differentials for the interest rate based fundamentals. For differences in the money stock growth and output growth we use the preceding 12-month trends in these variables. We do not detrend, filter or seasonally adjust the data. The output gaps for the Taylor-rule fundamentals are percentage deviations from "potential" output that was computed including (i) a linear trend, (ii) a quadratic trend, (iii) a linear and quadratic trends, or (iv) a Hodrick-Prescott filter using the data available prior to the date for which the output gap was calculated.

5. Results

5.1. Results for "classic" fundamentals

As discussed above, we use two samples with different exchange rate series (end-of-month and monthly-averages exchange rates), both present in the existing literature²¹. Our base results are shown in Tables 1–2 for the end-of-month and average exchange rate samples respectively. We show there the RMSE ×100 of the no-change prediction²² (column 1) and the Theil²³ ratios (columns 2, 5, 8 and 11) of predictions, as well as the corresponding p-values of the CW and DM tests. For the former, they are referred to as "CW p-value" and can be read in columns 3, 6, 9 and 12. As explained in Section 3.1.3, they should be used with extreme caution in our context: they correspond to fundamentals being useful (H'_1) or not (H'_0) in some existing underlying model – a matter of predictability not necessarily associated with better forecasting abilities. On the other hand, the p-values for the DM tests are more reliable in our context, as the DM test is a very general test for comparing forecasting abilities, that works under almost no assumption on the data (see Section 3.1.2 for details²⁴ and references). These p-values are called "DM p-value" (columns 4, 7, 10 and 13) and evaluate the hypothesis H_0 that the difference in the forecasting

²¹There is no reason to favor one over the other. For the end-of-month series the no change can be reinterpreted as a random walk with no drift prediction. However, most fundamentals available (and also used in the literature) are actually monthly averages, so for the end-of-month rates the random walk is given an advantage. What is more, the higher variance due to noise of the end-of-month exchange rates (for which there are numerous explanations, see for example Evans and Lyons [2005]) makes it inherently more difficult to uncover patterns even if they exist.

²²These may differ for the same currency pair for different fundamentals given that the latter are available for different periods (see the data Tables 8– 9 in the Appendix A), so the forecasted periods may vary.

²³In our context, this is the ratio of the RMSE predicted by the forecasting method of interest and that of the no-change prediction; a ratio below 1 means the particular method gave lower RMSE than the no-change prediction.

 $^{^{24}}$ We recall in particular that we chose the truncation lag H for each Diebold and Mariano [1995] test in a conservative and data-driven way, see the maximum in the denominator of (7). Realized values of the argument of this maximum are available upon request.

performance of the method under scrutiny is not significantly better than the one of the no-change prediction, against the alternative hypothesis H_1 that it is so.

The sets of columns correspond to the forecasting methods: rolling OLS, recursive OLS, sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors, respectively; while the sets of lines correspond to the sets of fundamentals discussed in 2.2: PPP fundamentals, UIRP fundamentals, monetary model fundamentals, and all these fundamentals altogether.

Monthly-averaged exchange rates sample. We start with discussing Table 1. We observe that both methods give better RMSE than the no change prediction for all currency pairs for all fundamentals considered. These improvements are statistically significant for the majority of cases. DM tests indicate statistically significant different results from the no–change prediction at the 10% level in 11 out of 12 cases for the PPP fundamentals or when using all considered fundamentals (lowest subtable²⁵ of Table 1). We obtain similar results for the exponentially weighted average strategy with discount factors method. The RMSE improvements in comparison to the no-change prediction are slightly better for sequential ridge regression with discount factors than for the exponentially weighted average strategy with discount factors, and for some currency pairs and forecasting equations reach up to 7% (for example, the Italian lira / the U.S. dollar for the monetary fundamentals) 26 .

End-of-month changes sample. We first notice looking at Table 2 that the quadratic variation of the end-of-month exchange rates (as measured by the RMSE) is higher than that of the average monthly exchange rates. Probably due to this larger noise the sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors methods are not able to forecast the exchange rates better in many circumstances in comparison to the no change prediction using the same fundamentals as before. Still, Table 2 shows that on this sample, we obtain better predictions than the no change prediction for the sequential ridge regression with

 $^{^{25}\}text{CW}$ tests show for sequential ridge regression with discount factors that all improvements are statistically significant at (at least) the $5\,\%$ level.

²⁶The only currency pair for which we cannot find a method and a set of fundamentals that improves the RMSE in a statistically significant way is the Portuguese Escudo / U.S. dollar; perhaps this is also due to the fact that the series available for this pair were the shortest in the sample (for example, the data on money aggregates was available only for 1979–1998, and 1983–1998 for interest rates) and it was not possible for the methods to fully establish their superior performance given the 120-month training period.

discount factors and the exponentially weighted average strategy with discount factors methods using the PPP or UIRP fundamentals. The improvements, however, are small (not higher than 1.3% in terms of the RMSE) and most of the time we cannot reject the hypothesis that the compared forecasting methods perform similarly. The exponentially weighted average strategy with discount factors is more successful as a method; both for the PPP and the UIRP it can improve upon the no change prediction for 10 out of 12 currency pairs. In the case of the PPP fundamentals, this performance is statistically superior at the 10% level in 5 out of 12 cases according to the DM tests (and in 9 out of 12 according to the CW tests). The sequential ridge regression with discount factors is less successful, being able to outperform the no-change prediction 7 out of 12 cases for the PPP fundamentals and 6 out of 12 for the UIRP fundamentals. There is no evidence that using the monetary fundamentals nor all fundamentals one could improve upon the no change prediction with the methods scrutinized.

Remarks about the performance of our methods. The conclusions from Tables 1-2 are striking. For the end-of-month sample we can claim forecastability of the exchange rates for the PPP and UIRP fundamentals using the exponentially weighted average strategy with discount factors, a feat that no study claimed so far at such a short range. In the average exchange rate sample, using the sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors we are able to beat the no change in terms of the RMSE improvement for *all* currency pairs for the period 1973-2014 for *all* fundamentals considered. The gains are not higher than ca. 7% in terms of the RMSE: we are not able to extract much more information from the fundamentals than what a no-change prediction contains. Yet, comparing to the existing literature as surveyed by Rossi [2013] these gains are substantial.

When we combine all fundamentals together – in the spirit of the information aggregation ability of machine learning techniques – we do not gain in terms of forecasting accuracy. Sometimes the forecasts turn to be actually less precise according to the RMSE criterion (cf. the end-of-month sample). This may be an indication that the fundamentals we study contain pretty much the same information²⁷. There is also a cost suffered by our methods while adding fundamentals that do not

²⁷Theoretically, with additional assumptions there are links between the theories from which our "classic" fundamentals are taken. Interest rates could be high in a high inflation environment, and with the same risk premium both UIRP and PPP would predict a depreciation of a such a currency. If inflation is primarily a monetary phenomenon, high monetary growth would be correlated with high inflation and interest rates.

perform well – hence parsimony seems to be preferred in our setting as found in other studies of exchange rate predictability using other methods. To rephrase this discussion in the terms used in Section 3.3, the potential decrease in the approximation error when considering all fundamentals is too small to be compensated by a larger sequential estimation error linked to dealing with more fundamentals.

Performance of the rolling and recursive OLS strategies. The same fundamentals seem not to have much forecasting power when evaluated using the classical methods considered in the literature, either the rolling or recursive OLS regressions, no matter what sample we use. For neither of the currency pairs we obtain an improvement in terms of RMSE upon the no change prediction (which is already rare) that is statistically significant at conventional levels using the DM tests. The sole difference in performance comes when all fundamentals and CW–tests are used for the monthly-averaged exchange rates sample. Even then, though, there are no improvements in the Theil ratios except for the PTE/USD exchange rate. These conclusions are in line with the common knowledge and the existing empirical literature documenting that fundamentals do not allow for systematic improvements in forecasting (see Rossi [2013]).

Direct comparison of the machine learning and OLS strategies. Instead of making an indirect comparison between our forecasting methods and the ones typically used in the literature we can directly compare their performance by comparing the Theil ratios (of the machine learning method RMSE with respect to the OLS methods) and perform an appropriate DM test. The results of this exercise are shown in Tables 3–4. The first conclusion is that both methods do globally better than the OLS methods while comparing the Theil ratios no matter what type of fundamentals on what sample are considered (so even for monetary models and all fundamentals for the end-of-month sample when the machine learning methods were not successful in beating the no change prediction). Interestingly, largest gains are obtained when all fundamentals are used (last subtable of each table) – up to 17 % in terms of RMSE (for the SEK/USD pair in the averaged data of sequential ridge regression with discount factors vs. rolling OLS). This is partly because of the weak performance of the OLS methods as the number of fundamentals used grows, which was recognized by the literature (Rossi [2013]), while the effect of this growing number is milder in the case of our machine learning methods (see, e.g., Mauricette et al. [2009]). The advantage

of both sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors is especially visible for the averaged exchange rate sample where the RMSE improvements are statistically significant at conventional levels for the vast majority of cases.

Finally, we note that some descriptive statistics show that this average good performance of our new methods does not come at the cost of local disasters, on the contrary: the forecasting errors seem to be uniformly better over time. Indeed, when computing the quantiles of the difference between the forecasting error of the no change minus the one of the methods under scrutiny, as is reported in Tables 11–12 of Appendix C, we see that these differences are uniformly much larger for the sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors than for the recursive or rolling OLS regressions when the methods exhibit forecastability for a given currency pair. By "uniformly" we mean here that quantiles of the same order for two series of differences are almost always ranked in the same manner, the ones corresponding to the sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors being larger than the ones for the rolling or recursive OLS regressions.

5.2. Robustness checks

We conducted several robustness checks to verify whether our results are not due to some quirk related to the samples that we constructed. This included several exercises²⁸.

First, we tried the "coupled" version of the fundamentals as well, that is, the forecasting equations in the form without taking into account the possibility that one of the country series pertaining to a particular fundamental would have a different weight in the linear combination. The results can be seen in Tables 13–14 of Appendix C. We also considered 6–month inflation, money stock or output changes in the decoupled version to see whether a different way of constructing fundamentals matters. All the above exercises did not change qualitatively our results.

Next, we reran the basic combinations of fundamentals and methods from Tables 1–2 on our data trimmed to shorter samples 1980–2014, post-Plaza accord period 1985–2014, the post ERM-crisis period 1992–2014 and 1973–2006 so as to see whether the inclusion of a particular period (e.g., high inflation period of 1970s, or post financial crisis period) drives the results. We show

²⁸All results not shown due to space limitations are available upon request.

the results for the 1980-2014 sample in Tables 17–18. We also conducted a separate exercise on the 1973–2006 data taken directly from Molodtsova and Papell [2009], shown in Table 19 (our averaged exchange rate sample is an extension of their data set till 2014, as noted in Section A). What we observe in general is that the good forecasting properties of our methods still hold, albeit the observed gains against the no-change prediction are typically smaller. Trimming at the beginning of the sample affects especially the European currencies that became part of the Euro. This may indicate that in practice, longer data series are in general beneficial for the (asymptotic) guarantees to kick in, as indicated in Section 3.3.

We also tried to work with what we call "absolute" fundamentals. This means that instead of inflation differentials we directly used price, money and output levels as substitutes for the PPP and monetary model fundamentals²⁹. The results here are weaker. We can still obtain improvements with the sequential ridge regression with discount factors. A problem arises, though, with the exponentially weighted average strategy with discount factors in that the forecast errors involving deviations of the exchange rates from the absolute fundamentals considered are much larger that those for relative fundamentals used in our original exercise; and the weights assigned to the absolute fundamentals by the method tend to be small (i.e., the predictions output by the method are close to the ones of no change). This is a documented fact: the method does not admit well fundamentals that err too much in comparison to one (here, the no-change prediction) that performs well. This is in contrast with the sequential ridge regression with discount factors, which can correct the fundamentals for proper scaling.

Finally, we tried the actually realized fundamentals. That is, instead of "predicting" the next period change in fundamentals by their past values, we fed the actual values that were realized in course of economic activity. As in the literature started by Meese and Rogoff [1983], these do not lead to good or better forecasts.

5.3. Taylor-rule fundamentals

A recent strand of literature identified Taylor-rule fundamentals as useful in achieving exchange rate predictability (which is not the same as forecastability; see Footnote 33). As Taylor-rule fundamentals can be created from the data in our possession, we also reran such forecasting equations

²⁹The UIRP-model inspired forecasting equation has no such counterpart. "Absolute" fundamentals A_t were created by adding price, money or output changes to an initial exchange rate in our data sets. Then the deviation of the actual nominal rate from its thus calculated "fundamental" value $A_t - s_t$ was used to predict $s_{t+1} - s_t$.

based on (5) with these fundamentals for our basic sample and present them in Tables 15–16 of Appendix C. We do not find forecastability on the end-of-month exchange rate sample. We find forecastability on the FRED average exchange rate sample but it is slightly worse than for the basic "classic" fundamentals. In Table 20 of Appendix C we use the Molodtsova and Papell [2009] data to allow a direct comparison with their results. There is no forecastability with either rolling or recursive OLS methods. However we find similar patterns as Molodtsova and Papell [2009] regarding predictability with the use of Clark-West (West [1996], Clark and West [2006, 2007]) tests designed for model comparison in several instances even for the OLS-based methods. But such identified predictability by the OLS methods does not help in forecasting exchange rates — most of the Theil ratios remain above 1.

5.4. Directional tests

A different, though secondary, measure of forecast quality that has been employed in the literature are "directional tests", i.e., tests whether the forecasting methods are able to predict the direction of change of the exchange rate better than a fair coin toss. Although the methods considered in this paper are not constructed to be efficient in this regard as their focus is on controlling RMSE (see Section 3.3), we test how the predictions obtained by these methods fare in this dimension. In Tables 5–6 we exhibit the percentage success of either of the methods (OLS regressions as well as sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors) in forecasting the direction of change of exchange rates with the corresponding DM p-value of a test of the difference with a fair coin toss.

The patterns that emerge are interesting. Machine learning methods are able to predict the correct direction in several cases 60 % of the time (with a maximum of 63.4 %) – improvements which are often statistically significant. Out of the two methods, the exponentially weighted average strategy with discount factors does better on both samples. The success in predicting RMSE goes typically hand in hand with the success of directional predictions. The worst performance is obtained for monetary fundamentals within the end-of-month sample. This time, the OLS-based methods achieve some success (especially the rolling regression) in improving the direction of the change of the exchange rate when all fundamentals are considered (though it is significant at conventional levels for at most 7 out of 12 currency pairs for the average exchange rate sample). However, the percentage improvements over a coin toss are typically small, almost always lower

than for the machine learning methods for the same currency pair, and never larger than $59.9\,\%$.

5.5. Economic evaluation

In Table 7 we report various economic criteria used to evaluate the economic benefits one would obtain using our forecasting models versus just a random walk. As this involves calculating investment returns we report the performance for the end-of-month exchange rates. We do this for the exponentially weighted average strategy with discount factors (EWA) for which we obtained also generally improvements in terms of the RMSE as shown in Table 2 to verify that detected improvements would also lead to superior performance from investor's perspective.

The idea is to assess the utility of the forecasts from the point of view of investors (see Corte et al. [2012] or Corte and Tsiakas [2012]) who build one-period portfolios using information known at that given period under standard mean-variance preferences with a target standard deviation. It is assumed in this context that investors decide over investing in money market rates of different countries taking into consideration the forecasted movements of the exchange rates. Then various performance measures of such portfolios are compared to see whether a forecasting method generates predictions that can be profitably used by investors over those of a no-change model. (See Appendix B for further details.)

We consider the standard measures in the literature: a direct comparison of annualized returns for a portfolio using forecasts and the random walk, annualized performance fee (in bps), annualized premium return (in bps), the annualized Sharpe and Sortino ratios. Given the oft discussed worry that optimal portfolio choice applied to investment problems with exchange rates generates frequently large portfolio weights (implying investors are taking very risky bets) we also study portfolio allocation when we restrict the weights $w_{1,t},\ldots,w_{K,t}$ on the K=7 currencies³⁰ so that $w_{1,t}^2+\ldots+w_{K,t}^2\leqslant c$, which is a standard constraint in the machine-learning literature. (See B for the technical details of how this can be achieved. When we put such a constraint, we provide the corresponding value for c; otherwise, we write "no c" in our tables.) For a benchmark comparison we also show the performance of a portfolio that results from a uniform allocation of funds across assets denominated in different currencies.

We see that in line with the results shown in Table 2 and obtained from RMSE comparisons of

³⁰We consider: GBP, JPY, CHF, CAD, SEK, DNK and AUD, but not the currencies that later merged to become the Euro as they would allow for a shorter overall time horizon.

forecasts, the usage of our methods would add value to investors using them over strategies based on the no-change predictions for the PPP or UIRP fundamentals. For the PPP, annualized returns are 13.48% for the EWA-based strategy while 11.83% for the random walk. But, the risk of the investments is also important and investors would be willing to pay a performance fee of 229 bps or would obtain a premium return of 277 bps by sticking to our forecasts. Both the Sharpe and the Sortino ratios would be higher for the EWA-forecast based portfolios versus no-change ones (0.84 vs. 0.64 and 1.08 vs. 0.68 respectively). The same holds for the UIRP-fundamentals. Returns, the Sharpe ratios and Sortino ratios for a portfolio based on EWA-algorithm forecasts are higher (12.13% vs. 11.44%, 0.78 vs. 0.62 and 1.38 vs. 0.65 respectively) and investors would be willing to pay a performance fee of 192 bps or obtain a premium return of 268 bps. This no longer holds true for the predictions based on "monetary model" fundamentals or when we use all fundamentals together. One observes that constraining the weights (with c=1 or c=2) improves greatly the results along all dimensions for all types of fundamentals. We achieve a premium return larger than 360 bps with the UIRP fundamental and the constraint c=2.

In Table 21 we repeat the exercise taking out the DNK/USD exchange rate from consideration which is highly correlated with the SEK/USD and CHF/USD; this correlation creates some instability in the weights and leads in our calculations to optimally short one of these currencies with a very high weight and go long in the others. We see the results of economic criterion evaluation greatly improve, especially for the "monetary model" fundamentals or when we use all fundamentals together. In Tables 22 and 23 we impose an additional constraint on the weights requiring that the sum of the weights on the currencies is smaller than 1, i.e., $w_{1,t} + \ldots + w_{K,t} \leq 1$.

We observe thus that overall, investors would benefit from the usage of our forecasts over the random walk prediction and would be willing to pay up to slightly more than 360 bps for our best predictions in our basic scenarios.

6. Related literature

Exchange rate forecasting literature. Rossi [2013] and Corte and Tsiakas [2012] provide comprehensive recent reviews of the exchange-rate forecasting literature so we keep ours at a bare minimum. There were different ways in which researchers tried to cope with the negative result of Meese and Rogoff [1983] that showed that the simple exchange rate models from the 1970s

(i.e., that proposed the fundamentals we study here) did poorly in comparison to a random walk without drift (a no-change prediction) in forecasting the exchange rates in the floating period after 1973 for short forecast horizons. One way, that we pursue here, was to use better tools to extract information from the data (better forecasting tools, not necessarily statistical/econometric tools). Some of the solutions proposed involved for example cointegration techniques as in Mark [1995] combined with the usage of panel data such as in Mark and Sul [2001], long samples in panels as in Rapach and Wohar [2002] and including a large set of countries as in Cerra and Saxena [2010]. These techniques typically are used for longer-horizon (above 1 year) forecasts when it is believed that the long-run relationship as modeled by the cointegrating equations kicks in. The cited studies obtain some success in demonstrating predictability and/or forecastability at longer horizons (typically longer than 2 years). We focus on short-term forecasts and use single-equation methods – a setup in which forecastability was not ascertained largely till today. For the short-run, Greenaway-Mcgrevy et al. [2012] obtain considerable success in outpredicting the no-change prediction using factor analysis extracting the factors from the exchange rates themselves (but not the economic fundamentals)³¹. Bianco et al. [2012] obtain forecastability for the Euro/U.S. dollar rate at weekly to monthly horizons using a stylized econometric model that mixes information from different fundamentals arriving at different frequencies. Our method is more general, directly geared for prediction, and importantly comes with some theoretical guarantees on the convergence of the RMSEs (see Section 3.3).

Sequential combination of fundamentals. The methodology used for forecasting in this paper consists of sequentially combining (aggregating) fundamentals. The distinguishing feature of our methodology is that it is sequential in nature: in contrast, all other approaches we know of in the forecasting of exchange rates use batch estimation or learning methods that need to be run in an incremental way (which usually leads to the loss of the theoretical guarantees associated with the batch case). These approaches can be divided into two groups: estimation methods of the statistics literature and learning methods of the machine-learning literature.

The second group contains, for instance, the studies of Li et al. [2015] and Wright [2008] for exchange rates or Bajari et al. [2015] for demand estimation. The first two references are perhaps the closest efforts to ours insofar as the idea of extracting information from fundamentals

³¹The problem is that factor analysis comes with no theoretical guarantees. Moreover, there is no theoretical reason (as opposed to regular economic models) why such relationships should hold.

to forecast exchange rates is concerned. In the introduction we already compared our approach and results to the ones of Li et al. [2015]: the end improvement in performance of the end-of-month forecasts with respect to no change are comparable though slightly better for us. But most importantly we obtain this improvement by considering only well-identified fundamentals (PPP or UIRP fundamentals; instead of following a "kitchen-sink" aggregation) and by implementing machine-learning methods that come with theoretical guarantees of performance (unlike the elastic net method used by Li et al. [2015]). Our approach thus conveys some economic meaning and is safe. The method at hand in Wright [2008] is called Bayesian model averaging. The building blocks for prediction therein are a great number of predictors, many of which do not come from the standard models considered in this paper (not only classic fundamentals are used). The approach of Wright [2008] does not improve the forecasts upon a no-change prediction in a statistically significant way in most cases he considers. Moreover, he does not explain how to properly choose the "shrinkage" parameter that guards the informativeness of priors not knowing the properties of the data ex ante. Our methods, in contrast, are entirely data driven³².

As for estimation methods, two approaches related to ours are estimations of nonlinear models and models with time-varying parameters. Rossi [2013] discusses that different nonlinear methods were not particularly successful in forecasting exchange rates while Rossi [2006] questions the robustness of the time-varying parameter models. Bacchetta et al. [2010] argue that the gain from using such an approach would be practically minimal. These authors find on simulated data that the benefits from using such models in terms of greater explanatory power are in practice outweighed by additional estimation errors of the time-varying parameters. Schinasi and Swamy [1989] reassess the study of Meese and Rogoff [1983] using various nonlinear methods, including an early version of a ridge regression. Engel [1994] documents a failure of a Markov-switching model to beat the no-change prediction in forecasting.

Other issues considered in exchange rate forecasting. Another way that researchers tried to improve the ability to forecast exchange rates from fundamentals is to consider different economic models with other fundamentals. It turned out in the most recent years that exchange rate models based on Taylor-rule fundamentals perform well in ascertaining the predictability³³ of the

³²Actually, some inspiration could be taken from them to perform a data-driven choice of the "shrinkage" parameter of Wright [2008].

³³ Predictability is a different concept than forecastability. In Molodtsova and Papell [2009] it is about testing whether the estimated coefficients of a considered model are jointly significantly different from zero when explaining changes

exchange rates at a short horizon (see Engel and West [2006], Molodtsova and Papell [2009], Molodtsova et al. [2008], Giacomini and Rossi [2010], Rossi and Inoue [2012]; though Rogoff and Stavrakeva [2008] disagree) but do not find that these perform much better in terms of forecasting and find sometimes that they give strictly worse results. For this reason we try out our methods on the same Taylor-rule fundamentals than the "classic" ones though we do not find them better (see Section 5.3).

Another successful fundamental was the behavior of net foreign assets as in Gourinchas and Rey [2007] or Corte et al. [2012]. The fundamentals to conduct these tests are available at 3–month frequencies resulting in fewer observations that can be used so we do not investigate them here. Other studies assessed the forecasting ability of exchange rate models of the 1990s such as Cheung et al. [2005], the differences in the term structures of forward premia such as Clarida et al. [2003], or the scapegoat model such as Fratzscher et al. [2015]. Given the scope of our exercise and the lacunae in some required data, we did not evaluate these models with our machine-learning methods, but it may be a useful research agenda for the future. Some other attempts involved work on what data was available at what instant ("real-time") while forming the forecasts. As including such data in our study is not feasible for the last 40 years, and these did not lead to a qualitative difference in the predictions obtained by Ince [2014], we leave the question whether such data leads to better forecasts using the methods we consider to future research.

7. Conclusions

In this paper we apply methods stemming from the field of machine learning – sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors – to the perennial problem of exchange rate forecasting. In doing so, we obtain gains in forecasting using the standardly applied RMSE criterion for PPP or UIRP fundamentals that were

in the exchange rate. It does not mean that a model that exhibits predictability necessarily provides better forecasts (in the literature typically it does not). The focus of these and many other attempts in general is rather to assess whether fundamentals play a role in exchange rate determination as it can be motivated theoretically why the forecasts they produce in terms of an evaluation criterion such as root mean square errors (RMSE) may fare worse than those of a forecast based on no change of the exchange rate. (See Rossi [2013] for a discussion. Special tests were designed by West [1996], Clark and West [2006, 2007] for this purpose. See also Section 3.1.3.) In this paper, however, we are not interested in such understood predictability – we actually are interested whether we can produce better forecasts of the exchange rates than the no-change prediction. We are also not interested in validating a particular model – that is trying to fit the coefficients and check whether the signs and magnitudes are those posited by theories; we simply try to extract from the fundamentals in question the information useful for the behavior of exchange rates.

not found using traditionally applied estimation methods based on OLS. The key is to use these machine-learning forecasting schemes to do what they are good for: produce forecasts – and not try to estimate some underlying model (if any such model exists) as was traditionally the case with more statistical methods. We conclude thus that a major problem of international economics – whether there is a short-term relationship between "classic" fundamentals and exchange rates – is answered in the affirmative under the condition that proper machine-learning techniques, e.g., the sequential ridge regression with discount factors or the exponentially weighted average strategy with discount factors, are applied. Our success points to a potential of such techniques for improving the evaluation of economic problems.

Machine learning techniques serve also to effectively aggregate information from many sources. A tempting exercise, beyond the scope of this paper, is to evaluate the forecasting performance including many more fundamentals than the "classic" ones considered here that were suggested by the literature – for example those based on productivity, interest rate yield curves, net foreign assets, etc. Venturing further one could consider many more series that are not typically associated with exchange rate forecasting in the true spirit of machine learning.

As with any new method applied to exchange rate forecasting, it remains to be seen whether our results could be replicated for different currencies, samples, forecasting periods and fundamentals. Given the robustness of the results shown in this paper, however, we hope that the application of these methods to exchange rate forecasting will stand the test of time and will allow for better predictions and decision making in the future.

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Table 1: RMSE and Theil ratios of forecasts for the average exchange rates sample.

