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# GLOBAL SOURCING UNDER IMPERFECT CAPITAL MARKETS

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## Abstract

We develop a simple model to study the interactions between a supplier's financial constraints and contract incompleteness in a vertical relationship. Production complexity increases the extent of contract incompleteness and the hold-up problem, which generates a cost when the supplier needs financial participation from the downstream firm. Vertical integration alleviates the impact of financial constraints but reduces the supplier's incentives. We apply the model to an analysis of multinational firms' sourcing strategies and predict that (1) complex and specific inputs are more likely to be sourced from financially developed countries and (2) multinationals are more likely to integrate suppliers located in countries with poor financial institutions, especially when trade involves complex goods. We examine and validate these predictions using firm-level trade data on multinational firms with operations in France. We provide evidence that financial development generates a comparative advantage in the supply of complex goods. Moreover, we find higher shares of intra-firm imports of complex inputs from countries with a lower level of financial development. The findings are robust to different measures of complexity and specificity, and are not driven by industry differences in fixed costs or traditional measures of external financial dependence. Quantitatively, we find that financial development is as important as contract enforcement in alleviating hold-up problems.

**Keywords:** sourcing, FDI, financial constraints, contractual frictions.

**JEL Classification:** F10, O16, L23.

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# 1 Introduction

The process of globalization is characterized by the growing fragmentation of production and the organization of firms' activities on a global scale. A substantial amount of empirical evidence has documented the importance of trade in intermediate goods in total world trade and the extent of vertical specialization (Yeats, 2001; Feenstra, 1998; Campa Goldberg, 1997; Hummels, Ishii and Yi, 2001). Moreover, multinational firms account for a large proportion of world trade, with affiliated or independent partners.<sup>1</sup>

Firms seeking to exploit cross-country differences in production costs interact with suppliers operating in environments with varying levels of institutional quality. One important institutional dimension is access to finance. Previous research has shown that it strongly impacts on investment, firm performance and exports (King and Levine, 1993; Rajan and Zingales, 1998; Manova, 2008). However, local financial markets are not the only source of finance for suppliers and there is also evidence of capital flows from multinational firms to affiliates as potential channels to overcome imperfections in local capital markets (Desai, Foley and Hines, 2004).

The purpose of this paper is to analyze how multinational firms' sourcing strategies are affected by suppliers' financial constraints and the financial development of the countries in which they operate. We show that financial constraints interact with the extent of contract incompleteness in vertical trade. The model predicts that the profitability of offshoring production increases, *ceteris paribus*, with the level of financial development of the destination country. In particular, this effect is stronger in industries with higher degrees of technological complexity and specificity, with greater dependence on non-contractible investments and tasks. We show that vertical integration of the supplier alleviates financial constraints, and thus intra-firm trade arises in complex industries especially when local capital markets are imperfect. The model's main empirical predictions are then tested and validated against firm-level trade data on the operations of multinational firms located in France.

Recent theoretical and empirical research has established that contractual frictions are strong determinants of trade patterns and the cross-border organization of firms (see Antras and Rossi-Hansberg, 2008). Wide variations are observed in the extent of contract incompleteness and its impact, depending on the type of product that is traded. In particular, the level of product complexity appears to be one of the main determinants of contract incompleteness.<sup>2</sup> In our model, complex products refer to goods involving a higher proportion of complex tasks and components that cannot be fully described in a contract. We focus on one supplier producing an input for one final good producer – the multinational firm. The greater the product complexity, the greater the hold-up problem *ex post* between both parties. A multinational sourcing complex inputs only secures a limited fraction of *ex-post* gains while the upstream party captures a positive rent. Therefore, the multinational refrains from financing the supplier's initial costs and may even ask *ex ante* for a compensating transfer (e.g. licensing fee). Because of this, the supplier's need for initial liquidity is inherently related to the level of input complexity. Supplying complex products calls for adequate financial capacity whereas the production of basic inputs can rely on co-financing from the multinational.

As is well known from the property rights theory of the firm (Grossman and Hart, 1986; Hart and

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<sup>1</sup>About 67% of French imports and 75% of French exports in 1999 concerned manufacturing groups controlling at least 50% of a foreign affiliate. Similar patterns are observed for the US. Moreover, one-third of world trade is intra-firm (UNCTAD 2002). This proportion is even higher for developed countries: nearly half of US trade occurs within firms' boundaries.

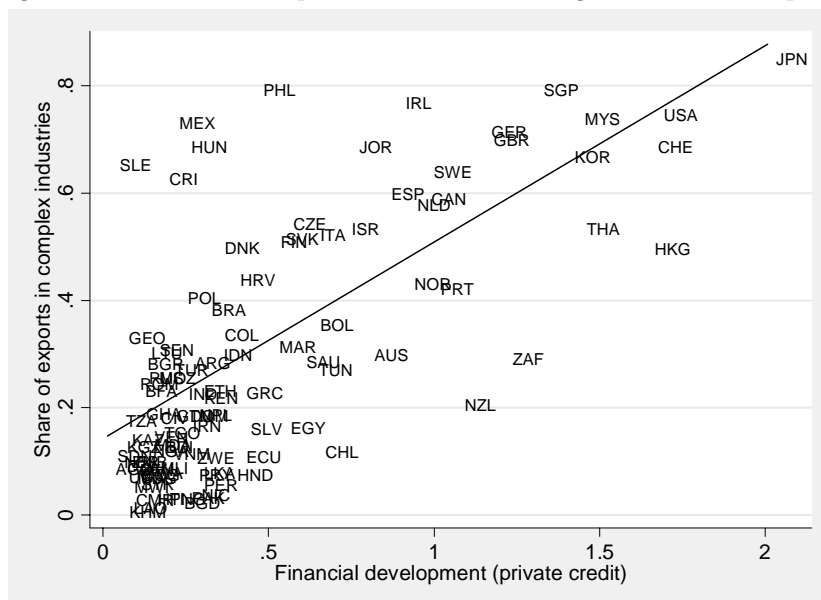
<sup>2</sup>Hart and Moore (1999) and Segal (1999) present theoretical grounds for the existence of incomplete contracts by incorporating the notion of complexity.

Moore, 1990), one way of solving the problems of incomplete contracts is with a suitable allocation of ownership. The theory posits that efficiency requires residual property rights to be assigned to the party whose non-contractible investments have the greatest marginal effect on the joint value. In the absence of credit constraints, the supplier should remain independent assuming that specific investments are essential to produce the intermediate input. Capital market frictions, however, limit the supplier's capacity to finance the initial cost of outsourcing with the multinational firm. In this case, we show that allocating residual property rights to the multinational firm reduces the need for supplier's financial participation. Therefore, vertical integration can arise as an equilibrium organizational form, even when it is suboptimal in terms of productive efficiency.

An important determinant of financial constraints is the level of financial development of the country in which the supplier is located. In the model, the upstream producer operates in an environment plagued with financial frictions that impose a bound on its access to liquidity. In addition to their initial contribution, we assume that suppliers can raise debt from local banks, depending on the level of financial development. A higher level of financial development is characterized by a larger proportion of pledgeable future income.

A first prediction of the model is that imperfect capital markets have a particular effect on complex input offshoring decisions. We look at differential effects across industries and show that financial development generates a comparative advantage in the supply of complex inputs. The grounds for this can be seen from Figure 1. The graph shows that the share of exports in complex industries (broadly speaking: machinery, equipment and chemicals) is highly correlated with the level of financial development.<sup>3</sup>

Figure 1: Financial development and the technological content of exports



The model also predicts that the share of intra-firm imports by multinational firms should be higher in countries with lower levels of financial development. Moreover, this effect should be more pronounced

<sup>3</sup>Figure 1: World trade flows in manufacturing goods in 1999 are drawn from CEPII at the 2-digit level of ISIC Revision 3. According to the OECD STAN definition of high and medium-high technology manufactures, complex industries correspond to codes 24 and 29 to 35 (more refined measures are used in the empirical section). Financial development is measured as the ratio of private credit to GDP (Beck *et al.*, 2000). We focus on the 100 largest countries in terms of population with data on private credit. The coefficient of the linear trend is 0.366 (std error: 0.033).

for complex inputs. Multinational firms should be more likely to integrate their suppliers as a means to alleviate their financial constraints and reduce exposure to opportunism in countries where financial development is low. Complex industries are those where the costs of outsourcing are higher for the reasons described above, and thus where financial constraints are more likely to bind.

The model's predictions are tested in the empirical section of the paper. We use data on imports by multinational firms located in France (manufacturing groups with at least one affiliate abroad), detailed by firm, product and country of origin. An important feature of these data is that they provide the proportion of intra-firm trade for each observation. Our econometric analysis shows that multinational firms are less likely to trade complex inputs with suppliers located in countries with lower levels of financial development. Furthermore, when financial development is low, imports of complex inputs are more likely to occur within the bounds of the firm, from an affiliate. These findings are robust to different measures of financial development and complexity: R&D intensity measured in France and the UK, the Lall (2000) index of technological content and the Rauch (1999) classification of differentiated versus homogeneous goods. Moreover, our results persist after the inclusion of firm fixed-effects and are not driven by reverse causality. Quantitatively, financial development has a statistically and economically significant impact on imports, depending on the complexity of the traded good. We estimate that a one standard deviation change in our index of financial development yields a 9% increase in the probability of imports of complex goods compared to basic goods (with a difference of one standard deviation in the index of complexity). This effect is larger than or comparable to the effect of contract enforcement on the composition of imports.

Our paper contributes to the literature on institutions and trade (Acemoglu et al., 2007; Levchenko, 2007; Costinot, 2007; Nunn, 2007) and more specifically on financial development and trade (Beck, 2003; Becker and Greenberg, 2005; Manova, 2006 and 2008; Svaleryd and Vlachos, 2005). The main idea behind most studies in this literature is that potential exporters with high initial costs need external finance, unless they have a sufficient amount of cash in hand (Beck 2002; Chaney, 2005; Manova, 2006; Becker and Greenberg, 2007).<sup>4</sup> This view, however, overlooks the fact that a large proportion of trade involves multinational firms that can directly finance their suppliers' initial costs, especially when the supplier is an affiliate. By contrast, in our model, the need for external finance is the result of a tension between suppliers and multinational firms when production relies on complex tasks and components that cannot be described in contracts. In very recent work, Basco (2008) studies the role of financial development in the product cycle in international trade and suggests that financial development interacts with the intensity in headquarter services to determine the location of production.<sup>5</sup> Our empirical results support, however, the notion put forward by our theory that financial development creates a comparative advantage in complex and specific goods. In our data, the impact of financial development is not driven by interactions with the traditional index of external financial dependence (Rajan and Zingales, 1998), alternative measures of fixed costs, or different indices of headquarter intensity. Moreover, our estimates suggest a stronger role of financial development compared to the rule of law and intellectual property rights.

Our paper also contributes to the literature on the determinants of intra-firm trade. It is the first to examine (theoretically and empirically) the role of financial constraints and financial development on the share of intra-firm trade. Previous theoretical studies have analyzed the role of intellectual property rights

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<sup>4</sup>Kletzer and Bardhan (1987) is the first theory on financial development and trade, where better financial institutions generate a comparative advantage in financially dependent sectors.

<sup>5</sup>The intensity in headquarter services is defined as in Antras (2005). However, Basco (2008) does not consider the choice between outsourcing and integration.

(Ethier and Markusen, 1996), capital endowments (Antras, 2003) and contract enforcement (Antras and Helpman, 2008),<sup>6</sup> but our empirical analysis suggests that financial development plays a comparable if not greater role. Moreover, our model combines predictions on comparative advantage and intra-firm trade in a simple way. Integration of the supplier eases financial constraints by taking formal control of the production of complex products when the supplier is located in a country with poor financial institutions. A different mechanism is at work in the recent model developed by Antras, Desai and Foley (2008) to analyze the role of weak investor protection in the horizontal FDI versus technology licensing decision (agents do not trade), where direct financial participation by the multinational is considered by the local bank to be a guarantee of monitoring (in line with Holmstrom and Tirole, 1997).

Recently, Machiavello (2007) and Acemoglu, Johnson and Mitton (*forthcoming*) have studied how contracting costs and financial development interact in the determination of firms' boundaries in a closed economy, using cross-country data.<sup>7</sup> Applying these questions to international trade turns up some interesting and proper identification strategies absent from closed economy studies. An analysis of the choice of import origin permits to identify the impact of the suppliers' environment independently from the location of the downstream firm. In addition, it permits to observe variations in the size of imports and sourcing modes (integration or outsourcing) by country of origin for the same downstream firm. Our empirical study thus draws on direct information on sourcing mode by firm, whereas the abovementioned studies have to rely on indirect measures of vertical integration such as those calculated on the basis of input-output matrices and the ratio of value added to total sales.

The paper is organized as follows. Section 2 develops a model that formalizes the interactions between a multinational firm and a supplier in an incomplete contract setting, where the latter has limited access to finance. We also discuss some extensions of the baseline model. Section 3 presents the empirical analysis based on firm-level trade data on French multinational firms. Section 4 concludes.

## 2 Theory

### 2.1 A simple model

Our model is based on a simplified version of the Antras and Helpman (2008) model of global sourcing, incorporating capital market frictions on the supplier's side.

#### **Production: basic and complex tasks**

We consider two agents in a vertical relationship: a downstream firm (referred to as the "multinational firm", labeled M) and the manager of a local upstream firm (the "supplier", labeled S). The supplier produces an intermediate good that is traded to the multinational firm, which then produces and sells the final good.

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<sup>6</sup>The main predictions made by Antras (2003) and Antras and Helpman (2004, 2008) have been empirically checked by Yeaple (2006), Nunn and Trefler (2008), and Bernard et al. (2008) using data on US multinational firms and by Defever and Toubal (2007) and Corcos, Irac, Mion and Verdier (2008) using the SESSI dataset on French multinational firms, which we also use in our empirical analysis. However, none of them considers the level of financial development.

<sup>7</sup>More broadly, our theory is also related to the literature on the interactions between financial constraints and control in corporate finance, first developed by Aghion and Bolton (1992) and later applied in Aghion and Tirole (1994) and Legros and Newman (2003), among others. Our model focuses on vertical relationships, but the mechanism presents similarities with the trade-off between debt and control, where the investor prefers to take control of the firm when the entrepreneur cannot raise enough debt.

Total revenues from final good sales equal  $Y = A^{1-\rho}Q^\rho$  where  $Q$  denotes the volume of production and  $A$  is a constant that reflects the perceived demand schedule. This expression can be derived from the Dixit-Stiglitz (1977) model of monopolistic competition with constant elasticity of substitution (CES) between final goods. We denote the elasticity of substitution by  $\sigma = \frac{1}{1-\rho} > 1$ , with  $0 < \rho < 1$  measuring the degree of product differentiation.

The multinational firm needs one unit of an input to produce one unit of a final good. In turn, the supplier needs to undertake a continuum of tasks to produce the intermediate good. We normalize to one the measure of the set of tasks. Each task can be taken to be the production of a single component of the intermediate good.

There are two types of task: complex and basic. Contrary to basic tasks, complex tasks are specific and cannot be fully described in a contract. They correspond to attributes and components of the intermediate good that are not verifiable by a third party. We denote by  $\theta$  the proportion of complex tasks required for the production of the input.  $\theta$  is thus a parameter that characterizes the input's industry.

Each task  $i$  demands an effort from the supplier, denoted  $x_i$ . The unit cost of effort put into each task is constant and denoted by  $c$ . We assume that the production technology is represented by a CES function where tasks are imperfect substitutes. To be more precise, the volume of production equals:

$$Q = \left[ \int_I x_i^{\rho'} di \right]^{\frac{1}{\rho}}$$

where  $I$  refers to the set of tasks and  $\frac{1}{1-\rho'} > 1$  to the elasticity of substitution. Symmetrically, the level of effort is the same for all complex tasks and all basic tasks respectively. We denote by  $e$  the level of effort in complex tasks and by  $q$  the level of effort in basic tasks.

We assume, for the sake of argument, that the elasticity of substitution between tasks is the same as it is between final goods ( $\rho = \rho'$ ), producing a simple linear form for the value of total sales. This assumption allows us to obtain a simple linear expression of total revenues as a function of the degree of complexity:<sup>8</sup>

$$Y = A^{1-\rho} [\theta e^\rho + (1-\theta)q^\rho]$$

Total costs sum up to:

$$C = \theta ce + (1-\theta)cq$$

Notice that we do not rely on the assumption of fixed costs, but they could simply be added to this framework.

### Organizational form and ex post bargaining

The inability to contract on the level of investment  $e$  in complex tasks results in ex-post bilateral bargaining over the surplus. Efforts in all tasks are implemented by the supplier before the bargaining stage. However, in the event that no agreement is reached in the bargaining stage, efforts in complex tasks can be wiped

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<sup>8</sup>Note that the linearity in  $\theta$  in the expression for  $Y$  comes from the equality between both elasticities of substitution. In Section 2.3 we discuss a more general framework and show that our main mechanisms remain. Nevertheless, in the general case it is not possible to obtain closed-form solutions. Similar linear forms have been put forward by Acemoglu, Aghion, Griffith and Zilibotti (2005) and Acemoglu, Johnson and Mitton (*forthcoming*). Alternatively, we can obtain such a linear expression by assuming that only one task has to be performed, characterized by a probability  $1 - \theta$  of being verifiable and a probability  $\theta$  of being non-verifiable. In this specification, the results would remain qualitatively similar.



out by the supplier, while efforts in basic tasks must follow the contract’s specifications.

We model the bargaining game using the Nash bargaining solution with symmetric bargaining shares. When negotiations are through, each party obtains half of the quasi-rents generated by the partnership plus its outside option, represented by the value of the better alternative available in the event of no agreement (threat point).

In keeping with Grossman and Hart (1986), outside options are assumed to be contingent on the prevailing organizational structure, determined by the party that is assigned the property rights over the input. We consider two forms of organizational structure. Under a non-integration, or “outsourcing” arrangement, the supplier retains residual property rights. Nevertheless, we suppose that efforts in complex tasks are specific and sunk, which implies that the supplier’s outside option is zero, irrespective of organizational form. Analytically, outside outsourcing options are given by the first line of the following table:<sup>9</sup>

Outside options	Supplier S	Multinational M
Outsourcing	0	$A^{1-\rho}(1-\theta)q^\rho$
Vertical Integration	0	$A^{1-\rho}(1-\delta)\theta e^\rho + A^{1-\rho}(1-\theta)q^\rho$

The alternative organizational arrangement is (backward) vertical integration, in which M incorporates S as an internal unit (affiliate), hires the manager as the head of this unit and retains ownership of the assets in the event of no agreement. In this case, then, the multinational could fire the manager and replace him with someone else. This constitutes its threat point in the ex-post negotiations. We assume that this would occur at the cost of losing part of the value of the complex components, which would be reduced by a proportion  $\delta < 1$ . As mentioned, the supplier’s outside option is zero under vertical integration. Analytically, we obtain the expressions in the second line of the above table.

At this point, it is worth noting that organizational form does not directly impact on the production function (there is no *ad-hoc* fixed cost of integration). It only directly affects the value of outside options.

### Financial constraints

Lastly, an important feature of our model is the fact that capital markets are assumed to work only imperfectly. In particular, the supplier may be financially constrained at the time of starting to do business with the multinational firm. The multinational firm, on the other hand, is assumed to have access to both foreign and domestic financial markets and thus not to face liquidity constraints at all.

Limited available funds reduce the supplier’s ability to make initial investments. More specifically, liquidity constraints impose a bound on the sum of initial costs  $C$  and the upfront payment  $T$  that can be asked initially by the multinational firm. Note that  $T$  may be either positive or negative. Negative transfers reflect the extent of co-financing whereas positive transfers represent a participation fee (e.g. royalties or licensing fees for the use of the multinational’s technology).

We assume that the supplier’s level of initial liquidity is made up of two elements. One is initial holdings of cash, summed up by parameter  $W$ . In addition, the supplier can raise debt  $L$  from local banks.

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<sup>9</sup>Our model would remain qualitatively unchanged if we were to relax the assumption of symmetry. In section 2.3, we relax the assumption that efforts in complex tasks are specific (e.g. efforts of technology assimilation), which yields similar conclusions to the baseline model. In a similar framework, Feenstra and Hanson (2005) estimate a bargaining power of 0.7 for the multinational firm and 0.3 for the local firm. These values, however, are not statistically different from 0.5.

The fact that both the initial transfer  $T$  and production costs  $C$  are limited by available liquidity implies the following liquidity constraint:

$$T + C \leq W + L$$

How much the supplier can borrow from local banks depends on the level of financial development, indexed by a parameter  $\kappa \in [0, 1]$ . Specifically, if the supplier's future revenues equal  $Y_S$ , the maximum amount of debt  $L$  that can be raised from local banks is limited by the financial constraint:

$$L \leq \kappa Y_S$$

In other words,  $\kappa$  is the ratio of pledgable income to total expected income. Thus, a value of  $\kappa = 1$  indicates perfect capital markets, whereas  $\kappa = 0$  indicates no capital market.

Various micro-economic theories can explain the presence of financial constraints. Models of credit rationing are based mainly on the non-observability of outcomes (for example, Aghion, Banerjee and Piketty, 1999; and Schneider and Tornell, 2004), investment decisions (Holmstrom and Tirole, 1997) and individual capabilities (Jaffe and Stiglitz, 1990). In keeping with ex-post moral hazard models, our variable  $\kappa$  can be interpreted as the cost of diverting the funds and not repaying the debt. Alternatively, we could model financial development as the efficiency of monitoring to prevent unproductive decisions (Antras, Desai and Foley, 2008). With either approach, the conclusions of our model would remain qualitatively similar.

In the baseline specification, we assume that repayment of the debt is not contingent on whether an agreement is reached with the multinational firm. In Section 2.3, we explore the alternative route and assume that repayment of the debt depends on the agreement between the supplier and the multinational.<sup>10</sup>

Note that we do not assume that vertical integration has a direct impact on the availability of external finance. For example, we could assume that stricter monitoring by the multinational firm allows for a higher level of debt from local banks. Such assumptions, however, yield very similar results regarding the choice between outsourcing and integration, as shown in Section 2.2.

## Timing

Before moving on to the solution of the model, we present the timing of events:

1. The multinational proposes a contract to the supplier specifying a triplet  $(Z, T, q)$  where  $Z$  is the organizational form,  $T$  an upfront monetary transfer from S to M and  $q$  the level of effort for the basic tasks. The upfront payment  $T$  is restricted by the supplier's initial level of liquidity (debt and initial capital).
2. Transfer  $T$  takes place.
3. The supplier decides its level of effort  $e$  and produces the input, under the constraint that total cost can not exceed available liquidity.
4. Nash bargaining on the value of joint production and repayment of the external debt take place.

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<sup>10</sup>Whether the debt is raised from the multinational firm or from local banks makes no difference in the baseline specification. However, raising debt from local is optimal when repayment of the debt is contingent on an agreement with the multinational firm.

## 2.2 Solution

### Maximization program

The multinational's total profits equal its ex-post revenues plus the ex-ante transfer. The contract chosen maximizes total profits subject to three constraints: a participation constraint imposing non-negative total profits on the supplier, the financial constraints discussed previously, and an incentive compatibility constraint reflecting the supplier's decision in terms of specific investments in complex tasks.

In formal terms, M chooses the triplet  $(Z, q, T)$  that solves the following constrained maximization problem:

$$\begin{aligned} \max_{Z, q, T} \quad & \Pi = Y_M^Z + T \\ \text{s.t.} \quad & T \leq Y_S^Z - C && \text{(Participation constraint, PC)} \\ & T \leq W + \kappa Y_S^Z - C && \text{(Financial constraint, FC)} \\ & e = \operatorname{argmax}_e \{ Y_S^Z - C \} && \text{(Incentive compatibility constraint, IC)} \end{aligned}$$

where  $C = \theta ce + (1-\theta)cq$  is the cost of effort, and  $Y_M^Z, Y_S^Z$  are revenues after bargaining. Ex-post revenues are functions of  $e, q$  and the organizational form  $Z \in \{O, I\}$  as we show in the next paragraph.

### Bargaining outcomes

In the case of an outsourcing arrangement, the multinational's outside option equals revenues related to basic tasks, as production related to complex tasks is lost in the event of disagreement. Given symmetric bargaining power, ex-post revenues for the multinational firm and the supplier (before repayment of the debt) equal respectively:

$$\begin{aligned} Y_M^o &= \frac{1}{2} A^{1-\rho} \theta e^\rho + A^{1-\rho} (1-\theta) q^\rho \\ Y_S^o &= \frac{1}{2} A^{1-\rho} \theta e^\rho \end{aligned}$$

Under vertical integration, ex-post payoffs to both parties change in line with the new allocation of residual property rights. Given the effort put into the complex and basic tasks, the multinational firm captures a larger proportion of the ex-post revenues under integration while the opposite holds for the supplier:

$$\begin{aligned} Y_M^I &= \left(1 - \frac{\delta}{2}\right) A^{1-\rho} \theta e^\rho + A^{1-\rho} (1-\theta) q^\rho \\ Y_S^I &= \frac{\delta}{2} A^{1-\rho} \theta e^\rho \end{aligned}$$

### Investments in complex and basic tasks

In this framework, investments in specific tasks are determined by the organizational form. Under the incentive compatibility constraint (IC), investments made by the supplier in complex tasks maximize its ex-post revenues  $Y_S^Z$  minus costs  $C$ , where  $Z \in \{I, O\}$  refers to the organizational form. Under outsourcing, the level of effort equals:

$$e^o(c) = a\rho(2c)^{-\sigma} \tag{1}$$

where  $a = A\rho^{\sigma-1}$  is a constant and coefficient 2 in brackets comes from the symmetric Nash bargaining solution. The fact that the supplier's bargaining power is strictly below one implies that it does not recover the full marginal return on its investments. Thus, the level of non-contractible effort is below the first best even in cases where liquidity constraints do not bind.

Under integration, the supplier only retains a fraction  $\frac{\delta}{2}$  of ex-post revenues in the bargaining stage. Given this, the optimal level of non-contractible effort is a function of the unit cost of effort and holds:

$$e^I(c) = a\rho\delta^\sigma(2c)^{-\sigma} < e^o(c)$$

Efforts are lower than under outsourcing. The underinvestment problem is exacerbated because the supplier captures a smaller share of the ex-post surplus.

Compared to complex tasks, efforts in basic tasks are contractible and it can be easily shown that the first-best level is achieved, irrespective of the organizational choice:

$$q^o(c) = q^I(c) = a\rho c^{-\sigma} \quad (2)$$

Note that, in this simple framework, efforts are not directly affected by financial constraints for a given organizational structure. Moreover, efforts in complex and basic tasks do not interact and do not depend on the proportion of complex tasks.<sup>11</sup>

Given these expressions of the level of effort, computing ex-post revenues for each party is a straightforward exercise. The supplier's ex-post revenues are strictly lower under integration than under outsourcing. This is due to the smaller share of revenues obtained by the supplier in the bargaining stage and the lower value of production (efforts are lower under integration). However, the multinational's ex-post revenues are not necessarily lower under integration as the multinational can capture a larger share of production value.<sup>12</sup>

### Initial transfers: co-financing versus licensing fees

The initial transfer  $T$  plays a key role. In our set-up,  $T$  is defined as a payment from the supplier to the multinational firm. Yet this payment can be negative, reflecting a monetary flow from the multinational to the supplier. In analytical terms,  $T$  is determined by the participation constraint (PC) and the financial constraint (FC), which impose an upper bound on the level of the ex-ante transfer  $T$ .

When the financial constraint is slack, the ex-ante transfer is directly determined by equality in the participation constraint. This equals the supplier's ex-post revenues net of the costs of all tasks, and can be expressed a function of marginal cost  $c$  and complexity  $\theta$ . Under outsourcing, we obtain:

$$\bar{T}^o(\theta, c) = a\theta \left( \frac{1-\rho}{2} \right) (2c)^{1-\sigma} - a(1-\theta)\rho c^{1-\sigma}$$

Under integration, the lower level of effort and the smaller share of revenues secured by the upstream firm call for lower ex-ante compensation to the multinational firm:

$$\bar{T}^I(\theta, c) = \delta^\sigma a\theta \left( \frac{1-\rho}{2} \right) (2c)^{1-\sigma} - a(1-\theta)\rho c^{1-\sigma} < \bar{T}^o(\theta, c)$$

Both  $\bar{T}^o$  and  $\bar{T}^I$  can be either positive or negative and increase with the degree of complexity  $\theta$ :

<sup>11</sup>The baseline model's main findings are robust to alternative structures. The first extension of the model in Section (2.3) allows for a positive outside option for the supplier. In some cases, where outside options are large, investments in complex efforts are constrained by available liquidity and do not maximize the supplier's ex-post revenues. The last extension of the model examines a more general framework and discusses a generalization of our results when efforts in complex and basic tasks are interdependent.

<sup>12</sup>The multinational's ex-post revenues are higher under integration when  $\rho < 1/2$  and  $\delta$  is close enough to one.

- When production relies essentially on basic tasks ( $\theta = 0$ ), the multinational firm can finance all of the supplier's initial costs because there is no hold-up problem and it is able to capture all ex-post revenues. In this case,  $T$  is strictly negative, reflecting an initial transfer from the multinational firm to the supplier (co-financing).
- Conversely, when production is complex and  $\theta$  is closer to one, the supplier retains a larger share of the ex-post surplus. Instead of co-financing, positive ex-ante transfers are used to extract the supplier's rent when it exceeds the cost of initial investments (when the production technology is new and specific to the multinational, this payment can be interpreted as a licensing fee).
- The more complex production is, the greater the difference between outsourcing and integration in terms of optimal transfer (ownership only affects revenues and investments related to complex tasks).

At this point, it is quite intuitive that the supplier's access to finance will have a larger impact when there is a larger proportion  $\theta$  of complex and non-contractible tasks. Moreover, financial constraints are less likely to affect the ex-ante transfer under integration compared to outsourcing, especially for complex tasks.

When are financial constraints binding? The financial constraint (FC) imposes an upper bound on the ex-ante transfer, depending on the debt from local banks and available liquidity. Formally, the financial constraint (FC) is binding when the optimal transfer  $\bar{T}^Z(\theta, c)$  does not satisfy the inequality in (FC). This occurs when initial capital is below a threshold  $\bar{W}^Z(\theta, c, \kappa)$ .