	No change		Rolling regression		Re	Recursive regression	uc		SRidge			FWA	
Currency pair	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
							PPP fundamental						
USD/GBP	2.4410	1.0005	0.116	0.518	1.0040	0.936	0.874	0.9735	0.001 ***	0.039 **	0.9887	0.004 ***	0.025 **
JPY/USD	2.7039	1.0173	0.863	0.947	1.0007	0.673	0.688	0.9816	0.000	0.013 **	0.9862	0.000	0.002 ***
CAPALISD	7.7960	0.9867	0.761 0.002 ***	0.891	0.0010	0.531	0.520	0.9848		* 1200	0.9882	0.000	0.003
SEK/USD	2.5633	1.0108	0.274	0.761	1.0020	0.275	0.622	0.9585	0.000	0.014 **	0.9865	0.000	0.003 ***
DNK/USD	2.5355	1.0138	0.852	0.959	1.0011	0.354	0.580	0.9817	*** 000.0	0.002 ***	0.9835	0.000	*** 000.0
USD/AUD	2.7434	6966.0	0.032 **	0.420	1.0015	0.453	0.625	0.9710	0.005 ***	0.048 **	0.9900	*** 600.0	0.056 *
FRF/USD	2.6411	1.0118	0.452	0.758	1.0018	0.283	0.556	0.9809	0.000 ***	0.002 ***	0.9829	0.001 ***	0.013 **
DEM/USD	2.7545	1.0148	0.895	0.837	1.0028	0.915	0.885	0.9780	*** 000.0	0.003 ***	0.9847	0.002 ***	0.010 ***
ITL/USD	2.6624	1.0087	0.509	0.702	1.0019	0.290	0.552	0.9627	0.000 ***	0.008 ***	0.9726	*** 000.0	0.050 **
NLG/USD	2.7517	1.0109	0.728	0.856	1.0022	0.961	0.918	0.9811	*** 000.0	0.001 ***	0.9839	0.001 ***	0.005 ***
PTE/USD	2.6956	0.9964	0.018 **	0.469	0.9859	0.011 **	0.375	0.9713	0.003 ***	0.235	0.9465	0.000	0.151
,				-			OINP IUIIUAIIIEIIIUAI		1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	***		1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	1
USD/GBP	2.4410	1.0241	0.502	0.852	1.0037	0.454	0.655	0.9688	* * * 000.0	0.034 **	0.9809	0.001 ***	0.037 **
CHE/LISD	2.7039	1.0034	* 1200	0.022	7660	0.062	0.332	0.979.0	*********	0.011	7979 0	** 000 0	0.032
CAD/USD	1.5091	1.0086	0.342	0.763	1.0013	0.473	0.643	0.9965	0.032 **	0.154	0.9986	0.134	0.306
SEK/USD	2.5626	1.0600	0.916	0.964	1.0136	0.727	0.911	0.9601	*** 0000	0.024 **	0.9733	*** 000.0	** 650:0
DNK/USD	2.5355	1.0220	0.759	0.942	1.0071	0.456	0.774	0.9776	*** 0000	0.001 ***	0.9784	*** 0000	0.002 ***
USD/AUD	2.7414	1.0142	0.197	0.762	1.0060	0.417	0.720	0.9746	0.002 ***	0.043 **	0.9861	*** 600.0	0.103
FRF/USD	2.6411	1.0212	0.818	0.895	1.0141	0.633	0.829	0.9687	0.000 ***	0.003 ***	0.9765	0.001 ***	0.024 **
DEM/USD	2.7545	1.0216	0.599	0.806	1.0049	0.551	0.676	0.9683	0.000 ***	0.001 ***	0.9716	0.000	0.004 ***
ITL/USD	2.6624	1.0143	0.303	0.707	1.0166	0.569	0.797	0.9419	*** 000.0	0.036 **	0.9578	0.000 ***	0.056 *
NLG/USD	2.7517	1.0424	0.283	0.760	6666'0	0.159	0.499	0.9605	0.000 ***	0.001 ***	0.9737	0.001 ***	0.019 **
PTE/USD	2.2481	0.9963	0.252	0.420	0.9991	0.301	0.484	0.9695	0.044 **	0.146	0.9617	0.022 **	0.157
				ŀ		Monetary	Monetary model fundamentals	ntals		}			
USD/GBP	2.4410	1.0349	0.527	0.972	1.0173	0.509	0.889	0.9647	*** 000.0	* 620.0	0.9757	*** 000.0	0.034 **
JPY/USD	2.7042	0.9999	0.000 ***	0.497	1.0016	0.005 ***	0.536	0.9683	*** 000.0	0.022 **	0.9692	*** 000.0	0.013 **
CHF/USD	2.8377	1.0219	0.306	0.885	1.0059	0.109	0.645	0.9815	*** 000.0	0.110	0.9787	*** 000.0	0.052 *
CAD/USD	1.5083	1.0366	0.093 *	0.887	1.0075	0.368	0.779	0.9746	0.001 ***	0.026 **	0.9872	0.025 **	0.154
SEK/USD	2.5626	1.0665	0.928	0.999	1.0289	0.838	0.948	0.9487	0.000	0.045 **	0.9714	*** 000.0	0.003 ***
DNK/USD	2.5553	1.0517	0.999	0.997	1.0252	1.000	0.949	0.9749	0.000	0.012 **	0.9689	0.000	0.013 **
USD/AUD FBE/HSD	7.7.547	1.0635	0.238	0.900	1.0098	0.692	0.890	0.9561	*** 900 0	0.035	0.9807	0.000	0.089
PRF/USD	2.3223	1.0013	0.822	0.361	1.0738	0.049	0.520	0.9787	***	0.126	0.9934	0.094	***
TI / USD	2 6458	1.0411	0.731	0 937	1.0165	0.572	0.798	0.9300	*** 000 0	** 920.0	0.9661	*** 000 0	0.000
NLG/USD	2.7517	1.0342	0.883	996:0	1.0076	0.430	0.699	0.9598	*** 0000	0.001 ***	0.9709	*** 0000	0.029 **
PTE/USD	2.2342	1.0103	0.151	0.574	1.0485	0.653	0.865	0.9590	0.041 **	0.205	0.9778	0.049 **	0.293
						IIA	l fundamentals						
USD/GBP	2.4410	1.0635	0.061 *	0.983	1.0330	0.493	0.927	0.9632	0.000 ***	0.027 **	0.9741	*** 0000	0.033 **
JPY/USD	2.7042	1.0424	0.003 ***	0.945	1.0129	0.000 ***	0.700	0.9626	*** 000.0	0.002 ***	0.9661	0.000 ***	0.005 ***
CHF/USD	2.8377	1.0650	0.038 **	0.975	1.0221	0.051 *	608.0	0.9797	0.000 ***	0.063 *	0.9751	0.000 ***	0.032 **
CAD/USD	1.5083	1.0052	0.001 ***	0.568	1.0068	0.121	0.765	0.9738	0.001 ***	0.032 **	0.9885	0.013 **	0.134
SEK/USD	2.5626	1.1506	0.682	0.950	1.0656	0.678	0.900	0.9557	*** 000.0	0.004 ***	0.9882	0.025 **	0.302
DNK/USD	2.5553	1.0859	0.562	966.0	1.0248	0.497	0.917	0.9691	0.000 ***	0.001 ***	0.9641	0.000 ***	0.003 ***
USD/AUD	2.7547	1.0209	0.008 ***	0.743	1.0224	0.621	0.900	0.9631	*** 000.0	0.027 **	0.9749	*** 0000	0.030 **
FRF/USD	2.5229	1.1137	0.480	0.980	1.0599	0.618	0.942	0.9736	0.005 ***	0.065 *	0.9786	0.011 **	0.186
DEM/USD	2.7624	1.0367	0.169	0.777	1.0157	0.058 *	0.663	0.9542	0.000	0.004 ***	0.9485	*** 000.0	0.002 ***
IIL/USD	2.6458	1.0054	0.016 **	0.556	1.0438	0.1/8	0.918	0.9351	0.000	0.040 **	0.9518	0.001	0.02/ **
NLG/USD PTE/LISD	2./51/	1.0251	0.006 ***	0.735	1.0182	0.049 ***	0.715	0.9614	0.000	0.001 ***	0.9654	0.000 ***	0.013 **
PIE/USD	7.5347	0.9571	0.000	0.197	1.0009	0.491	0.040	0.9649	0.049	0.130	0.9769	0.045	0.203

Table 2: RMSE and Theil ratios of forecasts for the end-of-month exchange rates sample.

	or section of A		Rolling regression		Re	Recursive regression			SRidge			ΕWA	
Currency pair	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
						рр f	PPP fundamental						
USD/GBP	2.9051	1.0025	0.212	0.615	1.0032		0.887	8666.0	0.365	0.452	0.9975	0.135	0.335
JPY/USD	3.1023	1.0144	0.890	0.956	1.0010		998.0	0.9999	0.169	0.183	0.9953	0.021 **	0.107
CHF/USD	3.3428	1.0117	0.919	696.0	1.0009		0.631	0.9994	0.167	0.205	0.9947	0.020 **	0.063 *
CAD/USD	1.9982	0.9988	0.133	0.438	0.9991		0.351	1.0003	0.750	0.696	1.0011	0.997	0.987
SEK/USD	3.2290	1.0089	0.445	0.835	1.0009	0.299	0.590	1.0001	0.426	0.527	0.9969	0.050 **	0.204
OSD/AND	3.1200 2.2796	1.0112	0.090	0.979	1,0008		0.500	1 0007	0.232	0.250	1 0001	0.012	0.031
USD/AUD	3.3700	10120	0.073	0.442	1.0009		0.600	1,0007	0.722	0.731	T.0001	0.045	0.312
DFM/USD	3.136	1.0120	0.031	0.854	1.0035		0.008	0.9994	0.383	0.002	0.9915	0.033	0.137
ITL/USD	3.1907	1.0070	0.533	0.727	1.0019		0.574	1.0002	0.303	0.510	0.9891	0.015 **	* 660.0
NLG/USD	3.3319	1.0096	0.807	0.910	1.0023		0.980	0.9995	0.161	0.172	0.9901	*** 900.0	0.023 **
PTE/USD	3.2207	1.0018	0.036 **	0.525	0.9936	*	0.408	0.9916	0.031 **	0.368	9886.0	0.021 **	0.348
							UIRP fundamental						
USD/GBP	2.9051	1.0222	0.594	0.887	1.0044	0.589	0.739	1.0028	0.668	0.742	1.0024	0.195	0.589
JPY/USD	3.1023	1.0060	0.132	0.675	0.9983		0.359	0.9993	0.207	0.266	0.9918	*** 600.0	0.153
CHF/USD	3.3747	1.0106	0.155	0.789	0.9985	0.086 *	0.442	0.9994	0.260	0.370	0.9925	0.013 **	0.075 *
CAD/USD	1.9982	1.0055	0.371	0.786	1.0009	0.533	0.678	1.0003	0.705	0.708	1.0010	0.979	0.941
SEK/USD	3.2263	1.0514	0.908	0.964	1.0110	0.709	0.912	0.9948	* 260.0	0.280	0.9890	0.051 *	0.255
DNK/USD	3.1200	1.0186	0.900	0.958	1.0042	0.649	0.799	1.0005	0.570	0.609	0.9954	0.029 **	0.255
USD/AUD	3.3793	1.0131	0.348	0.828	1.0058	0.556	0.757	1.0048	669.0	0.736	0.9980	0.196	0.400
FRF/USD	3.1978	1.0206	0.915	0.955	1.0112	0.734	0.875	1.0014	0.473	0.586	0.9951	0.083 *	0.328
DEM/USD	3.3136	1.0209	0.805	0.907	1.0052		0.723	9866.0	0.191	0.256	0.9888	0.023 **	0.141
ITL/USD	3.1907	1.0117	0.340	0.727	1.0116		0.790	1.0056	0.397	0.642	0.9867	0.011 **	0.222
NLG/USD	3.3319	1.0388	0.583	0.849	1.0036		0.697	0.9997	0.356	0.403	0.9903	0.035 **	0.185
PTE/USD	2.8024	0.9970	0.278	0.438	0.9968		0.434	0.9965	0.342	0.403	0.9932	0.290	0.353
						Jonetary	model fundamentals	ıtals					
USD/GBP	2.9051	1.0350	0.810	0.997	1.0165		0.932	1.0004	0.136	0.517	1.0008	0.091 *	0.531
JPY/USD	3.1046	1.0106	0.042 **	0.767	1.0046	* *	0.631	1.0025	0.127	0.613	0.9955	0.007 ***	0.330
CHF/USD	3.3887	1.0249	0.580	0.954	1.0103		0.785	1.0023	0.461	0.690	1.0102	0.785	0.932
CAD/USD	1.9994	1.0396	0.456	0.935	1.0104		0.893	1.0124	0.828	0.838	1.0070	0.895	0.899
SEK/USD	3.2263	1.0556	0.970	0.999	1.0222		0.970	1.0047	0.495	0.656	1.0076	0.401	0.801
UNK/USD	3.16/5	1.0435	0.999	766.0	1.0186	0.999	0.965	1.0041	0.957	0.844	1.0062	0.974	0.964
USD/AUD FRE/LISD	3.3920	1.0284	0.370	0.987	1.0108		0.948	1.001/	0.614	0.629	1,0000	0.471	0.755
DEM/USD	3 3286	1 0192	0.476	0.873	1.0121		0.50%	1 0026	0.542	0.665	1 0009	0.117	0.526
ITL/USD	3.2797	1.0233	0.538	0.876	1.0150		0.820	1.0003	0.388	0.515	1.0035	0.265	0.592
NLG/USD	3.3319	1.0305	0.698	0.945	1.0085		0.754	0.9978	0.107	0.234	0.9988	0.070 *	0.462
PTE/USD	2.7703	1.0053	0.200	0.548	1.0504	0.937	0.988	1.0071	0.504	0.606	1.0152	0.581	0.759
						All fu	fundamentals						
USD/GBP	2.9051	1.0715	0.303	0.999	1.0281		0.952	0.9984	0.213	0.386	1.0033	0.130	0.610
JPY/USD	3.1046	1.0478	0.055 *	0.964	1.0101	0.006 ***	0.724	1.0010	0.140	0.557	0.9971	0.011 **	0.380
CAD/LISD	1 9994	1.07.04	0.348	0.958	1.0244		0.505	1.0035	0.854	0.859	1.0018	0.997	0.991
SEK/USD	3,2263	1.1642	968.0	696:0	1.0691		0,883	1.0083	0.928	0.955	0.9963	0.052 *	0.406
DNK/USD	3.1675	1.0643	0.628	0.999	1.0195		0.951	1.0036	0.963	0.854	1.0053	0.992	0.949
USD/AUD	3.3926	1.0420	0.112	0.949	1.0197		0.913	1.0021	0.781	0.815	1.0029	0.211	609.0
FRF/USD	3.1041	1.0934	0.646	6.979	1.0407	0.657	0.908	1.0073	0.796	0.856	1.0065	0.282	0.635
DEM/USD	3.3286	1.0411	0.424	0.886	1.0281		0.846	1.0037	0.573	0.725	0.9952	* 650.0	0.375
ITL/USD	3.2797	1.0025	0.032 **	0.530	1.0392		0.883	1.0024	0.519	0.612	1.0043	0.181	0.588
NLG/USD	3.3319	1.0449	0.042 **	0.923	1.0292		0.901	0.9982	0.134	0.246	0.9920	0.023 **	0.276
PLE/USD	2.7703	T.U138	0.139	0.631	1.0580	U.863	0.968	1.0047	0.534	0.603	0.9955	0.319	0.397

Table 3: RMSE and Theil ratios of the machine learning methods vs. rolling and recursive regressions for the average exchange rates sample.

		SRidg	e vs.			EWA	vs.	
C	Rolling	regression		regression	Rolling i	regression		eregression
Currency pair	Ratio of		Ratio of	DM n value	Ratio of	DM n value	Ratio of	DM n value
	RMSEs	DM p-value	RMSEs	DM p-value	RMSEs	DM p-value	RMSEs	DM p-value
			PF	PP fundamenta				
USD/GBP	0.9730	0.056 *	0.9696	0.023 **	0.9883	0.148	0.9848	0.009 ***
JPY/USD	0.9648	0.002 ***	0.9808	0.009 ***	0.9693	0.003 ***	0.9854	0.001 ***
CHF/USD	0.9735	0.006 ***	0.9838	0.011 **	0.9769	0.004 ***	0.9872	0.002 ***
CAD/USD	1.0016	0.536	0.9917	0.116	1.0098	0.742	0.9999	0.488
SEK/USD	0.9482	0.049 **	0.9565	0.007 ***	0.9760	0.094 *	0.9845	0.013 **
DNK/USD	0.9683	0.001 ***	0.9806	0.003 ***	0.9701	0.000 ***	0.9824	0.008 ***
USD/AUD	0.9740	0.105	0.9695	0.029 **	0.9931	0.329	0.9885	0.017 **
FRF/USD	0.9695	0.039 **	0.9792	0.047 **	0.9715	0.036 **	0.9812	0.047 **
DEM/USD	0.9637	0.048 **	0.9753	0.005 ***	0.9703	0.033 **	0.9820	0.002 ***
ITL/USD	0.9545	0.007 ***	0.9609	0.002 ***	0.9642	0.002 ***	0.9708	0.000 ***
NLG/USD	0.9706	0.008 ***	0.9790	0.000 ***	0.9733	0.006 ***	0.9818	0.001 ***
PTE/USD	0.9748	0.035 **	0.9852	0.058 *	0.9498	0.002 ***	0.9600	0.001 ***
-				RP fundamenta		-		
USD/GBP	0.9460	0.045 **	0.9653	0.040 **	0.9578	0.068 *	0.9773	0.066 *
JPY/USD	0.9664	0.019 **	0.9742	0.025 **	0.9750	0.068 *	0.9828	0.087 *
CHF/USD	0.9659	0.047 **	0.9815	0.107	0.9666	0.047 **	0.9822	0.122
CAD/USD	0.9880	0.198	0.9952	0.162	0.9902	0.238	0.9974	0.293
SEK/USD	0.9057	0.017 **	0.9472	0.011 **	0.9182	0.019 **	0.9603	0.019 **
DNK/USD	0.9565	0.004 ***	0.9707	0.008 ***	0.9573	0.004 ***	0.9715	0.010 ***
USD/AUD	0.9610	0.060 *	0.9688	0.013 **	0.9723	0.101	0.9802	0.032 **
FRF/USD	0.9486	0.004 ***	0.9553	0.004 ***	0.9563	0.009 ***	0.9630	0.018 **
DEM/USD	0.9479	0.020 **	0.9636	0.007 ***	0.9511	0.030 **	0.9669	0.021 **
ITL/USD	0.9286	0.003 ***	0.9265	0.016 **	0.9443	0.003 ***	0.9421	0.011 **
NLG/USD	0.9214	0.096 *	0.9606	0.038 **	0.9341	0.132	0.9738	0.135
PTE/USD	0.9731	0.140	0.9703	0.113	0.9652	0.136	0.9625	0.125
				y model fundai				
USD/GBP	0.9322	0.014 **	0.9484	0.018 **	0.9428	0.005 ***	0.9592	0.012 **
JPY/USD	0.9685	0.027 **	0.9668	0.025 **	0.9693	0.020 **	0.9676	0.019 **
CHF/USD	0.9605	0.001 ***	0.9757	0.024 **	0.9577	0.002 ***	0.9730	0.021 **
CAD/USD	0.9401	0.026 **	0.9673	0.031 **	0.9523	0.078 *	0.9798	0.140
SEK/USD	0.8896	0.001 ***	0.9220	0.006 ***	0.9109	0.000 ***	0.9441	0.005 ***
DNK/USD	0.9270	0.000 ***	0.9510	0.004 ***	0.9213	0.003 ***	0.9451	0.017 **
USD/AUD	0.9420	0.003 ***	0.9567	0.013 **	0.9562	0.015 **	0.9712	0.029 **
FRF/USD	0.9220	0.007 ***	0.9114	0.004 ***	0.9377	0.007 ***	0.9270	0.004 ***
DEM/USD	0.9399	0.003 ***	0.9526	0.013 **	0.9351	0.007 ***	0.9478	0.020 **
ITL/USD	0.8933	0.001 ***	0.9149	0.041 **	0.9280	0.005 ***	0.9504	0.010 ***
NLG/USD	0.9280	0.000 ***	0.9525	0.004 ***	0.9388	0.000 ***	0.9635	0.010 ***
PTE/USD	0.9492	0.112	0.9146	0.013 **	0.9679	0.259	0.9326	0.090 *
				ll fundamentals				
USD/GBP	0.9057	0.001 ***	0.9324	0.023 **	0.9160	0.001 ***	0.9430	0.017 **
JPY/USD	0.9234	0.000 ***	0.9504	0.011 **	0.9268	0.002 ***	0.9539	0.010 ***
CHF/USD	0.9198	0.004 ***	0.9585	0.039 **	0.9156	0.002 ***	0.9540	0.025 **
CAD/USD	0.9688	0.113	0.9672	0.032 **	0.9834	0.279	0.9818	0.105
SEK/USD	0.8306	0.026 **	0.8969	0.031 **	0.8588	0.023 **	0.9274	0.026 **
DNK/USD	0.8924	0.000 ***	0.9456	0.003 ***	0.8878	0.000 ***	0.9408	0.002 ***
USD/AUD	0.9433	0.017 **	0.9420	0.007 ***	0.9549	0.050 *	0.9536	0.004 ***
FRF/USD	0.8742	0.012 **	0.9186	0.015 **	0.8787	0.021 **	0.9233	0.027 **
DEM/USD	0.9204	0.049 **	0.9394	0.068 *	0.9149	0.046 **	0.9338	0.063 *
ITL/USD	0.9301	0.010 **	0.8959	0.010 **	0.9467	0.039 **	0.9118	0.008 ***
NLG/USD	0.9378	0.042 **	0.9442	0.051 *	0.9417	0.062 *	0.9482	0.072 *
PTE/USD	1.0081	0.580	0.9095	0.029 **	1.0206	0.687	0.9208	0.051 *

Table 4: RMSE and Theil ratios of the machine learning methods vs. rolling and recursive regressions for the end-of-month exchange rates sample.

		SRidg			- III	EWA		
Currency pair	Rolling r Ratio of	regression	Recursive Ratio of	eregression	Rolling i Ratio of	regression	Recursive Ratio of	regression
	RMSEs	DM p-value	RMSEs	DM p-value	RMSEs	DM p-value	RMSEs	DM p-value
				PP fundamenta				
USD/GBP	0.9973	0.369	0.9966	0.091 *	0.9950	0.305	0.9943	0.178
JPY/USD	0.9857	0.043 **	0.9989	0.111	0.9812	0.014 **	0.9943	0.060 *
CHF/USD	0.9879	0.025 **	0.9986	0.292	0.9832	0.007 ***	0.9938	0.055 *
CAD/USD	1.0016	0.578	1.0012	0.720	1.0023	0.613	1.0019	0.808
SEK/USD	0.9913	0.183	0.9992	0.379	0.9881	0.150	0.9960	0.207
DNK/USD	0.9887	0.020 **	0.9989	0.388	0.9834	0.009 ***	0.9936	0.121
USD/AUD	1.0023	0.583	0.9998	0.470	1.0017	0.565	0.9993	0.427
FRF/USD	0.9888	0.108	0.9971	0.313	0.9839	0.079 *	0.9921	0.176
DEM/USD	0.9858	0.140	0.9960	0.037 **	0.9779	0.064 *	0.9880	0.006 ***
ITL/USD	0.9933	0.123	0.9983	0.197	0.9822	0.064 *	0.9872	0.066 *
NLG/USD	0.9900	0.076 *	0.9972	0.010 ***	0.9806	0.006 ***	0.9878	0.007 ***
PTE/USD	0.9898	0.081 *	0.9980	0.284	0.9867	0.143	0.9950	0.322
			UI	RP fundamenta	al			
USD/GBP	0.9810	0.155	0.9984	0.401	0.9806	0.203	0.9980	0.439
JPY/USD	0.9934	0.299	1.0010	0.588	0.9859	0.132	0.9935	0.236
CHF/USD	0.9889	0.188	1.0008	0.535	0.9821	0.104	0.9940	0.295
CAD/USD	0.9948	0.227	0.9993	0.359	0.9955	0.262	1.0001	0.516
SEK/USD	0.9462	0.032 **	0.9840	0.032 **	0.9407	0.053 *	0.9782	0.095 *
DNK/USD	0.9822	0.036 **	0.9963	0.156	0.9772	0.028 **	0.9912	0.140
USD/AUD	0.9917	0.244	0.9990	0.390	0.9850	0.155	0.9922	0.212
FRF/USD	0.9812	0.022 **	0.9903	0.135	0.9750	0.055 *	0.9841	0.131
DEM/USD	0.9782	0.079 *	0.9935	0.227	0.9685	0.061 *	0.9837	0.126
ITL/USD	0.9941	0.287	0.9941	0.155	0.9753	0.079 *	0.9754	0.077 *
NLG/USD	0.9623	0.150	0.9961	0.290	0.9533	0.121	0.9868	0.165
PTE/USD	0.9995	0.488	0.9996	0.487	0.9962	0.397	0.9964	0.367
				y model fundar				
USD/GBP	0.9665	0.017 **	0.9841	0.087 *	0.9669	0.036 **	0.9846	0.143
JPY/USD	0.9920	0.287	0.9979	0.422	0.9850	0.103	0.9909	0.205
CHF/USD	0.9780	0.033 **	0.9921	0.229	0.9857	0.142	0.9999	0.496
CAD/USD	0.9738	0.116	1.0020	0.555	0.9686	0.072 *	0.9966	0.347
SEK/USD	0.9517	0.000 ***	0.9828	0.058 *	0.9545	0.002 ***	0.9857	0.136
DNK/USD	0.9622	0.003 ***	0.9858	0.014 **	0.9643	0.007 ***	0.9879	0.067 *
USD/AUD	0.9740	0.044 **	0.9910	0.050 **	0.9782	0.080 *	0.9953	0.292
FRF/USD	0.9466	0.009 ***	0.9532	0.011 **	0.9551	0.025 **	0.9617	0.019 **
DEM/USD	0.9837	0.109 0.088 *	0.9906	0.162	0.9821	0.133	0.9890	0.234
ITL/USD NLG/USD	0.9775 0.9683	0.088 *	0.9855 0.9894	0.226 0.191	0.9806 0.9693	0.150 0.053 *	0.9887 0.9904	0.225 0.244
PTE/USD	1.0018	0.518	0.9588	0.191	1.0098	0.598	0.9665	0.244
PTE/USD	1.0016	0.516		II fundamentals		0.396	0.9003	0.133
USD/GBP	0.9318	0.001 ***	0.9711	0.039 **	0.9364	0.001 ***	0.9759	0.100 *
JPY/USD	0.9518	0.001	0.9909	0.039	0.9504	0.001	0.9871	0.100
CHF/USD	0.9353	0.031	0.9909	0.262	0.9310	0.016 ***	0.9871	0.138
CAD/USD	0.9505	0.008 ***	0.9930	0.103	0.9512	0.007 **	0.9803	0.105
SEK/USD	0.8660	0.047	0.9431	0.103	0.8557	0.030	0.9319	0.133
DNK/USD	0.8000	0.032	0.9844	0.131	0.8337	0.041	0.9861	0.123
USD/AUD	0.9617	0.054 *	0.9828	0.084	0.9625	0.002	0.9836	0.112
FRF/USD	0.9213	0.034	0.9679	0.083	0.9023	0.044	0.9671	0.113
DEM/USD	0.9213	0.020	0.9763	0.141	0.9559	0.017	0.9680	0.102
ITL/USD	0.9999	0.112	0.9646	0.130	1.0018	0.524	0.9665	0.127
NLG/USD	0.9552	0.455	0.9699	0.132	0.9493	0.045 **	0.9639	0.112
PTE/USD	0.9332	0.402	0.9496	0.080	0.9493	0.313	0.9409	0.045 **
1 11/030	0.3310	0.402	U.J430	0.047	0.3013	0.513	U.J 4 UJ	0.043

Table 5: Directional tests: Percentages of changes predicted for the average exchange rates sample.