Under outsourcing, this threshold is:

$$\bar{W}^O(\theta, c, \kappa) = \frac{1-\kappa}{2} a\theta (2c)^{1-\sigma} \quad (3)$$

The possibility of financial constraints being binding arises when the threshold is strictly positive, i.e. when the production of the input involves complex tasks ( $\theta > 0$ ) and the financial markets are imperfect ( $\kappa < 1$ ). The higher the complexity  $\theta$  and the lower the level of financial development  $\kappa$ , the greater the likelihood of the financial constraint being binding, for a given  $W$  – this is consistent with prior remarks on the level of optimal transfer.

Under integration, we obtain a smaller threshold:

$$\bar{W}^I(\theta, c, \kappa) = \delta^\sigma \bar{W}^O(\theta, c, \kappa) < \bar{W}^O(\theta, c, \kappa) \quad (4)$$

Financial constraints are less likely under integration because the multinational can retain a larger share of ex-post revenues and asks for less compensation. In addition, the value of production is lower than under outsourcing.

Both wealth thresholds under integration and outsourcing decrease with the level of financial development:

$$\frac{\partial \bar{W}^Z}{\partial \kappa} < 0 \quad \text{for } Z \in \{I, O\}$$

for  $Z \in \{I, O\}$ . Intuitively, financial constraints are less likely to be binding for suppliers in financially developed countries. This effect, however, depends on the type of input that is supplier. We obtain the following proposition:

**Proposition 1.** (i) *The effect of financial development on the wealth threshold  $\overline{W}^Z(\theta, c, \kappa)$  is stronger when production involves complex inputs:*

$$\frac{\partial^2 \overline{W}^Z}{\partial \kappa \partial \theta} < 0 \quad \text{for } Z \in \{I, O\}$$

(ii) *The threshold  $\overline{W}^Z(\theta, c, \kappa)$  is higher under outsourcing:  $\overline{W}^O(\theta, c, \kappa) > \overline{W}^I(\theta, c, \kappa)$ .*

In words, financial constraints are more likely to be binding for a low level of financial development, especially when production involves complex inputs. Moreover, financial constraints are more likely to be binding under outsourcing, compared to integration.

When the financial constraint is binding, the participation constraint (PC) is slack and the initial transfer  $T$  is determined by the constraint (FC). Hence, when  $W < \overline{W}^O(\theta, c, \kappa)$  under outsourcing, initial transfer  $T$  depends on initial capital  $W$  and the level of financial development  $\kappa$ :

$$T^O(W, \kappa, \theta, c) = W + a\theta \left( \frac{\kappa - \rho}{2} \right) (2c)^{1-\sigma} - a(1-\theta)\rho c^{1-\sigma} < \overline{T}^O(\theta, c)$$

Similarly, under integration we obtain:

$$T^I(W, \kappa, \theta, c) = W + \delta^\sigma a\theta \left( \frac{\kappa - \rho}{2} \right) (2c)^{1-\sigma} - a(1-\theta)\rho c^{1-\sigma} < \overline{T}^I(\theta, c)$$

Naturally, when both initial capital  $W$  and financial development  $\kappa$  are low, we observe a lower transfer  $T$ . When  $T$  becomes negative, it reflects the extent of co-financing when the supplier has little access to finance. Moreover,  $T$  is more likely to be negative under integration, in which case  $T$  can be interpreted as foreign direct investment. This result is consistent with Desai, Foley and Hines (2004) who find larger financial participation by multinationals in affiliates located in countries with poor financial markets.

### Multinational firm's profits

In the benchmark case where  $W$  is large and capital markets are frictionless, the financial constraint (FC) is slack. The only restraint on the actions of the multinational firm is the necessity to ensure the supplier's participation (ensuring that the supplier gets non-negative profits). In this case, the multinational is able to extract the entire surplus from the relationship and total profits equal total revenues  $Y$  minus total costs  $C$ :

$$\overline{\Pi}^O(\theta, c) = a\theta \left(1 - \frac{\rho}{2}\right) (2c)^{1-\sigma} + a(1-\theta)(1-\rho) c^{1-\sigma} \quad (5)$$

$$\overline{\Pi}^I(\theta, c) = \delta^{\sigma-1} a\theta \left(1 - \frac{\delta\rho}{2}\right) (2c)^{1-\sigma} + a(1-\theta)(1-\rho) c^{1-\sigma} \quad (6)$$

Note that:  $\overline{\Pi}^O(\theta, c) > \overline{\Pi}^I(\theta, c)$ . In this case, the preferred organizational form is the one that maximizes non-contractible efforts by the upstream unit. Ex-post marginal benefits for the supplier are maximal when it remains independent, making outsourcing the optimal choice.

When the financial constraint is binding, the multinational is unable to extract the entire surplus. Profits equal the sum of ex-post revenues and initial transfer  $T$ , which is affected by the lack of initial capital  $W$  or poor financial markets  $\kappa$ . Under outsourcing, the financial constraint is more likely to be

binding (when  $W < \bar{W}^o(\theta, c, \kappa)$ ), in which case we obtain:

$$\Pi^o(W, \kappa, \theta, c) = W + a\theta \left( \frac{1+\kappa-\rho}{2} \right) (2c)^{1-\sigma} + a(1-\theta)(1-\rho) c^{1-\sigma} \quad (7)$$

Under integration, financial constraints may also be binding (when  $W < \bar{W}^I(\theta, c, \kappa)$ ) and we obtain:

$$\Pi^I(W, \kappa, \theta, c) = W + \delta^{\sigma-1} a\theta \left( \frac{2-\delta+\delta\kappa-\delta\rho}{2} \right) (2c)^{1-\sigma} + a(1-\theta)(1-\rho) c^{1-\sigma} \quad (8)$$

When financial constraints are binding, profits depend on both  $\kappa$  and  $\theta$ . These profits are strictly lower than maximal profits when financial constraints are slack. Moreover, we can note that only revenues related to complex tasks are affected by financial constraints. Hence, the effects of financial constraints on the multinational firm's profits increase with the complexity of the tasks. Formally, we obtain the following proposition:

**Proposition 2.** *When the financial constraint is binding, the multinational firm's profits increase with the level of financial development of the country in which the supplier operates, especially when production involves complex tasks:*

$$\frac{\partial \Pi^Z}{\partial \kappa} > 0 \quad \text{and} \quad \frac{\partial^2 \Pi^Z}{\partial \kappa \partial \theta} > 0 \quad \text{for } Z \in \{I, O\}$$

Proposition 2 complements Proposition 1, which also emphasizes the interaction between  $\kappa$  and  $\theta$ , by highlighting how financial development and complexity affect the profitability of offshoring production. Financial development allows the multinational firm to recover ex ante the proportion of profits that are retained by the supplier. This proportion is higher the higher the complexity of production. Therefore, an increase in the level of financial development has a stronger impact on the multinational's profits when the intermediate good is more complex. Interestingly, financial development can have a negative impact on the supplier's profits because it reduces her rents, which are positive only when the financial constraint is binding.

### Financial constraints and organizational choice

We can now roll back the clock to the moment when the multinational firm makes its decision about the optimal organizational form. Vertical integration is preferred when profits under outsourcing are lower compared to profits under integration.

Outsourcing is preferred to integration when financial constraints do not bind:  $\bar{\Pi}^o(\theta, c) > \bar{\Pi}^I(\theta, c)$ , whatever  $c$  and  $\theta$ . As already discussed, ownership delivers greater incentives to the supplier, and the value of the joint surplus can be recovered by the multinational firm through ex-ante payments, which are maximal in this case.

However, the possibility of obtaining ex-ante compensation vanishes as financial constraints start to bind under outsourcing. Conversely, financial constraints are less likely to bind under integration. When the financial constraint under integration is not binding ( $W < \bar{W}^I(\theta, c, \kappa)$ ) but initial capital is below a threshold  $W^{I/O}(\kappa, \theta, c)$  defined implicitly by:

$$\Pi^o(W^{I/O}, \kappa, \theta, c) = \bar{\Pi}^I(\theta, c)$$

vertical integration becomes the preferred organizational form. By comparing expressions (5) and (7), we find the following expression for the threshold  $W^{I/O}$ :

$$W^{I/O}(\theta, c, \kappa) = a\theta (2c)^{1-\sigma} \left[ \delta^{\sigma-1} \left( 1 - \frac{\delta\rho}{2} \right) - \frac{1+\kappa-\rho}{2} \right] \quad (9)$$

Moreover, we can show that the financial constraint is not binding under integration when the multinational firm is indifferent between outsourcing and integration.<sup>13</sup> Formally, it implies that  $W^{I/O}$  lies between  $\overline{W}^I$  and  $\overline{W}^o$ :

$$\overline{W}^I(\theta, c, \kappa) \leq W^{I/O}(\theta, c, \kappa) < \overline{W}^o(\theta, c, \kappa)$$

When initial capital is below  $W^{I/O}$ , vertical integration is chosen because its costs in terms of productive efficiency are outweighed by the larger fraction of ex-post surplus that the multinational retains under integration. Vertical integration thus emerges as a device to alleviate the negative effects of the supplier's financial constraints.

As shown in Propositions 1 and 2, financial constraints are strongest when the level of financial development is low and the degree of complexity is high. Hence, we find that  $W^{I/O}$  is affected both by the level of financial development  $\kappa$  and the complexity of production  $\theta$ , and also by their interaction. This property is a key feature of our model:

**Proposition 3** *The effect of financial development on the organizational threshold  $W^{I/O}(\theta, c, \kappa)$  is negative and stronger when production involves complex inputs:*

$$\frac{\partial W^{I/O}}{\partial \kappa} < 0 \quad \text{and} \quad \frac{\partial^2 W^{I/O}}{\partial \kappa \partial \theta} < 0$$

The main intuition behind Proposition 3 is that vertical integration enables the multinational firm to retain a larger share of the ex-post surplus. Therefore, it reduces the need for ex-ante compensation and lessens the impact of the supplier's financial constraint when financial development is low and product complexity is high.

Note that vertical integration is more likely to arise in complex industries, because financial constraints are more likely to bind for complex goods:

$$\frac{\partial W^{I/O}}{\partial \theta} > 0$$

Lastly, notice that our results are not sensitive to hypotheses about the financial structure under integration. For various specifications about the availability of credit to the affiliated supplier, integration emerges in order to mitigate negative effects of financial constraints under outsourcing. For example, the same threshold  $W^{I/O}$  is obtained if we assume perfect internal capital markets in the case of vertical integration (no financial constraint). Similarly, Proposition 3 remains true if we consider that the supplier is not able to raise any debt in the case of integration, since it is just an employee of the affiliated firm.

<sup>13</sup>In extreme cases of high financial development with a sharp drop in the level of efforts under integration, integration does not arise regardless of  $W$ . Formally, this happens when  $\delta^{\sigma-1} \frac{2-\delta+\delta\kappa-\delta\rho}{2} < \frac{1+\kappa-\rho}{2}$ .



## 2.3 Robustness and generalizations

### Constrained investments

In the baseline model the outside option of the supplier is assumed to be nil even in the case where she retains residual property rights. In this extension, we relax this assumption and explore the case where the supplier can derive profits from investments in non-contractible tasks and components in the event of disagreement. To be more precise, we denote by  $\alpha > 0$  the proportion of profits from complex tasks that the supplier could secure if negotiations break down.<sup>14</sup>

At the bargaining stage, a positive outside option for the supplier provides a higher treat point. This implies that the share of ex-post surplus received by the supplier is increased and the share of ex-post surplus received by the multinational is reduced. Formally, the share of the value of complex tasks obtained ex post by the supplier equals  $\frac{1+\alpha}{2} > \frac{1}{2}$  (under symmetric Nash bargaining), whereas the multinational retains a fraction  $1 - \frac{1+\alpha}{2} < \frac{1}{2}$  after bargaining. When financial constraints are not binding, the multinational can recover the entire surplus by asking for adequate ex-ante transfers. Moreover, higher  $\alpha$  provides higher incentives to invest and thus increases the total value of the relationship. Nevertheless, a new mechanism appears in the case of binding financial constraints, by which efforts under outsourcing can be directly affected by the level of financial constraints.

When initial liquidity is scarce, financing higher levels of investments reduces the capacity of the supplier to make compensating ex ante transfers to the multinational. As investments are increasing in the value of outside options, they may become excessive from the multinational's point of view when liquidity is low. In such a situation, the multinational firm may find it optimal to restrict the supplier's available funds by asking larger initial transfer, at the cost of reducing the supplier's level of investment. We refer to this as a "constrained investments" situation.

Such a case arises when the level of investment that maximizes the multinational's total profits is lower than the level of investment that maximizes the supplier's ex post profits. Effort is constrained when the total share of revenues from complex tasks received by the multinational is lower than the share of these revenues secured by the supplier through bargaining. While the supplier's share of ex-post revenues equals  $\frac{1+\alpha}{2}$ , the multinational's total share of revenues is the sum of ex-post revenues and the share of profits that can be extracted ex ante through the debt contract  $\kappa(\frac{1+\alpha}{2})$ . We find that a constrained effort situation arises when:

$$1 - \left(\frac{1+\alpha}{2}\right) + \kappa\left(\frac{1+\alpha}{2}\right) < \frac{1+\alpha}{2}$$

which is equivalent to  $\kappa < \frac{\alpha}{1+\alpha}$ . This condition can be satisfied when the level of financial development is low and the supplier's outside option is large. However, when  $\alpha = 0$  (the special case described in the baseline model), this condition never holds.

This potential effect on investments does not alter the results from previous sections. In Appendix A, we derive multinational's profits  $\Pi^o$  and wealth thresholds  $\bar{W}^o$  and  $W^{I/O}$ . In particular, we find that the effects of financial constraints increase with  $\alpha$  because positive outside options exacerbate the hold-up problem from the multinational's point of view.

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<sup>14</sup>This parameter can be determined by technological conditions (i.e. the specificity of the production technology), or by characteristics of the environment where the supplier operates such as the presence of other downstream firms in the market (the "thickness" of the market), and the extent of intellectual property rights protection (the extent to which the non-licensed technologies can be used by the supplier for other purposes).

## Contingent debt repayment and bargaining

We now discuss the case in which debt repayment by the supplier is conditional on negotiations with the multinational being successful. This situation arises when the outside option of the supplier is nil and the only source of revenues is given by the outcome of the negotiations. Thus, in the event of disagreement, the supplier defaults on the debt.<sup>15</sup>

When this situation is internalized by both parties, it interacts with the bargaining process and yields different bargaining outcomes. The interesting point is that the debt  $L$  affects outside options and the surplus generated by the relationship, and thus affects ex post revenues. We find that the supplier's ex-post revenues are increased by  $\frac{L}{2}$  compared to the baseline model. Symmetrically, the multinational's ex-post revenues are decreased by  $\frac{L}{2}$ . As the debt cannot exceed  $\kappa$  times the supplier's ex-post revenues, we find that the maximum amount of debt that can be contracted is determined by:

$$\bar{L} = \kappa \left( \frac{1}{2} A\theta e^\rho + \frac{1}{2} \bar{L} \right)$$

We obtain that  $\bar{L} = \frac{\kappa}{2-\kappa} A\theta e^\rho$ , which is strictly higher than the maximum level of debt that can be raised in the baseline case and given by  $\frac{\kappa}{2} A\theta e^\rho$ .

Thus, contingent repayment affects the results in two ways: it induces a higher debt level but lower ex-post revenues for the multinational, especially given a high level of financial development. However, the impact of financial development on total profits remains positive, especially for complex inputs (we show in Appendix A that Propositions 1 to 3 still hold).

Interestingly, whether the debt is contracted with a local bank or the multinational firm matters in such a case. In the event of default, the multinational's position is worsened if it is the creditor, which implies that it is optimal to sign the debt contract with an external agent rather than between the supplier and the multinational.

## A general production function

In this last extension, we examine a more general setting. We now consider total revenues as a function  $Y(e, q)$  of non-contractible efforts  $e$  and contractible efforts  $q$ .<sup>16</sup> We assume that costs are linear in both types of efforts:  $C = c_e \cdot e + c_q \cdot q$ . As in the baseline model, the multinational firm's outside options under outsourcing correspond to revenues when non-contractible tasks equal zero,  $Y(0, q)$ , while the supplier's outside option is nil.

In this general framework, we find that optimal investments in complex tasks are determined by the supplier by a function  $e^*(q)$  of investments in basic tasks. In turn, investments in basic tasks are chosen by the multinational firm as a function  $q^*(\kappa, W)$  of financial development and initial capital  $W$ , by anticipating the supplier's choice. The reader is referred to Appendix A for the characterization of optimal choices in terms of non-contractible and contractible tasks.

How does financial development affect optimal choices and profits? First, by comparing both constraints, we can show that the financial constraint starts binding under outsourcing when initial capital is

<sup>15</sup>For simplicity, we assume in what follows that the supplier can default on the debt at no cost when the negotiation breaks down.

<sup>16</sup>We suppose that  $Y$  is increasing and concave in both types of efforts  $e$  and  $q$ . The baseline model would correspond to the case where  $\frac{\partial^2 Y}{\partial e \partial q} = 0$ .

below the threshold  $\bar{W}^o$  defined by:

$$\bar{W}^o = \left(\frac{1-\kappa}{2}\right) [Y(e, q) - Y(0, q)]$$

where  $q = q^*(\kappa, \infty) \equiv q^*$  and  $e = e^*(q^*)$  are optimal investments when the financial constraint is not binding. The term  $[Y(e, q) - Y(0, q)]$  reflects the extent to which revenues rely on complex tasks. As in the baseline model, this equality shows that financial development has a negative impact on the wealth threshold (financial constraints are more likely to bind for a low level of financial development) and that the effect is stronger when production relies on complex tasks (Proposition 1).<sup>17</sup>

When the supplier has no initial capital ( $W = 0$ ), only the financial constraint is binding. In this case, we can show that financial development has a positive impact on the multinational's profits and that this effect is more pronounced when production is more dependent on complex tasks (Proposition 2):

$$\frac{d\Pi^o}{d\kappa} = \frac{1}{2} [Y(e, q) - Y(0, q)] > 0$$

where  $q = q^*(\kappa, 0)$  and  $e = e^*(q^*(\kappa, 0))$ .

Similarly, we can analyze the choice between outsourcing and integration, assuming that the multinational only retains  $Y((1 - \delta)e, q)$  in the event of disagreement with the supplier. As in the baseline model, we obtain that outsourcing is preferred to integration unless financial constraints are strong and financial development is low. The magnitude of the effect however depends on whether both the participation constraint and the financial constraint are binding at the point where the multinational is indifferent between outsourcing and integration. Assuming that in this case the financial constraint is not binding under integration and the participation constraint is not binding under outsourcing, we obtain:

$$\frac{dW^{I/O}}{d\kappa} = -\frac{1}{2} [Y(e, q) - Y(0, q)] < 0$$

where  $q = q^*(\kappa, 0)$  and  $e = e^*(q^*(\kappa, 0))$ . This also suggests that the impact of financial development is strongest when production relies on complex tasks (Proposition 3).

## 2.4 Empirical predictions

The model's findings have direct implications in terms of optimal sourcing strategies. One of the main factors that lead firms to offshore production is the search for low labor costs, or low production costs in general. However, our model suggests that, while lower production costs point to choosing locations in developing countries, the profitability of this strategy could be affected by the financial capacity of suppliers and credit rationing by financial institutions. By emphasizing differential effects across industries, depending on the proportion of complex tasks to be undertaken by the supplier, our model shows that the effects of financial development are stronger in industries that are relatively more exposed to opportunism.

In the following, we illustrate some simple implications of the model in terms of sourcing strategies by considering the decision to import a product from a foreign country depending on the level of financial development, degree of complexity, and relative cost of producing the good in the foreign country.

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<sup>17</sup>However, the extent to which production depends on complex tasks is endogenous to both the technology and the financial constraints. It is not possible to derive simple comparative statics on an exogenous technological parameter  $\theta$  in the general case when  $\frac{\partial^2 Y}{\partial e \partial q} \neq 0$ .

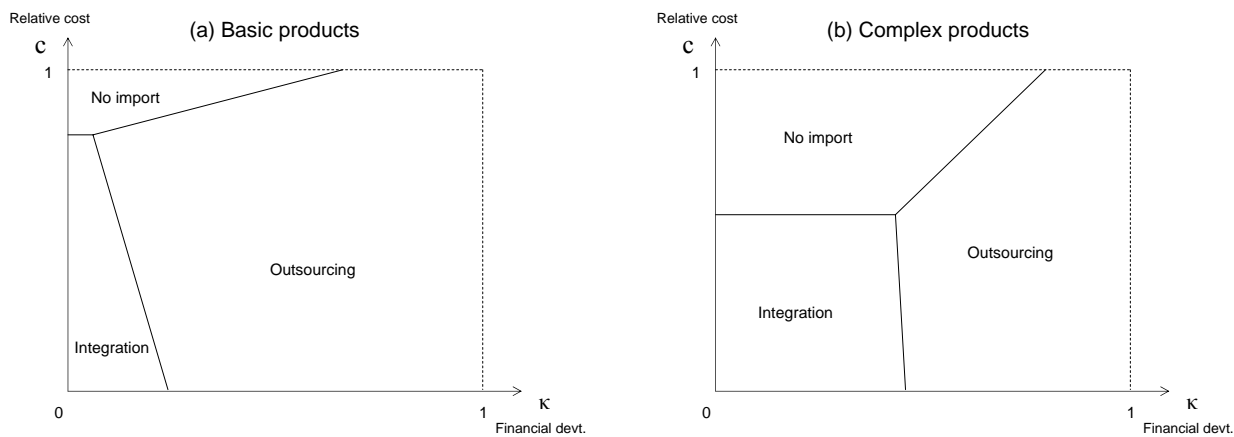
Formally, we suppose that the multinational has the choice between two suppliers: one at home and one in a foreign country. We assume an imperfect capital market in the foreign country  $\kappa < 1$  but no financial friction at home. We normalize to one the unit cost of the home supplier and denote by  $c < 1$  the relative cost of the foreign supplier's effort. Hence, the home supplier has the advantage of no financial friction whereas the foreign supplier has lower effort costs.

In this framework, the multinational firm can make one of three choices: (i) no import (outsourcing from home supplier);<sup>18</sup> (ii) outsourcing to a foreign supplier; (iii) sourcing from an integrated foreign supplier. Optimal strategies can be easily derived from expressions (5), (6) and (7). These choices are illustrated in Graphs 2a and 2b.

When the traded good relies mainly on basic tasks (Figure 2a), the foreign supplier's financial constraints are less likely to be binding and affect the multinational firm's profits (Propositions 1 and 2). Hence the multinational prefers to source inputs from abroad where production costs are lower even if the supplier has little access to finance.

Conversely, when production is intensive in complex tasks (Figure 2b), financial constraints are more likely to affect the multinational firm's profits (Propositions 1 and 2). When the foreign supplier is located in a country with poor financial institutions, the multinational will prefer to source inputs from home, unless there is a very large difference in production costs. Thus, we derive the following empirical prediction:

Figure 2: Sourcing mode and financial development



**Prediction 1** *Multinational firms are more likely to import inputs from a country with a higher level of financial development. This effect is stronger for complex inputs.*

By emphasizing differences across industries, our model shows that financial development generates a comparative advantage in the supply of complex inputs.

This prediction contrasts with previous theoretical studies (e.g. Beck, 2002; Manova, 2008; Becker and Greenberg, 2007), suggesting that financial development generates a comparative advantage in sectors depending more heavily on external finance. The notion of dependence on external finance (Rajan and

<sup>18</sup>As financial constraints are not binding for the home supplier, outsourcing is preferable. More general results could be obtained by assuming imperfect capital markets at home, but the main predictions would hold as long as the home supplier has better access to finance.

Zingales, 1998) is generally conceived as the need to finance initial costs before profits are generated, as opposed to sectors where initial costs are small and generated cash flows are sufficient to finance most investment. In our model, we account for the relationship between exporters and their customers who, for two-thirds of international trade flows, are multinational firms with broad access to liquidity. When production is intensive in basic tasks and there is no risk of hold-up from the supplier, financial development has little effect regardless of the level of initial costs.

Furthermore, our model can be used to make predictions about the sourcing mode: intra-firm or arm’s length. In the case of basic goods (Figure 2a), outsourcing is the preferred organizational form for a wide range of parameters. In the case of complex goods (Figure 2b), outsourcing is the optimal strategy only when the foreign supplier is located in a country with high financial development. When production costs are very low but capital markets are poorly developed, it may still be profitable to import, but integration is chosen in order to alleviate the impact of financial constraints.

**Prediction 2** *Intra-firm trade is more likely when the supplier is located in a country with a lower level of financial development. This effect is stronger for complex inputs.*

Last of all, note that our model makes the assumption that the supplier’s investments in basic and complex tasks are essential to the production of the intermediate good. Unlike Antras (2003 and 2005) and Antras and Helpman (2004 and 2008), there is no specific investment to be simultaneously implemented by the multinational: the technology is decided on before the supplier is contracted and the multinational transforms the intermediate good only after it is traded. Interestingly, our model can explain the emergence of vertical integration and intra-firm trade without the assumption that there are simultaneous efforts to be made by the multinational firm which are more important than the supplier’s efforts.<sup>19</sup>

### 3 Empirical Analysis

In this section, we first describe the data and then test the main predictions of the model using a detailed firm-level dataset on international operations by multinational firms in France. First we analyze how financial development differently impacts the number of multinationals sourcing complex or basic inputs across countries. Secondly, we study how the scale of imports in complex and basic goods responds to the level of financial development. Finally, we examine whether imports tend to occur within the firm depending on a country’s financial development and product complexity.

#### 3.1 Data description

##### SESSI Database

Our analysis is based on a firm-level survey in France by SESSI (French Office of Industrial Studies and Statistics). The survey covers firms that trade more than one million euros and belong to manufacturing groups holding at least 50% of a foreign affiliate’s equity capital.

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<sup>19</sup>If we introduce specific investments to be made by the multinational firm into our framework, we find that integration is always optimal in sectors that are intensive in investment by the parent company, irrespective of the extent of the supplier’s financial constraints. Financial constraints impact on the organizational form only in sectors that are intensive in supplier-specific investment.

The data provide trade flows detailed by firm, product and country of origin or destination. Products are classified using the four-digit CPA (Classification of Products by Activity) and four-digit harmonized system. However, we work at the three-digit level of the CPA classification which corresponds precisely to the NACE Rev1 classification (also closer to the ISIC Rev3 three-digit classification). The data cover 55% of total French imports, spread among 4,305 firms.<sup>20</sup> For our analysis, we focus solely on manufacturing imports and countries with available data on financial development and our main controls (Table 1 provides the list of countries corresponding to positive imports). We obtain a dataset on 3,957 importing firms, 52,617 cells by firm, input and country with positive imports.<sup>21</sup>

In addition to total value, the data provide information on the proportion of trade flows with an independent supplier or affiliate firm for each multinational, country and product. By definition, the trading partner is considered to be a subsidiary when the multinational controls at least 50% of the equity capital. This can be compared to US Census and Customs data, which consider the partner to be a “related party” when the multinational has 6% of the shares but does not necessarily control the firm. The data show that around half of French imports are intra-firm, which is approximatively the same as the share of intra-firm imports in the US.<sup>22</sup>

For the purpose of our analysis, a series of indicators by country and industry completes the SESSI survey. Table 3 provides the mean and standard deviations of the main variables used in our analysis.

## Country variables

Our main explanatory variable is the level of financial development, measured by the amount of credit from banks and other financial institutions to the private sector as a share of GDP (private credit). This index reflects the financial depth of the economy and the capacity to provide external financing. The variable is drawn from Beck *et al.* (2000). It ranges from 1.3% for Angola to 201% for Japan (see Figure 1). It is highly correlated with alternative measures of financial development (stock market capitalization, accounting standards, interest rate margin, bank regulation, investor protection, etc.). It has been used extensively as a measure of financial development (Rajan and Zingales, 1998; King and Levine, 1993). Alternatively, we use the net interest margin as an alternative measure of the efficiency of the banking sector (source: Beck *et al.*, 2000). We also use an index of accounting standards (La Porta *et al.*, 1998) as an example of a measure more directly related to regulations.

We complete our set of country variables with a number of controls. One important control is judicial quality (“rule of law”), reflecting the contracting environment. The rule of law index is drawn from Kaufmann, Kraay and Mastruzzi (2003). According to Levchenko (2007) and Nunn (2007), the rule of law provides a comparative advantage in industries that are more intensive in contracts. As described by Antras and Helpman (2008), it can also affect the sourcing mode. However, judicial quality is also one of the determinants of financial development, as better enforcement of creditor rights reduces the risk of default. This point is reflected by the strong correlation between both indices (correlation above 0.5).

In addition to the rule of law, we control for intellectual property rights, in particular patent protection

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<sup>20</sup>The survey covers 82% of trade flows involving multinational firms operating in France.

<sup>21</sup>Indices on skill and capital intensity are not available for food and beverages (ISIC 15). Note that our findings are not sensitive to the inclusion of this sector when these variables are excluded from the regressions.

<sup>22</sup>Note also that the quality of the SESSI survey has been checked by statistical offices using data from other sources. The trade flows are consistent with customs data and the intra-firm trade flows consistent with data on the location of the French affiliate (INSEE Financial Links Survey (LIFI), Bank of France and French General Treasury and Economic Policy Directorate (DGTPE) data).

as taken from Ginarte and Park (1997, updated through to 2005). It can affect both sourcing mode and destination: firms might be loathe to outsource intermediate goods with high technological content when IPR are not enforced. In the analysis of sourcing modes, we also use the index of FDI restrictions developed by the Heritage Foundation for the year 2000. In keeping with traditional theories of trade, we need to control for physical capital endowments and skilled labor endowments, as drawn from Hall and Jones, 1999 (we obtain similar results using endowments from Antweiler and Treffer, 2002). Lastly, we use a variable constructed by Do and Levchenko (2007), which predicts the level of financial development based on geographical determinants. In particular, the authors show that the external finance need of exports (EFNX) predicted by the country’s geographical characteristics (e.g. distance to other countries) has a strong effect on the level of financial development. We use this variable as a control.