	Rolling r	egression	Recursive	regression	SRi	dge	EV	NA
	Proportion of		Proportion of		Proportion of		Proportion of	
Currency pair	changes	DM p-value	changes	DM p-value	changes	DM p-value	changes	DM p-value
	predicted	·	predicted	·	predicted	·	predicted	·
				PPP fundamenta	ı			
USD/GBP	0.514	0.300	0.488	0.661	0.585	0.002 ***	0.562	0.009 ***
JPY/USD	0.496	0.552	0.475	0.761	0.593	0.001 ***	0.559	0.055 *
CHF/USD	0.551	0.071 *	0.491	0.586	0.598	0.002 ***	0.598	0.001 ***
CAD/USD	0.543	0.119	0.514	0.348	0.543	0.112	0.522	0.268
SEK/USD	0.499	0.515	0.496	0.540	0.633	0.000 ***	0.598	0.000 ***
DNK/USD	0.496	0.547	0.549	0.099 *	0.598	0.000 ***	0.577	0.002 ***
USD/AUD	0.520	0.337	0.463	0.825	0.587	0.002 ***	0.522	0.249
FRF/USD	0.455	0.826	0.503	0.480	0.577	0.017 **	0.593	0.005 ***
DEM/USD	0.466	0.758	0.418	0.949	0.608	0.001 ***	0.608	0.001 ***
ITL/USD	0.534	0.255	0.503	0.478	0.598	0.003 ***	0.566	0.090 *
NLG/USD	0.460	0.822	0.418	0.966	0.603	0.002 ***	0.598	0.003 ***
PTE/USD	0.481	0.637	0.513	0.397	0.524	0.312	0.577	0.031 **
				IIRP fundamenta	-			
USD/GBP	0.507	0.411	0.541	0.080 *	0.562	0.021 **	0.538	0.068 *
JPY/USD	0.570	0.021 **	0.530	0.155	0.614	0.000 ***	0.612	0.000 ***
CHF/USD	0.596	0.002 ***	0.604	0.001 ***	0.617	0.000 ***	0.582	0.003 ***
CAD/USD	0.551	0.072 *	0.509	0.396	0.522	0.241	0.472	0.765
SEK/USD	0.497	0.529	0.487	0.623	0.611	0.000 ***	0.616	0.000 ***
DNK/USD	0.483	0.706	0.517	0.317	0.609	0.000 ***	0.575	0.002 ***
USD/AUD	0.517	0.330	0.517	0.328	0.588	0.001 ***	0.522	0.240
FRF/USD	0.471	0.696	0.460	0.764	0.593	0.005 ***	0.598	0.003 ***
DEM/USD	0.513	0.408	0.508	0.439	0.603	0.002 ***	0.614	0.001 ***
ITL/USD	0.497	0.520	0.503	0.478	0.603	0.002 ***	0.603	0.002 ***
NLG/USD	0.524	0.303	0.561	0.079 *	0.608	0.001 ***	0.598	0.003 ***
PTE/USD	0.479	0.604	0.521	0.400	0.563	0.141	0.634	0.010 ***
				ry model funda				
USD/GBP	0.501	0.482	0.514	0.330	0.598	0.000 ***	0.575	0.004 ***
JPY/USD	0.587	0.001 ***	0.555	0.045 **	0.605	0.000 ***	0.611	0.000 ***
CHF/USD	0.522	0.289	0.536	0.178	0.599	0.001 ***	0.602	0.000 ***
CAD/USD	0.526	0.232	0.497	0.528	0.561	0.022 **	0.529	0.136
SEK/USD	0.505	0.439	0.471	0.783	0.632	0.000 ***	0.624	0.000 ***
DNK/USD	0.440	0.977	0.454	0.917	0.577	0.004 ***	0.577	0.007 ***
USD/AUD	0.509	0.409	0.494	0.574	0.585	0.002 ***	0.582	0.005 ***
FRF/USD	0.508	0.447	0.525	0.322	0.575	0.048 **	0.600	0.013 **
DEM/USD	0.571	0.038 **	0.593	0.009 ***	0.627	0.000 ***	0.605	0.003 ***
ITL/USD	0.468	0.684	0.506	0.445	0.596	0.007 ***	0.615	0.002 ***
NLG/USD	0.455	0.867	0.497	0.523	0.598	0.003 ***	0.587	0.007 ***
PTE/USD	0.471	0.648	0.457	0.736	0.586	0.073 *	0.600	0.060 *
USD/GBP	0.567	0.008 ***	0.533	All fundamentals 0.119	0.585	0.001 ***	0.577	0.004 ***
JPY/USD	0.582	0.008	0.589	0.119	0.613	0.001	0.603	0.004
CHF/USD	0.571	0.003	0.538	0.169	0.596	0.000	0.610	0.000
CAD/USD	0.582	0.010	0.505	0.437	0.571	0.002	0.516	0.275
SEK/USD	0.547	0.063 *	0.553	0.063 *	0.626	0.000 ***	0.629	0.000 ***
DNK/USD	0.506	0.425	0.551	0.051 *	0.583	0.001 ***	0.589	0.000 ***
USD/AUD	0.554	0.041 **	0.506	0.430	0.582	0.004 ***	0.565	0.018 **
FRF/USD	0.500	0.500	0.550	0.143	0.583	0.032 **	0.583	0.032 **
DEM/USD	0.554	0.107	0.599	0.005 ***	0.616	0.001 ***	0.610	0.001 ***
							0.590	0.011 **
ITL/USD	0.551	0.107	0.526	0.282	0.596	0.007 ***	0.550	0.011
ITL/USD NLG/USD	0.551 0.571	0.107 0.068 *	0.526 0.571	0.282 0.046 **	0.596 0.603	0.007 *** 0.002 ***	0.603	0.002 ***

Table 6: Directional tests: Percentages of changes predicted for the end-of-month exchange rates sample.

	Rolling r	egression	Recursive	regression	SRi	dge	EV	WA
	Proportion of		Proportion of		Proportion of		Proportion of	
Currency pair	changes	DM p-value	changes	DM p-value	changes	DM p-value	changes	DM p-value
	predicted	•	predicted	•	predicted	·	predicted	•
				PPP fundamenta				
USD/GBP	0.512	0.334	0.493	0.599	0.470	0.835	0.522	0.273
JPY/USD	0.470	0.874	0.462	0.906	0.546	0.043 **	0.543	0.067 *
CHF/USD	0.525	0.208	0.488	0.635	0.556	0.024 **	0.541	0.085 *
CAD/USD	0.528	0.174	0.496	0.548	0.496	0.546	0.441	0.987
SEK/USD	0.501	0.484	0.480	0.719	0.493	0.573	0.512	0.343
DNK/USD	0.499	0.517	0.525	0.257	0.564	0.015 **	0.559	0.011 **
USD/AUD	0.494	0.562	0.497	0.530	0.483	0.696	0.480	0.747
FRF/USD	0.481	0.650	0.540	0.268	0.481	0.659	0.571	0.029 **
DEM/USD	0.455	0.847	0.481	0.672	0.603	0.010 ***	0.608	0.005 ***
ITL/USD	0.487	0.591	0.466	0.737	0.466	0.737	0.577	0.094 *
NLG/USD	0.519	0.357	0.455	0.818	0.587	0.025 **	0.593	0.025 **
PTE/USD	0.481	0.613	0.497	0.517	0.497	0.517	0.550	0.171
			·	JIRP fundament	al		•	
USD/GBP	0.496	0.546	0.522	0.216	0.480	0.727	0.517	0.298
JPY/USD	0.546	0.072 *	0.522	0.214	0.562	0.008 ***	0.570	0.003 ***
CHF/USD	0.566	0.008 ***	0.547	0.044 **	0.550	0.040 **	0.563	0.023 **
CAD/USD	0.551	0.024 **	0.522	0.220	0.493	0.575	0.483	0.697
SEK/USD	0.497	0.534	0.471	0.821	0.508	0.388	0.539	0.075 *
DNK/USD	0.501	0.483	0.483	0.689	0.512	0.345	0.564	0.006 ***
USD/AUD	0.500	0.500	0.514	0.332	0.470	0.837	0.492	0.614
FRF/USD	0.497	0.516	0.460	0.760	0.561	0.046 **	0.593	0.005 ***
DEM/USD	0.508	0.447	0.540	0.209	0.593	0.028 **	0.608	0.005 ***
ITL/USD	0.497	0.519	0.466	0.742	0.450	0.827	0.603	0.016 **
, NLG/USD	0.556	0.100 *	0.561	0.068 *	0.593	0.017 **	0.593	0.014 **
PTE/USD	0.479	0.639	0.437	0.797	0.437	0.859	0.507	0.453
				ry model funda				
USD/GBP	0.501	0.480	0.535	0.138	0.509	0.383	0.517	0.333
JPY/USD	0.532	0.125	0.524	0.227	0.542	0.082 *	0.558	0.017 **
CHF/USD	0.497	0.536	0.541	0.079 *	0.541	0.092 *	0.533	0.115
CAD/USD	0.513	0.304	0.489	0.646	0.487	0.659	0.487	0.688
SEK/USD	0.432	0.993	0.474	0.759	0.537	0.076 *	0.524	0.228
DNK/USD	0.437	0.977	0.480	0.701	0.480	0.710	0.489	0.638
USD/AUD	0.526	0.212	0.489	0.665	0.477	0.754	0.497	0.537
FRF/USD	0.525	0.334	0.467	0.721	0.558	0.099 *	0.550	0.135
DEM/USD	0.559	0.056 *	0.576	0.035 **	0.559	0.142	0.548	0.114
ITL/USD	0.449	0.813	0.442	0.877	0.513	0.405	0.538	0.197
NLG/USD	0.434	0.892	0.513	0.394	0.582	0.043 **	0.534	0.180
PTE/USD	0.457	0.717	0.386	0.972	0.371	0.986	0.471	0.669
			,	All fundamental	S			
USD/GBP	0.556	0.019 **	0.491	0.627	0.496	0.550	0.488	0.623
JPY/USD	0.537	0.078 *	0.537	0.137	0.566	0.014 **	0.558	0.012 **
CHF/USD	0.525	0.207	0.514	0.328	0.541	0.084 *	0.547	0.055 *
CAD/USD	0.558	0.032 **	0.503	0.463	0.495	0.561	0.455	0.926
SEK/USD	0.505	0.428	0.495	0.564	0.487	0.650	0.542	0.050 **
DNK/USD	0.529	0.191	0.511	0.358	0.443	0.948	0.483	0.700
USD/AUD	0.517	0.285	0.520	0.264	0.472	0.800	0.500	0.500
FRF/USD	0.567	0.078 *	0.550	0.172	0.500	0.500	0.533	0.239
DEM/USD	0.537	0.231	0.508	0.424	0.542	0.142	0.576	0.064 *
ITL/USD	0.513	0.398	0.513	0.394	0.429	0.901	0.558	0.073 *
NLG/USD	0.513	0.385	0.540	0.174	0.582	0.044 **	0.561	0.106
PTE/USD	0.486	0.578	0.414	0.908	0.471	0.684	0.471	0.682

Table 7: Economic criterion: Basic results for the EWA algorithm for the end-of-month exchange rates sample. Note: the results for the portfolios based on the no-change exchange rate predictions vary because of the constraints on weights and because for different sets of fundamentals different time periods are considered depending on the availability of data.

				Annualized	alized			
Portfolio /	Returns (in %)	(in %)	Performance fee	Premium return	Sharpe ratios	atios	Sortino ratios	ratios
r constraint		(2)	of portfolio w.r.t.	of portfolio w.r.t.)))))			
	No change (with c constraint)	Portfolio	no change (in bps)	no change (in bps)	No change (with c constraint)	Portfolio	No change (with c constraint)	Portfolio
				PPP fundamental				
EWA / c = 0.5	11.07	12.03	153	190	0.65	08.0	29'0	1.13
EWA/c=1	12.06	13.24	213	269	0.65	0.83	0.67	1.09
EWA/c = 2	11.76	13.67	270	319	0.63	0.85	0.64	1.05
EWA/c = 5	11.86	13.44	224	271	0.64	0.83	0.68	1.06
EWA / c = 10	11.83	13.48	229	277	0.64	0.84	0.68	1.08
EWA / no c	11.83	13.48	229	277	0.64	0.84	0.68	1.08
011110111111 / 110 C	11.03	1.04	-302	-230	40.0	0.24	0.00	0.33
				UIRP fundamental				
EWA / c = 0.5	10.62	12.12	202	251	0.61	0.81	0.64	1.31
EWA/c = 1	11.67	13.18	249	332	0.63	0.84	0.65	1.36
EWA/c = 2	11.37	13.04	270	362	09.0	0.83	0.62	1.44
EWA/c = 5	11.47	12.16	184	263	0.62	0.77	0.65	1.35
EWA / c = 10	11.44	12.13	192	268	0.62	0.78	0.65	1.38
EWA / no c	11.44	12.13	192	268	0.62	0.78	0.65	1.38
Uniform / no c	11.44	5.29	-301	-238	0.62	0.19	0.65	0.28
			Monet	Monetary models fundamentals	entals			
EWA / c = 0.5	11.02	8.89	-124	68-	0.64	0.51	29'0	0.67
EWA/c = 1	11.95	9.59	-82	-23	0.64	0.54	99.0	0.70
EWA/c=2	11.68	8.95	-131	99-	0.62	0.47	0.64	0.63
EWA/c = 5	11.73	8.54	-182	-114	0.63	0.44	0.67	0.62
EWA / c = 10	11.70	8.49	-189	-120	0.63	0.43	0.67	0.62
EWA / no c	11.70	8.49	-189	-120	0.63	0.43	0.67	0.62
Uniform / no c	11.70	5.78	-275	-211	0.63	0.25	0.67	0.37
				All fundamentals				
EWA / c = 0.5	11.02	8.99	08-	-37	0.64	0.55	0.67	0.89
EWA/c=1	11.95	9.76	-43	30	0.64	0.58	99.0	0.93
EWA/c=2	11.68	9.37	-83	2	0.62	0.52	0.64	0.91
EWA/c = 5	11.73	8.70	-169	-85	0.63	0.45	0.67	0.83
EWA / c = 10	11.70	8.62	-179	-94	0.63	0.44	0.67	0.82
EWA / no c	11.70	8.62	-179	-94	0.63	0.44	0.67	0.82
Uniform / no c	11.70	5.78	-275	-211	0.63	0.25	0.67	0.37

Appendix for online publication

- A. Detailed data description
- B. Notes about the computations and forecasts
- C. Additional graphs and tables

A. Detailed data description

The original dataset of Molodtsova and Papell [2009] was extended to the period 1973–2014 whenever possible – when the series were not discontinued. Series that had to be substituted have Datastream as the principal source. For those, as there were typically more series available without a clear advantage over one another, only the ones that were eventually used to generate predictions presented in the paper are listed (we did not get qualitatively different results using the alternatives). Data was obtained through Datastream on 30/01/2014. When the code from the original data series was known it was given instead. All series are monthly unless noted.

Quarterly series were transformed into monthly ones by a local quadratic interpolation. This means that a local quadratic polynomial was fit for each set of three adjacent quarterly observations; then, the monthly observations right before and after the center quarterly observation (the second one in the set) were filled in using the value of the quadratic polynomial. The output gap was estimated for each country with at least 24 data points. For the Hodrick-Prescott filter we used $\lambda = 129600$ as advocated by Ravn and Uhlig [2002].

The exact data series (and the periods for which they could be used) are described in detail in the following table.

Table 8: Data series description, part I.

				Earliest date after	
Description	Series name	Intermediate source (if any)	Original source	March 1973 when the	End date
				series is available	
Nominal exchange rate, U.S. – U.K.	EXUSUK		FRED	Mar-73	Dec-14
Nominal exchange rate, Japan U.S.	EXJPUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Switzerland – U.S	EXSZUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Canada U.S.	EXCAUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Sweden – U.S.	EXSDUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Denmark – U.S.	EXDNUS		FRED	Mar-73	Dec-14
Nominal exchange rate, U.S. – Australia	EXUSAL		FRED	Mar-73	Dec-14
Nominal exchange rate, France – U.S.	EXFRUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Germany – U.S.	EXGEUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Italy – U.S.	EXITUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Netherlands – U.S.	EXNEUS		FRED	Mar-73	Dec-14
Nominal exchange rate, Portugal – U.S.	EXPOUS		FRED	Mar-73	Dec-14
End-of-month nominal exchange rate, U.SU.K	UKIAG.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Japan U.S.	JPIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Switzerland U.S.	SWIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Canada U.S.	CNIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Sweden U.S.	SDIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Denmark U.S.	DKIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, U.S Australia	AUIAG.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, France U.S.	FRIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Germany U.S.	BDIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Italy U.S.	ITIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Netherlands U.S.	NLIAE.	Datastream	IFS	Mar-73	Dec-14
End-of-month nominal exchange rate, Portugal U.S.	PTIAE.	Datastream	IFS	Mar-73	Dec-14
M1 money supply, n.s.a., U.S.	USI59MA.A	Datastream	IFS	Mar-73	Nov-14
Notes and coins in circulation outside the Bank of England, n.s.a., U.K.	UKAVAA	Datastream	Bank of England	Mar-73	Dec-14
M1 money supply, billions of yens, n.s.a, Japan	JPI59MA.A	Datastream	IFS	Mar-73	Nov-14
Narrow money, billions of Swiss franks, n.s.a., Switzerland	SWI34A	Datastream	IFS	Mar-73	Oct-14
M1 money supply, billions of Canadian dollars, n.s.a., Canada	CNI59MADA	Datastream	IFS	Mar-73	Oct-14
M0 money supply, millions of Swedish kronors, n.s.a., Sweden	SDM0A	Datastream	Statistics Sweden	Mar-73	Dec-14
M1 money supply, millions of Danish kroners, n.s.a., Denmark	DKOMA027A	Datastream	MEI, OECD	Mar-73	Feb-14
M1 money supply, millions of Australian dollars, n.s.a., Australia	AUOMA027A	Datastream	MEI, OECD	Mar-73	Feb-14
M1 money supply, billions of French franks, n.s.a., France	13259MA.ZF	Molodtsova and Papell (2009)	IFS	Dec-77	Dec-98
M1 money supply, billions of Deutsche marks, s.a., Germany	13459MACZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
M2 money supply, billions of Italian liras, n.s.a., Italy	13659MB.ZF	Molodtsova and Papell (2009)	IFS	Dec-74	Dec-98
M2 money supply, billions of Dutch guilders, n.s.a., Netherlands	13859MB.ZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
M1 money supply, billions of Portuguese escudos, n.s.a., Portugal	18259MA.ZF	Molodtsova and Papell (2009)	IFS	Dec-79	Dec-98

Table 9: Data series description, part II.

Probability of the probabili					Earliest date after	
USIGOB. Datistream FIS Mar-73 UKIGOB. Datastream FIS Mar-73 UKIGOB. Datastream FIS Mar-73 SWOINGT. Datastream FIS Mar-73 CONIGOB. Datastream FIS Mar-73 DKKGGB. Datastream FIS Mar-73 DKKGGB. Datastream FIS Mar-73 DKKGGB. Datastream FIS Mar-73 DKKGGB. Datastream FIS Mar-73 13460BTr. Molocitsova and Papell (2009) FIS Mar-73 UKIGE.CE Datastream FIS Mar-73 DKIGE.CE Molocitsova and Papell (2009) FIS Mar-73 DKIGE.CE Molocitsova and Papell (2009) FIS Mar-73	Description	Series name	Intermediate source (if any)	Original source	March 1973 when the	
USIGOB Datastream IFS Mar-73 UKIGOB Datastream IFS Mar-73 PIOGO Datastream IFS Mar-73 SUNGCC Datastream IFS Mar-73 SUNGG Datastream IFS Mar-73 AUGGB Datastream IFS Mar-73 AUGGB Datastream IFS Mar-73 AUGGB Molodstowa and Papell (2009) IFS Mar-73 1356B Molodstowa and Papell (2009) IFS Mar-73 1352GB Molodstowa and Papell (2009) IFS Mar-73 135A Molodstowa and Papell (2009) IFS Mar-73 135GB Molodstowa and Papell (2009) IFS Mar-73 15GB Datastream IFS Mar-73 15GB.C Molodstowa and Papell (2009) IFS Mar-73 15GB Datastream IFS Mar-73 135GC.C Molodstowa and Papell (2009) IFS Mar-73 15					series is available	End date
UKIGOB. Datastream IFS Mar-73 9 WIGOB. Datastream IFS Mar-73 1 WIGOC. Datastream IFS Mar-73 1 SWOIROTS. Datastream IFS Mar-73 1 SWOIROS. Datastream IFS Mar-73 1 SOBGOB. Datastream IFS Mar-73 1 1 SOBG. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 3560B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 3560B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 3560B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 3560B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 3560B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 356B. J.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 356B. G.T. Moloctsova and Papell (2009) IFS Mar-73 1 1 356B. G.T. Moloctsova and Papell (2009) IFS Mar-73 1 3 356B. G.T.	Federal Funds Rate, United States	USI60B	Datastream	IFS	Mar-73	Dec-14
Piecibs.	Money Market Rate, U.K.	UKI60B	Datastream	IFS	Mar-73	Nov-14
SWORNDYSR Datastream MEI, OECD Jan-74 CNIGOB Datastream FFS Mar-73 DKIGOB Datastream FFS Mar-73 AULIGOB Datastream FFS Mar-73 AULIGOB Datastream FFS Mar-73 AULIGOB Moloctsova and Papell (2009) FFS Mar-73 13460B.ZF Moloctsova and Papell (2009) FFS Mar-73 13460B.ZF Moloctsova and Papell (2009) FFS Mar-73 13460B.ZF Moloctsova and Papell (2009) FFS Mar-73 UKIGC Datastream FFS Mar-73 UKIGF Datastream FFS Mar-73 UKIGF Datastream FFS Mar-73 UKIGF Datastream	Money Market Rate, Japan	JP160B	Datastream	IFS	Mar-73	Dec-14
CNINGO Datastream FIS NAR-73	3-Month Euro Deposits, Switzerland	SWOIR075R	Datastream	MEI, OECD	Jan-74	Nov-14
Detactream	Treasury Bills rate, Canada	CNI60C	Datastream	IFS	Mar-73	Dec-14
AdulGobb Datastream FIS Mar-73 4JUIGOB Datastream FIS Mar-73 1340BZr Molodsova and Papell (2009) FIS Mar-73 1346BZr Molodstova and Papell (2009) FIS Mar-73 1366BZr Molodstova and Papell (2009) FIS Mar-73 1366BZr Molodstova and Papell (2009) FIS Mar-73 UKIBÉCE Datastream FIS Mar-73 UKIBÉCE Datastream FIS Mar-73 DIN Sa., Sweden SDOPRISG Datastream FIS Mar-73 DIN Sa., Sweden SDOPRISG Datastream FIS Mar-73 DIN Sa., Sweden SDOPRISG Datastream FIS Mar-73 AUKGEH Datastream FIS Mar-73 DIN Sa., Sweden SDOPRISG Datastream FIS Mar-73 LISAGETE Molodstova and Papell (2009) FIS Mar-73 J.K. UKIG4BF Datastream FIS Mar-73	Money Market Rate, Sweden	SDI60B	Datastream	IFS	Mar-73	Oct-14
132608.2F Molodstova and Papell (2009) FIS Mar-73 134608.2F Molodstova and Papell (2009) FIS Mar-73 134608.2F Molodstova and Papell (2009) FIS Mar-73 138608.2F Molodstova and Papell (2009) FIS Mar-73 13860.2F Datastream FIS Mar-73 1966.CE Datastream FIS Mar-73 1966.CE Datastream FIS Mar-73 1966.CE Datastream FIS Mar-73 13466.CF Molodstova and Papell (2009) FIS Mar-73 1366.CF Molodstova and Papell (2009) FIS Mar-73 13866.CF Datastream FIS Mar-73 13866.CF Datastream FIS Mar-73 13866.CF Molodstova and Papell (2009) FIS Mar-73 13866.CF Molodstova and Papell (2009) FIS Mar-73 13864.2F Mo	Money Market Rate, Denmark	DKI60B	Datastream	IFS	Mar-73	Dec-14
1350B. ZF Molodstova and Papell (2009) IFS Mar-73 1360B. ZF Molodstova and Papell (2009) IFS Mar-73 1360B. ZF Molodstova and Papell (2009) IFS Mar-73 1386B. ZF Molodstova and Papell (2009) IFS Mar-73 USIGO. ZF. Molodstova and Papell (2009) IFS Mar-73 UKIGO. EB Datastream IFS Mar-73 UKIGO. BH Datastream IFS Mar-73 DKIGO. BH Molodstova and Papell (2009) IFS Mar-73 J3866. ZF Molodstova and Papell (2009) IFS Mar-73 J466. ZF Molodstova and Papell (2009) IFS Mar-73 JK. UKIG4B. F Datastream IFS Mar-73 JAMIGA. F Datastream	Money Market Rate, Australia	AUI60B	Datastream	IFS	Mar-73	Dec-14
13460BZF Molodtsova and Papell (2009) FS Mar-73 1360BZF Molodtsova and Papell (2009) FS Mar-73 1360BZF Molodtsova and Papell (2009) FS Mar-73 UKIGG.CE Datastream FS Mar-73 UKIGG.CE Datastream FS Mar-73 UKIGG.CE Datastream FS Mar-73 DNG.CE Datastream FS Mar-73 DNG.CE Datastream FS Mar-73 DNG.CE Datastream FS Mar-73 Erly ALOGG.CE Datastream FS Mar-73 BNG.C.ZF Molodtsova and Papell (2009) FS Mar-73 J.K. UKIGGE Datastream FS Mar-73 J.K. UKIGBF Datastream FS Mar-73 J.K. UKIGBF Datastream FS Mar-73 J.K. UKIGBF Datastream FS Mar-73 J.K. Datastream FS Mar-73	Money Market Rate, France	13260BZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
136608Zr Molodtsova and Papell (2009) FS Mar-73	Money Market Rate, Germany	13460BZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
13860BZT Molodtsova and Papell (2009) IFS Mar-73 18260BZT Molodtsova and Papell (2009) IFS Mar-73 UKI66CE Datastream IFS Mar-73 UKI66CE Datastream IFS Mar-73 IPIG6CE Datastream IFS Mar-73 Dn, s.a., Sweden SDOPRI35G Datastream ME, OECD Mar-73 PUKI66CE Datastream ME, OECD Mar-73 PUKI66RH Datastream ME, OECD Mar-73 PUKI66RH Datastream ME, OECD Mar-73 1346ZE Molodtsova and Papell (2009) IFS Mar-73 1346ZE Molodtsova and Papell (2009) IFS Mar-73 J.K. UKI64F Datastream IFS Mar-73 J.B.GF.	Money Market Rate, Italy	13660BZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
18260BZF Molodtsova and Papell (2009) IFS Jan-83 UNG6CE Datastream IFS Mar-73 UNG6CE Datastream IFS Mar-73 anterly SWQ6BH Datastream IFS Mar-73 DN, s.a., Sweden SWQ6BH Datastream IFS Mar-73 DN, s.a., Sweden SDOPRI3G Datastream IFS Mar-73 Erly AUQ6CE Datastream IFS Mar-73 Erly AUQ6CE Datastream IFS Mar-73 13466CEL Molodtsova and Papell (2009) IFS Mar-73 13466CEL Molodtsova and Papell (2009) IFS Mar-73 13866CEL Molodtsova and Papell (2009) IFS Mar-73 13866CEL Molodtsova and Papell (2009) IFS Mar-73 1386F Datastream IFS Mar-73 14.64F Datastream IFS Mar-73 1406F Datastream IFS Mar-73 1	Money Market Rate, Netherlands	13860BZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
USIGE.CE Datastream IFS Mar-73 UKIGE.CE Datastream IFS Mar-73 anterly SWQGGBH Datastream IFS Mar-73 on, s.a., Sweden SOPRISSG Datastream IFS Mar-73 erly CNIGGCE Datastream IFS Mar-73 DKIGGBH Datastream IFS Mar-73 Erly AUQGGCE Datastream IFS Mar-73 134GGCZF Molodtsova and Papell (2009) IFS Mar-73 138GGCZF Molodtsova and Papell (2009) IFS Mar-73 138GGZF Molodtsova and Papell (2009) IFS Mar-73 13.8GGZF Molodtsova and Papell (2009) IFS Mar-73 J.K. UKIG4BF Datastream IFS Mar-73 SWIG4F Datastream IFS Mar-73 AULGGF Datastream IFS Mar-73 AUGGF Datastream IFS Mar-73 AUGGF Datastre	Money Market Rate, Portugal	18260BZF	Molodtsova and Papell (2009)	IFS	Jan-83	Dec-98
UKIGG.CE Datastream IFS Mar-73 arterty SWGG.RH Datastream IFS Mar-73 on, s.a., Sweden SWGG.CBH Datastream IFS Mar-73 on, s.a., Sweden SDOPRI35G Datastream IFS Mar-73 erly AUGG.CC Molodsova and Papell (2009) IFS Mar-73 1326G.CZF Molodsova and Papell (2009) IFS Mar-73 1326G.CZF Molodsova and Papell (2009) IFS Mar-73 1326G.CZF Molodsova and Papell (2009) IFS Mar-73 1386G.CZF Molodsova and Papell (2009) IFS Mar-73 J.K. UKIG4B.F Datastream IFS Mar-73 J.K. UKIG4B.F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUG4F Datastream IFS Mar-73 AUG4F Datastream IFS Mar-73	Industrial production, s.a., U.S.	USI66CE	Datastream	IFS	Mar-73	Dec-14
arterly SW066.BH Datastream IFS Mar-73 CN066.BH Datastream IFS Mar-73 CN066.CE Datastream IFS Mar-73 CN066.CE Datastream IFS Mar-73 CN066.CE Datastream IFS Mar-73 CN066.CE Datastream IFS Mar-73 1366.CZF Molodtsova and Papell (2009) IFS Mar-73 1366.CZF Datastream IFS Mar-73 J.K. UK164B.F Datastream IFS Mar-73 CN164F Datastream IFS Mar-73 CONSumer prices: all AU164F Datastream IFS Mar-73 CONSumer prices: all AU164F Datastream IFS Mar-73 I3364ZF Molodtsova and Papell (2009) IFS Mar-73 I3364ZF Molodtsova and Papell	Industrial production, s.a., U.K.	UKI66CE	Datastream	IFS	Mar-73	Nov-14
arterly SWQ66.BH Datastream IF5 Mar-73 CNIG6.CE Datastream IF5 Mar-73 CNIG6.CE Datastream IF5 Mar-73 CNIG6.CE Datastream MEI, OCCD Mar-73 CNIG6.CE Datastream IF5 Mar-73 AUG6.CE Datastream IF5 Mar-73 13266.CZF Molodtsova and Papell (2009) IF5 Mar-73 USIGAF Datastream IF5 Mar-73 SWIG4F Datastream IF5 Mar-73 SWIG4F Datastream IF5 Mar-73 CNIG4F Datastream IF5 Mar-73 CNIG4F Datastream IF5 Mar-73 AUIG4F Datastream IF5 Mar-73 COnsumer prices: all 13264F. Molodtsova and Papell (2009) IF5 Mar-73 13264F. Molodt	Industrial production, s.a., Japan	JP166CE	Datastream	IFS	Mar-73	Oct-14
On, S.a., Sweden CNIGG.CE Datastream IFS Mar-73 Por, S.a., Sweden SDOPRISSG Datastream MRI, OECD Mar-73 erly AUGG6.CE Datastream IFS Mar-73 13266ZF Molodtsova and Papell (2009) IFS Mar-73 13666ZF Molodtsova and Papell (2009) IFS Mar-73 13666ZF Molodtsova and Papell (2009) IFS Mar-73 13666ZF Molodtsova and Papell (2009) IFS Mar-73 J.K. UKI648F Datastream IFS Mar-73 J.K. UKI648F Datastream IFS Mar-73 J.K. UKI648F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 Consumer prices: all Items, 2000-100 Molodtsova and Papell (2009) <	Industrial production, s.a., Switzerland, quarterly	SWQ66BH	Datastream	IFS		
Pon, s.a., Sweden SDOPRIBSG Datastream MEI, OECD Mar-73 erly Datastream IFS Jan-74 Erly AUG66CE Datastream IFS Mar-73 13466CEF Molodtsova and Papell (2009) IFS Mar-73 13866CEF Molodtsova and Papell (2009) IFS Mar-73 13866CEF Molodtsova and Papell (2009) IFS Mar-73 JISAGECEF Molodtsova and Papell (2009) IFS Mar-73 JIK. UKIG4BF Datastream IFS Mar-73 JIK. UKIG4BF Datastream IFS Mar-73 JIK. Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 Consumer prices: all Items, 2000-100 Molodtsova and Papell (2009) IFS Mar-73	Industrial production, s.a., Canada	CNI66CE	Datastream	IFS	Mar-73	Oct-14
erly DKI66BH Datastream IFS Jan-74 4U066CE Datastream IFS Mar-73 13266CZF Molodtsova and Papell (2009) IFS Mar-73 13466CZF Molodtsova and Papell (2009) IFS Mar-73 13866CZF Molodtsova and Papell (2009) IFS Mar-73 USI64F Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 J.K. Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SDI66F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73	Industrial production excluding construction, s.a., Sweden	SDOPRI35G	Datastream	MEI, OECD	Mar-73	Oct-14
erly AUQ66CE Datastream IFS Mar-73 13266CZF Molodtsova and Papell (2009) IFS Mar-73 13466CZF Molodtsova and Papell (2009) IFS Mar-73 13666CZF Molodtsova and Papell (2009) IFS Mar-73 13666CZF Molodtsova and Papell (2009) IFS Mar-73 J.K. UKI648F Datastream IFS Mar-73 J.K. UKI648F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 Consumer prices: all IFS Mar-73 consumer prices: all Items, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and	Industrial production, s.a., Denmark	DKI66BH	Datastream	IFS	Jan-74	Oct-14
1326CZF Molodtsova and Papell (2009) IFS Mar-73 13466CZF Molodtsova and Papell (2009) IFS Mar-73 13466CZF Molodtsova and Papell (2009) IFS Mar-73 13826BZF Molodtsova and Papell (2009) IFS Mar-73 UKIG4BF Datastream IFS Mar-73 J.K. UKIG4BF Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SUI64F Datastream IFS Mar-73 AUI64F Molodtsova and Papell (2009) IFS Mar-73 13864ZF </td <td>Industrial production, s.a., Australia, quarterly</td> <td>AUQ66CE</td> <td>Datastream</td> <td>IFS</td> <td></td> <td></td>	Industrial production, s.a., Australia, quarterly	AUQ66CE	Datastream	IFS		
13466CZF Molodtsova and Papell (2009) IFS Mar-73 13666CZF Molodtsova and Papell (2009) IFS Mar-73 13866CZF Molodtsova and Papell (2009) IFS Mar-73 18266BZF Molodtsova and Papell (2009) IFS Mar-73 UKIG4F Datastream IFS Mar-73 JPIG4F Datastream IFS Mar-73 SWIG4F Datastream IFS Mar-73 CNIG4F Datastream IFS Mar-73 AUIG4F Molodtsova and Papell (2009) IFS Mar-73 13664F Molodtsova	Industrial production, s.a., France	13266CZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
13666CZF Molodtsova and Papell (2009) IFS Mar-73 13866CZF Molodtsova and Papell (2009) IFS Mar-73 13266BZF Molodtsova and Papell (2009) IFS Mar-73 UKI64BF Datastream IFS Mar-73 JPI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 SDI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 Consumer prices: all terms, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 1364ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) <t< td=""><td>Industrial production, s.a., Germany</td><td>13466CZF</td><td>Molodtsova and Papell (2009)</td><td>IFS</td><td>Mar-73</td><td>Dec-98</td></t<>	Industrial production, s.a., Germany	13466CZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
13866.CZF Molodtsova and Papell (2009) IFS Mar-73 18266.BZF Molodtsova and Papell (2009) IFS Mar-73 USI64F Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 CNI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 Consumer prices: all Items, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and	Industrial production, s.a., Italy	13666CZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
1.6.6.BZF Molodtsova and Papell (2009) IFS Mar-73 USI64F Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 CNI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 Consumer prices: all Items, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 1364ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73 Mar-73 Mar-73 <t< td=""><td>Industrial production, s.a., Netherlands</td><td>13866CZF</td><td>Molodtsova and Papell (2009)</td><td>IFS</td><td>Mar-73</td><td>Dec-98</td></t<>	Industrial production, s.a., Netherlands	13866CZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
USI64F Datastream IFS Mar-73 J.K. UKI64BF Datastream IFS Mar-73 JPI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 CNI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Molodtsova and Papell (2009) IFS Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Industrial production, s.a., Portugal	18266BZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
J.K. UKI64BF Datastream IFS Mar-73 JPI64F Datastream IFS Mar-73 SWI64F Datastream IFS Mar-73 CNI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Molodtsova and Papell (2009) IFS Mar-73 1364ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 13264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, U.S.	USI64F	Datastream	IFS	Mar-73	Dec-14
JP164F Datastream IFS Mar-73 SW164F Datastream IFS Mar-73 CN164F Datastream IFS Mar-73 SD164F Datastream IFS Mar-73 DK164F Datastream IFS Mar-73 AU164F Molodtsova and Papell (2009) IFS Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index (retail price index), U.K.	UK164BF	Datastream	IFS	Mar-73	Nov-14
SWI64F Datastream IFS Mar-73 CNI64F Datastream IFS Mar-73 SDI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Molodtsova and Papell (2009) IFS Mar-73 Isems, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 I3864ZF Molodtsova and Papell (2009) IFS Mar-73 I3264ZF Molodtsova and Papell (2009) IFS Mar-73 I3264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Japan	JP164F	Datastream	IFS	Mar-73	Nov-14
CNIG4F Datastream IFS Mar-73 SDIG4F Datastream IFS Mar-73 DXIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 AUIG4F Datastream IFS Mar-73 Consumer prices: all items, 2000=100 Molodtsova and Papell (2009) IFS Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Switzerland	SWI64F	Datastream	IFS	Mar-73	Dec-14
SDI64F Datastream IFS Mar-73 DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 Consumer prices: all items, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Canada	CNI64F	Datastream	IFS	Mar-73	Nov-14
DKI64F Datastream IFS Mar-73 AUI64F Datastream IFS Mar-73 13264ZF Molodtsova and Papell (2009) IFS Mar-73 Consumer prices: all items, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Sweden	SDI64F	Datastream	IFS	Mar-73	Dec-14
AUI64F Datastream IFS Mar-73 13264ZF Molodtsova and Papell (2009) IFS Mar-73 Consumer prices: all items, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Denmark	DK164F	Datastream	IFS	Mar-73	Dec-14
13264ZF Molodtsova and Papell (2009) IFS Mar-73 Consumer prices: all items, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, quarterly, Australia	AUI64F	Datastream	IFS	Mar-73	Aug-14
y tems, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 ands 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, France	13264ZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
y items, 2000=100 Molodtsova and Papell (2009) MEI, OECD Mar-73 13664ZF Molodtsova and Papell (2009) IFS Mar-73 ands 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73		Consumer prices: all				
13664ZF Molodtsova and Papell (2009) IFS Mar-73 13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Germany	items, 2000=100	Molodtsova and Papell (2009)	MEI, OECD	Mar-73	Dec-98
13864ZF Molodtsova and Papell (2009) IFS Mar-73 18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Italy	13664ZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
18264ZF Molodtsova and Papell (2009) IFS Mar-73	Consumer price index, Netherlands	13864ZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98
	Consumer price index, Portugal	18264ZF	Molodtsova and Papell (2009)	IFS	Mar-73	Dec-98

B. Technical complements

We provide in this section

- some additional details on the numerical implementation of our machine-learning methods (Section B.1);
- a more complete description of the economic evaluation of our forecasting strategies, including a presentation of how to put constraints on the currency weights (Section B.2).

B.1. Notes about the computations and forecasts

Calculations were performed with the Scilab 5.5.0 software; the codes are available upon request for the reviewers and will be publicly posted on a website upon publication of the present article in a journal.

The "training" period for our methods was set at 120 months, which is standard in the literature. Correspondingly, the rolling OLS regressions were estimated with 120 months of past data at each instance. The computed coefficients of our forecasts vary greatly over time and it is difficult to discern any time patterns – a feature known in the literature. This is not a surprise as our forecasting methods do not aim to estimate some model, they merely output efficient forecasts. The regularization and discount terms that are computed from past data are nonzero most of the time and they differ substantially between currencies, fundamentals used and time periods. Most of the time the best discount factors are high, above 10, which means that short-term trends are weighted most strongly during estimation.

We tried different grids from the ones in Section 3.3. All grids were logarithmic as is recommended by machine learning theory. We found that a larger grid typically yields small improvements in terms of RMSE over the initial grid we used – but not always as the RMSE of predictions can get worse. With a finer grid we may overfit the regularization and the discounting terms themselves. Smaller obtained RMSEs do not guarantee either obtaining better DM tests: all depends on the errors of individual predictions. It appears that there is a large set of said parameters where the quality of predictions (measured by the size of the RMSE) are qualitatively very similar and improve upon the OLS methods at hand.

B.2. Economic evaluation of forecasting strategies

We recall here the methodology introduced by Della Corte et al. [2009, 2011] and well-summarized in Li et al. [2015]. We followed it scrupulously to produce our main results for the economic evaluation, namely, the ones reported in Table 7 in the "no c" lines. We also considered two additional approaches that deal with known problems of economic criterion evaluation – by constraining the weights, which led to the lines with " $c=\ldots$ " in the main table and in the other tables. These other tables, Tables 22 and 23, report result under a summation-to-less-than-1 constraint for the weights. In the sequel, we only recall the minimal elements of the methodology for economic evaluation to be able to describe the intimated two modifications.

Minimal description of the basic methodology for economic evaluation. We consider one domestic (US) bond, with risk-free return rate $r_{\text{free},t}$, and K foreign bonds, indexed by $k=1,\ldots,K$, with risk-free return rates in their local currencies (that would be used as UIRP fundamentals) but with risky returns $r_{k,t}$ in US dollars because (and only because) of the exchange-rate risk. The subscripts t index the time instances, namely, the months in our case. At each month t-1, the investor forecasts the evolutions of the exchange rates using our methods and based on these forecasts, picks weights $w_t = (w_{1,t}, \ldots, w_{K,t})^T \in \mathbb{R}^K$ determining how to invest her money in the foreign bonds for month t; these weights are real numbers. Her return in month t equals

$$\sum_{k=1}^{K} w_{k,t} r_{k,t} + \left(1 - \sum_{k=1}^{K} w_{k,t}\right) r_{\text{free},t} = r_{\text{free},t} + \sum_{k=1}^{K} w_{k,t} \left(r_{k,t} - r_{\text{free},t}\right).$$

Now, the returns $r_{k,t}$ are not known yet in month t-1 but their conditional means $\overline{r}_{k,t}$ and variance-covariance matrix Σ_t can be assumed to be known (where the conditioning is with respect to the $r_{k,t-1}$ values for month t-1). Indeed, the covariance matrix Σ_t used follows from empirical estimation based on the past covariances, while the means $\overline{r}_{k,t}$ are precisely given by our forecasts. A target conditional volatility $\overline{\sigma}$ is set and the weights are picked to optimize the following mean-variance problem:

$$\begin{aligned} & \max_{w_t \in \mathbb{R}^K} & r_{\text{free},t} + \sum_{k=1}^K w_{k,t} \big(\overline{r}_{k,t} - r_{\text{free},t}\big) \\ & \text{subject to} & w_t^\mathsf{T} \, \Sigma_t \, w_t \leqslant \overline{\sigma}^2 \end{aligned}$$

A closed-form solution exists when no further constraint is imposed on the weights.

Imposing further constraints on the weights. As indicated in the main text, the weights chosen as described above have often large components – especially if the returns of some currencies are highly correlated – and are unstable over time. A good way to avoid these issues is to add a squared regularization; given a parameter c > 0, we want to perform the optimization

$$\begin{aligned} \max_{w_t \in \mathbb{R}^K} \quad & r_{\text{free},t} + \sum_{k=1}^K w_{k,t} \big(\overline{r}_{k,t} - r_{\text{free},t} \big) \\ \text{subject to} \quad & w_t^{\text{T}} \, \Sigma_t \, w_t \leqslant \overline{\sigma}^2 \\ \text{and} \quad & \sum_{k=1}^K w_{k,t}^2 \leqslant c \end{aligned}$$

We provide no closed-form solution for this problem and rather solve it numerically to get our results. The constrained maximization problem at hand being a particular case of what is called a quadratically constrained quadratic program (QCQP), its solution is equal to the one of its Lagrangian dual problem, which is of the form: minimize a (nonlinear) function of finitely many variables $\lambda_1, \lambda_2, \ldots$ subject to positivity constraints only on the arguments λ_i . See [Boyd and Vandenberghe, 2009, Section 5.2.4] for details on this equality. Now, the dual problem is easily solved numerically thanks to toolboxes of various softwares; in our case, we resorted to the optim routine of Scilab.

Further constraining the problem as

$$\begin{aligned} \max_{w_t \in \mathbb{R}^K} \quad & r_{\text{free},t} + \sum_{k=1}^K w_{k,t} \big(\overline{r}_{k,t} - r_{\text{free},t} \big) \\ \text{subject to} \quad & w_t^{\text{T}} \, \Sigma_t \, w_t \leqslant \overline{\sigma}^2 \\ \text{and} \quad & \sum_{k=1}^K w_{k,t}^2 \leqslant c \\ \text{and} \quad & \sum_{k=1}^K w_{k,t} \leqslant 1 \end{aligned}$$

makes it still a QCQP, so that the methodology indicated above applies as well.

Details on these calculations (exact statements of the dual problems, codes written to solve them, etc.) are available upon request.

C. Additional graphs and tables

This section provides complementary studies as well as some tables omitted from the main part.

Machine learning methods and the detection of structural breaks. We chose the time period around the Plaza agreement (September 1985) in Table 10 to compare the performance of sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors with the OLS-based methods for the major currency pairs (USD/GBP, JPY/USD, DEM/USD and FRF/USD) that were the subject of the coordinated currency interventions that followed. This episode was picked because it involved several currencies and may represent a rare "known" structural break. Forecasts using the UIRP- and flexible-prices monetary fundamentals-based sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors react strongly to this "regime" change where major central banks started to intervene in currency markets to bring down the value of the U.S. dollar. In contrast, the predictions from the OLS methods or sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors based on the PPP fundamentals do so much more weakly and slowly (if at all). This substantiates our claim that our algorithm can accommodate quickly structural breaks that might be present in the data³⁴. We obtain similar results for the end-of-month data.

Other tables with results discussed in text.

- Tables 11–12 report the quantiles of the differences in forecasting error between the nochange prediction and the methods under scrutiny.
- Tables 13–14 show the results when using the fundamentals in their "coupled" form.
- Tables 15–16 show the results for the Taylor-rule based exchange rate fundamentals in their "decoupled" (heterogenous) form to compare with the results from Molodtsova and Papell [2009].
- Tables 17–18 give the results for a subsample spanning all the observations since January 1980 till the end of our sample. Note: some results are not available for the PTE/USD pair as data on interest rates starts in 1983.

³⁴We would like to thank Charles Engel for suggesting to perform such an exercise.

- Tables 19–20 report the results on the original Molodtsova and Papell [2009] 1973-2006 sample.
- Tables 21–23 show robustness results for the economic evaluation of the prediction performance of the exponentially weighted average strategy with discount factors: with or without the Danish Krone, and with or without a summation-to-less-than-one constraint on the currency weights.

Other graphs with results discussed in text. By construction, the exponentially weighted average strategy with discount factors has a natural interpretation as being able to identify the "scapegoat" fundamentals considered by traders in some periods as in the model of Bacchetta and Wincoop [2004]. This is because based on past data and discounting, the fundamentals that are best at predicting exchange rates in a given period are given more weight while forming predictions. It is insightful therefore to investigate the behavior of these weights throughout time. For the sake of brevity, we show the results in Figures 1–6 only for the case of predictions with all fundamentals for the JPY/USD, USD/GBP and CHF/USD end-of-month and average exchange rates. The codes for the various fundamentals are the following: PPP refers to inflation rates; UIRP to the moneymarket interest rates; MONEY to money stocks, Y to industrial production, while INTCH denotes the changes in the money market interest rates. Finally, "no-change" refers to the no-change forecast that also is assigned a weight in the exponentially weighted average strategy with discount factors.

The conclusion from these figures is striking: typically only one fundamental is given a predominant weight in a given period; and most of the time for very short time intervals. In most cases, the fundamentals that are given most weight pertain to money market interest rates and money supply, and for one country only. Rarely we observe this behavior for inflation or industrial production. This behavior would be consistent with the scapegoat model and fickle nature of expectations about the fundamentals driving the exchange rate. This is, however, open to yet another interpretation. It also may point out that different shocks to different fundamentals (news) could be driving exchange rates at different times, and (given their prevalence in our data) monetary policy and financial market shocks are most important.

Table 10: Actual monthly exchange rate changes versus those forecasted by rolling, recursive and sequential ridge regression with discount factors and the exponentially weighted average strategy with discount factors for the PPP, UIRP and monetary-model fundamentals around the Plaza agreement (September 1985) for major currencies for the average exchange rate sample.

	multiness canton cuto				to:Poor!	cool lossitos be	, opendo matin	t the menth.	tor opendano	001			
Ë	monthly change (relative to	Rolling	Recursive	C P C C C C C C C C C C C C C C C C C C	redict.	eu natural loga Rolling	Predicted flatural logarithm change of the monthly exchange rate x too	Spides	excilatinge rate	Rolling	Recursive	7	4,440
ש ב	the preceding month) in the	regression	regression	agninc	EWA	regression	regression	agnive	EWA.	regression	regression	agniuc	EWA.
	exchange rate x 100		PPP experts	perts			UIRP experts	xperts		Wo	Monetary model with flexible prices	ith flexible pr	ces
1	1			6		USD/GBP							
avr-85	9.5205	-0.2245	-0.1878	0.1300	0.0000	0.6009	-0.5556	0.5697	-0.2635	-0.3014	-0.4503	0.1275	0.4713
iuin-85	0.8328	-0.2597	-0.2376	0.5773	0.2584	-0.2131	-0.2350	3.0233	0.6390	0.4835	0.0737	0.7916	0.5921
juil-85	7.5106	-0.2610	-0.2465	0.7263	0.2908	0.0003	-0.1359	1.4618	0.6050	0.6230	0.1734	1.2340	0.7348
août-85	0.2459	-0.2011	-0.2239	1.7780	0.2918	0.1249	-0.0409	2.9214	0.6321	0.3266	0.0288	3.3319	0.5578
sept-85	-1.4482	-0.1551	-0.1974	0.4860	0.1668	0.1252	-0.0453	0.9238	0.6336	0.1935	-0.0007	0.9612	0.6559
oct-85	4.1144	-0.1442	-0.1916	-0.0468	-0.0004	0.0819	-0.0776	-0.0736	0.0000	-0.1220	-0.1709	-0.1574	-0.0034
nov-85	1.2653	-0.1120	-0.1548	0.6791	0.2546	0.1226	-0.0376	1.3292	0.6406	-0.4584	-0.3716	1.4193	0.9174
dec-85	0.3536	-0.0990	-0.1253	0.3956	0.0796	0.1/8/	0.0008	0.8302	0.6452	-0.0612	-0.0322	0.9773	0.9325
févr-86	0.3714	-0.0937	-0.1068	-0.1562	-0.0356	0.1904	0.0032	-0.2210	-0.0396	-1.3067	-0.8689	-0.3564	-0.2003
mars-86	2.6027	-0.1106	-0.1362	0.0452	0.0000	0.2246	0.0301	0.1626	0.0000	-0.6994	-0.5346	0.0668	0.3105
	,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	0000	7610	0000	2000	JP1/USD			0000	200	0	0000	1,00
avr-85	-2.3836	0.0880	0.0536	0.0089	0.0003	-0.0657	0.0890	0.0232	0.0000	-0.3014	0.0503	0.0263	0.0471
inin-85	1 1540.	0.0990	0.0330	-0.1043	-0.2942	-0.0909	0.0868	-0.3200	0.000	0.3339	-0.1444	-0.2314	-0.4621
iuil-85	-3.1447	-0.0586	0.0270	-0.0635	-0.2846	-0.2424	0.0068	-0.1729	-0.5807	0.6230	-0.068	-0.2266	-0.3658
août-85	-1.5360	-0.0765	0.0223	-0.1832	-0.2918	-0.3132	-0.0073	-0.5047	-0.6321	0.3266	-0.1244	-0.6441	-0.7113
sept-85	-0.3938	-0.2414	0.0032	-0.1340	-0.2738	-0.3131	0.0070	-0.3585	-0.6336	0.1935	-0.2458	-0.4689	-0.7649
oct-85	-9.6914	0.0289	0.0277	-0.0509	-0.2356	-0.4152	-0.0216	-0.1736	-0.5555	-0.1220	-0.3639	-0.2181	-0.4503
nov-85	-5.0670	-0.0387	0.0101	-0.3811	-0.2549	-0.5683	-0.0791	-2.1351	-0.6406	-0.4584	-0.6294	-2.5959	-0.9174
déc-85	-0.6320	0.0815	0.0134	-0.3125	-0.2883	-0.8785	-0.1904	-1.7159	-0.6452	-0.0612	-0.7502	-2.0464	-0.9325
janv-86	-1.4392	0.1049	0.0142	-0.1351	-0.3049	-1.1638	-0.2740	-0.7296	-0.6621	-1.0470	-0.7856	-0.8405	-0.9569
févr-86	-7.8216	0.1267	0.0125	-0.1354	-0.3209	-0.8020	-0.1464	-0.6299	-0.6521	-1.3067	-0.7382	-0.7615	-0.9201
mars-86	-3.3880	-0.051/	-0.0129	-0.3811	-0.2533	-0.5572	-0.079	-1.8235	-0.6305	-0.6994	-0.9359	-2.2024	-0.8723
	0.41	0,00	,000	0000	0100	DEINI) OSD	0	**000		0,000	70070	0 7	0000
avr-85 mai-85	-5.3/18 0.4739	0.2340	0.0284	0.0866 -0.3201	0.2058	0.3786	0.2450	0.2044 -0.5840	0.2222	0.0319	0.2284	0.1272	-0.0247
juin-85	-1.4807	0.1905	0.0210	-0.0332	-0.0685	0.3027	0.1758	-0.0557	0.0098	0.0452	0.1746	-0.0381	-0.3916
juil-85	-5.1988	0.1295	0.0287	-0.0909	-0.1791	0.2780	0.1433	-0.1557	0.0017	-1.6376	-0.9217	-1.3419	-1.3289
août-85	-4.0236	0.1193	0.0142	-0.2892	-0.2916	0.2463	0.1951	-0.5157	-0.0333	-0.1205	-0.0294	-0.9543	-0.4445
sept-85	1.5768	0.0849	0.0079	-0.2629	-0.2750	0.2377	0.2166	-0.5722	-0.6242	0.1954	0.1787	-0.7839	-0.3161
oct-85	-7.0615	0.0869	0.0061	-0.0069	0.0000	0.2492	0.2438	0.0005	-0.0231	0.3880	0.3409	0.2028	-0.2443
nov-85	-1.8779	0.0282	0.0150	-0.3010	-0.2653	0.2198	0.2170	-0.6637	-0.6394	-0.1748	-0.1585	-1.4088	-0.9174
: dec-85	-3.2582	0.0082	0.0197	-0.1691	-0.2934	0.2066	0.2038	-0.4058	-0.6003	-0.3469	-0.31/4	-0.9595	-0.9231
févr-86	-4.4744	-0.0532	0.0179	-0.2210	-0.3181	0.1688	0.1762	-0.3821	-0.6307	-0.0780	-0.1218	-1.0797	-0.9201
mars-86	-2.4530	-0.1108	0.0072	-0.2372	-0.2610	0.1359	0.1323	-0.6320	-0.6261	-0.4612	-0.4187	-1.4225	-0.8723
						FRF/USD							
avr-85	-6.5073	0.6470	0.5062	0.1995	0.3387	0.6574	0.4625	0.5326	0.7729	0.6422	0.6422	0.7009	0.8180
mai-85	0.4259	0.6503	0.5037	-0.4586	-0.1805	0.5712	0.3858	-0.4238	-0.3422	0.2692	0.2692	0.2155	-0.2007
juin-85	-1.5045	0.6476	0.4842	0.0022	0.0115	0.5770	0.3649	0.1462	0.1279	0.8395	0.8395	0.8260	0.3556
co-line	2985.6-	0.6520	0.4616	2/60.0-	-0.0463	0.3130	0.27.80	0.0320	0.0288	1.0270	1.0279	0.3733	0.7157
sent-85	1.4856	0.5333	0.3661	97550-	-0 1541	0.4544	0.3150	-0.8138	-0.2601	0.8491	0.8491	0.2373	-0.8223
oct-85	-7.1293	0.4452	0.3421	0.0318	0.0270	0.4687	0.3461	0.1882	0.1680	0.7210	0.7210	0.7453	-0.0298
nov-85	-1.9358	0.3624	0.2652	-0.4338	-0.1777	0.4614	0.3552	-1.0426	-0.6030	-0.0966	-0.0966	-1.3423	-0.9174
déc-85	-2.8807	0.2838	0.1972	-0.2094	-0.1975	0.4862	0.3985	-0.5892	-0.4577	-1.2851	-1.2851	-1.4672	-0.9325
janv-86	-2.6744	0.2359	0.1591	-0.2329	-0.2256	0.4982	0.4255	-0.6462	-0.4962	-1.2759	-1.2759	-1.4634	-0.9569
févr-86	-4.4353	0.1333	0.0757	-0.2137	-0.3181	0.4608	0.3873	-0.6202	-0.4808	-1.2891	-1.2891	-1.4046	-0.9201
mars-86	-2.2765	0.09/1	0.0516	-0.2543	-0.2610	0.4079	0.3350	-0.8690	-0.5541	-0.118/	-0.118/	-0.6030	-0.8/23

Table 11: Quantiles of the differences in forecasting error between the no-change prediction and the method under scrutiny for the average exchange rates sample. Distributions of errors that are shifted towards positive values indicate methods that outperfom the no-change prediction.