Table 1: List of countries

Countries trading with multinational firms in France, with available data for private credit and other controls included in the benchmark regression (column 1 of Table 5):

Algeria	Finland	Korea, South	Saudi Arabia
Angola	Gabon	Madagascar	Senegal
Argentina	Germany	Malaysia	Sierra Leone
Australia	Ghana	Malta	Singapore
Bangladesh	Greece	Mauritius	South Africa
Benin	Guatemala	Mexico	Spain
Bolivia	Guyana	Morocco	Sri Lanka
Brazil	Haiti	Mozambique	Sweden
Burkina Faso	Honduras	Netherlands	Switzerland
Cameroon	Hong Kong	New Zealand	Syria
Canada	Hungary	Nicaragua	Tanzania
Central African Republic	Iceland	Niger	Thailand
Chad	India	Nigeria	Togo
Chile	Indonesia	Norway	Trinidad and Tobago
Colombia	Iran	Pakistan	Tunisia
Congo	Ireland	Panama	Turkey
Costa Rica	Israel	Papua New Guinea	Uganda
Cyprus	Italy	Paraguay	United Kingdom
Denmark	Ivory Coast	Peru	United States
Ecuador	Jamaica	Philippines	Uruguay
Egypt	Japan	Poland	Venezuela
El Salvador	Jordan	Portugal	Zambia
Fiji	Kenya	Romania	Zimbabwe

## Industry variables

Our main measure of product complexity is an index of R&D intensity. Since we focus on French firms’ offshoring decisions, we measure R&D intensity for the corresponding French industries. The use of R&D intensity indices on French industries implicitly assumes either that the technology can be transferred to the supplier (and is thus similar to the technologies in the French industries) or that industry ranking in terms of technological content varies little from one country to the next. Firm-level data on R&D expenditure are taken from the 1999 Third Community Innovation Survey (CIS 3). Indices are constructed for each three-digit NACE industry with at least ten observations. We can use three measures: (i) total industry R&D

expenditure to total sales; (ii) the percentage of firms in the industry with positive R&D expenditure; (iii) the 75th percentile of R&D expenditure to sales for firms in the industry. We focus on the latter because it is less sensitive to measurement errors, but we checked that all measures yield similar results.<sup>23</sup> Table 2 provides the list of industries associated with the highest R&D intensities. Not surprisingly, computers, precision instruments and aircraft correspond to the most complex industries. In opposition, few industries have a zero index of R&D intensity (e.g. cutting and shaping of stone, manufacture of wooden containers, preparation and spinning of textile fibres).

Table 2: Industries with the highest indices of R&D intensity (France)

	NACE Industry classification	R&D int.
244	Manuf. of pharmaceuticals, medicinal chemicals and botanical products	5.2%
321	Manuf. of electronic valves and tubes and other electronic components	5.8%
323	Manuf. of television and radio receivers, sound or video recording or reproducing apparatus	6.2%
322	Manuf. of television and radio transmitters and apparatus for line telephony and telegraphy	7.3%
353	Manuf. of aircraft and spacecraft	8.9%
332	Manuf. of instruments and appliances for measuring, checking, testing and navigating	9.7%
300	Manuf. of office machinery and computers	10.5%

To make sure that we are not picking up on a French industry particularity generated by specific technologies, regulations or limited competition, we use R&D intensity constructed in the same way from CIS data on firms in the United Kingdom. The French and English indices are correlated at 57% (the Spearman rank correlation is 63%).

For our robustness checks, we take the Lall (2000) classification of industries. There are five main classes: high-tech, medium-tech, low-tech, resource-based and primary goods. We eliminate primary and resource-based industries from our analysis because they can be strongly driven by natural endowments, as reported by Lall (2000). Moreover, some resource-based industries may also involve skilled labor and advanced technologies, which make them hard to compare with low-tech and high-tech industries. The index is initially constructed using the three-digit SITC Revision 2 classification. Given the correspondence between SITC and NACE industries, the new classification clearly differentiates medium-tech industries from low-tech industries, but not from the few high-tech industries. We thus use an index that equals one for high-tech and medium-tech industries and zero for low-tech industries.

Alternatively, we also use the Rauch (1999) classification of products. In Rauch (1999), products are either differentiated, traded on organized exchanges or reference priced. The latter two categories are often referred to as homogeneous goods and the former as heterogeneous goods. We thus construct an index that equals one for differentiated goods and zero otherwise.<sup>24</sup> This index can be used to some extent to differentiate complex from basic products, but it is often interpreted as a measure of contract intensity: differentiated products (which are neither traded on organized exchanges nor reference priced) cannot be easily substituted and are relationship-specific. Taking this interpretation, Nunn (2007) shows that better contract enforcement, measured by the rule of law variable, generates a comparative advantage in industries that rely intensively on differentiated inputs. As we focus on the comparative advantage in

<sup>23</sup>The 75th percentile is preferred to the median to reduce the number of zeros (industries where over half the firms do not do R&D).

<sup>24</sup>As in Nunn (2007), we follow the “liberal” classification as opposed to the “conservative” classification (see Rauch, 1999).



Table 3: Mean and standard deviation of main variables

Variables	Mean	Std. deviation	Obs.
<i>- Aggregate values:</i>			
Number of importers	8.288	39.00	6348
Log(1 + Number of importers)	0.684	1.259	6348
Log aggregated imports	7.901	2.728	2021
<i>- Firm-level values:</i>			
Log firm-level imports	4.984	2.196	52617
Share of intra-firm trade	0.338	0.453	52617
<i>- Across countries:</i>			
Private Credit over GDP	0.509	0.467	92
Net interest margin	5.633	4.106	81
Accounting standards	6.078	1.399	37
Predicted demand in external finance (Pred EFNX)	27.03	0.995	91
Patent Protection	3.128	1.010	92
Rule of law	0.552	0.209	92
Skill endowment	0.621	0.284	92
Capital endowment	9.378	1.535	92
Log GDP per capita	9.377	1.136	89
FDI restrictions	9.377	1.136	89
<i>- Across industries</i>			
R&D intensity (France)	0.021	0.023	69
R&D intensity (UK)	0.010	0.013	58
Lall (2000) index	0.615	0.491	52
Rauch (1999) index	0.738	0.344	69
External financial dependence	0.237	0.227	69
Skill intensity	4.992	0.153	69
Capital intensity	5.290	0.604	69
Median firm size	3.989	0.440	69
Headquarter intensity	0.380	0.062	69

Notes: “aggregate values”: by country and product; “firm-level values”: by firm, country and product.

supplying inputs rather than final goods, a notable difference with Nunn (2007) is that we directly interact the rule of law variable with the Rauch index using the traded input’s product classification.<sup>25</sup>

A great deal of attention has been paid to the interaction between financial development and dependence on external finance since Rajan and Zingales’ (1998) work on financial dependence and growth. The index of financial dependence we use is drawn from Klapper, Leaven and Rajan (2007). It is constructed using the ISIC Revision 3 classification from Compustat data on US firms, in keeping with the Rajan and Zingales (1998) methodology. Financial dependence is defined as one minus the ratio of cash flow to investment, taking the average across years for a given firm (the 1990s) and the median across firms in the sector. The same index has been used by Beck (2003) and Manova (2006, 2008) based on the ISIC Revision 2 classification.

Skill and physical capital endowments are interacted with indices of skill and capital intensity. Assuming perfect labor markets and homogeneous firms, differences in skill intensities across industries are reflected by differences in wages. Using the Annual Manufacturing Survey on French firms (*Enquête Annuelle d’Entreprises*) which covers all firms with more than 20 employees, we construct our index of skill intensity

<sup>25</sup>Nunn (2007) focuses on comparative advantage in terms of output. He thus constructs an alternative index equivalent to the Rauch (1999) index weighted by the coefficient of the US input-output matrix.

as the median (in log) of the distribution of average wages across firms.<sup>26</sup> Similarly, we use the median (in log) of the distribution of the capital labor ratio across firms in the industry as a proxy for capital intensity. Notice that capital intensity could also be interpreted as a proxy for the size of fixed costs and headquarter intensity (Antras, 2003).

In addition, we use the Annual Manufacturing Survey to construct two additional indices. For each industry, we take the median of firm size in the industry (in log) as well as the median of value added over sales. The latter is commonly interpreted as an index of vertical integration or headquarter intensity.

## 3.2 Financial development and the geography of imports

### Empirical strategy

The model predicts that, other things being equal, multinational firms prefer to import complex inputs from financially developed countries (Prediction 1). This points to a differential effect between complex and basic goods and naturally calls for a difference-in-difference approach (pioneered by Rajan and Zingales, 1998).

In the model, complex goods are defined as goods with a large range of tasks that are firm-specific and cannot be described in a contract. Complexity lies at the heart of the incomplete contracts literature, as established by the seminal work of Grossman and Hart (1986).<sup>27</sup> In this regard, our preferred measure is R&D intensity for several reasons. First, R&D is associated with the production of complex products and processes. This makes codifiability and verifiability of information a much harder and more expensive task. Hence, firms in R&D-intensive sectors are more likely to be unable to describe all specifications in contracts. Second, a R&D intensive product is often new and not standardized, which makes it difficult to specify in advance. A firm may also be reluctant to do so for strategic reasons.<sup>28</sup> Moreover, one of the main motives for R&D is to accumulate specific knowledge helping firms to create differentiated products. Therefore, R&D intensity may also reflect product specificity. Lastly, R&D may not be performed in order to create new products or processes, but to adapt existing technologies to different contexts (Cohen and Levinthal, 1989). Even in this case, higher R&D intensity reflects the difficulty to assimilate the technology and closely relates to efforts in complex tasks in the model.

Other indices of complexity can be used. A second index of technological content is that constructed from Lall (2000) classification, which differentiates between low-tech versus medium- and high-tech goods. A third index is drawn from Rauch (1999). It is somewhat different and essentially captures product specificity in terms of pricing. Non-specific products are those that can be traded on integrated markets or referenced. Those which are not referenced are considered to be specific to one firm or a small number of firms and cannot be priced uniformly. In our empirical analysis, we mainly focus on R&D intensity and consider these two alternative measures as robustness checks.

Table 4 presents the basic statistics drawn from the data, revealing strong correlations between the country of origin's level of financial development (private credit) and the number of firms importing from this country. The correlation is stronger for R&D-intensive goods, in line with the differentiated effect

<sup>26</sup>The Annual Manufacturing Survey does not provide a skills breakdown.

<sup>27</sup>The model developed by Grossman and Hart (1986) is based around the key assumption that "[production decisions]...are sufficiently complex that they cannot be specified completely in an initial contract between the firms". In later work, Hart and Moore (1999) and Segal (1999) develop this idea further by providing theoretical foundation to incomplete contracts, around the notions of complexity and renegotiation.

<sup>28</sup>Specification of the characteristics of new product and processes makes them easier to imitate as codification facilitates technological leakage at the innovator's expense (Caves, 2007).

described in the model: over 50% for products with high R&D intensity (above the median) compared to 38% for other products. Conversely, the correlation between the number of importers and level of complexity is stronger for countries above the financial development median than for countries below the median. Similar differences in correlation coefficients are observed when looking at aggregated imports instead of the number of importers. These statistics suggest that financial development impacts on the composition of exports to multinational firms depending on R&D intensity.

Table 4: Correlation between import variables and financial development (A) or product complexity (B)

Correlation:	A) With Financial development (across countries)		B) With R&D intensity (across industries)	
	High R&D int.	Low R&D int.	High fin devt	Low fin devt
Sample:				
Log(1 + Nb importers)	0.525	0.384	0.134	0.031
Log(Imports)	0.344	0.176	0.159	-0.024
Share of intra-firm trade	0.058	0.138	0.105	0.154

*Notes:* correlation between import variables (number of importers, import by firm, intra-firm imports) and either private credit (Panel A) or R&D intensity (Panel B). Samples: “high R&D intensity”: imports of products with R&D intensity above median; “high financial development”: imports from countries with financial development above median.

In the following econometric investigation, we analyze the determinants of the number of firms  $N_{cs}$  that import inputs from country  $c$  in sector  $s$  (based on classification  $s$  of imports). In particular, we want to test whether the financial development of country  $c$  has a differential effect depending on the complexity of traded good  $s$ . The difference-in-difference approach includes country dummies that control for any observable or unobservable country characteristics that do not affect the export’s level of complexity. It also includes product dummies to control for product characteristics that do not affect the origin of the imports.

Ordinary Least Square (OLS) regressions are not suitable for count data, especially when a significant number of industry/country cells have no importer (two-thirds of our observations).<sup>29</sup> We model the distribution of sourcing choices using the Negative Binomial distribution (NegBin II) described in Cameron and Trivedi (1998). Negative Binomial differs from the Poisson distribution in that it accounts for overdispersion of the residual and often provides a better fit (it encompasses the Poisson distribution as a particular case). Moreover, we can test and check that specifying a Poisson distribution is not appropriate.

We assume that firms’ choices depend on a set of variables  $X_{cs}$ . These variables include a set  $\tilde{X}_{cs}$  of observed industry and country characteristics, but we also allow for unobserved determinants. Specifically, we suppose that the probability that a firm imports a product  $s$  from country  $c$  is an exponential function of country fixed effects, industry fixed effects, an interaction term between financial development and input complexity, a set of controls and the unobserved effect. Taking the aggregate number of importers by country and product, it follows that:

$$E[N_{cs}|X_{cs}] = \exp[\beta.(FinDev_t * Complexity_s) + \gamma.Controls_{cs} + \eta_c + \zeta_s + \lambda_{cs}] \quad (10)$$

where  $FinDev_t$  is the level of financial development of country  $c$ ,  $Complexity_s$  the level of complexity of product  $s$ ,  $\eta_c$  a country fixed effect,  $\zeta_s$  an industry fixed effect and  $\lambda_{cs}$  an unobserved effect.

<sup>29</sup>Santos Silva and Teneyro (2006) find OLS inconsistency may also apply to traditional gravity equations.

Conditionally on the mean  $E[N_{cs}|X_{cs}]$ , the number of importers follows a Poisson distribution. The Poisson distribution can be justified by taking the limit of a Bernoulli law for a large number of firms. However, conditionally on the set of observed characteristics  $\tilde{X}_{cs}$ , the distribution deviates from the Poisson distribution because of the unobserved component  $\lambda_{cs}$ . Assuming that this term is an i.i.d. random effect following a gamma distribution with coefficient  $1/\alpha$  (with variance  $\alpha$ ), it can be shown that the number of importers follows a Negative Binomial distribution (see Cameron and Trivedi, 1998) with law:

$$P(N_{cs} = n|\tilde{X}_{cs}) = \frac{\Gamma(n + \alpha^{-1})}{n! \Gamma(\alpha^{-1})} \frac{\alpha^{-(1/\alpha)} \mu^n}{(\mu + \alpha^{-1})^{n+\alpha^{-1}}}$$

where  $\mu = E[N_{cs}|\tilde{X}_{cs}]$ . Moreover, the conditional variance is given by  $Var[N_{cs}|\tilde{X}_{cs}] = \mu(1 + \alpha\mu)$ . When the random effect  $\lambda$  has zero variance ( $\alpha = 0$ ), the variance of  $N_{cs}$  equals the (conditional) mean and the distribution is Poisson. When  $\alpha$  is strictly positive, the distribution is said to be overdispersed.