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Currency pair	10%	KOIII 25%	Kolling regression 50%	ion 75%	%06	10%	Recur 25%	Kecursive regression % 50% 75	sion 75%	%06	10%	25%	skidge 50%	75%	%06	10%	25%	EWA 50%	75%	%06
									PPP fu	ındamental										
USD/GBP	-1.30	-0.39	-0.01	0.33	1.12	-0.41	-0.13	0.00	0.10	0.34	-1.02	-0.22	0.03	0.36	1.68	-0.83	-0.16	0.00	0.26	1.25
JPY/USD	-2.05	-0.81	-0.03	0.40	1.36	-0.32	-0.14	-0.01	0.12	0.32	-1.16	-0.29	0.02	99.0	1.98	-0.82	-0.04	0.00	0.31	1.30
CHF/USD	-1.54	-0.43	0.01	0.41	1.22	-0.69	-0.26	-0.01	0.28	69.0	-0.73	-0.12	0.04	0.42	1.65	-0.88	-0.15	0.02	0.56	1.47
CAD/USD	-0.52	-0.14	0.01	0.22	99.0	-0.21	-0.10	0.01	0.11	0.23	-0.14	-0.05	0.01	0.08	0.24	-0.07	-0.02	0.00	0.02	60.0
SEK/USD	-2.41	-0.91	-0.10	0.76	2.42	-1.01	-0.39	-0.02	0.34	0.97	-0.98	-0.19	0.08	0.72	1.98	-0.51	-0.13	0.03	0.36	1.10
DNK/USD	-2.04	-0.73	-0.03	0.52	1.40	-1.10	-0.41	0.06	0.43	0.89	-0.53	-0.12	0.04	0.48	1.26	-0.81	-0.26	0.02	0.73	1.59
USD/AUD	-2.07	-0.77	-0.03	0.76	2.30	-0.68	-0.29	-0.03	0.20	0.71	-0.77	-0.18	0.03	0.42	1.54	-1.01	-0.27	0.00	0.40	1.29
FRF/USD	-1.96	-0.73	-0.09	0.49	2.05	-1.17	-0.41	-0.02	0.33	1.44	-0.59	-0.16	0.05	0.58	1.65	-1.13	-0.39	0.17	0.91	2.04
DEM/USD	-1.95	0.71	-0.05	0.46	1.26	-0.45	-0.16	-0.03	0.10	0.32	1.10	-0.25	0.10	0.68	1.86	-1.16	-0.26	0.13	0.80	1.83
NLG/USD	-2.17	-0.66	-0.06	0.35	1.26	-1.01	60.0-	-0.04	0.05	0.15	-0.75	-0.20	0.09	0.70	1.67	-1.04	-0.30	0.07	0.77	1.72
PTE/USD	-4.92	-1.93	-0.15	1.34	4.96	-4.02	-1.54	0.00	1.21	4.59	-2.59	-0.90	0.01	0.94	4.39	-2.91	-0.28	0.00	0.82	00.9
									UIRP fu	undamental										
USD/GBP	-1.99	-0.51	-0.02	0.49	2.22	-0.84	-0.21	0.01	0.26	1.15	-1.08	-0.19	0.01	0.35	2.85	-1.22	-0.16	0.00	0.40	2.06
JPY/USD	-2.20	-0.72	0.05	0.84	2.73	-1.14	-0.26	0.00	0.48	1.66	-1.09	-0.11	0.00	0.47	1.64	-1.19	0.00	0.00	0.24	1.88
CHF/USD	-2.97	-0.44	0.03	0.68	2.10	-1.75	-0.22	0.02	0.46	1.96	-1.71	-0.16	0.01	69.0	2.80	-1.59	-0.14	0.00	0.80	2.40
CAD/USD	-0.63	-0.29	0.02	0.21	0.56	-0.23	-0.08	0.00	0.07	0.22	-0.17	-0.05	0.00	0.02	0.21	-0.08	-0.03	0.00	0.03	0.11
SEK/USD	-2.90	-0.90	-0.01	0.63	1.95	-1.21	-0.40	-0.03	0.38	0.84	-1.14	-0.21	0.03	0.69	2.23	-1.02	-0.16	0.03	0.52	1.92
DNK/USD	-2.79	-0.60	-0.02	0.56	1.99	-1.97	-0.50	0.00	0.47	1.60	96:0-	-0.13	0.01	09.0	2.23	-1.12	-0.17	0.00	0.68	2.44
USD/AUD	-2.95	-0.91	-0.02	0.90	2.86	-1.31	-0.50	0.00	0.40	1.04	-1.08	-0.26	0.02	0.54	1.77	-0.93	-0.21	0.00	0.28	1.03
FRF/USD	-2.65	-1.07	-0.04	0.56	1.92	-2.72	-0.99	-0.12	0.72	2.09	-1.42	-0.47	0.16	1.45	3.15	-2.03	-0.51	0.19	1.47	2.79
DEM/USD	-2.26	-1.03	-0.01	0.79	2.23	-1.51	-0.24	0.00	0.39	1.34	-1.79	-0.47	0.16	1.46	3.12	-2.01	-0.43	0.05	1.36	3.13
ITL/USD	-4.37	-1.36	-0.02	1.12	3.50	-3.21	-1.13	-0.08	1.10	2.22	-2.61	-0.58	0.17	2.11	4.52	-1.99	-0.25	0.07	1.32	3.52
NLG/USD	-4.10	-1.11	-0.03	1.52	4.03	-2.34	-0.40	0.10	0.97	2.65	-1.84	-0.55	0.20	1.47	3.53	-2.15	-0.48	0.01	1.32	3.38
PTE/USD	-1.32	-0.49	-0.Ub	0.58	1.04	-1.44	-0.00	0.04 M	U.D.	1.40	-1.49	-0.22	0.08	0.75	2.33	-4.59	-0./0	0.02	1.23	3.08
000/001	77 د	777	900	60.0	20.0	777	0.43	2	Olicial y III	1 76	2.72	72.0	90.0	0 70	, ,,,	1.66	04.0		77. 0	,
USU/GBP	6/.6-	01.1-	0.00	1.01	2.03	-2.14 07.0	-0.73	0.01	0.42	1.70 2.04	67.7-	,o.o,	0.00	1.76	3.21 7.7E	2.17	0.40	0.02	7.7	20.0
CHE/LISD	-4.20	-1.39	0.02	1.43	2.75	2.70	-1.12 -0 95	0.01	1.10	2.54	-3.02 -2.85	20.0-	0.03	1.30	2 86	-2.14	-0.04	0.05	1.21	2 84
CAD/USD	-1.18	-0.49	-0.02	0.32	1.00	-0.45	-0.20	0.00	0.15	0.46	-0.47	-0.12	0.01	0.24	0.72	-0.23	-0.04	0.00	0.05	0.24
SEK/USD	-4.61	-1.67	-0.11	0.71	2.36	-2.03	-0.67	-0.04	0.27	1.00	-2.20	-0.60	0.13	1.43	3.20	-2.02	-0.40	0.10	1.23	2.95
DNK/USD	-2.88	-1.30	-0.20	0.39	1.21	-1.76	-0.57	-0.02	0.25	0.89	-1.81	-0.43	0.02	1.11	3.39	-2.57	-0.92	0.00	1.72	3.99
USD/AUD	-4.05	-1.32	-0.08	0.71	2.47	-1.64	-0.62	-0.03	0.32	1.13	-2.20	-0.52	0.02	0.98	3.49	-2.44	-0.55	0.03	1.03	3.16
FRF/USD	-5.79	-2.04	-0.08	1.08	2.94	-6.37	-1.42	-0.01	69.0	2.83	-2.49	-0.72	0.15	1.32	3.45	-2.09	-0.51	60.0	96.0	2.11
DEM/USD	-3.52	-1.31	0.10	1.37	3.63	-2.68	-0.95	0.12	1.19	3.04	-2.94	-1.07	0.31	1.82	5.31	-1.91	-0.55	0.18	1.72	3.63
ITL/USD	-4.90	-1.99	-0.10	0.82	2.69	-3.05	-0.92	-0.01	0.82	2.09	-2.44	-0.53	0.10	1.56	5.02	-2.40	-0.57	0.07	1.73	3.68
NLG/USD PTF/USD	-4.05	-1.39	-0.20	1.21	2.13	-2.75	-0.94	-0.01	0.79	1.90	-2.00	-0.67	0.25	1.80	4.35	-1./3	-0.62	0.05	1.38	3.20
									All fur	ndamentals										
USD/GBP	-6.24	-1.92	-0.11	1.43	4.00	-2.69	-0.79	-0.01	0.58	2.39	-2.14	-0.51	0.03	69.0	2.94	-2.21	-0.76	0.04	0.95	3.47
JPY/USD	-7.32	-2.65	-0.09	1.49	4.44	-6.23	-2.27	0.00	1.62	5.28	-2.34	-0.61	0.02	1.13	3.64	-1.88	-0.63	0.03	1.18	4.01
CHF/USD	-8.60	-2.73	-0.07	1.73	4.73	-4.83	-1.59	-0.01	1.50	4.28	-2.56	-0.82	0.13	1.65	3.72	-3.15	-1.00	0.07	1.96	4.25
CAD/USD	-1.71	-0.60	0.00	0.59	1.71	-0.63	-0.21	0.00	0.17	0.59	-0.36	-0.09	0.02	0.20	0.59	-0.18	-0.06	0.00	0.08	0.27
SEK/USD	-7.11	-2.23	-0.07	1.43	4.12	-3.13	-1.03	0.01	0.84	2.82	-1.69	-0.40	0.12	1.11	2.67	-2.17	-0.39	90.0	1.13	3.18
DNK/USD	-6.01	-2.44	-0.12	1.31	3.93	-2.79	-0.81	0.04	0.65	2.10	-1.41	-0.54	0.07	1.22	3.30	-2.91	-0.92	0.08	1.91	4.22
USD/AUD	-6.42	-1.99	-0.08	1.55	5.19	-3.06	-1.11	-0.02	0.62	2.17	-1.97	-0.44	0.02	0.83	2.70	-1.37	-0.37	0.02	0.62	2.33
FRF/USD	-8.71	-2.89	-0.11	1.75	5.16	-5.08	-2.56	0.02	1.86	4.34	-2.08	-0.72	0.15	1.29	3.63	-2.95	-1.31	0.14	1.92	4.06
DEM/USD	-6.28	-1.52	0.03	2.10	4.62	-6.13	-1.29	0.12	2.19	4.53	-2.35	-0.85	0.21	1.61	4.88	-2.71	-1.15	0.24	2.34	4.72
IIL/USD	77.60	-1.44 2.15	70.0	L.35	4.55	10.0-	-1.//	-0.02	1.05	20.0	107	-0.41	0.03	1.00	00.4	-1.00	0.50	0.00	1.24	9.04
NLG/USD PTE/LISD	00.7-	1.26	0.00	1 07	2.00	96.0-	1.40	50.0	1./3	4.00 10.0	1.97	0.00	0.19	1./4	7.40	2.53	100	0.20	1.6/	3.78
r r/ 0.50	. J.v.	J.30	-U.U-	1.0,1	0.00	-0.00	CC:T-	رد.۷-	0.0	۲.3	-T./ C	-U.34	0.10	20.0	04.7	-72-	-1.00	0.00	1.47	ļ. ļ.

Table 12: Quantiles of the differences in forecasting error between the no-change prediction and the method under scrutiny for the end-of-month exchange rates sample. Distributions of errors that are shifted towards positive values indicate methods that outperfom the no-change prediction.

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Currency pair	10%	KOIII 25%	Kolling regression 50%	ion 75%	%06	10%	Recur 25%	Kecursive regression % 50% 75	sion 75%	%06	10%	25%	skidge 50%	75%	%06	10%	25%	EWA 50%	75%	%06
									PPP fu	ındamental										
USD/GBP	-1.32	-0.45	-0.01	0.41	1.00	-0.55	-0.14	0.00	0.12	0.50	-0.21	-0.06	0.00	0.04	0.25	-1.47	-0.38	0.00	0.37	1.65
JPY/USD	-2.32	-0.78	-0.06	0.53	1.75	-0.30	-0.13	-0.01	0.10	0.26	-0.02	0.00	0.00	0.01	0.02	-0.90	-0.11	0.00	0.24	1.06
CHF/USD	-1.81	-0.51	0.00	0.43	1.48	-0.78	-0.25	0.00	0.26	0.85	-0.14	-0.03	0.00	90.0	0.21	-1.34	-0.29	0.00	0.42	1.76
CAD/USD	-0.65	-0.21	0.00	0.19	0.76	-0.26	-0.11	-0.01	0.10	0.27	-0.06	-0.02	0.00	0.02	0.05	-0.05	-0.01	0.00	0.01	0.04
SEK/USD DAIK/USD	-2.84 97.6	-1.02	-0.07	0.90	2.78	-1.11	-0.49	-0.03	0.42	1.28	-0.33	-0.09	0.00	0.08	0.37	-1.19	-0.23	0.00	0.32	1.41
USD/ALID	07.2- 20.2-	-0.02	-0.02	0.03	2.56	-1.27 -0.97	0.00	0.01	0.45	1.24	-0.00	0.01	8.6	0.03	0.10	-1.19	0.50	0.00	0.40	1.57 0.48
FRE/LISD	-2.39	-1.13	-0.03	0.52	2.06	-1.17	-0.30	0.01	0.20	1,11	-0.24	-0.08	0.00	0.04	0.10	-1.00	-0.55	0.00	0.03	1 48
DEM/USD	-2.05	-0.71	-0.04	0.41	1.50	-0.44	-0.18	0.00	0.09	0.27	-0.11	-0.02	0.01	0.06	0.16	-1.53	-0.33	0.09	0.98	1.76
ITL/USD	-2.38	-0.83	-0.04	0.47	2.28	-2.17	-0.76	-0.16	0.66	2.39	-1.67	-0.57	-0.11	0.60	1.86	-1.46	-0.35	0.03	0.83	3.15
NLG/USD	-2.24	-0.53	-0.01	0.32	1.30	-0.29	-0.11	-0.01	90.0	0.15	-0.11	-0.02	0.00	0.05	0.14	-1.46	-0.39	0.02	0.90	1.85
PTE/USD	-5.02	-2.09	-0.21	1.54	4.88	-4.43	-1.38	-0.09	0.99	4.45	-3.72	-1.24	-0.13	1.12	4.33	-2.72	-0.70	0.00	0.75	3.94
									UIRP ti	undamenta										
USD/GBP	-2.58	-0.64	-0.03	0.54	2.59	-0.92	-0.25	0.00	0.26	1.36	-0.29	-0.06	0.00	0.05	0.42	-2.26	-0.49	0.00	0.35	2.46
JPY/USD	-2.52	-0.92	0.03	0.97	3.19	-1.14	-0.27	0.00	0.39	1.64	-0.23	-0.02	0.00	0.04	0.25	-1.57	-0.08	0.00	0.16	1.78
CHF/USD	-3.41	-0.68	0.05	0.79	2.76	-1.76	-0.32	0.01	0.65	2.44	-0.41	-0.06	0.00	0.12	0.54	-1.67	-0.02	0.00	0.00	2.45
CAD/USD	-0.83	-0.26	0.02	0.29	69.0	-0.28	-0.10	0.00	0.10	0.24	-0.02	0.00	0.00	0.00	0.01	-0.05	-0.01	0.00	0.00	0.02
SEK/USD	-3.89	-1.40	-0.05	0.76	2.76	-2.09	-0.55	-0.03	0.21	1.55	-0.68	-0.12	0.00	0.10	0.72	-2.21	-0.57	0.00	0.38	2.06
UNK/USD	99.7-	-0.68	-0.02	0.47	1.98	-1.49	-0.39	-0.01	95.0	1.3/	50.0	-0.01	0.00	0.01	0.08	-1.92	97.0	0.00	79.0	2.33
USD/AUD	-3.84	-1.13	-0.03	1.04 2.04	3.13	-1.50	-U.58	0.00	0.45 0.45	1.23	-0.63	-0.19	0.00	0.12	0.36	-0.26	0.00	0.00	0.00	0.08
FRF/USD DEMANISD	-3.35	-1.38	-0.01	0.62	2.26	-2.64	-1.02	60.0-	0.93	1.37	-0.92	-0.15	0.02	0.25	9.0	-3.5/	-1.19	0.21	1.29	3.21
DEINI/ USD	-2.13 7.11	1 60	40.0	1.06	2 80	-1.10	1 20	20.0	0.43	1.20	0.40	0.10	50.0	0.23	0.09	2.00	1 20	0.12	1.72 1.60	0.30
NLG/USD	-4.07	-1.02	0.00	1.04	3.35	-0.83	-0.21	0.04	0.39	1.07	-0.28	-0.05	0.01	0.15	0.31	-3.18	-1.01	0.02	1.61	3.64
PTE/USD	-2.10	-0.92	-0.11	0.67	2.60	-1.45	-0.78	-0.08	0.45	1.35	-0.70	-0.30	-0.02	0.26	0.74	-0.80	-0.28	0.00	0.22	0.57
									ΙĒ	odel fundan	nentals									
USD/GBP	-4.29	-1.28	-0.09	99.0	2.76	-2.29	-0.71	0.01	0.45	2.22	-1.32	-0.27	0.00	0.25	1.28	-2.49	-0.87	0.00	0.81	2.58
JPY/USD	-4.37	-1.04	-0.01	1.12	4.69	-3.85	-1.12	-0.01	1.00	4.25	-2.31	-0.54	0.01	0.67	2.47	-3.23	-0.90	0.00	1.18	3.61
CHF/USD	-5.61	-1.74	-0.09	1.44	4.82	-4.71	-1.25	0.01	1.43	4.21	-1.33	-0.37	0.01	0.50	1.36	-1.76	-0.48	0.01	0.58	1.45
CAD/USD	-1.86	-0.64	-0.01	0.36	1.09	-0.71	-0.26	-0.02	0.18	0.59	-0.11	-0.01	0.00	0.01	0.09	-0.09	-0.03	0.00	0.02	90.0
SEK/USD	-6.77	-2.37	-0.35	0.65	2.93	-3.05	-0.83	-0.05	0.42	1.35	-0.85	-0.11	0.00	0.17	0.72	-2.41	-0.79	0.00	0.71	2.45
DNK/USD	-5.15	-1.51	-0.14	0.52	1.76	-2.26	-0.66	-0.01	0.37	0.99	-0.02	-0.01	0.00	0.01	0.01	-0.80	-0.28	-0.01	0.21	0.67
USD/AUD	-4.72	-1.57	-0.01	1.03	3.66	-2.00	-0.67	-0.02	0.45	1.50	-0.24	-0.02	0.00	0.02	0.08	-1.67	-0.46	-0.01	0.43	1.49
PRF/USD DEM/HSD	0.40	-4.57	71.0-	1.10	2.20	-5.51	77.7-	-0.1/ 0.08	1.73	3.30 3.15	-1.50	0.30	0.02	0.33	1.18	76.7-	-1.10	0.01	1.75	7.71
ITL/USD	-4.61	-1.98	-0.40	0.95	3.87	-3.23	-1.41	-0.16	0.83	3.42	-1.20	-0.41	0.00	0.26	0.83	-4.10	-1.15	0.00	1.09	3.50
NLG/USD	-6.38	-2.11	-0.31	0.89	3.88	-3.81	-1.03	0.00	0.82	3.10	-0.57	-0.09	0.02	0.24	0.62	-3.00	-0.96	0.00	1.05	3.33
PTE/USD	-6.51	-1.60	-0.21	1.38	4.32	-3.80	-1.54	-0.29	0.42	1.30	-1.77	-0.41	-0.06	0.14	0.87	-2.08	-0.54	-0.01	99.0	1.76
									All fu	ndamentals										
USD/GBP	-9.18	-2.66	-0.16	1.94	5.54	-3.11	-0.84	-0.07	99.0	2.94	-0.64	-0.16	0.00	0.19	0.84	-3.09	-0.93	-0.03	69:0	3.44
JPY/USD	-7.77	-2.95	-0.08	2.01	7.04	-5.78	-2.35	-0.05	1.99	5.71	-1.58	-0.38	0.02	0.55	2.03	-3.62	-1.08	0.01	1.21	3.46
CHF/USD	-11.91	-4.06	-0.32	2.19	6.89	-7.11	-2.06	-0.04	1.68	5.74	-0.70	-0.14	0.01	0.15	0.52	-3.87	-1.23	0.01	1.39	4.18
CAD/USD	-2.65	-0.88	0.00	0.64	1.88	-0.82	-0.28	-0.01	0.22	0.65	-0.06	-0.01	0.00	0.01	0.02	-0.09	-0.03	0.00	0.01	0.02
SEK/USD	-9.47	-3.49	-0.20	1.70	5.17	-4.71	-1.63	-0.14	1.18	4.07	-0.35	-0.08	0.00	90.0	0.34	-3.39	-1.40	0.01	1.37	3.72
DNK/USD	-7.65	-2.45	-0.13	1.43	5.47	-3.15	-0.95	-0.01	0.73	2.79	-0.04	-0.01	0.00	0.01	0.02	-0.53	-0.18	-0.01	0.12	0.33
USD/AUD	-9.14	-3.09	-0.15	2.00	5.79	-4.24	-1.31	0.00	0.86	7.52	-0.46	-0.11	0.00	0.07	0.26	-2.80	-0.26	0.00	1.77	1.44
DFM/USD	-7.86 -7.86	-5.10	-0.00	1.40	5.20	-5.68	-2.19	0.00	1.80	5.06	-1.35	-0.09	00.00	0.10	0.92	-4.64	-1.39	0.01	1.47	4.32
ITL/USD	-6.21	-2.47	-0.16	1.24	6.67	-5.77	-1.98	-0.16	1.25	6.08	-1.55	-0.54	-0.06	0.33	1.33	-5.39	-1.67	0.03	1.38	6.05
NLG/USD	-9.23	-3.09	-0.45	2.36	8.71	-8.09	-2.47	0.00	1.49	5.67	-0.47	-0.08	0.02	0.23	0.73	-4.11	-1.01	0.02	1.67	4.23
PTE/USD	-4.84	-2.53	-0.25	1.63	4.40	-4.28	-1.87	-0.31	0.51	2.07	-0.68	-0.26	-0.02	0.15	0.57	-0.83	-0.36	-0.01	0.22	0.92
																				1

Table 13: RMSE and Theil ratios of forecasts using coupled fundamentals for the average exchange rates sample.