Equation (10) is estimated under Maximum Likelihood (see Cameron and Trivedi, 1998). This can also be used to estimate  $\alpha$ , assuming that it does not depend on the expected mean  $\mu$  (NegBin II specification).

As in OLS regressions, the identification of our main coefficient relies on the assumption of orthogonality between the interaction term and the residual  $N_{cs} - E[\hat{N}_{cs}|\tilde{X}_{cs}]$ . However, given the inclusion of the full set of country and industry dummies, this estimation strategy is not sensitive to the exclusion of controls that are not correlated with financial development or do not affect trade differently depending on the complexity of inputs. Therefore, we are mainly concerned with variables potentially correlated with financial development and that may impact on complex goods industries in particular.

In all the tables, we report robust standard errors corrected for clusters by country because our main variable  $FinDevt_c$  only varies across countries. The coefficients can be interpreted as semi-elasticities, as they measure the impact of the right-hand-side variable on the log of the (expected) number of importers.<sup>30</sup>

## Results on the number of importers

In Table 5, we analyze the interaction term between complexity and financial development for different measures of complexity, including our main controls. In column (1), financial development is interacted with our benchmark index of complexity: R&D intensity measured for French industries. This yields a positive and significant coefficient, a result in line with Prediction 1: financial development is a source of comparative advantage in the supply of complex inputs.

Our control variables include capital and skill endowments interacted with capital and skill intensity. They enter positively and significantly, in keeping with the Heckscher-Ohlin model of trade. Two other important controls, rule of law and patent protection, are positively correlated with financial development and may potentially affect sourcing strategies. In line with the recent literature on institutions and trade (Nunn, 2007; Levchenko, 2007; Acemoglu, Antras and Helpman, 2007), judicial quality and better contract enforcement generate a comparative advantage in complex industries that rely more intensively on contracts. In our regression, rule of law is interacted with the Rauch (1999) index of differentiated goods versus homogeneous goods, often interpreted as an index of contract intensity. We find a positive interaction term, which confirms the results obtained by Nunn (2007). Patent protection, our proxy for

<sup>30</sup>The interaction term should be interpreted as a cross derivative of the logarithm of the predicted number of importers, which is linear in the right-hand-side variables. Note also that our fixed effects subsume all direct effects of country- and industry-level variables, and thus we do not need to control for direct effects.

intellectual property rights, is interacted with the index of R&D intensity, as R&D-intensive industries rely more intensively on patents. However, we do not find that stronger intellectual property rights favor trade in R&D-intensive inputs.

Potential problems may affect our main measure of R&D intensity. Given that it is computed from data on French industries, it may capture some particularities among the French firms in these industries, which could also impact on trade. For example, the distribution of firm size, imperfect competition and specific regulations may simultaneously affect R&D investment and sourcing strategies. To avoid these potential biases, we use an index of R&D intensity measured from similar data on firms in the UK. In column (2), we obtain a strong positive coefficient with the UK index. The coefficient is higher, but once it is normalized by the respective standard deviation (see Table 3), the amplitude of the latter estimation is slightly lower compared to column (1).

Our result is also robust to the alternative measures of complexity. In column (3), financial development is interacted with the Lall (2000) index of high-/medium-tech inputs versus low-tech inputs. We find a positive and significant coefficient, showing that countries with a higher level of financial development are more likely to supply multinational firms with high- and medium-tech inputs. In this case, the index of patent protection is also interacted with the Lall index and yields a stronger coefficient compared to the other columns. In column (4), we use the Rauch (1999) index of differentiated goods versus homogeneous goods (contract intensity), which also yields a positive and significant coefficient for the interaction term. Interestingly, the rule of law coefficient becomes insignificant, suggesting that financial development has the strongest impact.<sup>31</sup>

Beck (2003) and Manova (2006 and 2008) use world trade data by country and sector to show that financial development generates a comparative advantage in sectors that are more dependent on external finance. These studies use the measure of financial dependence developed by Rajan and Zingales (1998), constructed as the share of investment not financed by cash flows, taking the median across US listed firms in the sectors. Our model, however, suggests that the dependence of multinationals' suppliers on external finance is specifically related to contract incompleteness and increases with their production complexity. As both indices of financial dependence and technological complexity are positively correlated (at 20%), we need to check that our results are not driven by the Rajan and Zingales (1998) index of financial dependence. In column (5), we therefore control for the interaction term between financial development and financial dependence. We also control for any interactions between financial development and capital intensity and firm size (in the corresponding sector in France). Both indices may depend on the size of fixed costs and be positively correlated with the index of R&D intensity (larger firms are more likely to invest in R&D). Alternatively, capital intensity may be interpreted as an index of headquarter intensity (Antras, 2003). Another index of headquarter intensity (value added over sales across industries) is also included into the regression. In line with our model, only the interaction term between financial development and complexity is positive and significant while the other interaction terms taking in financial development are not statistically significant. Qualitatively, the same results are obtained using the Lall (2000) and Rauch (1999) indices instead of R&D intensity.

In Table 6, we perform further robustness checks. In column (1), we drop the controls for rule of law and intellectual property rights, and compare the coefficient to column (1) of Table 5. Although both variables are positively correlated with financial development, we check that their inclusion or exclusion

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<sup>31</sup>This result, however, needs to be interpreted with caution as the correlation between both indices of financial development and rule of law is over 50%.

Table 5: Imports, product complexity and financial development

Dependent variable:	Number of importers, with $E[N] = \exp(bX)$				
	(1)	(2)	(3)	(4)	(5)
(Fin devt) * (R&D intensity)	9.417 [3.263]***	16.30 [5.140]***			9.586 [3.480]***
(Fin devt) * (Lall index)			0.357 [0.150]**		
(Fin devt) * (Rauch index)				0.682 [0.326]**	
(Fin devt) * (Ext fin dep)					-0.337 [0.251]
(Fin devt) * (Capital intensity)					-0.330 [0.184]*
(Fin devt) * (Firm size)					-0.110 [0.121]
(Fin devt) * (HQ intensity)					0.985 [0.733]
(Patent prot) * (Lall index)			0.301 [0.101]***		
(Patent prot) * (R&D intensity)	-0.228 [2.004]	-2.859 [4.408]		1.677 [1.979]	-0.203 [1.986]
(Rule of law) * (Rauch index)	2.121 [0.609]***	1.863 [0.580]***	0.836 [0.515]	1.074 [0.914]	1.794 [0.620]***
(H/L) * (Skill intensity)	4.261 [1.151]***	3.267 [1.262]***	3.706 [1.195]***	4.561 [1.207]***	4.682 [1.195]***
(K/L) * (Capital intensity)	0.221 [0.065]***	0.281 [0.067]***	0.290 [0.078]***	0.207 [0.063]***	0.296 [0.067]***
$\alpha$	0.532	0.549	0.455	0.542	0.516
Country dummies	yes	yes	yes	yes	yes
Product dummies	yes	yes	yes	yes	yes
Measure of R&D intensity	France	UK	France	France	France
Observations	6348	5336	5704	6348	6348
Log pseudo-likelihood	-7447	-6669	-6251	-7454	-7431

Notes: dependent variable: number of multinationals with positive imports, by country and product. Negative binomial regressions. Main regressors: interactions between country variables and product characteristics. “Fin devt”: private credit over GDP; “R&D intensity”: measured either in UK (column 2) or France (other columns); “Lall index”: equal to one for high- and medium-tech products, zero for low-tech products (Lall, 2000); “Rauch index”: equal to one for goods that are not traded on integrated market or reference priced (Rauch, 1999); “Ext fin dep”: dependence in external finance (Rajan and Zingales, 1998); “Firm size”: median firm size in the corresponding French industry; “HQ intensity”: headquarter intensity; “Patent prot”: patent protection.  $\alpha$ : see equation (10) in the text. Fixed effects by country and by product. Robust standard errors into brackets, corrected for clusters by country; \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.

does not alter the coefficient for financial development.

In columns (2), we check that our results are robust to the inclusion of other controls. In particular, we might still be concerned with whether financial development captures the effect of an alternative country characteristic with a different impact on complex and basic products. We thus interact rule of law with R&D intensity in the place of the Rauch index. The new interaction term with rule of law is not significant. It confirms that rule of law yields the strongest effect under the Rauch (1999) measure of contract intensity. We control for the interaction of capital and skill endowments with the index of R&D intensity in addition to the interaction with capital and skill intensity, but the corresponding coefficients are not significant. We can also include an interaction between R&D intensity and per capital GDP (in log). Although financial

development is positively correlated with income level, the new interaction term is not significant and the coefficient for financial development remains unchanged. Thus, our results are not generated by the level of overall development rather than financial development.<sup>32</sup>

In columns (3) and (4), we check that our results are robust to alternative measures of financial development. We use a measure of the net interest rate margin in column (3) and an index of accounting standards in column (4). The former provides an alternative measure of the (in)efficiency of the banking sector while the latter directly reflects financial regulations. The results indicate strong interaction terms between complexity and both indicators, with the expected signs.

Although our data are particularly rich from many points of view, the cross-sectional nature of this study may raise concerns about the direction of causality. It could be argued that exporting opportunities for local firms in sectors dependent on external finance increase the demand for credit and hence foster the development of the financial sector. Do and Levchenko (2007) rationalize this mechanism. They build a simple model to show that greater foreign demand in financially dependent industries leads to larger volumes of credit and more developed financial markets. Based on the estimated effect of geographic variables on trade volumes across sectors, they construct a country-level variable (“predicted EFNX”) that predicts the value of external finance need of exports. They report a substantial correlation between this variable and financial development. In column (5), we interact this variable with R&D intensity. We find that the coefficient is not significant, irrespective of whether the interaction between private credit and R&D intensity is included or not. This suggests that our results are not driven by observed geographical determinants of foreign demand.

Another argument that can be advanced is that countries with technological advantages in sectors that rely relatively more on external finance will tend to develop their financial systems to exploit these advantages. Our work, however, focus on multinationals’ suppliers. There is extensive evidence that these firms benefit from technology transfers by multinationals (Caves, 2007) and operate different technologies from local firms.<sup>33</sup> Hence, omitting controls on the local technology is unlikely to bias our results. Arguably, the costs of technology transfers can vary across countries depending on the extent of their absorptive capacity but we account for this in our regressions by controlling for human capital interacted with R&D intensity (see column 2).

With cross-sectional data, one way to check the direction of causality is to use an instrument for financial development which does not directly impact on exports. Legal origin provides an exogenous variable that meets these requirements. In particular, common-law countries are better financially developed than civil-law countries. Nevertheless, it has also been shown to have a strong effect on various other aspects of institutional quality (Djankov, MacLiesh and Shleifer, 2007; La Porta *et al.* 2008). Nunn (2007) proposes a simple way to isolate the effect on financial development by matching civil and common law countries

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<sup>32</sup>Moreover, we should note that the inclusion of R&D intensity index interactions with distance, language and colonial links does not affect our main result (not reported).

<sup>33</sup>In Carluccio and Fally (2008) we analyze the effects of technological incompatibilities between foreign and domestic technologies on spillovers and technology adoption.

Table 6: Number of importers, financial development and input complexity: other robustness checks

Dependent variable:	Number of importers, with $E[N] = \exp(bX)$					Log(1+N) (7)	
	(1)	(2)	(3)	(4)	(5)		(6)
Estimation:	NegBin	NegBin	NegBin	NegBin	NegBin	Poisson	OLS
Measure of Fin devt:	Private credit		Net interest margin	Accounting standards	Private credit		
(Fin devt) * (R&D intensity)	9.744 [3.159]***	12.137 [3.520]***	-0.985 [0.320]***	5.795 [0.161]***	9.440 [3.401]**	12.786 [3.259]***	8.967 [2.068]***
(Rule of law) * (R&D intensity)		-27.325 [25.565]					
(K/L) * (R&D intensity)		4.328 [4.670]					
(H/L) * (R&D intensity)		11.901 [16.588]					
(Log CGDP) * (R&D intensity)		-1.003 [8.124]					
(Pred EFNX) * (R&D intensity)					-0.019 [0.357]		
(Patent prot) * (R&D intensity)		-2.658	1.263	-4.323*	-0.275	-3.658*	2.241***
(Rule of law) * (Rauch index)		2.097***	1.947***	1.599*	2.039***	0.863	0.239*
(H/L) * (Skill intensity)	4.116***	4.026***	4.127***	3.923***	4.234***	4.479***	2.122***
(K/L) * (Capital intensity)	0.129**	0.237***	0.235***	0.233**	0.223***	0.231***	0.052***
$\alpha$	0.543	0.525	0.537	0.458	0.530	/	/
Country and product dummies	yes	yes	yes	yes	yes	yes	yes
Observations	6624	6141	5589	2553	6279	6348	6348
Log pseudo-likelihood & R2	-7502	-7260	-7367	-5755	-7445	-10439	0.79

Notes: dependent variable: number of multinationals with positive imports, by country and product. Negative binomial regressions (columns 1 to 5), QMLE Poisson regression (column 6) or OLS regression (column 7). Main regressors: interactions between country variables and product characteristics. "Log CGDP": log per capita GDP; "Pred EFNX": predicted share of exports in financially dependent sectors.  $\alpha$ : see equation (10) in the text. Fixed effects by country and by product. Robust standard errors into brackets, corrected for clusters by country; \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.



with similar characteristics in control variables. In unreported estimations, we apply a similar method and find a strong and positive effect of common law on the composition of exports in complex products, after matching countries by rule of law, level of intellectual property rights protection and factor endowments (results available upon request).

In all regressions specifying a Negative Binomial distribution for the residual, the estimated coefficient  $\alpha$  is approximately 0.5. The estimated standard error is 0.08 and the likelihood-ratio test for  $\alpha = 0$  (corresponding to the Poisson distribution) is rejected.<sup>34</sup> It shows that the residual's variance is strictly larger than the mean, which implies that the residual does not follow a Poisson distribution (overdispersion). In columns (6) and (7), we check alternative specifications to verify that our results do not depend on our assumptions regarding the error term distribution. In column (6), we estimate a Poisson regression with quasi-maximum likelihood (QMLE). Gourieroux, Monfort and Trognon (1984) show that Poisson estimations are consistent under misspecifications of the error structure. Our estimated coefficients are very similar, with the main interaction term being slightly higher. The only difference concerns the rule of law coefficient, which is no longer significant. However, note that the log-likelihood is much lower for the Poisson specification compared to the Negative Binomial – it illustrates the better fit obtained with the former. In column (7), we run a simple OLS regression but, in order to accommodate for zeros, the dependent variable is the logarithm of one plus the number of importers. We obtain similar coefficients for the interaction terms between financial development and complexity. However, several controls yield different coefficients: patent protection now enters positively and significantly, whereas the estimated impact of rule of law is less significant compared to negative binomial specifications.

We perform additional robustness checks in Appendix B. In particular, we test whether our results are sensible to the identification of intermediate goods, as opposed to goods that are not further processed by the importer. Moreover, we test whether our results are not driven by the presence of export platforms by restricting to groups with headquarters in France.

Finally, we have checked that our results are not sensitive to minor industry or country sample changes. For example, the coefficient remains equally similar when we drop countries with less than one million inhabitants, African countries, OECD countries or Asian tigers. Neither is the coefficient sensitive to dropping private credit or R&D intensity percentiles either end of the scale.