Currency pair			Kolling regression			sion		;	Skidge		;	EWA	
	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
							PPP fundamental						
USD/GBP	2.4410	1.0048	0.824	0.911	1.0045		806.0	0.9945	0.015 **	0.328	0.9998	0.331	0.447
JPY/USD	2.7039	1.0067	0.658	0.795	1.0008		0.812	0.9813	0.000	0.041 **	0.9911	0.000	0.005 ***
CHF/USD	2.7960	1.0029	0.629	0.762	1.0012		0.760	9066:0	0.030 **	0.194	0.9967	0.029 **	0.051 *
CAD/USD	1.5091	0.9877	0.003 ***	0.160	0.9972	*	0.139	0.9844	0.003 ***	0.042 **	1.0018	0.818	0.844
SEK/USD	2.5633	1.0008	0.084 *	0.534	1.0056		0.850	0.9778	*** 0000	*** 900.0	0.9975	0.034 **	0.134
DNK/USD	2.5355	0.9999	0.153	0.493	0.9988	*	0.251	0.9853	0.000	0.001 ***	0.9973	0.027 **	0.081 *
USD/AUD	2.7434	0.9965	0.044 **	0.405	1.0049		0.876	0.9946	0.003 ***	0.031 **	0.9984	0.184	0.279
FRF/USD	2.6411	1.0055	0.315	0.635	1.0003	0.212 0.	0.510	9626.0	0.001 ***	0.041 **	0.9932	0.018 **	0.140
DEM/USD	2.7545	1.0084	0.989	0.879	1.0007	0.678 0.	0.673	6996.0	0.000 ***	*** 600.0	0.9931	0.006 ***	0.070 *
ITL/USD	2.6624	1.0019	0.211	0.544	1.0011	0.229 0.	0.527	0.9720	0.000 ***	0.023 **	0.9891	0.011 **	0.142
NLG/USD	2.7517	1.0076	0.999	0.976	1.0016	0.904 0.	0.846	96260	0.001 ***	0.055 *	0.9932	0.010 **	0.076 *
PTE/USD	2.6956	0.9884	0.012 **	0.405	0.9843	0.010 ** 0.	0.369	0.9787	0.008 ***	0.328	0.9607	0.001 ***	0.184
						UIRP fi	UIRP fundamental						
USD/GBP	2.4410	1.0137	0.251	0.753	1.0007	0.238 0.	0.525	0.9651	0.000	* 680.0	0.9923	0.013 **	0.085 *
dsu/ydr	2.7039	1.0072	0.691	0.800	1.0025		0.704	0.9857	0.002 ***	0.183	0.9923	0.013 **	0.083 *
CHF/USD	2.8195	1.0005	0.276	0.554	1.0005		909.0	0.9824	0.000	0.013 **	0.9952	0.005 ***	0.057 *
CAD/USD	1.5091	1.0032	0.271	0.611	1.0007		0.726	0.9919	0.003 ***	0.038 **	69660	0.019 **	0.163
SEK/USD	2.5626	1.0280	0.933	0.977	1.0045		0.767	0.9561	*** 0000	0.026 **	0.9855	0.005 ***	0.098 *
DNK/USD	2.5355	1.0125	0.777	0.888	0.9998		0.485	0.9933	0.023 **	0.226	9966.0	0.033 **	0.091 *
USD/AUD	2.7414	1.0105	0.316	0.732	1.0075		0.825	0.9650	0.005 ***	0.108	0.9917	0.016 **	0.085 *
EBE/LISD	2 6411	1 0066	0.749	0.772	1 0045		0.677	0.9715	*** 0000	** 2000	0.9937	0.047 **	0.113
DEM/USD	2 7545	1 0047	0.762	0.767	1 0028		0.672	0.9760	*** 0000	***	0.9904	***	0.032 **
TI /USD	2,6624	1.0057	0.406	0.637	0.9982		0.460	0.9416	0.001	0.051 *	0.9694	***	0.066 *
NIG/USD	2.7517	1.0043	0.380	0.641	1,0007		0.527	0.9754	*** 0000	0.050 **	0.9940	0.025 **	0.063 *
PTE/USD	2.2481	0.9848	* 080.0	0.206	0.9920		0.375	0.9839	0.142	0.348	72.0	0.076 *	0.196
						Jonetary	model fundamentals	tals					
USD/GRP	2,4410	1,0200	0.706	0.874	1,0007		0.526	0.9802	0.004 ***	0.190	0.9919	0.010 ***	0.220
JEV/USD	2.7042	0.9982	0.003 ***	0.451	1.0030	*	0.585	9708	*** 0000	0.025 **	0.9794	*** 0000	0.035 **
CHE/LISD	2 8377	1 0042	0.023	0.644	1 0052		0.665	0.9833	*** 0000	* 250.0	1776.0	*** 0000	** 9800
CAD/USD	1.5083	1.0254	0.344	0.776	1.0028		0.688	0.9794	0.003	0.053 *	0.9978	0.193	0.429
SFK/HSD	2 5626	1 0311	7 8 0	67.60	1 0046		0.820	0.9402	*** 0000	0.056 *	7679.0	*** 000 0	0.014 **
DNK/USD	2.5553	1.0164	0.923	0.961	1.0084		0.941	0.9794	*** 0000	0.011 **	0.9873	*** 0000	0.123
USD/AUD	2.7547	1.0185	0.559	0.782	1.0033		0.858	0.9720	0.001 ***	* 560.0	0.9930	0.007	0.243
FRF/USD	2.5229	1.0536	0.915	0.965	1.0578		0.973	1.0120	0.056 *	0.663	0.9998	0.132	0.495
DEM/USD	2.7624	1.0353	0.793	0.980	1.0127	0.750 0.	0.883	0.9660	0.000 ***	* 6.00	0.9558	*** 0000	0.007 ***
ITL/USD	2.6458	1.0114	0.291	0.671	1.0069	0.440 0.	0.667	0.9523	0.004 ***	0.076 *	0.9730	0.004 ***	0.046 **
NLG/USD	2.7517	1.0010	0.276	0.557	1.0020	0.518 0.	0.665	0.9826	0.001 ***	0.024 **	1.0052	0.230	0.708
PTE/USD	2.2342	0.9800	0.066 *	0.347	0.9862	0.102 0.	0.368	0.9445	0.010 **	0.162	0.9667	0.021 **	0.209
						All fur	fundamentals						
USD/GBP	2.4410	1.0550	0.823	0.995	1.0169	0.668 0.	0.954	0.9737	0.001 ***	0.136	0.9823	0.000 ***	0.068 *
JPY/USD	2.7042	1.0143	0.004 ***	0.718	1.0079	0.085 * 0.	0.732	6896.0	0.000 ***	0.032 **	99260	0.000 ***	0.018 **
CHF/USD	2.8377	1.0211	0.320	0.923	1.0096	0.129 0.	0.753	0.9814	0.000 ***	0.076 *	0.9750	*** 000.0	0.029 **
CAD/USD	1.5083	1.0137	0.032 **	0.645	1.0059	0.677 0.	0.856	0.9804	0.003 ***	.0029	0.9988	0.255	0.463
SEK/USD	2.5626	1.1397	0.846	0.902	1.0747	0.839 0.	0.843	0.9619	0.000 ***	0.026 **	0.9924	0.117	0.358
DNK/USD	2.5553	1.0447	0.323	0.945	1.0104		0.889	79767	0.000 ***	*** 900.0	0.9793	0.000 ***	0.015 **
USD/AUD	2.7547	1.0279	* 680.0	0.821	1.0146	0.796 0.	0.924	8796.0	0.000 ***	0.061 *	0.9898	0.003 ***	0.189
FRF/USD	2.5229	1.0935	0.623	0.934	1.0637		0.873	1.0134	0.039 **	0.648	0.9903	0.040 **	0.321
DEM/USD	2.7624	1.0608	0.962	0.993	1.0232		0.922	0.9589	0.000 ***	0.029 **	0.9539	*** 0000	0.004 ***
ITL/USD	2.6458	1.0378	0.226	0.794	1.0381		0.841	0.9339	0.002 ***	0.057 *	0.9625	0.003 ***	0.046 **
NLG/USD	2.7517	1.0172	0.373	0.762	1.0053		0.659	0.9753	*** 0000	0.018 **	1.0051	0.063 *	0.703
P1E/USU	2.2342	0.9855	0.068	0.362	1.0377	0.845 0.	0.909	0.9590	0.031 **	0.134	0.9396	0.004 ***	0.103

Table 14: RMSE and Theil ratios of forecasts using coupled fundamentals for the end-of-month exchange rates sample.

			: : : : : : : : : : : : : : : : : : :				1		7:00			8747.	
Currency pair	No change		Kolling regression			Recursive regression	lion President	i	Skidge		Ī	EWA	
	KIVISE X 100	ineli ratio	cw p-value	DINI p-value	I nell ratio	cw p-value	DIVI p-value	Inell ratio	cw p-value	DIVI p-value	I nell ratio	Cw p-value	DINI p-value
440/401	7.00	0.00			2000		o o a a	7000		0.70	1000		000
USD/GBP	2.9051	1.0050	0.913	0.965	1.0032	0.937	0.917	1.0006	0.525	0.618	1.0005	0.531	0.609
USD/YAL	3.1U23	1.0037	0.725	0.826	1.0010	0.994	0.976	0.9999	0.407	0.437	0.9973	0.060	0.143
287,030	3.3420	1.002/	* 0800	0.047	1.0003	0.00	0.017	1.0000	0.020	0.903	0.000	0.278	0.302
CEK/USD	2 2 2 2 9 0 0	1 0015	0.082	0.280	1,000	0.21/	0.321	1,0000	0.516	0.330	1,0000	0.338	0.991
JEK) O3D DNIK/LISD	3.2230	1,0001	0.176	0.533	1.0003	0.754	0.847	0.000.1	0.087	0.782	0000.T	0.114	0.453
USD/AUD	3.3786	0.9976	0.086 *	0.405	1.0009	0.827	0.866	1.0006	0.905	0.852	1.0009	0.583	0.644
ERE/LISD	3 1978	1 0072	0.386	0.749	1.0035	0.326	0.538	1 0005	0.308	0.527	0.9981	0.176	0.319
DEM/IISD	3.37.8	1.0072	0.976	0.891	1.0035	0.340	0.014	1,0000	0.308	0.527	1966.0	0.170	0.319
ITL/USD	3.1907	1.0026	0.297	0.589	1.0019	0.312	0.566	1.0011	0.296	0.547	0.9959	0.073 *	0.287
NIG/IISD	3 3319	1 0056	0 993	0 984	1 0023	996 0	0.928	1 0000	0.971	0 940	9266 0	0.115	0.172
PTE/USD	3.2207	0.9961	0.028 **	0.449	0.9936	0.023 **	0.406	0.9925	0.023 **	0.397	0.9821	0.010 ***	0.265
							UIRP fundamental						
USD/GRP	2 9051	1 0132	0 344	0.813	1 0012	0.323	0.562	1 0045	0 338	0.634	9266	0.189	0.310
DSU/YQL	3.1023	1.0062	0.707	0.816	1.0020	0.604	0.736	1.0012	0.919	0.816	0.9993	0.277	0.400
CHF/USD	3.3747	1.0009	0.387	0,641	1.0005	0.709	0.725	1,0000	0.709	0.691	0666.0	0.191	0.292
CAD/USD	1.9982	1.0027	0.331	0.667	1.0006	0.749	0.747	1.0001	0.894	0.863	1.0020	0.863	0.923
SEK/USD	3.2263	1.0242	0.912	0.975	1.0037	0.602	0.732	0.9907	0.077 *	0.211	0.9868	0.071 *	0.172
DNK/USD	3.1200	1.0119	0.912	0.931	1.0006	0.578	0.620	1.0002	0.847	0.839	1.0012	0.531	0.677
USD/AUD	3.3793	1.0097	0.471	0.793	1.0067	0.740	0.835	1.0075	0.763	0.826	0.9999	0.328	0.495
FRF/USD	3.1978	1.0085	806.0	0.858	1.0039	0.685	0.731	1.0019	0.413	0.621	0.9984	0.195	0.306
DEM/USD	3.3136	1.0036	0.849	0.866	1.0026	0.778	0.753	1.0001	0.841	0.841	0.9962	* 690.0	0.149
ITL/USD	3.1907	1.0060	0.488	0.680	9666.0	0.211	0.488	1.0055	0.487	0.661	0.9869	0.023 **	0.151
NLG/USD	3.3319	1.0044	0.641	0.757	1.0014	0.658	0.701	1.0022	0.281	0.594	0.9978	0.140	0.234
PTE/USD	2.8024	0.9874	0.117	0.223	0.9917	0.196	0.351	0.9918	0.273	0.347	0.9899	0.212	0.288
				•		Moneta	Monetary model fundamentals	ntals		•			
USD/GBP	2.9051	1.0201	0.853	0.953	1.0021	0.486	0.653	1.0050	0.309	969.0	1.0034	0.309	0.664
JPY/USD	3.1046	1.0042	0.071 *	0.669	1.0042	0.220	0.687	1.0065	0.523	0.780	0.9965	0.029 **	0.348
CHF/USD	3.3887	1.0029	0.104	0.622	1.0063	0.152	0.723	1.0015	0.204	0.590	1.0005	0.046 **	0.520
CAD/USD	1.9994	1.0254	0.591	0.822	1.0032	0.697	0.751	1.0065	0.839	0.849	1.0074	0.939	0.911
SEK/USD	3.2263	1.0283	0.956	0.981	1.0047	0.725	0.843	1.0138	0.424	0.841	1.0040	0.357	0.735
DNK/USD	3.1675	1.0151	0.805	0.948	1.0059	0.923	0.870	1.0008	0.796	0.800	1.0086	0.961	0.930
USD/AUD	3.3926	1.0180	0.692	0.845	1.0044	0.891	0.939	1.0008	0.683	0.673	1.0016	0.136	0.565
FRF/USD PFM4/15D	3.1041	1.0460	0.91/	0.983	1.0439	0.929	0.979	1.0280	0.952	0.982	1.011/	0.524	0.772
DEINI/USD ITI / IED	5.3260	1.0312	0.935	0.994	1.0169	0.970	0.965	1,0001	0.000	0.360	0.9827	0.024	0.133
NI G/LISD	3 3 3 1 9	1500.1	0.238	0.602	T.0030	0.443	0.872	0.9976	0.204	0.380	1 0022	0.173	0.529
PTE/USD	2.7703	0.9864	0.123	0.332	0.9965	0.247	0.414	1.0045	0.635	0.696	0.9991	0.285	0.487
							All fundamentals						
USD/GBP	2.9051	1.0515	0.895	666.0	1.0154	0.787	0.980	1.0000	0.188	0.500	0.9977	0.078 *	0.402
JPY/USD	3.1046	1.0190	* 0.000	0.844	1.0072	0.285	0.812	1.0039	0.522	0.739	0.9951	0.020 **	0.285
CHF/USD	3.3887	1.0190	0.429	0.958	1.0100	0.194	0.811	1.0020	0.334	0.643	0.9975	0.029 **	0.402
CAD/USD	1.9994	1.0274	0.183	0.860	1.0043	0.560	0.802	1.0060	0.822	0.831	1.0072	0.958	0.943
SEK/USD	3.2263	1.1676	0.902	0.902	1.0753	0.860	0.859	1.0060	0.847	0.847	0.9852	0.021 **	0.169
DNK/USD	3.1675	1.0307	0.439	0.977	1.0081	0.945	0.948	1.0005	0.780	0.788	1.0051	0.994	0.989
USD/AUD	3.3926	1.0310	0.266	0.913	1.0138	0.846	0.939	1.0021	0.835	0.840	1.0014	0.121	0.555
FRF/USD	3.1041	1.0761	0.813	0.978	1.0516	0.781	0.926	1.0291	0.870	0.950	1.0061	0.382	0.649
DEM/USD	3.3286	1.0501	0.984	0.996	1.0265	0.970	0.983	0.9995	0.193	0.384	0.9815	0.019 **	0.101
ITL/USD	3.2797	1.0226	0.212	0.768	1.0306	0.753	0.889	0.9988	0.332	0.455	0.9955	0.184	0.363
NLG/USD	3.3319	1.0125	0.255	0.726	1.0025	0.255	0.581	0.9939	0.025 **	0.179	0.9978	0.108	0.201
FIE/USD	2.7703	1.01/0	0.514	0.732	T.0430	0.977	776.0	1.0023	0.424	0.343	0.5550	0.202	0.403

Table 15: RMSE and Theil ratios of forecasts using Taylor-rule fundamentals for the average exchange rates sample.

			acinomor pailled			Signature Children	-		CDidas			V/V/2	Ī
Currency pair	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
					Outp	Output gap fundamentals: deviations	als: deviations fr	from a linear trend					
USD/GBP	2.4410	1.0588	0.110	066.0	1.0313	0.332	0.921	0.9751	0.001 ***	0.046 **	0.9798	*** 000.0	0.034 **
JPY/USD CHF/USD	2.7042	1.0262	0.013 **	0.995	1.0026	0.006 ***	0.549	0.9700	*** 000.0	0.001 ***	0.9758	*** 000.0	0.002 ***
CAD/USD	1.5083	0.9926	0.001 ***	0.319	0.9909	0.019 **	0.295	0.9816	0.002 ***	0.042 **	0.9939	0.001 ***	0.022 **
SEK/USD	2.5626	1.1004	0.549	0.908	1.0309	0.434	0.969	0.9419	*** 000.0	0.029 **	0.9786	*** 000.0	0.002 ***
DNK/USD	2.4891	1.0261	0.073 *	0.847	1.0168	0.284	0.863	0.9755	*** 000.0	0.001 ***	0.9730	*** 000.0	0.002 ***
USD/AUD	2.7547	0.9970	*** 0000	0.451	1.0145	0.407	0.817	0.9731	0.001 ***	0.033 **	0.9851	0.001 ***	0.024 **
FRF/USD	2.6411	1.0321	0.101	0.807	1.0236	0.182	0.812	0.9684	*** 0000	0.002 ***	0.9712	* * * 0000	0.034 **
DEM/USD ITL/USD	2./545 2.6624	1.0032	0.334	0.539	1.0250	0.573	0.843	0.9686	*** 0000	0.001 ***	0.9803	0.002 ***	0.037 **
NLG/USD	2.7517	1.0335	* 090.0	0.783	1.0141	0.195	0.736	0.9662	*** 000:0	*** 0000	0.9735	*** 0000	*** 900.0
PTE/USD	2.2342	1.0337	0.453	0.838	1.0922	0.604	0.941	0.9637	0.042 **	0.132	0.9857	0.043 **	0.148
					Output	Output gap fundamentals: deviations from a quadratic trend	s: deviations froi	n a quadratic tr	end				
USD/GBP	2.4410	1.0516	0.108	0.983	1.0349	0.471	0.951	0.9772	0.001 ***	* 990.0	0.9799	0.000 ***	0.034 **
JPY/USD	2.7042	1.0191	0.005 ***	0.744	1.0040	0.012 **	0.574	0.9702	0.000 ***	0.001 ***	0.9799	*** 0000	0.002 ***
CHF/USD	2.8377	1.0703	0.133	0.993	1.0230	0.498	0.924	09260	0.000 ***	0.004 ***	0.9773	*** 0000	0.001 ***
CAD/USD	1.5083	0.9961	0.002 ***	0.438	9066.0	0.018 **	0.245	0.9831	0.007 ***	* 260.0	0.9921	*** 0000	0.014 **
SEK/USD	2.5626	1.0912	0.577	0.882	1.0399	0.706	0.951	0.9405	*** 000.0	0.027 **	0.9778	*** 0000	0.002 ***
DNK/USD	2.4891	1.0216	0.026 **	0.840	1.0170	0.703	0.961	0.9749	*** 000.0	0.001 ***	0.9726	*** 0000	0.001 ***
USD/AUD	2.7547	0.9967	0.000	0.449	1.0110	0.163	0.745	0.9704	0.000	0.028 **	0.9844	0.001 ***	0.021 **
FRF/USD DEM/LISD	2.6411 2.7545	1.0283	0.085	0.787	1.0233	0.217	0.813	0.9684	*** 0000	0.002 ***	0.9698	0.000	0.027 **
TL/USD	2.6624	1.0005	0.018 **	0.506	1.0176	0.582	0.811	0.9441	*** 0000	0.014 **	0.9882	0.022 **	0.200
NLG/USD	2.7517	1.0407	0.062 *	0.826	1.0066	0.214	0.650	0.9662	0.000 ***	*** 000.0	0.9746	*** 000.0	0.006
PTE/USD	2.2342	1.0104	0.246	0.618	1.0760	0.617	0.970	0.9637	0.043 **	0.133	0.9848	0.036 **	0.127
					Output gap fu	Output gap fundamentals: deviations from a linear and a quadratic trend	ations from a lin	ear and a quad	atic trend				
USD/GBP	2.4410	1.0450	0.197	926.0	1.0296	0.708	0.967	0.9759	0.000 ***	0.051 *	0.9794	*** 0000	0.032 **
JPY/USD	2.7042	1.0232	0.007 ***	0.831	1.0013	*** 800.0	0.529	0.9728	*** 000.0	0.002 ***	0.9760	*** 0000	0.003 ***
CHF/USD	2.8377	1.0491	0.070 *	0.972	1.0170	0.361	0.850	0.9761	0.000	0.004 ***	0.9776	* * * 0000	0.001 ***
CAD/USD SEV/ISD	1.5083	1.0757	0.000	0.332	0.9905	0.022 **	0.252	0.9891	0.0IO **	0.072 **	0.9924	* * * 0000	0.009 ***
SEN/USD DNK/USD	2.4891	1.0240	0.032 **	0.876	1.0146	0.697	0.941	0.9754	*** 000.0	0.002 ***	0.9734	*** 000.0	0.002 ***
USD/AUD	2.7547	0.9918	*** 0000	0.376	1.0058	0.129	0.631	0.9708	*** 0000	0.028 **	0.9836	0.001 ***	0.017 **
FRF/USD	2.6411	1.0303	0.029 **	0.757	1.0218	0.213	0.797	0.9684	0.000 ***	0.002 ***	0.9720	0.001 ***	0.039 **
DEM/USD	2.7545	1.0475	0.427	0.915	1.0098	0.174	0.644	0.9684	*** 000.0	0.001 ***	0.9794	0.002 ***	0.030 **
ITL/USD	2.6624	1.0049	0.011 **	0.553	1.0123	0.350	0.712	0.9442	*** 0000	0.014 **	0.9885	0.024 **	0.204
NLG/USD PTE/USD	2.7517	1.0351	0.080 *	0.804	1.0067	0.209	0.644	0.9662	0.000 ***	0.000 ***	0.9735	0.000 ***	0.006 ***
1			!		æ	als: devi	ations from a Hodrick-Prescott	drick-Prescott fil	tered trend				
USD/GBP	2.4410	1.0527	0.354	726.0	1.0217	0.522	0.888	0.9753	0.001 ***	0.047 **	0.9794	*** 0000	0.031 **
JPY/USD	2.7042	1.0231	0.040 **	0.815	1.0013	0.003 ***	0.526	0.9733	*** 0000	0.002 ***	9776	*** 0000	0.005 ***
CHF/USD	2.8377	1.0554	0.381	0.990	1.0231	0.555	0.910	0.9764	0.000 ***	0.005 ***	0.9790	*** 0000	0.002 ***
CAD/USD	1.5083	0.9812	0.001 ***	0.201	1.0022	0.377	0.671	0.9863	0.001 ***	0.054 *	0.9941	*** 0000	0.010 **
SEK/USD	2.5626	1.0878	0.587	696.0	1.0399	0.727	0.952	0.9423	*** 000.0	0.030 **	0.9783	*** 000.0	0.002 ***
DNK/USD	2.4891	1.0335	0.152	0.943	1.0098	0.475	0.817	0.9764	*** 000.0	0.002 ***	0.9730	*** 000.0	0.001 ***
USD/AUD	2.7547	0.9897	*** 0000	0.376	1.0146	0.440	0.814	0.9727	* * * 000.0	0.031 **	0.9833	0.001 **	0.016 **
FRF/USD DEM / ISD	2.5411	1.0536	0.068	0.796	1.0264	0.327	0.835	0.9685	***************************************	0.002	0.9712	***	0.034 ::
DEIMI/USD ITL/USD	2.7343	0.9990	0.013 **	0.488	1.0148	0.494	0.765	0.9442	** 000.0	0.001	0.9880	0.002	0.030
NLG/USD	2.7517	1.0394	0.127	0.838	1.0123	0.399	0.746	0.9663	0.000	*** 0000	0.9733	*** 0000	0.005 ***
PTE/USD	2.2342	1.0049	* 260.0	0.535	1.0609	0.579	0.924	0.9638	0.043 **	0.132	0.9860	0.046 **	0.154

Table 16: RMSE and Theil ratios of forecasts using Taylor-rule fundamentals for the end-of-month exchange rates sample.

			a dinagraphy a dillo		200	acinocanos cuinanto o	2		00000			67470	
Currency pair	No change RMSF x 100	Theil ratio	CW n-value	DM n-value	net Theil ratio	CW n-value	DM n-value	Theil ratio	CW n-value	oM n-value	Theil ratio	CW n-value	OM n-value
					Output	gap fundamenta	Output gap fundamentals: deviations from a linear trend	ım a linear tren				<u> </u>	L
USD/GBP	2.9051	1.0430	0.192	0.984	1.0276	0.566	0.953	1.0011	0.493	0.607	1.0030	0.185	0.622
JPY/USD	3.1046	1.0297	* 860.0	0.915	1.0048	0.025 **	0.624	9666.0	0.204	0.228	0.9937	0.016 **	0.187
CHF/USD	3.3887	1.0644	0.535	766.0	1.0216	0.508	0.941	0.9986	* 260.0	0.141	0.9941	0.028 **	0.143
CAD/USD	1.9994	1.0145	0.287	0.782	0.9988	0.137	0.455	1.0002	0.640	0.646	0.9995	0.201	0.315
SEK/USD	3.2263	1.1008	0.837	0.965	1.0241	0.650	0.992	0.9971	0.097 *	0.387	0.9954	0.012 **	0.211
DNK/USD	3.1037	1.0323	0.265	0.937	1.0136	0.361	0.890	1.0023	0.895	0.894	0.9939	0.028 **	0.219
USD/AUD	3.3926	1.0073	0.007	0.642	1.0126	0.533	0.852	1.0043	0.752	0.820	1.0009	0.375	0.562
FRF/USD	3.1978	1.0359	0.430	0.915	1.0277	0.510	0.909	1.0005	0.443	0.539	0.9961	0.085 *	0.382
DEM/USD	3.3136	1.0444	0.658	0.938	1.0303	0.612	0.931	0.9987	0.213	0.287	0.9914	0.042 **	0.210
IIL/USD	3.1907	1.0082	0.116	0.631	1.0216	0.814	0.944	1.0024	0.339	0.578	1.0067	0.276	0.665
NLG/USD PTE/1SD	3.3319	1.0341	0.256	0.870	1.0212	0.551	0.900	0.9998	0.394	0.444	0.9883	0.022 **	0.143
r 1 L/ 03D	5.7.703	1:0000	0.714	0.318	Output g	3p fundamentals	Output gap fundamentals: deviations from a quadratic trend	ם a quadratic tre	pu;	0.410	10000	0.00	0.44.0
USD/GBP	2.9051	1.0387	0.157	0.973	1.0298	0.681	0.975	1.0010	0.482	0.597	1.0027	0.155	0.610
JPY/USD	3.1046	1.0238	0.056 *	0.870	1.0063	0.048 **	0.664	0.9995	0.156	0.177	0.9943	0.022 **	0.214
CHF/USD	3.3887	1.0601	0.333	0.989	1.0223	0.713	0.973	0.9986	* 660.0	0.143	0.9927	0.016 **	0.097 *
CAD/USD	1.9994	1.0184	0.246	0.799	0.9991	0.172	0.465	1.0002	0.636	0.643	0.9993	0.136	0.216
SEK/USD	3.2263	1.0756	0.757	0.943	1.0314	0.853	0.982	9866.0	0.132	0.445	0.9951	0.010 **	0.198
DNK/USD	3.1037	1.0270	0.152	0.870	1.0154	0.723	696.0	1.0005	0.892	0.847	0.9933	0.024 **	0.198
USD/AUD	3.3926	1.0063	0.003 ***	0.615	1.0096	0.242	0.782	1.0043	0.747	0.817	1.0012	0.429	0.590
FRF/USD	3.1978	1.0333	0.406	606:0	1.0279	0.566	0.915	1.0005	0.444	0.539	0.9964	* 280.0	0.390
DEM/USD	3.3136	1.0456	0.674	0.949	1.0219	0.599	0.891	0.9987	0.213	0.286	8066.0	0.037 **	0.192
ITL/USD	3.1907	1.0029	* 680.0	0.549	1.0206	0.851	0.938	1.0024	0.339	0.578	1.0047	0.238	0.619
NLG/USD	3.3319	1.0356	0.272	0.883	1.0146	0.676	0.886	0.9998	0.395	0.445	0.9883	0.022 **	0.144
PIE/USD	2.7703	1,0001	0.000	0.022	Dutant gas fus	U.736	1.0040 0.730 0.300 0.300 0.300 0.300 0.300 0.300 0.300	0.9970	0.300	0.413	0.9960	0.331	0.429
UISD/GRD	2 9051	1 0353	0.254	0200	1 0262	O 846	0 987	1 0010	0.488	0 602	1 001/	0120	0.557
IBV/ISD	3 1046	1.0303	0.234	0.57.0	1.0202	0.040	0.587	96660	0.207	0.331	0.9943	0.120	0.337
CHF/USD	3.3887	1.0461	0.257	0.968	1.0178	0.557	0.923	0.9986	* 660.0	0.144	0.9930	0.018 **	0.104
CAD/USD	1.9994	1.0167	0.211	0.777	0.9982	0.177	0.430	1.0002	0.643	0.648	0.9990	0.074 *	0.163
SEK/USD	3.2263	1.0638	0.837	0.988	1.0272	0.868	0.991	0.9982	0.125	0.429	0.9958	0.014 **	0.231
DNK/USD	3.1037	1.0298	0.167	0.895	1.0136	0.627	0.935	1.0004	0.882	0.835	0.9926	0.020 **	0.175
USD/AUD	3.3926	1.0029	0.003 ***	0.553	1.0052	0.181	0.663	1.0042	0.741	0.813	1.0021	0.472	0.638
FRF/USD	3.1978	1.0363	0.243	0.892	1.0263	0.570	0.903	1.0005	0.443	0.539	0.9951	0.073 *	0.354
DEM/USD	3.3136	1.0431	0.665	0.944	1.0174	0.371	0.796	0.9987	0.213	0.287	0.9905	0.035 **	0.185
ITL/USD	3.1907	1.0160	0.118	0.718	1.0182	0.732	0.883	1.0025	0.344	0.581	1.0075	0.294	0.684
NLG/USD PTE/USD	5.3319	1.0489	0.587	0.880	1.0811	0.727	0.961	0.9970	0.365	0.445	0.9987	0.339	0.426
					Output gap fund	amentals: deviat	Output gap fundamentals: deviations from a Hodrick-Prescott filtered trend	rick-Prescott filt	ered trend				
USD/GBP	2.9051	1.0427	0.555	0.988	1.0203	0.750	0.951	1.0011	0.495	609.0	1.0026	0.153	0.606
JPY/USD	3.1046	1.0233	0.150	0.901	1.0034	0.015 **	0.590	9666.0	0.175	0.196	0.9944	0.022 **	0.219
CHF/USD	3.3887	1.0504	0.673	866.0	1.0220	0.757	0.961	9866.0	0.094 *	0.137	0.9938	0.025 **	0.134
CAD/USD	1.9994	1.0167	0.274	0.890	1.0041	0.650	0.826	1.0002	0.643	0.648	0.9992	0.078 *	0.168
SEK/USD	3.2263	1.0679	908.0	0.987	1.0300	0.841	0.981	0.9972	* 660.0	0.388	0.9954	0.012 **	0.210
DNK/USD	3.1037	1.0322	0.343	0.977	1.0086	0.447	0.805	1.0004	0.881	0.832	0.9928	0.022 **	0.181
USD/AUD	3.3926	1.0041	0.007 ***	0.570	1.0121	0.520	0.841	1.0043	0.747	0.817	1.0019	0.468	0.633
FRF/USD	3.1978	1.0361	0.374	0.922	1.0283	0.638	0.913	1.0005	0.444	0.539	0.9951	0.072 *	0.354
DEM/USD	3.3136	1.0482	0.835	0.980	1.0239	0.704	0.895	0.9987	0.214	0.287	0.9913	0.041 **	0.207
11L/USD	3.1907	1.0190	0.117	0.752	1.0218	0.843	0.926	1.0024	0.340	0.578	1.0077	0.296	0.687
NLG/USD PTE/11SD	3.3319	1.0381	0.393	0.909	1.0174	0.754	0.895	0.9998	0.396 0.365	0.445	0.9889	0.025 **	0.155
7 1 L/ U3D	2.7.103	T.040.0	0.000	0.011	T.U.S.O.	0./.0	0.500	0.2000	0.300	0.410	0.5301	0.540	0.44.7

Table 17: RMSE and Theil ratios of the forecasts for a shorter-period sample (1980 onwards, average exchange rates sample).

Currency pair	No change		Rolling regression			Recursive regression	ion		SRidge			EWA	
	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
							PPP tundamental						
USD/GBP	2.2980	0.9955	0.063 *	0.386	0.9978	0.031 **	0.450	0.9913	0.007 ***	0.091 *	0.9905	0.020 **	0.094 *
JPY/USD	2.6567	1.0123	0.616 0.72r	0.832	1.0033	0.689	0.733	0.9822	0.001 ***	0.127	0.9893	0.001 ***	0.01/ **
CAF/USD	2.0808	1.0141	0.725	0.8/1	1.0048	0.522	0.093	0.9846	0.002 ***	0.044 *	0.9903	0.004	0.016
CAD/USD SEK/LISD	7,6802	0.9645	0.001	0.129	0.9930	** 0000	0.00/	0.9962	*** OOO O	** 170	0.9930	0.154 0.000 ***	0.242
JEN/O3D DNIK/HSD	2.0802	1.0073	0,639	0.071	1 0011	0.022	0.533	0.9381	*** 0000	** 0000	0.9853	** 000.0	0.017
050 (XIVI)	2.4383	0.9922	0.028	0.879	1.0011	0.362	0.372	0.9612	0.000	* 00.0	0.9608	0.000 * TIOO	* 0.000
CDE//15D	2.07.38	1.0162	0.014	0.338	1 0076	0.000	0.183	20960	0.000	0.030	9066.0	* 690.0	0.070
DEM/USD	2.4003	1.0215	0.021	0.783	1,0070	0.976	0.931	0.9651	***	0.023	0.9900	0.003	0.103
DEIM) 03D	2.3381	1.0313	0.383	0.307	0.9995	0.873	0.631	0.9377	0.003	0.045	0.9755	0.032	* 8900
NIG/ISD	2 5501	1 01 82	0.450	9000	0000	0.200	0.400	0.550	*** 9000	* 9900	02.600	** 0000	* 1000
NLG/USD PTE/HSD	2.5501	1.0182	0.906	0.906	1.0229	0.790	0.708	0.9670	0.006 ***	0.066	0.9870	0.038 ::	0.087
750/31	2:326.3	T:0253	0.01	0.00	1:0110		UIRP fundamental	6666	0.010	0:01	10.000	0.020	0010
da5/USI I	7 2080	1 0/30	777 0	0.018	1 0387	0770	0 00	72900	* 7700	0.318	77800	0.01¤ **	0.103
DY/USD	2.6567	1.0430	0.250	0.918	1.0084	0.740	0.668	0.9813	0.003	0.035 **	0.9847	0.005 ***	0.046 **
CHEALISD	2 6868	1 0255	0.218	0.816	1 0273	0.214	0.808	0.9855	***	** 860 0	0.9858	0.000	0.035 **
CAD/IISD	1 6306	1.0233	0.599	0.834	1.0273	0.559	0.821	0.9841	0.001	* 890.0	1 0001	0.001	0.525
SEK/USD	2.6797	1.0600	0.904	0.943	1.0350	0.767	0.840	0.9463	*** 0000	0.041 **	0.9748	0.001 ***	0.064 *
DNK/USD	2.4385	1.0258	0.839	0.935	1.0131	0.636	0.841	0.9688	0.001 ***	0.021 **	0.9804	0.001 ***	0.010 **
UISD/AUD	2 6728	1 0112	0.174	0.694	9666	0.262	0.463	0.9572	0.000	* 850 0	0.9915	0.032 **	0.138
FRE/USD	2 4663	1 0247	0.856	0.912	1 0333	0.697	0.841	0.9813	***	0.053 *	0.9735	0.012 **	* 8800
DEM/IISD	2 5581	1.0450	0.652	0.924	1.0331	0.681	0.242	0.9670	** 7000	0.033	77700	** 0100	* 9000
ITI /IISD	2.3381	1.0450	0.303	0.606	1.0331	0.031	0.876	0.9582	***	0.032	0.9450	0.010	* 1200
NI 6/115P	2.7032	1,1305	2000	920.0	1.0342	7.6.0	0.070	20000	****	* 010.0	0.5130	**	* 7900
NEG/USD PTE/LISD	7.3301	1.1303	0.030	0.070	T.0939	0.002	U.0/1	6/96/9	0.004	60.0	0.9744	0.011	0.004
r1E/03D						Monetar	J Monetary model fundamentals	ntals					
USD/GBP	2.2980	1.0330	0.552	0.962	1.0183	0.558	0.895	0.9627	0.003 ***	0.143	0.9888	*** 600.0	0.210
dsu/yqr	2.6570	1.0118	0.004 ***	0.741	1.0170	0.093 *	0.837	0.9713	*** 0000	* 690.0	0.9804	0.000	0,077 *
CHF/USD	2.7041	1.0246	0.314	0.857	1.0086	0.317	0.761	0.9964	0.013 **	0.425	0.9913	0.004 ***	0,284
CAD/USD	1,6300	1.0381	0.109	0.876	1.0063	0.322	0.682	0.9741	0.002 ***	0.037 **	9966.0	0.031 **	0,399
SEK/USD	2.6797	1.0502	0.883	0.995	1.0257	0.900	0.967	0.9402	*** 0000	0.072 *	0.9759	*** 0000	0.017 **
DNK/USD	2.4619	1.0410	0.970	0.994	1.0094	0.926	0.943	0.9814	0.002 ***	0.110	0.9768	*** 0000	* 650.0
USD/AUD	2.6871	1.0240	0.227	0.854	1.0113	0.856	0.932	0.9660	0.001 ***	* 990.0	0.9785	*** 0000	0.118
FRF/USD	2.4663	1.0413	0.517	0.897	1.0581	0.662	0.935	0.9784	*** 600.0	0.154	0.9879	0.035 **	0.272
DEM/USD	2.5581	1.0305	0.288	0.823	1.0207	0.234	0.745	0.9652	*** 900.0	0.167	0.9493	0.001 ***	0.020 **
ITL/USD	2.7052	1.0258	0.423	0.763	1.0158	0.366	0.703	0.9259	0.003 ***	0.083 *	0.9725	0.010 ***	0.065 *
NLG/USD PTE/11SD	2.5501	1.0494	0.856	0.925	1.0065	0.514	0.655	0.9640	0.001 ***	0.048 **	93260	0.011 **	0.076 *
250/311							All fundamentals						
USD/GBP	2.2980	1.0849	0.168	0.990	1.0479	0.502	0.931	0.9594	0.002 ***	* 290.0	0.9835	0.005 ***	0.144
JPY/USD	2.6570	1.0558	0.036 **	0.970	1.0336	* 690.0	0.904	0.9665	*** 0000	0.018 **	0.9746	0.000 ***	0.039 **
CHF/USD	2.7041	1.0965	0.210	0.989	1.0275	0.071 *	0.837	9066.0	0.005 ***	0.277	0.9891	0.003 ***	0.244
CAD/USD	1.6300	1.0026	0.001 ***	0.532	1.0059	0.075 *	0.605	0.9728	0.002 ***	0.039 **	0.9959	0.049 **	0.390
SEK/USD	2.6797	1.1499	0.721	0.921	1.1129	0.798	0.862	0.9578	*** 0000	0.015 **	0.9945	0.151	0.415
DNK/USD	2.4619	1.0976	0.525	0.993	1.0331	0.518	0.894	0.9710	*** 0000	0.013 **	0.9638	0.000 ***	0.007 ***
USD/AUD	2.6871	1.0170	*** 600.0	0.674	1.0241	0.619	0.887	0.9604	0.001 ***	0.078 *	0.9814	0.001 ***	0.134
FRF/USD	2.4663	1.1292	0.479	0.981	1.0676	0.630	0.945	90.60	0.005 ***	* 890.0	0.9749	0.010 **	0.177
DEM/USD	2.5581	1.1006	0.588	0.945	1.0596	0.467	0.864	0.9597	0.007 ***	0.079 *	0.9508	0.002 ***	0.028 **
ITL/USD	2.7052	1.0123	* 260.0	0.602	1.0424	0.353	0.815	0.9387	0.004 ***	0.074 *	0.9618	0.019 **	0.138
NLG/USD	2.5501	1.1023	0.478	0.939	1.0880	0.749	0.933	0.9646	0.002 ***	0.034 **	0.9733	0.006 ***	0.063 *
PTE/USD													

Table 18: RMSE and Theil ratios of the forecasts for a shorter-period sample (1980 onwards, end-of-month exchange rates sample).

					d								
Currency pair	No change		Kolling regression	0.10	Thoil ratio	Recursive regression	onless of Mo	Thoil	Skidge		ci+cr lio4T	CW P Value	2010
	NIVISE A 100	וובוו ומנוס	Con Praga	Divi p-value		PP 1	PPP fundamental		Cw p-value	Divi p-value		cw p-value	Dist p-value
USD/GBP	2.6343	0.9999	0.132	0.497	0.9981	0.056 * 0	0.444	0.9981	0.078 *	0.197	0.9981	0.202	0.408
JPY/USD	3.0508	1.0098	0.718	0.855	1.0036	0.843 0	0.852	66660	0.338	0.353	0.9980	0.125	0.221
CHF/USD	3.2016	1.0140	0.891	0.958	1.0036	0.550 0	0.718	0.9992	0.198	0.272	0.9951	0.039 **	0.092 *
CAD/USD	2.1689	0.9979	0.122	0.403	0.9975		0.160	1.0000	0.584	0.569	1.0015	0.990	0.968
SEK/USD	3.3955	1.0067	0.394	0.753	0.9964	*	0.284	1.0006	0.462	0.567	0.9951	0.037 **	0.149
DNK/USD	3.0006	1.0097	0.729	0.926	1.0004		0.542	0.9985	0.106	0.177	0.9918	0.005	0.024 **
USD/AUD	3.2036	0.9952	0.044 **	0.369	0.9962	* *	0.180	0.9977	0.216	0.319	0.9977	0.183	0.269
FRF/USD	3.0352	1.0136	0.702	0.867	1.0056		0.941	1.0045	0.865	0.843	0.9944	0.123	0.218
DEM/USD	3.0482	1.0290	0.978	906.0	1.0147		0.890	0.9983	0.182	0.205	0.9897	0.030 **	* 620.0
ITL/USD	3.2141	1.0025	0.370	0.627	0.9994		0.476	9866.0	0.292	0.384	0.9959	0.214	0.335
NLG/USD	3.0634	1.0143	0.885	0.914	1.0156		0.823	0.9978	0.240	0.290	0.9912	0.050 *	0.104
PIE/USD	3.1025	1.0164	0.549	0.834	T.0093	U.SIS U	U.700 UIRP fundamental	1.0063	0.483	0.65/	0.9986	797.0	0.446
995/0511	2 62/2	1 0000	0.804	9000	1 0244	0 750	0.018	1 0184	8890	0.825	1 0025	775 0	0.611
JPY/USD	3.0508	1.0151	0.317	0.832	1.0056		0.657	1.0004	0.471	0.578	0.9951	0.059 *	0.176
CHE/USD	3.2016	1.0183	0.307	0.844	1.0170		0.813	08660	* 980'0	0.127	0.9933	0.022 **	0.081 *
CAD/USD	2.1689	1.0066	0.530	0.820	1.0031		0.767	1.0000	0.717	0.708	1.0003	0.621	0.640
SEK/USD	3.3927	1.0504	0.886	0.940	1.0298	0.786 0	0.840	0.9910	0.058 *	0.160	0.9833	0.050 *	0.187
DNK/USD	3.0006	1.0208	0.902	0.943	1.0068		0.814	9866.0	0.100 *	0.135	0.9938	0.035 **	0.184
USD/AUD	3.2081	1.0110	0.259	0.763	1.0002	0.366 0	0.528	7666.0	0.366	0.433	9866.0	0.227	0.368
FRF/USD	3.0352	1.0234	0.907	0.926	1.0207		0.835	1.0001	0.395	0.505	0.9902	0.088 *	0.266
DEM/USD	3.0482	1.0358	0.742	0.867	1.0226	0.807 0	0.848	1.0000	0.405	0.500	0.9853	0.039 **	0.136
ITL/USD	3.2141	1.0044	0.283	0.571	1.0175	0.771 0	0.860	6066.0	0.111	0.212	0.9789	0.031 **	0.170
NLG/USD	3.0634	1.0884	0.678	0.883	1.0524	0.646 0	0.859	0.9961	0.177	0.246	0.9877	0.063 *	0.184
PTE/USD													
						Monetary m	Monetary model fundamentals	ıtals					
USD/GBP	2.6343	1.0335	0.752	0.991	1.0154		806.0	0.9900	0.033 **	0.264	1.0058	0.310	0.691
JPY/USD	3.0536	1.0216	0.234	0.918	1.0180		0.886	1.0058	0.287	0.707	0.9999	0.032 **	0.496
CHF/USD	3.2169	1.0305	0.644	0.947	1.0137		0.911	1.0046	0.613	0.824	1.0091	0.414	0.808
CAD/USD	2.1708	1.0395	0.472	0.920	1.0100		0.813	1.0131	0.827	0.832	1.0090	0.919	0.927
SEK/USD	3.3927	1.0467	0.941	0.993	1.0195		0.978	1.0087	0.376	0.668	1.0057	0.333	0.704
DNK/USD	3.0252	1.0407	0.979	0.991	1.0087	0.944 0	0.961	1.0010	0.740	0.753	1.0062	0.499	0.766
USD/AUD EPE/HSD	3.2191	1.0277	0.512	0.918	1.0109		0.960	1.0014	0.299	0.630	1.0028	0.426	0.67
DFM/USD	3.0482	1.0220	0.370	0.784	1.0132		0.698	0.9993	0.109	0.464	1.0019	0.156	0.540
ITL/USD	3.2141	1.0179	0.381	0.733	1.0189		0.773	0.9937	0.148	0.306	1.0004	0.307	0.508
NLG/USD	3.0634	1.0459	0.771	0.923	1.0076		269.0	0.9948	0.140	0.239	1.0012	0.308	0.531
PTE/USD						-							
440/451	2,000	7000	100.0		0000		All fundamentals	7,000	**	.0.0	0.00		
USD/GBP	2.0343	1.0001	0.207	0.999	1.0369		0.930	0.9934	0.039	0.203	1.0050	0.500	0.693
JPY/USD CHE/LISD	3.0536 3.2169	1.0604	0.183 0.678	0.979	1.0281	0.164 0	0.930	1.0029	0.332	0.652	0.99/8 1 0093	0.020 ***	0.414
CAD/USD	2 1708	1 0377	0.275	0.936	1 0090		969 0	1,0035	0.845	0.845	1 0023	0.432	0.934
SEK/USD	3.3927	1.1693	0.893	0.954	1.1165		0.873	1.0104	906.0	0.914	0.9895	0.044 **	0.298
DNK/USD	3.0252	1.0763	0.669	0.998	1.0293		0.940	0.9999	0.389	0.476	1.0055	0.523	0.762
USD/AUD	3.2191	1.0404	0.087 *	0.906	1.0224		0.940	1.0002	0.441	0.520	0.9981	0.130	0.440
FRF/USD	3.0352	1.1011	0.595	0.971	1.0475		0.914	0.9992	0.336	0.438	1.0127	0.420	0.748
DEM/USD	3.0482	1.0721	0.571	0.911	1.0496		0.866	0.9948	0.081 *	0.233	0.9938	0.100 *	0.378
ITL/USD	3.2141	0.9967	0.063 *	0.474	1.0259		0.742	0.9911	0.134	0.219	0.9905	0.126	0.340
NLG/USD	3.0634	1.0656	0.348	0.944	1.0579	0.760 0	0.958	0.9947	0.138	0.221	0.9942	0.152	0.382
PTE/USD													

Table 19: RMSE and Theil ratios of the forecasts with classic fundamentals using Molodtsova and Papell [2009] data set.

	Operator Oly		Rolling regression			Recursive regression	nois		SBidge			ΕWA	
Currency pair	RMSE x 100	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value	Theil ratio	CW p-value	DM p-value
						_	ppp f						
USD/GBP	2.4868	1.0032	0.181	0.586	1.0059	696.0	0.905	0.9805	*** 900.0	0.093 *	0.9888	0.010 **	0.046 **
JPY/USD CHF/USD	2.7955	1.0050	0.482	0.672	1.0025	0.576	0.739	0.9731	0.002 ***	0.004 ***	0.9864	0.000	0.007 ***
CAD/USD	1.2612	0.9869	*** 600.0	0.186	1.0050	0.823	0.801	92660	0.145	0.256	1.0021	0.923	0.917
SEK/USD	2.4678	1.0135	0.427	0.839	1.0082	909.0	0.808	0.9700	0.000 ***	0.007 ***	0.9847	0.000 ***	0.009 ***
DNK/USD	2.5955	1.0197	0.977	0.985	1.0055	0.797	0.840	0.9832	*** 0000	*** 000.0	0.9819	*** 0000	*** 0000
USD/AUD	2.4487	1.0104	0.455	0.836	1.0096	0.826	0.907	0.9905	0.006 ***	0.030 **	0.9927	0.049 **	0.169
FRF/USD	2.6311	1.0137	0.515	0.796	1.0036	0.354	0.617	0.9814	0.000	0.002 ***	0.9838	0.002 ***	0.017 **
DEM/USD	2.7605	1.0147	0.892	0.835	1.0028	0.912	0.882	0.9771	* * * 00000	0.002 ***	0.9846	0.002 ***	0.010 ***
NI GAISD	2.7535	1.0092	0.539	0.716	1.0025	0.520	0.366	0.9630	***	***	0.9727	** 1000	0.045
PTE/USD	2.7333	0.9995	0.092 *	0.491	0.9988	0.068 *	0.480	1.0003	* 060.0	0.505	0.9895	0.049 **	0.345
							UIRP fundamental						
USD/GBP	2.4868	1.0283	0.443	0.823	0.9992	0.231	0.473	0.9618	0.000 ***	0.040 **	0.9782	0.002 ***	0.052 *
JPY/USD	2.7803	1.0033	0.057 *	0.560	0.9950	0.054 *	0.281	0.9634	0.000 ***	0.005 ***	0.9782	0.000 ***	0.025 **
CHF/USD	2.7955	1.0062	0.029 **	0.584	1.0129	0.044 **	0.649	0.9752	0.000 ***	0.083 *	0.9710	*** 0000	*** 900.0
CAD/USD	1.2612	1.0023	0.117	0.561	0.9987	* 860.0	0.452	8766.0	0.121	0.391	1.0032	0.895	0.937
SEK/USD	2.4678	1.0685	0.809	0.914	1.0222	0.764	0.923	0.9464	*** 0000	0.038 **	0.9656	*** 0000	* 990.0
DNK/USD	2.5955	1.0156	0.526	0.860	1.0047	0.289	0.654	0.9738	0.000	0.002 ***	0.9735	*** 000.0	0.002 ***
USD/AUD	2.4487	1.0047	0.043 **	0.573	1.0028	0.195	0.567	0.9795	0.007 ***	0.106	0.9844	0.015 **	0.152
FRF/USD	2.6311	1.0215	0.820	0.897	1.0143	0.635	0.830	0.9704	0.000	0.005 ***	0.9766	0.001 ***	0.024 **
DEM/USD	2.7605	1.0217	0.602	0.807	1.0051	0.526	0.669	0.9690	0.000	0.001 ***	0.9699	0.000	0.004 ***
11L/USD	2.66/0	1.0148	0.31/	0./15	1.0163	0.5/3	0.797	0.9428	0.000	0.039 **	0.9584	0.001	0.05/*
NLG/USD PTE/IISD	2./535	1.0431	0.290	0.763	1.0003	0.1//	0.507	0.9593	0.000 ***	0.001 ***	0.9710	0.000 ***	0.013 **
150/25	7.00.7	1,000.0	0.102	0.514	0.001	Monetary	٠,	o.5527	0.000	660.0	00000	550.0	0.123
IISD/GRP	2 1868	1 0274	0.240	0.881	1 0125	995		0.9662	***	0.107	0.9742	*** 000 0	** 1700
dsu/ydr	2.7803	0.9974	0.001 ***	0.440	0.9815	0.000	0.124	0.9639	*** 000.0	0.021 **	0.9696	0.000	0.042 **
CHF/USD	2.7955	1.0090	0.085 *	0.678	0.9970	0.053 *	0.424	9696.0	0.000	0.027 **	0.9710	0.000	0.035 **
CAD/USD	1.2612	1.0451	0.603	0.995	1.0161	0.736	0.913	0.9982	0.094 *	0.431	0.9962	0.072 *	0.273
SEK/USD	2.4678	1.0826	0.937	0.997	1.0383	0.885	0.946	0.9267	0.000 ***	0.046 **	0.9667	0.000 ***	0.002 ***
DNK/USD	2.5955	1.0438	0.989	0.977	1.0201	0.975	968.0	929670	0.000 ***	0.004 ***	0.9586	0.000 ***	0.001 ***
USD/AUD	2.4487	1.0197	0.462	0.883	1.0191	0.588	0.900	0.9840	0.003 ***	0.173	0.9928	0.011 **	0.328
FRF/USD	2.5229	1.0615	0.822	0.981	1.0738	0.849	0.982	0.9787	0.006 ***	0.126	0.9954	0.094 *	0.390
DEM/USD	2.7624	1.0174	0.203	0.770	1.0038	0.159	0.579	0.9562	*** 0000	0.020 **	0.9514	*** 0000	*** 800.0
ITL/USD	2.6458	1.0411	0.731	0.932	1.0165	0.572	0.798	0.9300	0.000	0.036 **	0.9661	0.000	0.019 **
NEG/USD PTE/USD	2.7535	1.0537	0.873	0.963	1.0079	0.421	0.897	0.9902	0.000	0.364	0.9777	0.000	0.026
							All fundamentals						
USD/GBP	2.4868	1.0693	0.083 *	0.968	1.0232	0.221	0.796	0.9644	0.001 ***	0.044 **	0.9692	0.000	0.035 **
JPY/USD	2.7803	1.0303	0.002 ***	0.853	0.9942	0.000 ***	0.418	0.9612	0.000 ***	0.003 ***	0.9649	0.000 ***	0.017 **
CHF/USD	2.7955	1.0609	*** 600.0	0.916	1.0244	0.045 **	0.744	0.9671	0.000	0.019 **	0.9645	0.000	0.015 **
CAD/USD	1.2612	1.0084	0.001 ***	0.611	1.0154	0.182	0.828	0.9962	0.064 *	0.313	0.9978	0.133	0.261
SEK/USD	2.4678	1.2283	0.862	0.929	1.1167	0.792	0.931	0.9513	*** 00000	0.015 **	1.0009	0.206	0.511
DNK/USD	2.5955	1.0746	0.392	0.979	1.0268	0.472	0.901	0.9662	0.000	0.001 ***	0.9636	*** 0000	0.003 ***
USD/AUD	2.4487	1.0375	0.076 *	0.870	1.0145	0.219	0.755	0.9772	0.001 ***	0.046 **	0.9877	0.011 **	0.233
PRF/USD	677C7	1.113/	0.480	0.980	1.0599	0.018	0.942	0.9736	*** 0000	0.065	0.9786	** ** OOO O	0.180 **
DEINI/03D	2.7024	1.0307	0.016 **	0.556	1.0438	0.038	0.918	0.9351	***	0.004	0.9518	0.000	0.027 **
NLG/USD	2.7535	1.0254	*** 900.0	0.737	1.0198	* 090.0	0.735	0.9583	*** 0000	0.001 ***	0.9660	*** 0000	0.015 **
PTE/USD	2.3017	1.1243	0.933	0.955	1.0162	0.386	0.696	0.9895	0.151	0.306	0.9690	0.033 **	0.133

Table 20: RMSE and Theil ratios of the forecasts with Taylor-rule fundamentals using Molodtsova and Papell [2009] data set.