All of these results confirm that the country of origin's financial development is strongly correlated with the number of multinational firms importing complex inputs compared to basic inputs.

But how sizeable is the effect of financial development on the number of multinational firms importing complex inputs? In view of our estimation approach, we consider the effect of financial development by comparing two industries that differ in one standard deviation in terms of R&D intensity. Using the coefficient of the interaction term, we estimate the differential effect of a one standard deviation change in financial development. Given the coefficient in column (1) of Table 5, we obtain a differential effect of 9% on the number of importers. This effect is large and is comparable to the effect of skill endowment on the number of importers (differential effect of 18% by comparing low- and high-skill-intensive industries). Moreover, the estimated effect of financial development is stable across the different specifications in columns (2) through (4) of Table 5 (10% with the Rauch index, 8% with the Lall index and 8% with R&D intensity calculated from the UK data). By comparison, the rule of law effect is larger if we take the column (1) estimation, but smaller taking column (3) in Table 5.

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<sup>34</sup>This test does not account for heteroskedasticity, but the P-value is lower than 0.001.

### 3.3 Financial development and the scale of imports

So far, we have tested Prediction 1 by looking at the impact of financial development and input complexity on the number of multinational firms by country and industry with positive imports. In Table 7, we examine whether the financial development effect is also observed using the aggregate value of imports by sector and country as well as firm-level imports. Yet the total value of imports by product and country of origin could well capture various trade characteristic aspects. Firstly, the value depends on the number of multinationals importing. However, we can eliminate this effect by taking firm-level imports and analyzing within-firm variations. As we do not observe trade flows for each supplier, the value depends on the number of suppliers. It also depends on the scale of production that is offshored. Lastly, it is affected by the unit value of the traded goods as we do not have trade flows in quantities.

As firm-level data on trade is often not available, most studies perform regressions on the log of aggregate trade flows (Anderson and Wincoop, 2004). For the first two columns of Table 7, we first specify the following equation for aggregate imports  $M_{cs}$  by product  $s$  and country  $c$ :

$$\text{Log}(M_{cs}) = \beta.(FinDev_t_c * Complexity_s) + \gamma.Controls_{cs} + \eta_c + \zeta_s + \varepsilon_{cs} \quad (11)$$

Note that observations of zero trade flows are naturally excluded from this regression – taking zeros into account by estimating Poisson or Negative Binomial regressions (as in Silva Santos and Teryeno, 2006) provides results close to our estimations of the number of importers.

We also use our firm-level data to test whether financial development has a differential impact on firm-level imports of complex products. To some extent, effects on firm-level trade flows correspond to the intensive margin of imports. Specifically, we can estimate:

$$\text{Log}(M_{csi}) = \beta.(FinDev_t_c * Complexity_s) + \gamma.Controls_{cs} + \eta_c + \zeta_s + \phi_i + \varepsilon_{csi} \quad (12)$$

where  $\phi_i$  is a fixed effect by firm. This is estimated for all observations by firm, country and product with positive trade flows. Specifications (11) and (12) are estimated by a simple OLS. As financial development only varies by country, we report robust standard errors corrected for clusters by country.

Table 7 presents our findings on import value. In columns (1) and (2), we observe that financial constraints interacted with the R&D intensity index have a strong effect on the aggregate value of imports, controlling for factor endowments. The coefficient is not sensitive to the inclusion of other controls (column 2). Qualitatively, the results in column (2) are in line with the results in column (1) of Table 5: rule of law has a significant coefficient, but not intellectual property rights. Again, these results are robust to the interaction of financial development with the three alternative indices (not reported).

In columns (3) through (7), the dependent variable is the value of imports at firm level. There is no firm dummy in column (3), but firm fixed effects are systematically included in columns (4) to (7). The results find that financial development has a significant impact on firm-level imports. A comparison of columns (3) and (4) shows that controlling for firm dummies does not affect the financial development coefficient.

Interestingly, the impact of financial development on firm-level exports roughly corresponds to the difference between the estimated impact on aggregate imports (column 1) and the estimated impact on the number of importers (column 1 of Table 5), which is consistent with the simple equality between total imports and the product of the average of firm level imports with the number of importers.

Table 7: Import values, financial development and input complexity

Dependent variable:	Log aggregate imports			Log firm-level imports			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
(Fin devt) * (R&D intensity)	17.593 [4.949]***	16.650 [5.018]***	6.377 [1.828]***	7.522 [1.691]***	9.073 [2.164]***	6.954 [2.123]***	7.082 [1.607]***
(Fin devt) * (Ext fin dep)					-0.158 [0.162]		
(Fin devt) * (Capital intensity)					0.210 [0.086]**		
(Fin devt) * (Firm size)					-0.203 [0.099]**		
(Fin devt) * (HQ intensity)					0.928 [0.849]		
(Rule of law) * (R&D intensity)						2.745 [6.202]	
(Log CGDP) * (R&D intensity)						2.181 [5.054]	
(Pred EFNX) * (R&D intensity)							0.349 [0.163]**
(Patent prot) * (R&D intensity)		1.659	-2.005*	-2.259**	-2.220**	3.527	-2.283
(Rule of law) * (Rauch index)		2.269**	-0.442	-0.864	-0.712	-0.958*	-0.841*
(H/L) * (Skil intensity)	8.480***	8.257***	2.442***	2.456***	2.405***	2.515***	2.567***
(K/L) * (Capital intensity)	0.184**	0.302***	0.008	0.022	0.018	0.014	0.016
Firm dummies	no	no	no	yes	yes	yes	yes
Country and product dummies	yes	yes	yes	yes	yes	yes	yes
Observations	2027	2021	52617	52617	52617	52164	52613
R-squared	0.60	0.61	0.07	0.25	0.25	0.25	0.25

Notes: dependent variable: value of imports, by country and product (columns 1 and 2) or by firm, country and product (columns 3 to 7). OLS regressions. Main regressors: interactions between country variables and product characteristics. “Fin devt”: private credit over GDP; “R&D intensity”: measured in France; “Ext fin dep”: dependence in external finance (Rajan and Zingales, 1998); “Firm size”: median firm size in the corresponding French industry; “HQ intensity”: headquarter intensity; “Log CGDP”: log per capita GDP; “Pred EFNX”: predicted share of exports in financially dependent sectors based on geographical characteristics; “Patent prot”: patent protection. Fixed effects by country and by product; fixed effects by firm (columns 4 to 7). Robust standard errors into brackets, corrected for clusters by country; \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.

We perform the same robustness checks as for the previous tables. In column (5), we check that our financial development findings are not driven by the interaction with external finance dependence, capital intensity, firm size or headquarter intensity measured in the corresponding sector in France. Although the interactions with capital intensity and firm size are marginally significant, the main coefficient remains unchanged. In column (6), we interact rule of law with the R&D intensity index in the place of the Rauch index of contract intensity: the coefficient for rule of law becomes insignificant whereas the coefficient for credit remains. Moreover, our results are robust to controlling for the interaction term between per capita GDP and R&D intensity.<sup>35</sup> Finally, in column (7) we control for R&D intensity interacted with Do and Levchenko (2007) variable EFNX (prediction of financial development based on geographic characteristics). The coefficient is now significant at 5% but our interaction is not affected.

The Table 7 results show that the impact of financial development on the import of complex inputs carries over different specifications and different margins. It affects not only the number of importers, but also total imports and firm-level imports. Quantitatively, financial development has less of an effect on firm-level imports (column 4 of Table 7) than on the number of importers (column 1 of Table 5). Consistently, the sum of these effects equals the effect on total imports (column 2 of Table 7).

### 3.4 Financial development and sourcing mode

We now investigate the model’s theoretical predictions with regards to the relation between financial constraints and optimal organizational mode. Our theory predicts that integration is preferred to trade with independent suppliers when financial constraints are strong and trade involves complex products with inputs requiring a larger proportion of specific, non-contractible investments (Prediction 2). To provide evidence of this statement, the following econometric analysis uses as its dependent variable the share of intra-firm imports by firm for each product-country pair.<sup>36</sup>

Table 4 presents some descriptive statistics. It shows that the correlation between our measure of financial development and the share of intra-firm imports is substantially stronger in the case of complex goods (Panel A). Conversely, the correlation between the average share of intra-firm imports and R&D intensity is stronger in countries with low financial development. However, these correlations might be driven by the effects of variables other than financial development, such the general level of contract enforcement. So the above simple correlations, although informative, could be misleading.<sup>37</sup>

In order to overcome this problem, we take a similar approach to the previous section and focus our analysis on interaction terms. In this case, we also look at variations in financial development across countries and variations in complexity across industries. As already mentioned, this strategy is appealing because it has the advantage of controlling for all industry- and country-specific characteristics that we do not observe and that might affect intra-firm shares. In order to be sure that we are picking up the effect of financial development, we control for several alternative interaction terms.

The empirical strategy we employ is in line with the difference-in-difference approach taken in the

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<sup>35</sup>In addition, we can control for interaction terms of skill and capital endowments with R&D intensity but it does not affect our results.

<sup>36</sup>Our data is disaggregated at product level, so we are unable to observe the value and sourcing mode governing the relationship with each particular supplier. Our product-level data consequently contains “mixed” trade flows, where the same product has been imported by the same firm by means of both intra-firm trade and outsourcing.

<sup>37</sup>Acemoglu *et al.* (*forthcoming*) show that simple cross-country comparisons may be driven by industry composition. Once industry fixed effects are included in their regressions, the direct effects of institutional quality on the organizational mode are generally not significant.

previous section. More specifically, our estimating equation is given by:

$$I_{csi} = \beta.(FinDevt_c * Complexity_s) + \gamma.Controls_{cs} + \alpha_i + \eta_c + \zeta_s + \varepsilon_{csi} \quad (13)$$

where  $FinDevt_c$  is the financial development variable for country  $c$ ,  $Complexity_s$  is the complexity index for traded input  $s$ , and  $\alpha_i$ ,  $\eta_c$  and  $\zeta_s$  are fixed effects by firm, country and industry respectively. Thus, our identification strategy focuses on within-firm variations in sourcing modes across industry-country cells. We estimate the above equation using Ordinarily Least Squares (OLS) and report robust standard errors with correction for clusters by country. One potential problem with estimating (15) by OLS is that our dependent variable is defined as a share and is therefore bounded between zero and one. To account for this, we can estimate the model under a fractional Logit specification and obtain very similar results (available upon request). As a robustness check, we also provide results from the estimation of a conditional Logit model using only “pure” organizational strategies (i.e. when the share of intra-firm trade in an industry-country cell is either 100% or zero). Note, however, that such an approach potentially loses valuable information, as around 13% of our observations are mixed.

Results are presented in Table 8. The coefficient of main interest is, as before, the interaction between financial development and complexity. As can be seen from column (1), the coefficient is negative and significant. This result is robust to the inclusion a full set of country, industry and firm fixed effects, which are included in all regressions presented in this section.

Note that we obtain very similar coefficients for financial development whether firm dummies are included or not. It implies that the effects we are picking up are not due to any unobserved firm characteristics that may impact systematically on the choice between intra-firm and arm’s length trade – for example, Antras and Helpman (2004 and 2008) predict that more productive firms tend to trade relatively more within the firm than less productive firms.<sup>38</sup>

In addition, as can be seen from the table, all the regressions include the following set of controls. First, the index of patent protection is interacted with R&D intensity. Multinational firms may prefer to source complex inputs from an affiliate in order to reduce technological imitation and appropriation by local competitors, and these forces might be stronger in environments where intellectual property rights are weakly enforced. This coefficient is only significant in the case where complexity is used using the UK data, and does not affect the sign and significance of our main coefficient. Secondly, we control for judicial quality, interacted with the Rauch index. Foreign firms’ internalization decisions may be affected by the quality of judicial system; rule-of-law effects could arguably be greater in contract-dependent industries. This coefficient is positive in all specifications, consistent with the findings of Defever and Toubal (2008) for French firms. In addition, its inclusion does not affect the relevance of the financial development variable. Lastly, all of our regressions include controls for traditional comparative advantage determinants, interacting capital intensity and skill intensity at industry level with capital and skill endowments at country level respectively. Note that we have chosen to report results from regressions including all of these controls together in order to limit the number of columns. Yet our results remain the same when we add one control at a time, and are similar when we exclude firm fixed effects.<sup>39</sup>

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<sup>38</sup>Defever and Toubal (2008) and Corcos *et al.* (2008) test this prediction using the same dataset.

<sup>39</sup>In unreported estimates, we also control for the interaction between per capita GDP and complexity to check whether we are indeed picking up the level of overall development rather than a financial channel. This interaction term comes out positive, but not significant.

Table 8: Intra-firm imports, financial development and input complexity

Dependent variable:	Share of intra-firm imports						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	OLS	OLS	OLS	OLS	OLS	OLS	Logit
Estimation:							
(Fin devt) * (R&D intensity)	-0.702 [0.162]***	-1.579 [0.118]***			-0.788 [0.202]***	-0.790 [0.162]***	-7.370 [2.446]***
(Fin devt) * (Lall index)			-0.037 [0.013]***				
(Fin devt) * (Rauch index)				-0.037 [0.013]***			
(Fin devt) * (Ext fin dep)					0.029 [0.021]		
(Fin devt) * (Capital intensity)					0.009 [0.014]		
(Fin devt) * (Firm size)					-0.007 [0.016]		
(Fin devt) * (HQ intensity)					0.059 [0.113]		
(Rule of law) * (R&D intensity)						0.657 [1.002]	
(Rule of law) * (Capital intensity)						-0.067 [0.042]	
(Pred EFNX) * (R&D intensity)						0.012 [0.024]	
(FDI restrictions) * (R&D intensity)						0.005 [0.012]	
(Patent prot) * (Lall index)			0.018*				
(Patent prot) * (R&D intensity)	0.211	0.595***		0.109	0.225*	0.075	0.884
(Rule of law) * (Rauch index)	0.118***	0.119***	0.145***	0.170***	0.127***	0.046	1.514***
(H/L) * (Skill intensity)	0.039	0.044	-0.013	-0.017	0.015	0.064	0.492
(K/L) * (Capital intensity)	0.010**	0.011**	0.003	0.011**	0.009*	0.015**	0.088
Firm dummies	yes	yes	yes	yes	yes	yes	yes
Country and product dummies	yes	yes	yes	yes	yes	yes	yes
Measure of R&D intensity	France	UK	France	France	France	France	France
Observations	52617	49147	43548	52617	52617	52613	50159
R-squared or Log likelihood	0.65	0.65	0.67	0.65	0.65	0.65	-9672

Notes: dependent variable: share of intra-firm imports, by firm, country and product. OLS regressions except in column 7. Column 7: Logit regression, dropping mixed strategies (keeping observations where the dependent variable is either 0 or 1). Main regressors: interactions between country variables and product characteristics. "Fin devt": private credit over GDP; "R&D intensity": measured either in UK (column 2) or France (other columns); "Lall index": equal to one for high- and medium-tech products, zero for low-tech products (Lall, 2000); "Rauch index": equal to one for goods that are not traded on integrated market or reference priced (Rauch, 1999); "Ext fin dep": dependence in external finance (Rajan and Zingales, 1998); "Firm size": median firm size in the corresponding French industry; "HQ intensity": headquarter intensity; "Log CGDP": log per capita GDP; "Pred EFNX": predicted share of exports in financially dependent sectors based on geographical characteristics; "Patent prot": patent protection. Fixed effects by firm, by country and by product. Robust standard errors into brackets, corrected for clusters by country; \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.

We perform various robustness checks to ensure that it is really the interaction of financial development and R&D intensity that drives the intra-firm share effects on which we report. In column (2), we use R&D intensity calculated for three-digit UK industries as a measure of product complexity. The sign of the interaction between this variable and financial development is negative, significant and greater than that in column (2). However, the magnitude of the effects of both measures is similar if we normalize the coefficient by the respective standard deviation. In columns (3) and (4), we again use the Lall and Rauch indices and obtain significant coefficients with the expected signs. These results indicate that trade is more likely to occur with an affiliate when financial development is low and the level of complexity is high, as predicted by the theoretical analysis.

It might be argued that if complex industries are more dependent on external finance, then our results might be picking up the effect of external finance dependence instead of complexity. In column (5), the inclusion of the interaction between financial development and the external finance index confirms that this is not the case. In the same column, we include interaction between financial development and two alternative measures of fixed costs constructed at industry level: capital intensity and firm size. These controls are included to ensure that R&D intensity reflects technological complexity and not the size of the industry's fixed costs. None of these controls turns out to be significant.

In column (6), we address other important questions by including four additional controls. We first include an interaction between the rule-of-law measure and R&D intensity. This coefficient is not significant, while the interaction between financial development and R&D intensity remains negative and significant. Therefore, our results on the impact of financial development on the share of intra-firm trade are not driven by the overall level of contract enforcement. Secondly, we interact rule of law with capital intensity at industry level. The coefficient is negative, but not statistically different from zero; note that its sign is in line with the prediction by Antras and Helpman (2008) and the empirical results of Bernard *et al.* (2008). Thirdly, as in tables 6 and 7, we control for the interaction between R&D intensity and EFNX (which predicts the level of financial development based on geographic characteristics). Fourthly, we may be concerned with whether financial development, measured by private credit to GDP, is endogenous to the intensity of foreign direct investment in the country. However, we show that the inclusion of a measure of restrictions on foreign investors does not alter our result. Note that we obtain a similar result when controlling directly for FDI inflows.

Finally, column (7) presents the results of the conditional Logit estimation, including firm fixed effects and other controls, from which mixed strategies are dropped (the dependent variable is thus either equal to zero or unity). Again, the sign and level of significance of our coefficient of interest are in line with Prediction 2.

We also perform robustness checks on the estimating sample. We run the above regressions minus Asian, African and small countries (less than one million inhabitants) and obtain the same results. Our results are robust to the elimination of one country at a time (obtaining even stronger results when we drop Japan and the USA, which have the highest ratios of private credit to GDP). In Appendix B, we examine narrower definitions of intermediate imports and restrict our sample to groups with headquarters in France. We find similar and even stronger results.

## 4 Conclusion

In recent years, international trade literature has devoted substantial efforts to understand the role of institutions in shaping trade patterns. In this paper, we propose a new mechanism driven by an interaction between financial development and the extent of contract incompleteness. We argue that suppliers' financial constraints exacerbate the hold-up problem and impact on multinational firms' optimal sourcing strategies. These effects are stronger in complex industries characterized by higher degrees of contract incompleteness. Further, the model shows that vertical integration alleviates financial constraints, albeit at the cost of lower efforts and productivity.

The model generates two novel predictions. First, we predict that financial development creates a comparative advantage in the supply of complex products. Second, we should observe higher shares of intra-firm trade in complex products from countries with a low level of financial development. These predictions are tested using detailed firm-level data on the international operations of multinational firms located in France. We use a difference-in-difference approach, which allows for the inclusion of country and product fixed effects. A one standard deviation increase in the level of financial development yields an increase of 9% in the number of multinational firms importing complex goods compared to basic goods, an increase of 7% in the volume of firm-level imports and a decrease of 3% in the share of intra-firm trade.<sup>40</sup> These results are robust to different measures of the complexity and specificity of the traded goods, to alternative measures of the level of financial development and to the inclusion of firm fixed effects among other controls. The estimated effect of financial development appears to be as large as that of contract enforcement.

As pointed up by the model, the positive effect of financial development on international trade is not limited to isolated exporters but carries over into multinationals' suppliers. Multinational firms are known to play an important role in the transfer of technologies to suppliers located in less developed countries, where the adoption of technologies constitutes a major source of productivity growth. One implication of our model is that suppliers located in countries with poor financial institutions are less likely to acquire new technological knowledge to produce complex products. This mechanism may explain, for example, why the presence of multinational firms has a stronger impact on growth in countries with better financial institutions (Alfaro *et al*, 2004).

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<sup>40</sup>Based on a comparison of goods that differ by one standard deviation in their level of complexity.



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## Appendix A: Theoretical appendix on extensions

### Constrained investments

In this extension, we find that the financial constraint (FC) starts to be binding when initial capital  $W$  is below the threshold:

$$\bar{W}^o = \frac{1-\kappa}{2} a\theta (2c)^{1-\sigma} (1+\alpha)^\sigma$$

This threshold  $\bar{W}^o$  verifies the same properties as in Proposition 1: it is decreasing in  $\kappa$ , especially when  $\theta$  is large.

When only the financial constraint (FC) is binding, profits under outsourcing equal:

$$\Pi^o = W + a\theta(1-\rho) \left[ 1 - \left( \frac{1+\alpha}{2} \right) + \kappa \left( \frac{1+\alpha}{2} \right) \right]^\sigma c^{\sigma-1} + a(1-\theta)(1-\rho)c^{\sigma-1}$$

However, both the financial constraint (FC) and the participation constraint (PC) may be binding for intermediate values of  $W$ . In that case, we find that the level of effort depends directly on  $W$  and that profits equal:

$$\Pi^o = \left( \frac{2W}{(1-\kappa)(1+\alpha)} \right) - \theta c \left( \frac{2W}{(1-\kappa)(1+\alpha)\theta A^{1-\rho}} \right)^{\frac{1}{\rho}} + a(1-\theta)(1-\rho)c^{\sigma-1}$$

Both expressions yield results that confirm Proposition 2: the multinational’s profits increase in  $\kappa$ , especially for high  $\theta$ .

There are also two cases for the organizational threshold. At the point where the multinational is indifferent between outsourcing and integration, we can show that the financial constraint (FC) is binding under outsourcing and the financial constraint (FC) is not binding under integration.

However, the participation constraint may also be binding at this point under outsourcing. If this is the case, we obtain the following organizational threshold:

$$\bar{W}^{I/O} = a\theta(1-\kappa) (2c)^{1-\sigma} \delta^\sigma \left( \frac{1+\alpha}{2} \right)$$

If this is not the case, the organizational threshold is:

$$\bar{W}^{I/O} = a\theta (2c)^{1-\sigma} \left[ \delta^{\sigma-1} \left( 1 - \frac{\delta\rho}{2} \right) - \frac{1}{2-\kappa} + \frac{\rho}{2} \right]$$

Both thresholds are decreasing in  $\kappa$ , especially for high  $\theta$ , and verify Proposition 3.

### Contingent debt repayment and bargaining

In this extension of the model, we find that the financial constraint (FC) starts to be binding when initial capital is below the threshold:

$$\bar{W}^o = \frac{1-\kappa}{2-\kappa} a\theta (2c)^{1-\sigma}$$

This threshold is decreasing in  $\kappa$ , especially for high  $\theta$ , which confirms Proposition 1.

When the financial constraint is binding, we find that the multinational’s total profits equal:

$$\Pi^o = W + a\theta \left( \frac{1}{2-\kappa} - \frac{\rho}{2} \right) (2c)^{1-\sigma} + a(1-\theta)(1-\rho)c^{1-\sigma}$$

which is increasing in  $\kappa$ , especially for high  $\theta$ , which confirms Proposition 2.

The organizational threshold  $\overline{W}^{I/O}$  is now equal to:

$$\overline{W}^{I/O} = a\theta (2c)^{1-\sigma} \left[ \delta^{\sigma-1} \left(1 - \frac{\delta\rho}{2}\right) - \frac{1}{2-\kappa} + \frac{\rho}{2} \right]$$

which is decreasing in  $\kappa$ , especially for high  $\theta$ , and verifies Proposition 3.

### A general production function

In this extension, total revenues are given by a general function  $Y(e, q)$  that depends on non contractible efforts  $e$  and contractible efforts  $q$ . Costs are linear in both types of efforts:  $C = c_e \cdot e + c_q \cdot q$ . As in the baseline model, outside options of the multinational firm under outsourcing correspond to revenues  $Y(0, q)$  when non-contractible tasks equal zero, while the supplier's outside option is nil.

Revenues of the supplier after bargaining equal  $[\frac{1}{2}Y(e, q) - \frac{1}{2}Y(0, q)]$ . Investments in complex tasks are chosen by the supplier in order to maximize its ex post revenues minus the costs. Under outsourcing, we obtain investments in complex tasks  $e^*(q)$  as a function of investments in contractible tasks:<sup>41</sup>

$$e^*(q) = \arg \max_e \left\{ \frac{1}{2}Y(e, q) - c_e \cdot e \right\}$$

When investments in complex and basic tasks are strong complements ( $\frac{\partial^2 Y}{\partial e \partial q} > 0$ ), we find investments in complex tasks are increasing in investments in basic tasks.

Investments in contractible tasks are determined ex ante by the multinational firm, by anticipating the supplier's investments  $e = e^*(q)$ . When the financial constraint is slack, the participation constraint is binding and total profits of the multinational equal the entire surplus. In this case, the multinational chooses  $q$  to maximize

$$\Pi = Y(e, q) - c_e \cdot e - c_q \cdot q$$

In the opposite case where the initial capital of the supplier is zero ( $W = 0$ ), only the financial constraint is binding and the multinational's total profits are the sum of ex post revenues plus the debt that can be contracted by the supplier. We find that the multinational chooses  $q$  to maximize:

$$\Pi = \left[ \frac{1}{2}Y(e, q) + \frac{1}{2}Y(0, q) \right] + \kappa \left[ \frac{1}{2}Y(e, q) - \frac{1}{2}Y(0, q) \right] - c_e \cdot e - c_q \cdot q$$

where  $e = e^*(q)$ . The optimal choice in contractible investments is a function  $q^*(\kappa)$  of the level financial development  $\kappa$  where  $q^*(1)$  corresponds to the optimal choice when the financial constraint is not binding. When investments in complex and basic tasks are strong complements ( $\frac{\partial^2 Y}{\partial e \partial q} > 0$ ), we find investments in basic tasks are increasing in investments in financial development.

By applying the envelop theorem, the total derivative of the multinational's profits according to  $\kappa$  equals the partial derivative of the above expression taken at optimal choices. We find the expression described in the text:

$$\frac{d\Pi}{d\kappa} = \frac{1}{2} [Y(e, q) - Y(0, q)] > 0$$

Concerning the choice between outsourcing and integration, we can obtain a similar result if we assume that only the participation constraint is binding under integration and only the financial constraint is binding under outsourcing at the point where the multinational is indifferent. On the one hand, when only the financial constraint is binding under outsourcing, the multinational's profits are linear in the initial capital of the supplier. On the other hand, the multinational's profits under integration do not depend on financial development when the financial constraint is not binding. Therefore, we find that the effect of financial development on the organizational threshold is the opposite compared to the effect on profits under outsourcing:

$$\frac{dW^{I/O}}{d\kappa} = -\frac{1}{2} [Y(e, q) - Y(0, q)] < 0$$

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<sup>41</sup>With the symmetric Nash bargaining solution and zero outside option for the supplier, investments are not constrained in liquidity by the multinational, as in the baseline model.

## **Appendix B: Robustness checks on the identification of vertical trade and supply chains**

In this appendix we present results from robustness checks regarding the identification of vertical supply chains. Following the methodology adopted by Feenstra and Hanson (1996), we drop imports of goods that are classified in the same category as the main product of the importing business unit within the group (and thus we drop business units of manufacturing groups that are not primarily classified as manufacturing). As can be seen in the odd columns of Table 9, we observe that our main coefficient remains unchanged under this specification, and it sometimes become stronger.

In even columns, we propose further robustness checks by restricting our sample to firms with headquarters in France (i.e. French multinationals). As can be seen, this does not affect our results either.

Table 9: Appendix B - robustness checks: trade in intermediate goods, firms with headquarter in France

Dependent variable:	(1)		(2)		(3)		(4)		(5)		(6)		(7)		(8)		
	NegBin		Negbin		OLS		OLS		OLS		OLS		OLS		OLS		
Estimation:																	
(Fin devt) * (R&D intensity)	11.201 [3.531]***		11.238 [3.939]***		21.037 [6.360]***		18.743 [7.857]**		6.017 [1.714]***		8.145 [2.905]***		-0.818 [0.224]***		-0.838 [0.309]***		
(Patent prot) * (R&D intensity)	-2.495 [2.016]		-0.986 [2.286]		0.259 [3.691]		1.552 [4.326]		-2.891 [1.406]**		-2.145 [1.843]		0.415 [0.202]**		0.411 [0.390]		
(Rule of law) * (Rauch index)	1.812 [0.734]**		1.434 [0.737]*		2.061 [1.304]		0.829 [1.317]		-1.247 [0.708]*		-2.104 [0.832]**		0.139 [0.042]***		-0.029 [0.057]		
(H/L) * (Skil intensity)	5.535 [1.311]***		5.258 [1.204]***		6.047 [2.170]***		5.958 [1.757]***		1.778 [0.928]*		1.380 [0.977]		0.117 [0.095]		0.177 [0.120]		
(K/L) * (Capital intensity)	0.229 [0.088]***		0.197 [0.086]**		0.404 [0.121]***		0.302 [0.122]**		-0.024 [0.078]		0.023 [0.104]		0.009 [0.006]		-0.000 [0.007]		
Headquarter in France		yes		yes		yes		yes		yes		yes		yes		yes	
Firm dummies	no	no	no	no	no	no	no	no	yes	yes	yes	yes	yes	yes	yes	yes	yes
Country and product dummies	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Observations	4899	4556	4556	4556	1568	1568	1307	1307	31161	31161	14686	14686	31161	31161	14686	14686	14686
R-squared or Log likelihood	-5380	-4133	-4133	-4133	0.58	0.58	0.56	0.56	0.24	0.24	0.26	0.26	0.61	0.61	0.60	0.60	0.60

Notes: we drop imports in the same category as the main good produced by the importing business unit. In even columns, we restrict the sample to firms with headquarters in France. Dependent variable: number of multinationals with positive imports, by country and product (columns 1 and 2); value of imports, by country and product (columns 3 and 4) or by firm, country and product (columns 5 and 6); share of intra-firm imports by firm, country and product (columns 7 and 8). Negative binomial regressions and OLS regressions. Main regressors: interactions between country variables and product characteristics. "Fin devt": private credit over GDP; "R&D intensity": measured in France; "Patent prot": patent protection. Fixed effects by country and by product; fixed effects by firm (columns 5 to 8). Robust standard errors into brackets, corrected for clusters by country; \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.