Currency pair	No change		Kolling regression	2		Kecursive regression	on Dr Jr	(i d d F	Skidge	1	11 0 J	EWA	9
	KINISE X 100	lieli iatio	Cvv p-value	DIVI p-value	Outpr	outout gap fundamentals: deviations		rinelliratio from a linear trend	en p-vaine	DINI p-value	I I I I I I I I I I I I I I I I I I I	cvv p-value	DIVI p-value
USD/GBP	2.4868	1.0434	* 2000	0.925	1.0094	0.100		0.9754	0.002 ***	0.085 *	0.9760	0.001 ***	0.048 **
JPY/USD	2.7803	1.0125	0.005 ***	0.656	0.9913	0.001 ***	0.368	0.9678	0.000	0.002 ***	0.9712	0.000	0.003 ***
CHF/USD	2.7955	1.0829	0.186	0.986	1.0313	0.188	0.867	0.9702	*** 0000	0.007 ***	0.9719	0.000 ***	0.001 ***
CAD/USD	1.2612	1.0038	0.029 **	0.548	0.9749	0.003 ***	0.085 *	9266.0	0.204	0.340	0.9995	0.286	0.432
SEK/USD	2.4678	1.1360	0.722	0.881	1.0398	0.456	0.956	0.9293	0.001 ***	0.061 *	0.9776	*** 000.0	0.016 **
DNK/USD	2.5303	1.0210	0.072 *	0.792	1.0209	0.391	0.848	0.9738	0.000	0.003 ***	0.9711	0.000	0.007 ***
USD/AUD	2.4664	1.0133	0.020 **	0.650	1.0078	0.152	0.608	0.9821	0.013 **	0.166	0.9833	0.006 ***	0.065 *
FRF/USD	2.6311	1.0352	0.124	0.831	1.0265	0.220	0.839	0.9695	0.000	0.002 ***	0.9724	0.001 ***	0.042 **
DEM/USD	2.7605	1.0425	0.345	0.885	1.0272	0.424	0.866	0.9688	* * * 00000	0.001 ***	0.9799	0.002 ***	0.035 **
11L/ USD	2.0070	1.0040	0.022	0.349	1.0203	0.047	0.003	0.9450	***************************************	**************************************	0.9903	0.000	0.240
NLG/USD PTE/USD	2.7535	1.0351	0.067 *	0.794	1.01/5 1.0007	0.256	0.785	0.9651	0.000	0.000 * * * 0	0.9759	0.001 ***	0.013 ***
200/1					Output	Output gap fundamentals: deviations from a quadratic trend	s: deviations from	n a quadratic tr	end				
USD/GBP	2.4868	1.0374	0.092 *	0.913	1.0150	0.167	0.778	0.9749	0.002 ***	0.082 *	0.9755	0.001 ***	0.046 **
DSU/YAL	2.7803	1.0050	0.002 ***	0.561	0.9881	0.002 ***	0.296	0.9673	*** 000.0	0.002 ***	0.9755	*** 000.0	0.002 ***
CHF/USD	2.7955	1.0729	0.219	0.978	1.0336	0.299	0.897	0.9700	*** 00000	0.007 ***	0.9708	0.000	0.000 ***
CAD/USD	1.2612	0.9921	0.005 ***	0.397	0.9798	0.003 ***	* 080.0	0666.0	0.220	0.382	0.9980	0.100	0.201
SEK/USD	2.4678	1.1376	0.749	0.879	1.0599	0.751	0.944	0.9295	0.001 ***	0.061 *	0.9768	0.000	0.012 **
DNK/USD	2.5303	1.0194	* 880.0	0.787	1.0228	0.815	926.0	0.9741	*** 0000	0.004 ***	0.9708	0.000 ***	*** 900.0
USD/AUD	2.4664	1.0220	0.023 **	0.744	1.0126	0.170	0.685	0.9821	0.013 **	0.166	0.9846	0.009 ***	0.082 *
FRF/USD	2.6311	1.0313	0.105	0.814	1.0262	0.260	0.842	0.9695	0.000 ***	0.002 ***	0.9699	0.000 ***	0.029 **
DEM/USD	2.7605	1.0533	0.413	0.931	1.0168	0.386	0.773	6896.0	0.000 ***	0.001 ***	0.9797	0.002 ***	0.036 **
ITL/USD	2.6670	1.0013	0.020 **	0.517	1.0187	0.650	0.839	0.9454	*** 000.0	0.017 **	0.9888	0.026 **	0.207
NLG/USD	2.7535	1.0422	* 690.0	0.834	1.0093	0.296	0.710	0.9651	* * * 000.0	*** 000.0	0.9750	*** 000.0	*** 800.0
PIE/USD	7.301/	1.0683	0.791	0.964	1.0008	0.168	0.510	0.9857	0.028 **	0.052 *	0.9804	0.053 *	0.23/
					Output gap tu	Output gap fundamentals: deviations from a linear and a quadratic trend	lations from a lin	ear and a quad	atic trend				
USD/GBP	2.4868	1.0353	0.188	0.913	1.0135	0.298	0.786	0.9751	0.002 ***	0.083 *	0.9753	0.001 ***	0.047 **
JPY/USD	2.7803	1.0105	0.002 ***	0.648	0.9875	0.001 ***	0.243	0.9683	* * * 0000	0.002 ***	0.9719	0.000	0.003 ***
CHF/USD	2.7955	1.0541	0.131	0.945	1.0327	0.363	0.923	0.9699	0.000	0.006	0.9708	0.000	0.000 ***
CAD/USD SEK/HSD	7.4678	1 1005	0.004 ***	0.291	1.0638	0.025 ***	0.219	0.9991	0.229	0.390	0.9981	0.120	0.251
DNK/USD	2.5303	1.0230	0.130	0.841	1.0225	0.848	0.971	0.9741	***	0.004	0.9685	***	0.004 ***
USD/AUD	2.4664	1.0190	0.032 **	0.721	1.0088	0.169	0.640	0.9820	0.013 **	0.166	0.9848	0.010 ***	0.085 *
FRF/USD	2.6311	1.0329	0.036 **	0.775	1.0244	0.248	0.823	0.9695	*** 0000	0.002 ***	0.9721	0.001 ***	0.042 **
DEM/USD	2.7605	1.0479	0.435	0.917	1.0105	0.185	0.655	0.9694	0.000 ***	0.001 ***	0.9795	0.002 ***	0.031 **
ITL/USD	2.6670	1.0055	0.012 **	0.560	1.0131	0.389	0.731	0.9457	*** 0000	0.017 **	0.9901	0.031 **	0.238
NLG/USD	2.7535	1.0363	0.086 *	0.811	1.0082	0.261	0.679	0.9651	* * * 000.0	*** 000.0	0.9752	*** 000.0	0.011 **
PIE/USD	7.301/	1.0051	0.846	0.950	0.9955	0.1b1	10.429	0.9859	0.029 ***	0.053 *	0.9807	0.053 **	0.239
IISD/GRP	2 4868	1 0460	0.368	7160	J OOSE	1 0086 0.190 0.190 0.682	0 682		10 00 ***	0 081 *	0 9756	0.001 ***	0.046 **
usi //xdi	2 7803	1 0216	0.035 **	0.741	0.9877	*** 0000	0 296	78960	*** 0000	***	72750	*** 000 0	0.005
CHF/USD	2.7955	1.0587	0.289	0.971	1.0304	0.401	0.896	0.9704	* * * 000.0	0.007 ***	0.9708	*** 00000	*** 000.0
CAD/USD	1.2612	0.9917	0.014 **	0.378	1.0120	0.814	0.938	0.9992	0.234	0.395	0.9980	0.101	0.209
SEK/USD	2.4678	1.1056	0.799	0.951	1.0643	0.812	0.960	0.9296	0.001 ***	0.062 *	7220	0.000	0.015 **
DNK/USD	2.5303	1.0357	0.248	0.930	1.0167	0.668	968.0	0.9741	0.000 ***	0.004 ***	0.9704	0.000	0.006 ***
USD/AUD	2.4664	1.0190	0.039 **	0.710	1.0111	0.203	0.656	0.9822	0.013 **	0.168	0.9847	*** 600.0	0.082 *
FRF/USD	2.6311	1.0333	0.081 *	0.817	1.0290	0.372	0.859	0.9695	*** 000.0	0.002 ***	0.9724	0.001 ***	0.042 **
DEM/USD	2.7605	1.0539	0.720	0.966	1.0219	0.623	0.843	0.9694	* * * 000.0	0.001 ***	0.9791	0.002 ***	0.029 **
ITL/USD	2.6670	0.9997	0.014 **	0.497	1.0170	0.577	0.802	0.9456	*** 000.0	0.017 **	0.9898	0.031 **	0.230
NLG/USD	2.7535	1.0406	0.136	0.844	1.0149	0.502	0.797	0.9652	* * * 0000	0.000 * * *	0.9763	0.001 ***	0.014 **
PIE/USD	2.301/	1.0584	0.594	0.916	0.9991	0.180	0.488	0.9860	0.029 **	0.053 *	0.9804	0.052 *	0.236

Table 21: Economic criterion: Results for the EWA algorithm for the end-of-month exchange rates sample without the DNK/USD exchange rate. Note: the results for the portfolios based on the no-change exchange rate predictions vary because of the constraints on weights and because for different sets of fundamentals different time periods are considered depending on the availability of data.

				Annualized	alized			
Portfolio /	Returns (in %)	(in %)	Performance fee of portfolio w.r.t.	Premium return of portfolio w.r.t.	Sharpe ratios	atios	Sortino ratios	atios
c constraint	No change (with c constraint)	Portfolio	no change (in bps)	no change (in bps)	No change (with c constraint)	Portfolio	No change (with c constraint)	Portfolio
				PPP fundamental				
EWA / c = 0.5	10.36	11.38	146	177	09:0	0.75	69:0	1.06
EWA/c=1	11.43	12.83	241	289	0.61	0.81	0.63	1.04
EWA/c=2	10.87	12.36	226	256	0.56	0.75	0.61	0.94
EWA/c=5	11.11	12.33	193	223	0.58	0.74	0.62	0.90
EWA / $c = 10$	11.11	12.33	193	223	0.58	0.74	0.62	0.90
EWA / no c	11.11	12.33	193	223	0.58	0.74	0.62	0.90
Unitorm / no c	11.11	5.46	-776	-153	0.58	0.22	0.62	0.33
				UIRP fundamental				
EWA / c = 0.5	6.92	11.84	246	288	0.57	0.82	09'0	1.38
EWA/c=1	11.04	13.32	326	409	0.58	98.0	09:0	1.51
EWA/c=2	10.49	12.70	324	403	0.54	0.82	0.59	1.57
EWA/c=5	10.73	12.26	276	361	0.55	0.79	09.0	1.48
EWA / c = 10	10.73	12.26	276	361	0.55	0.79	09.0	1.48
EWA / no c	10.73	12.26	276	361	0.55	0.79	09.0	1.48
Uniform / no c	10.73	5.11	-226	-154	0.55	0.18	09:0	0.26
			Monet	Monetary models fundamentals	entals			
EWA / c = 0.5	10.33	9.38	98-	8-	0.59	0.55	0.62	0.72
EWA/c=1	11.40	10.75	73	121	09.0	0.64	0.62	0.81
EWA/c=2	10.85	9.70	12	41	0.56	0.54	09.0	0.67
EWA/c=5	11.08	9.82	-2	27	0.57	0.54	0.61	99.0
EWA / c = 10	11.08	9.82	-2	27	0.57	0.54	0.61	99.0
EWA / no c	11.08	9.82	-2	27	0.57	0.54	0.61	99.0
Uniform / no c	11.08	5.49	-218	-144	0.57	0.22	0.61	0.33
				All fundamentals				
EWA / c = 0.5	10.33	9.71	20	61	0.59	0.61	0.62	1.07
EWA/c=1	11.40	10.61	59	133	09.0	0.63	0.62	1.03
EWA/c=2	10.85	10.47	75	153	0.56	09:0	09.0	1.04
EWA/c=5	11.08	10.27	38	128	0.57	0.58	0.61	1.00
EWA / c = 10	11.08	10.27	38	128	0.57	0.58	0.61	1.00
EWA / no c	11.08	10.27	38	128	0.57	0.58	0.61	1.00
Uniform / no c	11.08	5.49	-218	-144	0.57	0.22	0.61	0.33

Table 22: Economic criterion: results for the EWA algorithm for the end-of-month exchange rates sample with an additional summation-to-less-than-1 constraint on the weights. Note: the results for the portfolios based on the no-change exchange rate predictions vary because of the constraints on weights and because for different sets of fundamentals different time periods are considered depending on the availability of data.

				Annualized	lized			
Portfolio /	Returns (in %)	(in %)	Performance fee of portfolio w.r.t.	Premium return of portfolio w.r.t.	Sharpe ratios	atios	Sortino ratios	atios
c constraint	No change (with c constraint)	Portfolio	no change (in bps)	no change (in bps)	No change (with c constraint)	Portfolio	No change (with c constraint)	Portfolio
				PPP fundamental				
EWA / c = 0.5	10.76	11.77	155	185	0.64	08.0	0.67	1.13
EWA/c=1	11.55	13.07	244	301	0.61	0.82	0.63	1.08
EWA/c=2	11.58	13.54	276	325	0.61	0.84	0.63	1.05
EWA/c = 5	11.72	13.35	231	278	0.63	0.82	0.67	1.06
EWA/c = 10	11.69	13.39	236	284	0.63	0.83	0.67	1.08
EWA / no c Uniform / no c	11.69	13.39	236 -302	284 -238	0.63	0.83	0.67	1.08
)			1	I IIRP fundamental		. !		
FW/A / c = 0 5	10.32	11 89	196	739	09 0	0.79	0.64	1 29
) (11.16	13.02	279	365	0.59	0.82	0.61	1.34
EWA/c=2	11.19	12.99	283	375	0.59	0.82	0.61	1.43
EWA/c=5	11.33	12.14	197	275	09:0	0.77	0.65	1.35
EWA/c = 10	11.30	12.11	204	280	09.0	0.78	0.65	1.38
EWA / no c	11.30	12.11	204	280	09.0	0.78	0.65	1.38
Uniform / no c	11.44	5.29	-301	-238	0.62	0.19	0.65	0.28
			Monet	Monetary models fundamentals	entals			
EWA / c = 0.5	10.70	8.54	-117	-87	0.63	0.50	0.67	0.67
EWA/c=1	11.46	9.34	-56	9	09:0	0.52	0.63	0.68
EWA/c=2	11.51	8.89	-114	-50	09:0	0.47	0.63	0.62
EWA/c=5	11.59	8.46	-168	-101	0.62	0.43	99.0	0.62
EWA / c = 10	11.55	8.41	-175	-108	0.62	0.43	99.0	0.61
EWA / no c	11.55	8.41	-175	-108	0.62	0.43	99.0	0.61
Uniform / no c	11.70	5.78	-275	-211	0.63	0.25	0.67	0.37
				All fundamentals		•		
EWA / c = 0.5	10.70	8.52	-100	-63	0.63	0.51	0.67	0.84
EWA/c=1	11.46	9.56	-12	62	09:0	0.56	0.63	0.91
EWA/c=2	11.51	9.20	-81	က	09:0	0.50	0.63	0.89
EWA/c = 5	11.59	8.57	-166	-82	0.62	0.44	99.0	08.0
EWA / c = 10	11.55	8.48	-175	-91	0.62	0.43	99.0	0.79
EWA / no c	11.55	8.48	-175	-91	0.62	0.43	0.66	0.79
UTILIOTEII / IIO C	11.70	5.78	6/7-	117-	0.03	0.25	0.0/	0.37

Table 23: Economic criterion: results for the EWA algorithm for the end-of-month exchange rates sample without the DNK/USD exchange rate but with an additional summation-to-less-than-1 constraint on the weights. Note: the results for the portfolios based on the no-change exchange rate predictions vary because of the constraints on weights and because for different sets of fundamentals different time periods are considered depending on the availability of data.

				Annualized	alized			
Portfolio /	Returns (in %)	(in %)	Performance fee	Premium return	Sharpe ratios	atios	Sortino ratios	atios
c constraint		(2)	of portfolio w.r.t.	of portfolio w.r.t.))))			
	No change (with c constraint)	Portfolio	no change (in bps)	no change (in bps)	No change (with c constraint)	Portfolio	No change (with c constraint)	Portfolio
				PPP fundamental				
EWA / c = 0.5	10.24	11.30	158	185	09:0	0.76	0.64	1.08
EWA/c=1	10.89	12.71	288	340	0.56	0.80	0.58	1.05
EWA/c=2	10.57	12.28	252	282	0.54	0.74	0.58	0.94
EWA/c=5	10.73	12.20	223	253	0.54	0.72	0.59	0.89
EWA / c = 10	10.73	12.20	223	253	0.54	0.72	0.59	0.89
EWA / no c	10.73	12.20	223	253	0.54	0.72	0.59	0.89
0111011111	11.11	0.40	-220	CCT-	0.30	0.22	0.07	0.33
				UIRP fundamental				
EWA / c = 0.5	9.81	11.66	237	275	0.56	0.81	09.0	1.37
EWA/c=1	10.51	13.17	367	452	0.53	0.84	0.56	1.48
EWA/c=2	10.19	12.49	338	418	0.51	0.80	0.56	1.53
EWA/c=5	10.35	12.13	304	391	0.52	0.77	0.57	1.46
EWA / c = 10	10.35	12.13	304	391	0.52	0.77	0.57	1.46
EWA / no c	10.35	12.13	304	391	0.52	0.77	0.57	1.46
Uniform / no c	10.73	5.11	-226	-154	0.55	0.18	09.0	0.26
			Monet	Monetary models fundamentals	entals			
EWA / c = 0.5	10.22	60.6	-35	6-	0.59	0.55	69:0	0.74
EWA/c=1	10.86	10.61	117	168	0.55	0.63	0.58	0.80
EWA/c=2	10.54	89.6	48	77	0.53	0.54	0.58	0.67
EWA/c=5	10.70	89.6	31	61	0.54	0.53	0.58	0.65
EWA / c = 10	10.70	89.6	31	61	0.54	0.53	0.58	0.65
EWA / no c	10.70	89.6	31	61	0.54	0.53	0.58	0.65
Uniform / no c	11.08	5.49	-218	-144	0.57	0.22	0.61	0.33
				All fundamentals				
EWA / c = 0.5	10.22	9.26	-11	28	0.59	0.58	69:0	1.02
EWA/c=1	10.86	10.39	93	171	0.55	0.61	0.58	1.02
EWA/c=2	10.54	10.28	06	168	0.53	0.58	0.58	1.01
EWA/c=5	10.70	10.12	64	156	0.54	0.57	0.58	0.98
EWA / c = 10	10.70	10.12	64	156	0.54	0.57	0.58	0.98
EWA / no c	10.70	10.12	64	156	0.54	0.57	0.58	0.98
Unitorm / no c	11.08	5.49	-218	-144	0.57	0.22	0.61	0.33

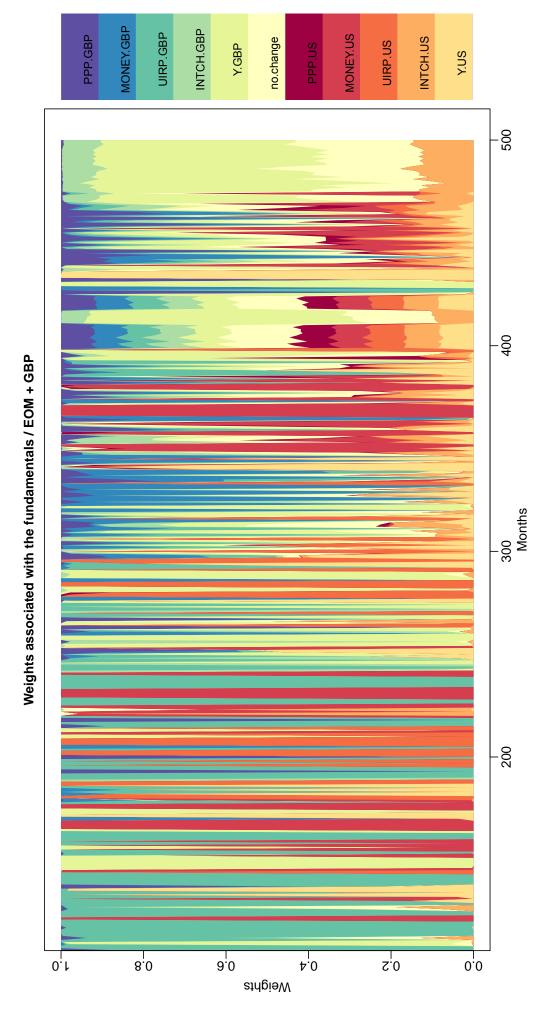


Figure 1: Weights of different fundamentals in prediction for the end-of-month exchange rates sample: USD/GBP case, all fundamentals considered together.

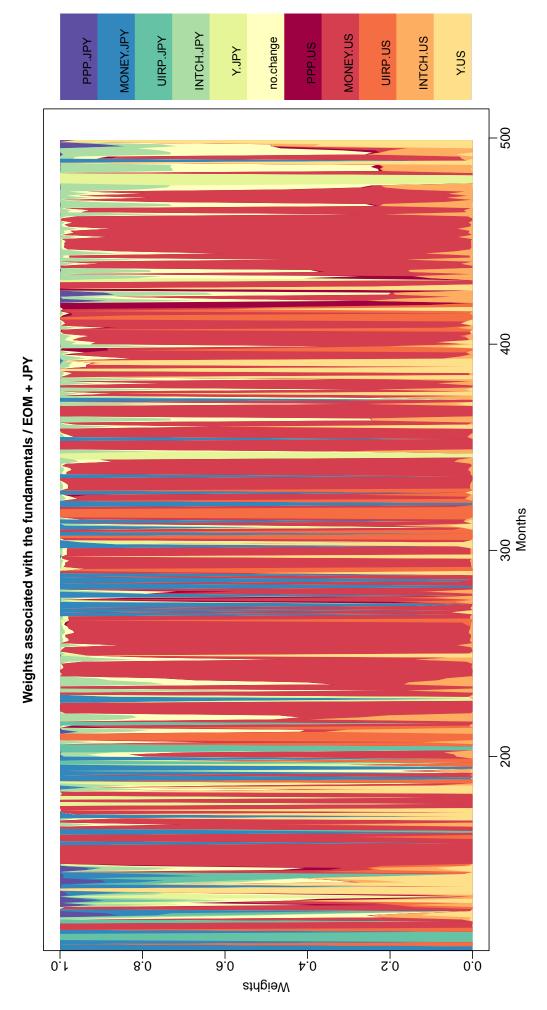


Figure 2: Weights of different fundamentals in prediction for the end-of-month exchange rates sample: JPY/USD case, all fundamentals considered together.

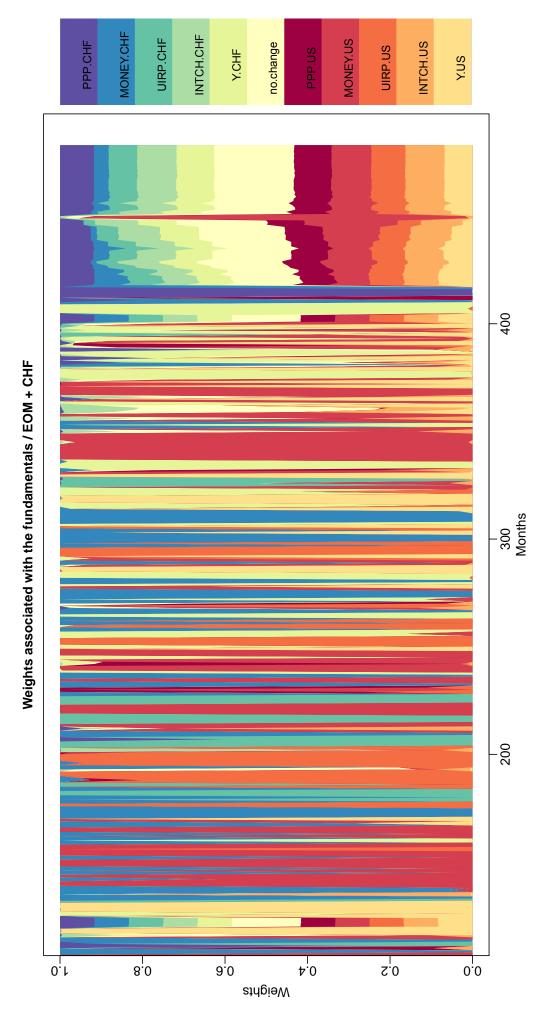


Figure 3: Weights of different fundamentals in prediction for the end-of-month exchange rates sample: CHF/USD case, all fundamentals considered together.

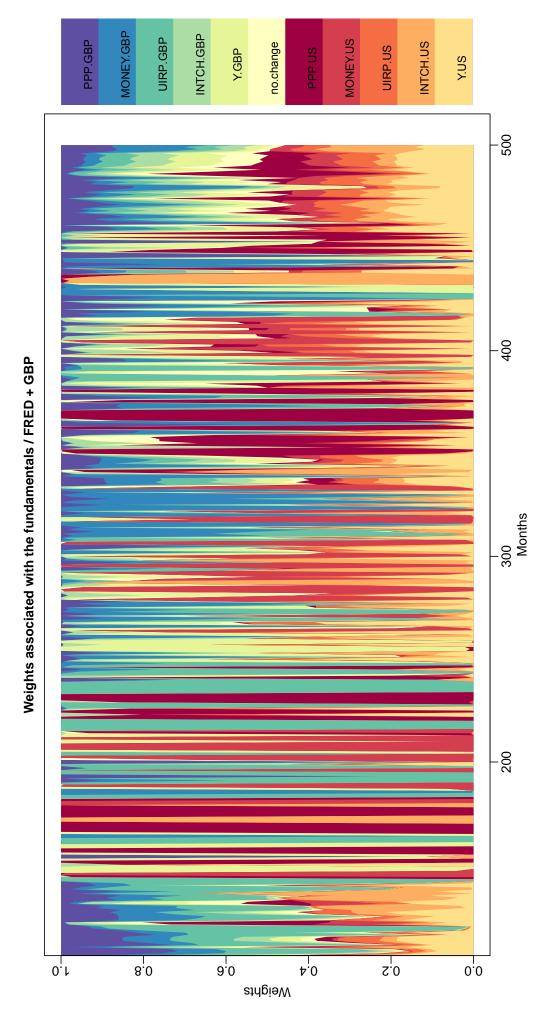


Figure 4: Weights of different fundamentals in prediction for the average exchange rates sample: USD/GBP case, all fundamentals considered together.

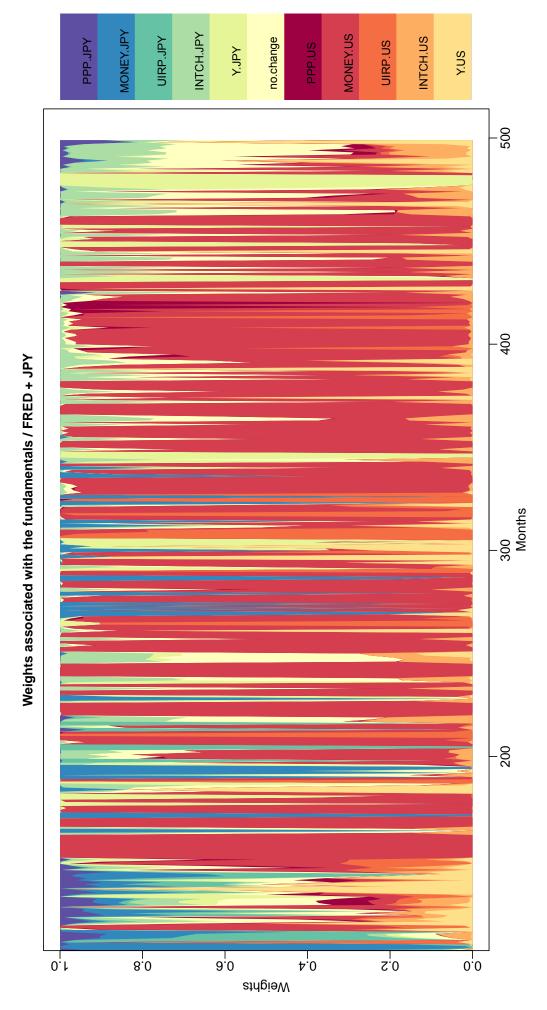


Figure 5: Weights of different fundamentals in prediction for the average exchange rates sample: JPY/USD case, all fundamentals considered together.

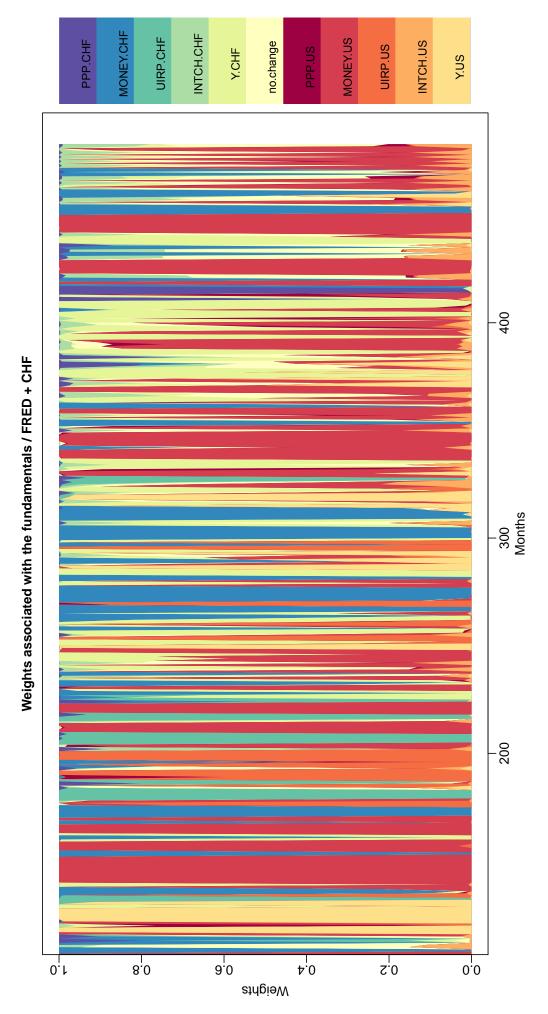


Figure 6: Weights of different fundamentals in prediction for the average exchange rates sample: CHF/USD case, all fundamentals considered together.